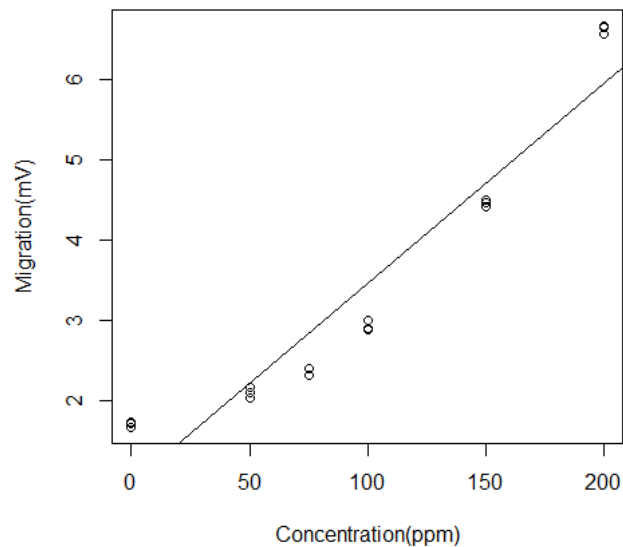


## Problem 1

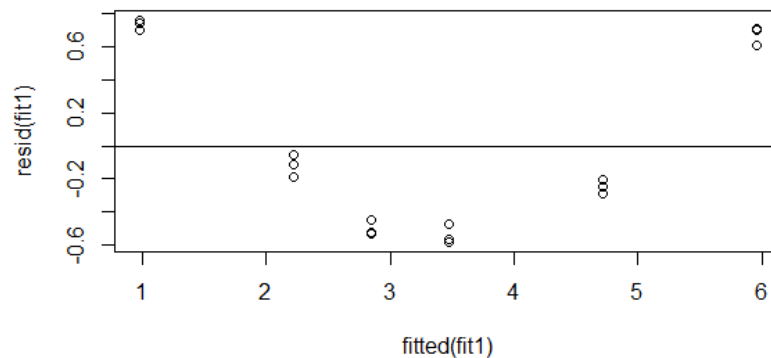
```
> ISEs=read.table(file="HW10_ISEs.dat", header=TRUE)
> colnames(ISEs)=c("concentration", "migration")
> head(ISEs, n=3)
  concentration migration
1              0      1.72
2              0      1.68
3              0      1.74
```

(a)

```
> fit1=lm(migration~concentration, data = ISEs)
> plot(ISEs$concentration, ISEs$migration, xlab = "Concentration(ppm)", ylab = "
Migration(mV)")
> abline(fit1)
```



```
> plot(fitted(fit1), resid(fit1))
> abline(h=0)
```



From both plots above, we can conclude that the linear model does **NOT** seem to be appropriate.

(b)

Quadratic model:  $migration_i = \beta_0 + \beta_1 * concentration_i + \beta_2 * (concentration_i^2) + e_i$

Null Hypothesis  $H_0$ :  $\beta_2 = 0$

```
> fit2=lm(migration~concentration+I(concentration^2), data = ISEs)
```

```
> anova(fit1, fit2)
```

Analysis of Variance Table

Model 1: migration ~ concentration

Model 2: migration ~ concentration + I(concentration^2)

	Res. Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	16	4.8323				
2	15	0.0603	1	4.772	1187	1.073e-15 ***

---

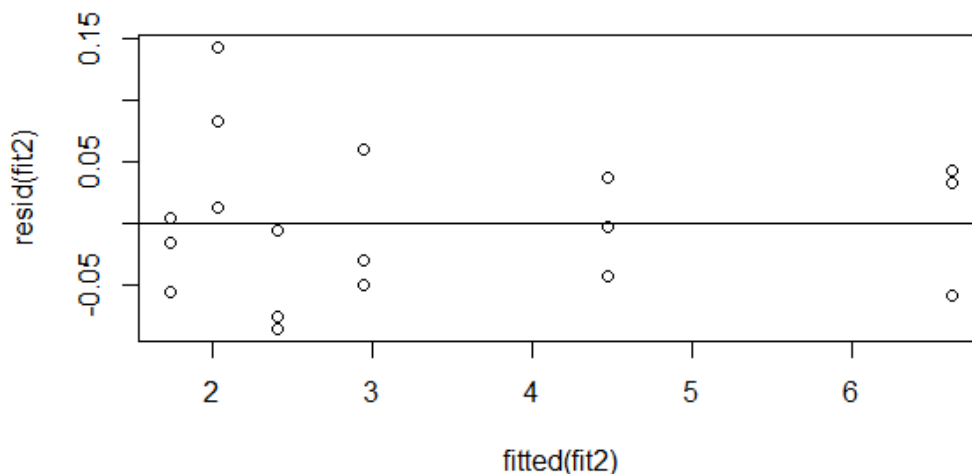
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

F test statistic = **1187**, and under the null hypothesis, the test statistic has an **F distribution** with **1** and **15** degrees freedom.

p-value = **1.073e-15** << 0.05, **Reject  $H_0$**  (Null Model) at  $\alpha = 0.05$ . Hence the new, quadratic model is much better at 5% or even much lower significance level.

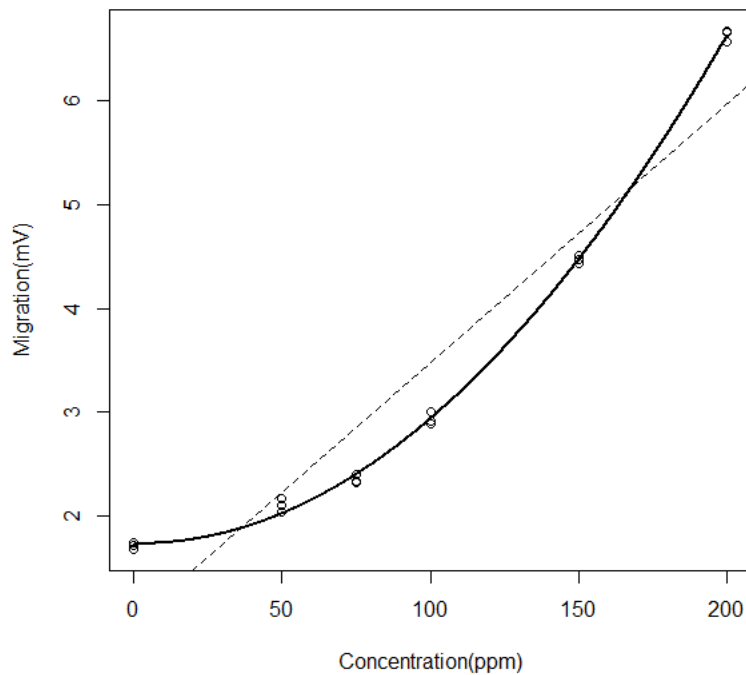
Also, we can see this result by looking at the plot of residuals vs. fitted values as shown below.

```
> plot(fitted(fit2), resid(fit2))
> abline(h=0)
```



(c)

```
> plot(ISEs$concentration, ISEs$migration, xlab = "Concentration(ppm)", ylab = "Migration(mV)")
> abline(fit1, lty=2)
> xplot=seq(0, 200, by=0.1)
> lines(xplot, predict(fit2, newdata=data.frame(concentration=xplot)), lwd=2)
```



## Problem 2

```
> library(faraway)
> data(odor)
> nrow(odor)
[1] 15
> odor[1:3, ]
  odor temp gas pack
1   66   -1  -1    0
2   39    1  -1    0
3   43   -1   1    0
```

(a)

```
> fit=lm(odor~temp+gas+pack+I(temp^2)+I(gas^2)+I(pack^2), data=odor)
> summary(fit)
```

Call:

```
lm(formula = odor ~ temp + gas + pack + I(temp^2) + I(gas^2) +
    I(pack^2), data = odor)
```

Residuals:

Min	1Q	Median	3Q	Max
-20.625	-9.625	-1.375	4.021	28.875

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-30.667	10.840	-2.829	0.0222 *
temp	-12.125	6.638	-1.827	0.1052
gas	-17.000	6.638	-2.561	0.0336 *
pack	-21.375	6.638	-3.220	0.0122 *
I(temp^2)	32.083	9.771	3.284	0.0111 *
I(gas^2)	47.833	9.771	4.896	0.0012 **
I(pack^2)	6.083	9.771	0.623	0.5509

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 18.77 on 8 degrees of freedom

Multiple R-squared: 0.8683, Adjusted R-squared: 0.7695

F-statistic: 8.789 on 6 and 8 DF, p-value: 0.003616

F test statistic = 8.789, and under the null hypothesis, the test statistic has an F distribution with 6 and 8 degrees freedom.

p-value = 0.003616 < 0.05, hence Reject  $H_0$  at  $\alpha = 0.05$ . Hence this quadratic model is significant at 5% significance level.

(b)

```
> names(summary(fit))
[1] "call" "terms" "residuals" "coefficients" "aliases"
[6] "sigma" "df" "r.squared" "adj.r.squared" "fstatistic"
[11] "cov.unscaled"
> summary(fit)$r.squared
[1] 0.8682799
```

86.83% of the observed variation of odor is explained by the model in part (a).

(c)

```
> summary(fit)$adj.r.squared
[1] 0.7694898
> extractAIC(fit)
[1] 7.00000 92.54619
```

### Problem 3

```
> library(faraway)
> data(prostate)
> nrow(prostate)
[1] 97
> prostate[1:3,]
      lcavol lweight age      lbph svi      lcp gleason pgg45      lpsa
1 -0.5798185  2.7695  50 -1.386294  0 -1.38629      6      0 -0.43078
2 -0.9942523  3.3196  58 -1.386294  0 -1.38629      6      0 -0.16252
3 -0.5108256  2.6912  74 -1.386294  0 -1.38629      7     20 -0.16252
```

```
> fit=lm(lpsa~., data=prostate)
> summary(fit)
```

Call:

```
lm(formula = lpsa ~ ., data = prostate)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.7331	-0.3713	-0.0170	0.4141	1.6381

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	0.669337	1.296387	0.516	0.60693
lcavol	0.587022	0.087920	6.677	2.11e-09 ***
lweight	0.454467	0.170012	2.673	0.00896 **
age	-0.019637	0.011173	-1.758	0.08229 .
lbph	0.107054	0.058449	1.832	0.07040 .
svi	0.766157	0.244309	3.136	0.00233 **
lcp	-0.105474	0.091013	-1.159	0.24964
gleason	0.045142	0.157465	0.287	0.77503
pgg45	0.004525	0.004421	1.024	0.30886

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.7084 on 88 degrees of freedom

Multiple R-squared: 0.6548, Adjusted R-squared: 0.6234

F-statistic: 20.86 on 8 and 88 DF, p-value: < 2.2e-16

(a)

(a-i) Backward AIC

```
> n=length(resid(fit))
> fit_bac_AIC=step(fit, direction = "backward"); fit_bac_AIC
Start: AIC=-58.32
lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
      pgg45
```

	Df	Sum of Sq	RSS	AIC
- gleason	1	0.0412	44.204	-60.231
- pgg45	1	0.5258	44.689	-59.174
- lcp	1	0.6740	44.837	-58.853
<none>			44.163	-58.322
- age	1	1.5503	45.713	-56.975
- lbph	1	1.6835	45.847	-56.693
- lweight	1	3.5861	47.749	-52.749
- svi	1	4.9355	49.099	-50.046
- lcavol	1	22.3721	66.535	-20.567

Step: AIC=- 60. 23

lpsa ~ lcavol + lweight + age + lbph + svi + lcp + pgg45

	Df	Sum of Sq	RSS	AIC
- lcp	1	0.6623	44.867	-60.789
<none>			44.204	-60.231
- pgg45	1	1.1920	45.396	-59.650
- age	1	1.5166	45.721	-58.959
- lbph	1	1.7053	45.910	-58.560
- lweight	1	3.5462	47.750	-54.746
- svi	1	4.8984	49.103	-52.037
- lcavol	1	23.5039	67.708	-20.872

Step: AIC=- 60. 79

lpsa ~ lcavol + lweight + age + lbph + svi + pgg45

	Df	Sum of Sq	RSS	AIC
- pgg45	1	0.6590	45.526	-61.374
<none>			44.867	-60.789
- age	1	1.2649	46.131	-60.092
- lbph	1	1.6465	46.513	-59.293
- lweight	1	3.5647	48.431	-55.373
- svi	1	4.2503	49.117	-54.009
- lcavol	1	25.4189	70.285	-19.248

Step: AIC=- 61. 37

lpsa ~ lcavol + lweight + age + lbph + svi

	Df	Sum of Sq	RSS	AIC
<none>			45.526	-61.374
- age	1	0.9592	46.485	-61.352
- lbph	1	1.8568	47.382	-59.497
- lweight	1	3.2251	48.751	-56.735
- svi	1	5.9517	51.477	-51.456
- lcavol	1	28.7665	74.292	-15.871

Call:

lm(formula = lpsa ~ lcavol + lweight + age + lbph + svi, data = prostate)

Coefficients:

(Intercept)	lcavol	lweight	age	lbph	svi
0.95100	0.56561	0.42369	-0.01489	0.11184	0.72095

**Best model: lpsa ~ lcavol + lweight + age + lbph + svi**

**(a-ii) Backward BIC**

> fit\_bac\_BIC=step(fit, direction = "backward", k=log(n)); fit\_bac\_BIC

Start: AIC=- 35. 15

lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +  
pgg45

	Df	Sum of Sq	RSS	AIC
- gleason	1	0.0412	44.204	-39.634
- pgg45	1	0.5258	44.689	-38.576
- lcp	1	0.6740	44.837	-38.255
- age	1	1.5503	45.713	-36.377

- lbph	1	1.6835	45.847	-36.095
<none>			44.163	-35.149
- lweight	1	3.5861	47.749	-32.151
- svi	1	4.9355	49.099	-29.448
- lcavol	1	22.3721	66.535	0.030

Step: AIC=- 39.63

lpsa ~ lcavol + lweight + age + lbph + svi + lcp + pgg45

	Df	Sum of Sq	RSS	AIC
- lcp	1	0.6623	44.867	-42.766
- pgg45	1	1.1920	45.396	-41.627
- age	1	1.5166	45.721	-40.936
- lbph	1	1.7053	45.910	-40.537
<none>			44.204	-39.634
- lweight	1	3.5462	47.750	-36.723
- svi	1	4.8984	49.103	-34.014
- lcavol	1	23.5039	67.708	-2.849

Step: AIC=- 42.77

lpsa ~ lcavol + lweight + age + lbph + svi + pgg45

	Df	Sum of Sq	RSS	AIC
- pgg45	1	0.6590	45.526	-45.926
- age	1	1.2649	46.131	-44.644
- lbph	1	1.6465	46.513	-43.844
<none>			44.867	-42.766
- lweight	1	3.5647	48.431	-39.925
- svi	1	4.2503	49.117	-38.561
- lcavol	1	25.4189	70.285	-3.800

Step: AIC=- 45.93

lpsa ~ lcavol + lweight + age + lbph + svi

	Df	Sum of Sq	RSS	AIC
- age	1	0.9592	46.485	-48.478
- lbph	1	1.8568	47.382	-46.623
<none>			45.526	-45.926
- lweight	1	3.2251	48.751	-43.862
- svi	1	5.9517	51.477	-38.583
- lcavol	1	28.7665	74.292	-2.997

Step: AIC=- 48.48

lpsa ~ lcavol + lweight + lbph + svi

	Df	Sum of Sq	RSS	AIC
- lbph	1	1.3001	47.785	-50.377
<none>			46.485	-48.478
- lweight	1	2.8014	49.286	-47.377
- svi	1	5.8063	52.291	-41.636
- lcavol	1	27.8298	74.315	-7.542

Step: AIC=- 50.38

lpsa ~ lcavol + lweight + svi

	Df	Sum of Sq	RSS	AIC
<none>			47.785	-50.377
- svi	1	5.1814	52.966	-44.966
- lweight	1	5.8924	53.677	-43.673
- lcavol	1	28.0445	75.829	-10.160

Call:

lm(formula = lpsa ~ lcavol + lweight + svi, data = prostate)

Coefficients:				
(Intercept)	lcavol	lweight		svi
-0.2681	0.5516	0.5085		0.6662

**Best model:  $\text{lpsa} \sim \text{lcavol} + \text{lweight} + \text{svi}$**

**(b)**

**(b-i) Forward AIC**

```
> n=length(resid(fit))
> fit_for_AIC=step(fit, direction = "forward"); fit_for_AIC
Start: AIC=-58.32
lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
pgg45
```

```
Call:
lm(formula = lpsa ~ lcavol + lweight + age + lbph + svi + lcp +
gleason + pgg45, data = prostate)
```

Coefficients:					
(Intercept)	lcavol	lweight	age	lbph	svi
lcp	gleason				
0.669337	0.587022	0.454467	-0.019637	0.107054	0.766157
-0.105474	0.045142				
pgg45					
0.004525					

**Best model:  $\text{lpsa} \sim \text{lcavol} + \text{lweight} + \text{age} + \text{lbph} + \text{svi} + \text{lcp} + \text{gleason} + \text{pgg45}$**

**(b-ii) Forward BIC**

```
> fit_for_BIC=step(fit, direction = "forward", k=log(n)); fit_for_BIC
Start: AIC=-35.15
lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
pgg45
```

```
Call:
lm(formula = lpsa ~ lcavol + lweight + age + lbph + svi + lcp +
gleason + pgg45, data = prostate)
```

Coefficients:					
(Intercept)	lcavol	lweight	age	lbph	svi
lcp	gleason				
0.669337	0.587022	0.454467	-0.019637	0.107054	0.766157
-0.105474	0.045142				
pgg45					
0.004525					

**Best model:  $\text{lpsa} \sim \text{lcavol} + \text{lweight} + \text{age} + \text{lbph} + \text{svi} + \text{lcp} + \text{gleason} + \text{pgg45}$**



(c)

```
> library(leaps)
> all_fits=regsubsets(lpsa~., data=prostate)
> all_fits_sum=summary(all_fits)
> all_fits_sum$which
(Intercept) lcavol lweight age lbph svi lcp gleason pgg45
1 TRUE TRUE FALSE FALSE FALSE FALSE FALSE FALSE FALSE
2 TRUE TRUE TRUE FALSE FALSE FALSE FALSE FALSE FALSE
3 TRUE TRUE TRUE FALSE FALSE TRUE FALSE FALSE FALSE
4 TRUE TRUE TRUE FALSE TRUE TRUE FALSE FALSE FALSE
5 TRUE TRUE TRUE TRUE TRUE TRUE TRUE FALSE FALSE FALSE
6 TRUE TRUE TRUE TRUE TRUE TRUE TRUE FALSE FALSE TRUE
7 TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE FALSE TRUE
8 TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE
```

(c-i) "leaps" AIC

```
> p=length(coef(fit))
> n=length(resid(fit))
> AIC=n*log(all_fits_sum$rss/n) + 2*(2:p)
> AIC
[1] -44.36608 -52.69043 -60.67621 -61.35179 -61.37439 -60.78867 -60.23130 -5
8.32184
> which.min(AIC)
[1] 5
```

**Best model #5:  $lpsa \sim lcavol + lweight + age + lbph + svi$**

(c-ii) "leaps" BIC

```
> BIC=n*log(all_fits_sum$rss/n) + log(n)*(2:p)
> BIC
[1] -39.21666 -44.96629 -50.37736 -48.47824 -45.92613 -42.76570 -39.63361 -3
5.14944
> which.min(BIC)
[1] 3
```

**Best model #3:  $lpsa \sim lcavol + lweight + svi$**

(c-iii) "leaps" Adjusted R<sup>2</sup>

```
> R_adj=all_fits_sum$adjr2
> R_adj
[1] 0.5345838 0.5771246 0.6143899 0.6208036 0.6245476 0.6258707 0.6272521 0.6
233681
> which.max(R_adj)
[1] 7
```

**Best model #7:  $lpsa \sim lcavol + lweight + age + lbph + svi + lcp + pgg45$**

(d)

The seven best models from above (a)-(c), some of them are the same:

Method	Model	Criteria
Forward AIC	lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason + pgg45	AIC=-58.32
Forward BIC	lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason + pgg45	BIC=-35.15
"leaps" Adjusted R <sup>2</sup>	lpsa ~ lcavol + lweight + age + lbph + svi + lcp + pgg45	62.73%
Backward AIC	lpsa ~ lcavol + lweight + age + lbph + svi	AIC=-61.37
"leaps" AIC	lpsa ~ lcavol + lweight + age + lbph + svi	AIC=-61.37
Backward BIC	lpsa ~ lcavol + lweight + svi	BIC=-50.38
"leaps" BIC	lpsa ~ lcavol + lweight + svi	BIC=-50.38

Hence we only to compare four models:

**Model 1:** lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason + pgg45, AIC=-58.32, BIC=-35.15

**Model 2:** lpsa ~ lcavol + lweight + age + lbph + svi + lcp + pgg45, Adjusted R<sup>2</sup>=62.73%

**Model 3:** lpsa ~ lcavol + lweight + age + lbph + svi, AIC=-61.37

**Model 4:** lpsa ~ lcavol + lweight + svi, BIC=-50.38

```
> fit1=l m(lpsa~. , data=prostate)
> fit2=l m(lpsa~. - gleason, data=prostate)
> fit3=l m(lpsa~lcavol+lwei ght+age+lbph+svi , data=prostate)
> fit4=l m(lpsa~lcavol+lwei ght+svi , data=prostate)
```

(1) AIC

```
> rbind(
+   extractAIC(fit1),
+   extractAIC(fit2),
+   extractAIC(fit3),
+   extractAIC(fit4)
+ )
      [, 1]      [, 2]
[1, ]    9 -58.32184
[2, ]    8 -60.23130
[3, ]    6 -61.37439
[4, ]    4 -60.67621
```

(2) Adjusted R<sup>2</sup>

```
> rbind(
+   summary(fit1)$adj. r. squared,
+   summary(fit2)$adj. r. squared,
+   summary(fit3)$adj. r. squared,
+   summary(fit4)$adj. r. squared
+ )
```

```

      [, 1]
[1, ] 0.6233681
[2, ] 0.6272521
[3, ] 0.6245476
[4, ] 0.6143899

```

### (3) PRESS

```

> rbind(
+   sum( (resid(fi t1) / (1 - hatval ues(fi t1))) ^2 ),
+   sum( (resid(fi t2) / (1 - hatval ues(fi t2))) ^2 ),
+   sum( (resid(fi t3) / (1 - hatval ues(fi t3))) ^2 ),
+   sum( (resid(fi t4) / (1 - hatval ues(fi t4))) ^2 )
+ )
      [, 1]
[1, ] 54.23246
[2, ] 53.27402
[3, ] 52.67252
[4, ] 52.84734

```

### Overall Analysis:

The criterion “AIC” and “PRESS” both shows that the Model 3 is the best model.

On the other hand, the criterion “Adjusted R<sup>2</sup>” shows that the Model 2 is the best model. However, the adjusted R squared value of Model 3 is similar with Model 2.

Hence, considering the overall effectiveness, the Model 3 is the best model as shown below:

**Model 3:     $lpsa \sim lcavol + lweight + age + lbph + svi$**