

## Problem 2

```
> library(faraway)
> data(prostate)
> nrow(prostate)
[1] 97
> save(prostate, file="prostate.Rdata")
```

(a)

```
> fit = lm(lpsa~l cavol +lwei ght+age+l bph+svi +l cp+gl eason+pgg45)
> summary(fit)
```

Call:

```
lm(formula = lpsa ~ l cavol + lweight + age + l bph + svi + l cp +
    gleason + pgg45)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.7331	-0.3713	-0.0170	0.4141	1.6381

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	0.669337	1.296387	0.516	0.60693
l cavol	0.587022	0.087920	6.677	2.11e-09 ***
lweight	0.454467	0.170012	2.673	0.00896 **
age	-0.019637	0.011173	-1.758	0.08229 .
l bph	0.107054	0.058449	1.832	0.07040 .
svi	0.766157	0.244309	3.136	0.00233 **
l cp	-0.105474	0.091013	-1.159	0.24964
gleason	0.045142	0.157465	0.287	0.77503
pgg45	0.004525	0.004421	1.024	0.30886

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.7084 on 88 degrees of freedom

Multiple R-squared: 0.6548, Adjusted R-squared: 0.6234

F-statistic: 20.86 on 8 and 88 DF, p-value: < 2.2e-16

(b)

Null hypothesis Ho:

Alternative hypothesis Ha:

```
> fit0 = lm(lpsa~1) # Null model
> anova(fit0, fit) # anova(Null model, Full model)
Analysis of Variance Table
```

Model 1: lpsa ~ 1

Model 2: lpsa ~ l cavol + lweight + age + l bph + svi + l cp + gleason + pgg45

	Res. Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	96	127.918				
2	88	44.163	8	83.755	20.861	< 2.2e-16 ***

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

F test statistic = 20.861, and  $p\text{-value} < 2.2 \times 10^{-16} < \alpha = 0.05$ .

Therefore, **Reject  $H_0$**  at  $\alpha = 5\%$  significance level.

(c)

90% Confidence Interval:

```
> confint(fit, "age", level=0.90)
           5 %          95 %
age -0.0382102 -0.001064151
```

95% Confidence Interval:

```
> confint(fit, "age", level=0.95)
           2.5 %          97.5 %
age -0.04184062  0.002566267
```

For regression summary:  $H_0: \text{Parameter}(\text{age}) = 0$

90% CI does not cover 0, Reject  $H_0$ , so  $p\text{-value} < 0.1$

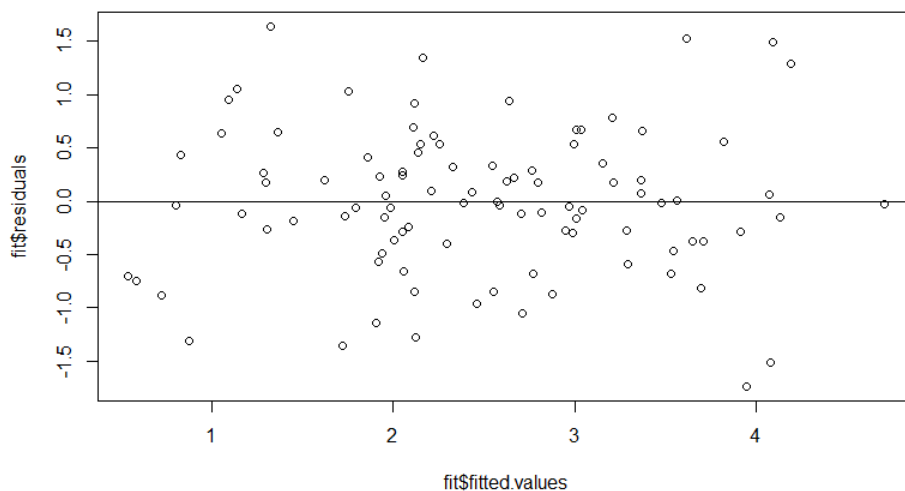
95% CI covers 0, Do Not Reject  $H_0$ , so  $p\text{-value} > 0.05$

Therefore, we can deduce that  $0.05 < p\text{-value} < 0.1$

Indeed, the regression summary shows that the parameter associated with age has  $p\text{-value} = 0.08229$ .

(d)

```
> plot(fit$fitted.values, fit$residuals)
> abline(h = 0)
```



```
> library(lmtest)
> bptest(fit)
```

studentized Breusch-Pagan test

```
data: fit
BP = 10.08, df = 8, p-value = 0.2594
```

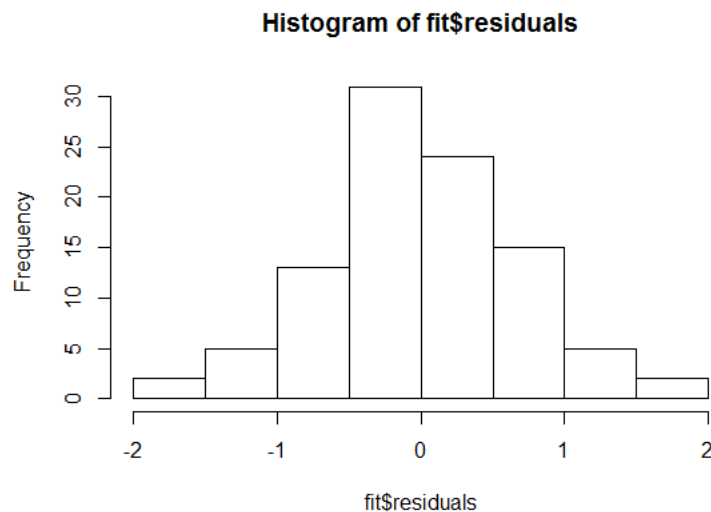
#### Result and Comment:

Fitted vs. Residual Plot does not seem to show any obvious pattern in the variance of the residuals, thus not showing clear evidence for a non-constant variance.

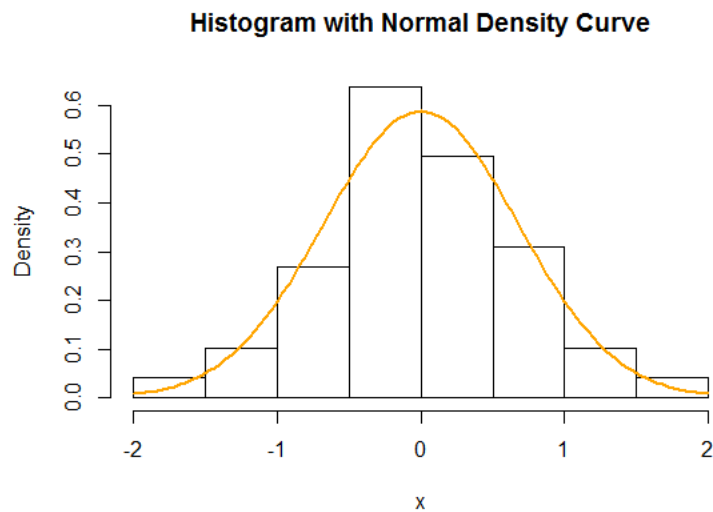
And the Breusch-Pagan test has a  $p\text{-value}=0.2594 > 0.1$ , **Do Not Reject Null Hypothesis (homoscedasticity)** at  $\alpha = 10\%$  significance level or smaller. Therefore, the residuals' variance is basically constant.

(e)

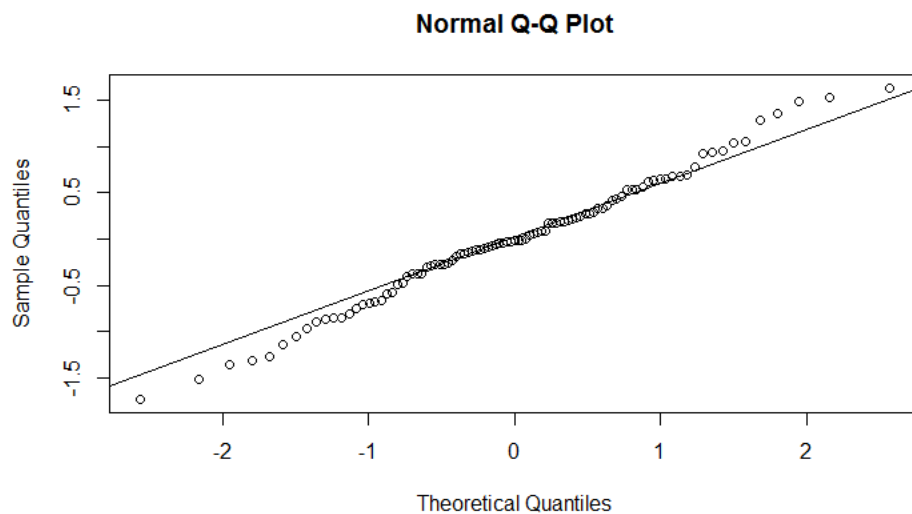
```
> hist(fit$residuals)
```



```
> histNorm <- function(x, densCol = "darkblue"){
+   m <- mean(x)
+   std <- sqrt(var(x))
+   h <- max(hist(x, plot=FALSE)$density)
+   d <- dnorm(x, mean=m, sd=std)
+   maxY <- max(h, d)
+   hist(x, prob=TRUE,
+        xlab="x", ylim=c(0, maxY),
+        main="Histogram with Normal Density Curve")
+   curve(dnorm(x, mean=m, sd=std),
+         col=densCol, lwd=2, add=TRUE)
+ }
> histNorm(fit$residuals, "orange")
```



```
> qqnorm(fit$residuals)
> qqline(fit$residuals)
```



```
> shapiro.test(fit$residuals)
```

Shapiro-Wilk normality test

data: fit\$residuals  
W = 0.99113, p-value = 0.7721

**Result and Comment:**

Histogram and Normal Q-Q Plot shows some evidence that the residuals could be from a normal distribution, but does not fit a normal distribution very well.

And the Shapiro-Wilk test has a p-value=0.7721 > 0.1, **Do Not Reject Null Hypothesis** at  $\alpha = 10\%$  significance level or smaller. Therefore, the residuals' data basically follow a normal distribution.

(f)

Two ways to calculate and find large leverages:

```
> diag_H = hatvalues(fit)
> sum(diag_H)
[1] 9
> 2*mean(diag_H)
[1] 0.185567
> sum(diag_H > 2*mean(diag_H))
[1] 5
> diag_H[which(diag_H > 2*mean(diag_H))]
      32      37      41      74      92
0.3304757 0.2184392 0.2410079 0.1912109 0.2092421
```

OR

```
> X = cbind(rep(1, 97), prostate[, 1:8])
> X = as.matrix(X)
> H = X %*% solve(t(X) %*% X) %*% t(X)
> sum(diag(H))
[1] 9
> 2*mean(diag(H))
[1] 0.185567
> sum(diag(H) > 2*mean(diag(H)))
[1] 5
> diag(H)[which(diag(H) > 2*mean(diag(H)))]
      32      37      41      74      92
0.3304757 0.2184392 0.2410079 0.1912109 0.2092421
```

Result:

There are five observations with large leverage, they are the 32<sup>th</sup>, 37<sup>th</sup>, 41<sup>th</sup>, 74<sup>th</sup>, and 92<sup>th</sup> points.

And their values are shown above.

(g)

Two ways to find potential outliers:

```
> stu_r = rstudent(fit)
> hist(stu_r)
> max(abs(stu_r))
[1] 2.61698
> n = length(stu_r)
> p = length(fit$coefficients)
> df = n - p - 1
> alpha = 0.05 ## Here we use 5% significance level to perform the t-test
```

Without Bonferroni adjustment:

```
> t = qt(1 - alpha/2, df)
```

```
> sum(abs(stu_r)>t)
[1] 6
> stu_r[abs(stu_r)>t]
      39      47      69      81      95      97
-2.616980 -2.376671  2.553530  1.987947  2.385070  2.293279
```

With Bonferroni adjustment:

```
> t_B = qt(1-(alpha/2)/n, df)
> sum(abs(stu_r)>t_B)
[1] 0
```

OR

```
> library(car)
> outlierTest(fit)
```

No Studentized residuals with Bonferonni  $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted	p-value	Bonferonni	p
39	-2.61698		0.01046		NA

Result:

There are 6 potential outliers without Bonferroni adjustment, and their values are shown above.

After Bonferroni adjustment, there are no outliers at  $\alpha = 5\%$  significance level.

(h)

```
> cook = cooks.distance(fit)
> sum(cook>4/n)
[1] 7
> cook[cook>4/n]
      32      39      47      69      95      96      97
0.12269771 0.05201916 0.10574362 0.10053751 0.09873809 0.05593862 0.07377558
```

Result:

There are seven influential observations with a large Cook's Distance.

And their Cook's Distances are shown above.

(i)

In the full model, only the parameters associated with predictors "lcaval", "lweight", and "svi" .

have p-value < 0.05, which are significant at  $\alpha = 5\%$  significance level. Therefore, set up the new, smaller model only with these three predictors.

```
> fit_new = lm(lpsa~lcavol+lweight+svi)
> anova(fit_new, fit)
```

Analysis of Variance Table

Model 1: lpsa ~ lcavol + lweight + svi

Model 2: lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +  
pgg45

	Res. Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	93	47.785				
2	88	44.163	5	3.6218	1.4434	0.2167

F test statistic = 1.4434, and p-value = 0.2167 > 0.1

Therefore, **Do not Reject  $H_0$**  (Null Model) at  $\alpha = 10\%$  or smaller significance level.

**The new, smaller model is preferred**, which is only explained by predictors "lcavol", "lweight", and "svi".