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Adapting Probabilistic Roadmaps to Handle Uncertain Maps

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***Abstract*— Randomized motion planning techniques are very good at solving high-dimensional motion planning problems. However, most planners assume complete knowledge of the environment, an assumption that can lead to collisions if there are errors in the world model due to uncertainty. We propose an extension of the Probabilistic Roadmap [1] algorithm that computes motion plans that are robust to uncertain maps. We show that the adapted PRM generates less collision-prone trajectories with fewer samples than the standard method.**

I. Introduction

Randomized motion planners, such as the Probabilistic Roadmap (PRM) [1] and the Rapidly-exploring Randomized Tree (RRT) [2] have been very successful in solving planning problems for robots with many degrees of freedom, problems that were previously considered intractable [3], [4]. However, these algorithms depend on having a complete and accurate model of the world.

In carefully engineered settings, such as a robot manipulator on the factory floor, the "perfect map" assumption is a reasonable one, as the environment can be precisely measured and tightly controlled. In contrast, major successes in robotics have been driven by appropriate models of uncertainty. Statistical inference techniques such as Markov localization and the Kalman filter [5], [6] have enabled mobile robots to navigate safely in populated, dynamic, uncertain environments without getting lost. For robots with few degrees of freedom, non-randomized motion planning algorithms have become more robust to sensor and model errors by using robust control techniques that incorporate both the cost of control errors and position uncertainty [7]. In order to allow the same robust operation for high-dimensional robots such as humanoid robots interacting with the real world, we need to incorporate knowledge of model uncertainty into the motion planning.

This paper extends the Probabilistic Roadmap (PRM) algorithm in two ways. Firstly, we modify randomized sampling of poses in order to minimize the number of samples required to express good plans. Our results show that good sampling strategies that respect map uncertainty can be used to substantially reduce the number of samples required to express good (less collision-prone) trajectories in uncertain worlds. Since the complexity of the trajectory search process is strongly dependent on the number of nodes, minimizing the set of samples is a major factor in keeping the planning problem tractable. Secondly, we show how to evaluate PRM actions efficiently in the context of uncertainty and generate motion plans with minimal expected cost.

II. The Probabilistic Roadmap Method

Given a map, robot dimensions, and the start and goal positions, the PRM method [1], [8] aims to produce a valid path from start to goal. We use $\mathcal{C}_{space}$ to denote the configuration space, the space of all robot poses [9], $\mathcal{C}_{free}$ as the set of all collision-free poses and $\mathcal{C}_{obst}$ as the set of poses resulting in collision with obstacles, so that $\mathcal{C} \equiv \mathcal{C}_{free} \cup \mathcal{C}_{obst}$. The PRM solves the planning problem in two stages [1], [10][1]:

1) The *preprocessing phase*, during which robot poses are sampled from $\mathcal{C}_{space}$, and only collision-free samples are retained. The retained samples, with the start and goal configurations constitute the nodes of a graph. The graph edges are found using a local planner that determines connectivity of each node to its k nearest neighbors. If the local planner finds a collision-free path between two nodes, then an edge between the nodes is added to the graph; the edge is labelled with an associated cost, such as distance.
2) The *query phase*, during which a graph search algorithm finds path from the start to the goal node.

The power of the PRM resides in the preprocessing stage, which exploits the fact that even if $\mathcal{C}_{free}$ cannot be tractably computed, it is relatively efficient to determine if an arbitrary pose lies in $\mathcal{C}_{free}$. The PRM learns a discrete graph approximation of $\mathcal{C}_{free}$ by sampling poses from $\mathcal{C}_{space}$ and rejecting samples that lie in $\mathcal{C}_{obst}$ (i.e., that collide with obstacles). If the sampling strategy is good, then a relatively small number of nodes suffices to approximate $\mathcal{C}_{free}$. The challenge with PRM planning is determining good sampling strategies that capture the topology of $\mathcal{C}_{free}$ with as few samples as possible, reducing the time to search for the path during the *query stage* of PRM. Although the PRM solution is complete in the limit of sampling an infinite number of nodes, the rate of convergence depends on the sampling strategy and difficulty of the environment [11]. An example of a challenging $\mathcal{C}_{free}$ configuration is a narrow passage where a naïve sampling strategy (such as uniform [12]) may require a large number of samples before it captures a trajectory through such a corridor. Three commonly used sampling strategies are:

Uniform sampling [12] samples poses with uniform probability from $\mathcal{C}_{space}$. The method's shortcoming is that no information about the map is used and unnecessarily many points are sampled in the empty regions (figure 1 left).

[1]Our work extends a classic variant of the PRM which tests for collisions during sampling and local planning as opposed to the Lazy PRM method [8].

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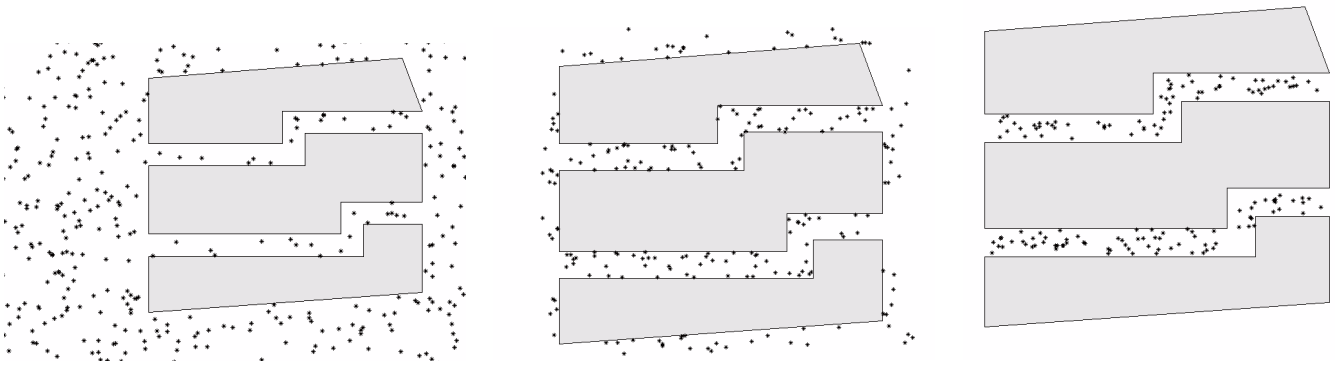

Fig. 1. A comparison of conventional sampling, from left to right: Uniform, Gaussian, Bridge

Gaussian obstacle-based sampling [13], [14] tries to capture the fact that the optimal length paths in $\mathcal{C}_{space}$ follow around and near the obstacles. It samples points near and only in the presence of obstacles (figure 1). A first sample, $\mathbf{s_1}$, is drawn using the uniform sampler. A distance d is sampled from a normal distribution with mean and covariance based on knowledge of obstacle density in $\mathcal{C}_{space}$. A second sample, $\mathbf{s_2}$ is then drawn from a uniform distribution of radius d centered at $\mathbf{s_1}$. If both samples are either in $\mathcal{C}_{free}$ or in $\mathcal{C}_{obst}$, they are discarded. Otherwise, the sample inside $\mathcal{C}_{free}$ is retained.

Bridge obstacle-based sampling [15] biases the sampling towards narrow passages where retaining connectivity of $\mathcal{C}_{free}$ is difficult. A first sample $\mathbf{s_1}$ is drawn using the uniform sampler and discarded if it lies in $\mathcal{C}_{free}$. Otherwise, a distance d is drawn from a normal distribution with mean and covariance based on *a priori* knowledge of the topology of $\mathcal{C}_{space}$. A second sample $\mathbf{s_2}$ is drawn from a uniform distribution with radius d centered at $\mathbf{s_1}$. If $\mathbf{s_2}$ lies in $\mathcal{C}_{free}$, both samples are discarded. If, however, $\mathbf{s_2}$ is in $\mathcal{C}_{obst}$, then the midpoint $\mathbf{m}$ between $\mathbf{s_1}$ and $\mathbf{s_2}$ is computed. If pose $\mathbf{m}$ falls within $\mathcal{C}_{free}$, then $\mathbf{m}$ is added to the graph. This sampler cannot generally be used alone since it samples only in corridors (figure 1 right).

Note that different sampling strategies can be used in combination, for example, the Bridge sampler cannot generally be used alone since it samples only in corridors.

III. Uncertain Worlds

A. Modelling world uncertainty

The sampling methods described so far assume that the world model is known exactly. However, this is an unreasonably optimistic assumption, since the statistical techniques used to estimate maps assume measurement errors. For example, the Kalman filter [5] maintains an explicit representation of the uncertainty of world features. In this paper, we represent the uncertainty in the map as multivariate Gaussians over obstacle features as provided by a feature-based EKF algorithm [5], [16]. Figure 2 (left) shows an example map in which the vertices of the polygons have uncertainty modelled with Gaussian distributions. The uncertainty of the location of each vertex is visualized with a one standard deviation uncertainty ellipse computed from the map covariance Σ around each vertex. Possible world models are shown in figure 2 (right). The features of the two "$\bigwedge$"-shaped obstacles are known with less precision and the trajectories above them are more likely to collide with the actual obstacles. The preferred path would go around these obstacles, not directly across to the goal.

In the following section, we adapt the sampling strategies to adjust for map uncertainty. In subsequent sections, we describe how to incorporate the obstacle uncertainty into motion costs and trajectory generation.

B. Proposed algorithms for sampling in uncertain worlds

In an uncertain world the actual positions of obstacles are unknown. It is unclear whether a sampled pose $\mathbf{s}$ collides with some obstacle or is in $\mathcal{C}_{free}$, thus whether we should accept or reject it. We propose that the decision to accept or reject $\mathbf{s}$ be a function of the probability that a pose is in $\mathcal{C}_{free}$.

Our method can be logically broken down into two steps:

1) Estimate collision probability for sampled pose $\mathbf{s}$.
2) Reject or accept $\mathbf{s}$ based on the collision probability.

C. Estimation of collision probability

Let us use $\delta(\mathbf{s}, \mathbf{W})$ to denote a function that determines whether or not $\mathbf{s}$ is in $\mathcal{C}_{free}$ of a possible model $\mathbf{W}$:

$$\delta(\mathbf{s}, \mathbf{W}) = \begin{cases} 1 & : \mathbf{s} \in \mathcal{C}_{free} \text{ of } \mathbf{W} \\ 0 & : \text{otherwise.} \end{cases} \tag{1}$$

Note that the world $\mathbf{W}$ is a single instance of sampled world from the distribution as shown in figure 2 (right).

The probability that a sample is in $\mathcal{C}_{free}$ can then be computed by

$$p\left(\mathbf{s} \in \mathcal{C}_{free}\right) = \int_{\mathbf{W}} \delta(s, \mathbf{W}) p(\mathbf{W}) d\mathbf{W} \tag{2}$$

where $p(\mathbf{W})$ is a multivariate probability density of the world model at the world $\mathbf{W}$, $V_{\mathbf{W}}$ is an integration variable being a volume in a configuration space of all degrees of freedom of all obstacles, and the world $\mathbf{W}$ can be uniquely characterized by the set of vertices of all obstacles.

To simplify the problem we assume that the obstacles are not correlated. Correlation among obstacle configurations requires domain knowledge which is usually not available from sensing and increases computational intractability. Assuming obstacle independence, we can approximate the likelihood that $\mathbf{s}$ is in $\mathcal{C}_{free}$ by a product of probabilities that $\mathbf{s}$ does not collide with any of the obstacles:

$$p(\mathbf{s} \in \mathcal{C}_{free}) \approx \prod_{i \in \text{obstacles}} \left[1 - p_{coll,i}(\mathbf{s})\right] \tag{3}$$

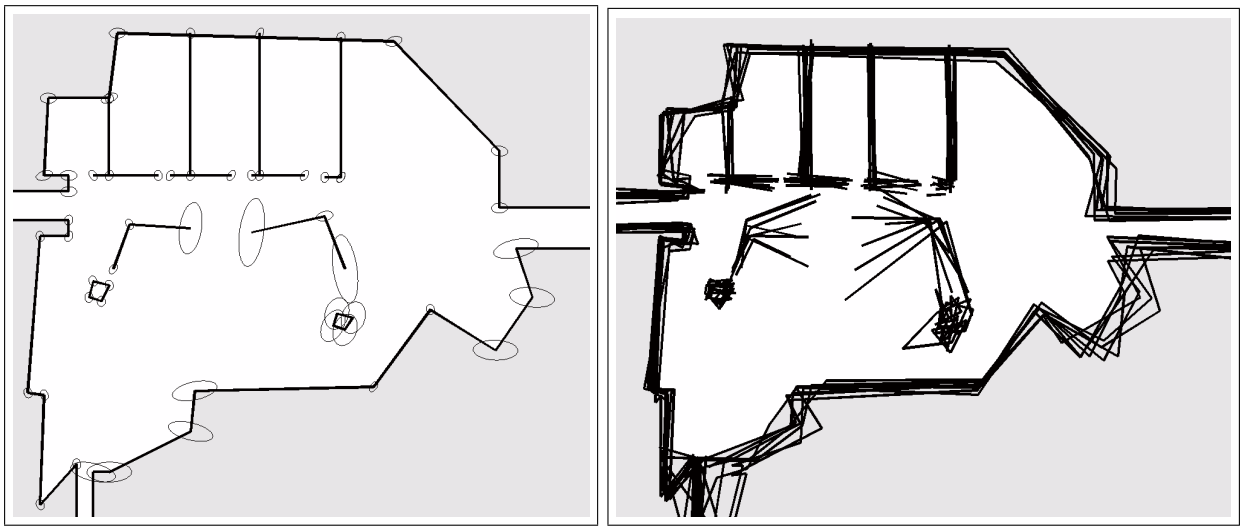

Fig. 2. Map of Robotics3 lab in Stata, MIT used as an example uncertain world in $\mathcal{R}^2$. Left: the maximum-likelihood map, shown with distributions of the features. Ellipses visualize one standard deviation from the mean. Right: example maps sampled from distributions.

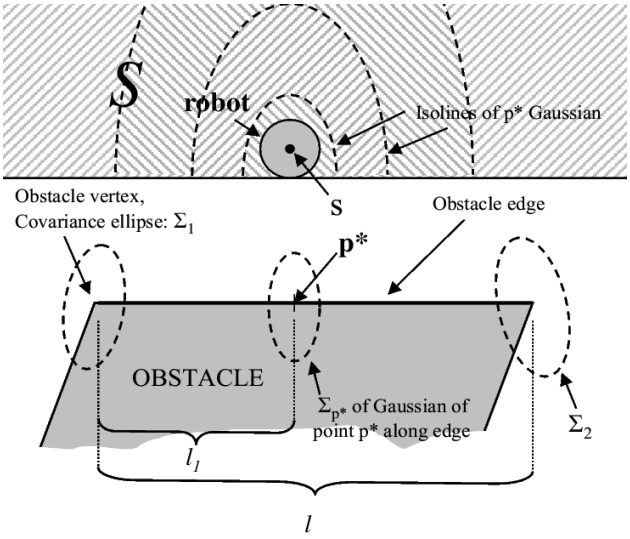

Fig. 3. The distribution of a point $\mathbf{p}^*$ on the obstacle edge which is closest to $\mathbf{s}$, can be computed from normal distributions of neighboring vertices. Integrating $\mathbf{p}^*$ distribution over the region S gives an estimate of the collision likelihood with a particular obstacle.

where $p_{coll,i}(\mathbf{s})$ is the probability of colliding with obstacle i. Given that, the total collision probability is given by:

$$p_{coll,total}(\mathbf{s}) = 1 - p(\mathbf{s} \in \mathcal{C}_{free}) = 1 - \prod_{i \in \text{obstacles}} [1 - p_{coll,i}(\mathbf{s})] \quad (4)$$

The remaining task is to compute the probability of collision with one obstacle, $p_{coll,i}(\mathbf{s})$.

IV. Efficient estimation of collision probabilities via Nearest Point method

On each obstacle in our world model, we can find some point $\mathbf{p}^*$ which is closest to the robot. Our estimate of the collision probability is based on the realization that the likelihood for a robot to collide with a particular obstacle is dominated by likelihood of colliding with $\mathbf{p}^*$. The robot will generally collide with the obstacle if the point $\mathbf{p}^*$ appears in the half-space bounded by the line tangent to the robot and parallel to the obstacle edge (the area S in figure 3). Due to the uncertainty of the obstacle, the position of $\mathbf{p}^*$ is also uncertain, but the probability of collision can be approximated using only knowledge of the distribution of $\mathbf{p}^*$. Assuming Gaussian distributions on the vertices, if $\mathbf{p}^*$ lies on some "obstacle edge"[2], $\mathbf{p}^*$ is distributed according to a Gaussian with covariance derived from the covariance matrices of the neighboring vertices:

$$\Sigma_{\mathbf{p}^*} = \frac{{l_1}^2}{l^2}\Sigma_2 + \frac{(l - l_1)^2}{l^2}\Sigma_1 \quad (5)$$

where l is the length of the edge, Σ_i are the corresponding covariances of vertices and l_1 is a distance from $\mathbf{p}^*$ to the vertex with covariance Σ_1 (see figure 3) [17]. In an uncertain world, $\mathbf{p}^*$ minimizes the Mahalanobis distance between the robot pose and the obstacle, but we cannot solve for $\mathbf{p}^*$ since we do not have a closed-form solution for the distribution of an entire 'edge' and in practice the statistical distance can be sufficiently approximated by the Euclidean case. In order to estimate the collision probability, we integrate a Gaussian function $\mathcal{N}(\mathbf{p}^*, \Sigma_{\mathbf{p}^*})$ over the half-space:

$$p_{coll,i}(\mathbf{s}) \approx \int_S \mathcal{N}(\mathbf{p}^*, \Sigma_{\mathbf{p}^*}) dV. \quad (6)$$

In order to validate this technique, we compared our approximation with collision probability estimates computed using a Monte Carlo technique. The probability that the pose $\mathbf{s}$ collides with some obstacle i can be approximated by sampling from the obstacle vertices distribution functions and counting the number of collisions with $\mathbf{s}$:

$$p_{coll,i}(\mathbf{s}) \approx \frac{N_{collision}}{N_{total}} \quad (7)$$

where $N_{collision}$ is the total number of pose collisions with obstacle i and N_{total} is the total number of trials. The difficulty

[2]The notion of an *'obstacle edge'* is not well-defined since our world model consists of polygons defined by distributions over vertices. We define an *'obstacle edge'* as the line connecting the means of a pair of vertices.

with the Monte Carlo approximation is that it requires many samples for an accurate estimate, leading to computational intractability. We used Monte Carlo estimates to validate our approximation, and Nearest Point method gave collision probabilities comparable to the Monte Carlo estimates.

A. Rejection function based on collision probability

Equation 4 provides a way to estimate the cumulative collision probability for a given pose efficiently; we reject or accept sampled poses based on this probability, $p_{coll,total}(\mathbf{s})$: the higher the collision likelihood of a sampled pose, the less likely our sampling algorithm will accept such pose. In practice, we shaped our rejection function to generally reject samples that lay inside the "nominal obstacle."[3] based on the notion that samples where robot is 50% or more likely to collide with obstacles are costly and unlikely to be used by the planning stage. Our experimental results will indicate that our sample measure outperforms conventional techniques.

B. Pose rejection algorithm

We summarize the resulting accept/reject decision in the following algorithm:

Algorithm 1: Accept or reject pose $\mathbf{s}$
Input: pose $\mathbf{s}$, all obstacles
Output: boolean s_{reject} describing whether to reject pose
(1) **foreach** obstacle i
(2) Compute $p_{coll,i}$, collision probability for $\mathbf{s}$ and i
(3) Compute the collision probability for all obstacles: $p_{coll,total}(\mathbf{s}) = 1 - \prod_i (1 - p_{coll,i}(\mathbf{s}))$.
(4) $p_{reject}(\mathbf{s}) \approx p_{coll,total}(\mathbf{s})$.
(5) s_{reject} = randomly sample from Bernoulli distribution where $p = p_{reject}(\mathbf{s})$.
(6) **return** s_{reject}

Adapted Sampling Strategies

We modify each sampling strategy to accept or reject samples stochastically, depending on the likelihood of pose $\mathbf{s}$ lying in $\mathcal{C}_{free}$.

Adapted Uniform Sampling: poses are drawn uniformly at random from $\mathcal{C}_{space}$ and each sample is retained or rejected based on collision likelihood as summarized in Algorithm 1 (figure 4 left).

Adapted Gaussian sampling: Gaussian sampling generates samples that are close to obstacles by generating pairs of samples that lie on opposite sides of an obstacle edge, one inside the obstacle, the other outside. Accounting for the uncertainty, we retain samples based on probability that the *pair* of samples lie on opposite sides of an edge. We first generate a sample $\mathbf{s}_1$ inside the "nominal obstacle," where the probability of rejection is high, in our setup $p_{reject}(\mathbf{s}) = 1$. Let us call this sample an *"anchor"*. Next, we sample a distance d, and generate a second sample $\mathbf{s}_2$ from a uniform distribution of radius d[4]. This sample is then retained according to Algorithm 1.

We require the *"anchor"* points to be inside the "nominal obstacle" to avoid sampling from poses too distant from obstacle. When we originally did not use this heuristic, many of the resulting points were very far from the obstacles. This accounts for the case when the first point is already outside the "nominal obstacle," and happens to be labelled as being in $\mathcal{C}_{obst}$ (since the process is stochastic), the second sample may end up far from the obstacle. Since $\mathbf{s_2}$ is far from an obstacle, $p_{reject}(\mathbf{s_2})$ will be small, and $\mathbf{s_2}$ will be readily accepted as a sample. By keeping the first sample anchored, we retain the *near-obstacle* feature of the original method. Figure 4, middle, shows Gaussian sampling in an uncertain world.

Adapted bridge sampling: Once again, in order to retain the *near-obstacle* characteristics of the original bridge method, the adapted bridge method secures anchoring of the first two sampled points inside the "nominal obstacle" by generating two samples that each have $p_{\text{reject}} = 1$. Next, the midpoint $\mathbf{m}$ is computed and retained based on Algorithm 1. As a result, narrow pathways characterized by higher certainty in the obstacle positions are favored, and the search for a path with lower collision likelihood is focused on those passages (see figure 4, right).

Once a set of samples has been generated using any hybrid of sampling strategies, we use a *local planner* to determine if pairs of samples can be connected. We have not modified this stage of the PRM, and we use k-nearest-neighbors planner using the Euclidean distance metric in a world defined by "nominal obstacles." We ignore uncertainty in the local planner, because we want a spatially interconnected graph.

We incorporate the world uncertainty into both stages of the PRM. This does not constitute double-counting; if we knew *a priori* the optimal path, then the optimal sampler would sample only poses on the path. By incorporating knowledge used in cost calculations, we can focus the sampling appropriately. For example, our adapted samplers are biased to sample closer to more certain sides of the corridor space. This allows the PRM planner to focus its search in regions where collisions are less likely.

V. Planning in Uncertain Worlds

To generate good trajectories, the *query phase* of PRM also needs to incorporate the world uncertainty. Standard planners find paths that minimize some quantity such as distance or travel time represented via a cost function associated with travelling each path segment. We use a minimum-collision-cost (*MCC*) planner to find trajectories that minimize the *expected* cost of collision when travelling edges in the graph. We define the expected cost of traversing a path segment between robot poses $\mathbf{s_1}$ and $\mathbf{s_2}$ as:

$$C(\mathbf{s_1}, \mathbf{s_2}) = p_{seg\_coll,total}(\mathbf{s_1}, \mathbf{s_2}) * C_{collision} + [1 - p_{seg\_coll,total}(\mathbf{s_1}, \mathbf{s_2})] ||\mathbf{s_1}, \mathbf{s_2}|| \quad (8)$$

[3]We define "nominal obstacle" as the obstacle with the vertices at the maximum likelihood positions, which are the means of Gaussian distributions.

[4]We need to keep in mind that, because our map is uncertain, the variance of normal distribution of d must be greater than the variance in map features/vertices. This is because the sampled points need to be placed sufficiently far from the obstacles in order not to have high rejection probability.

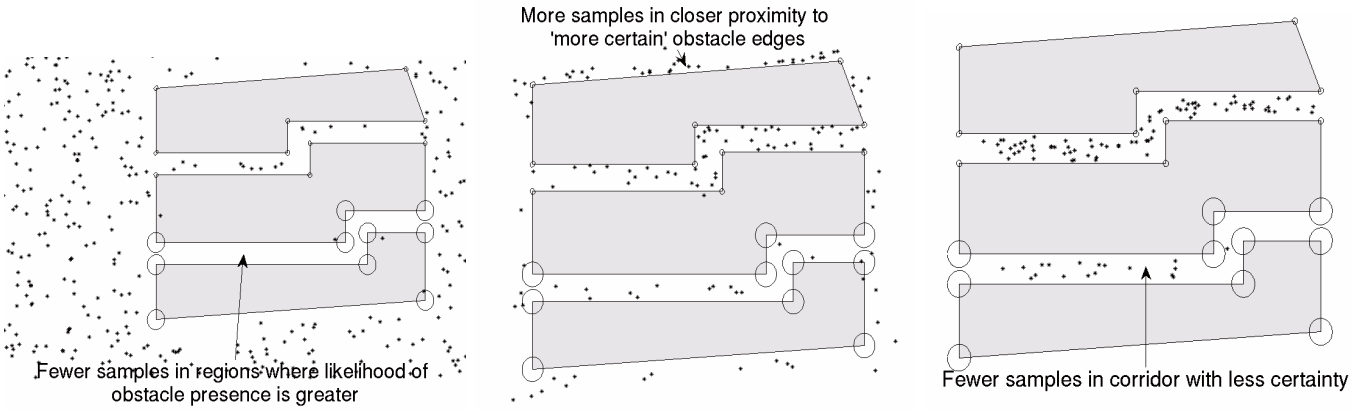

Fig. 4. A comparison of the sampling strategies adapted for uncertainty sampling, from left to right: Uniform, Gaussian, Bridge

where $p_{seg\_coll,total}(\mathbf{s_1}, \mathbf{s_2})$ is the probability that a robot collides with any obstacle while travelling on the line segment from $\mathbf{s_1}$ to $\mathbf{s_2}$, and $C_{collision}$ is some fixed estimate of how much it would cost when a robot collides with something. $C_{collision}$ can also be chosen to be a function of distance; for example, travelling 1 km less may be worth the risk of colliding with an obstacle.

Continuing the assumption that individual obstacle distributions are independent, the total probability of collision can be approximated as:

$$p_{seg\_coll,total}(\mathbf{s_1}, \mathbf{s_2}) = 1 - \prod_{i \in \text{obstacles}} [1 - p_{seg\_coll,i}(\mathbf{s_1}, \mathbf{s_2})] \quad (9)$$

where $p_{seg\_coll,i}(\mathbf{s_1}, \mathbf{s_2})$ is the probability of hitting obstacle i when travelling between $\mathbf{s_1}$ and $\mathbf{s_2}$. We cannot use the approximation technique of equation 6 to calculate the probability of collision, as this method computes the probability of a single pose. Integrating equation 6 along the edge would overestimate the probability of collision due to violated independence assumptions. We therefore use a Monte Carlo technique as in equation 7, but for any collision during simulated motion along the edge.[5]. We compute the joint probability $p_{seg\_coll,total}(\mathbf{s_1}, \mathbf{s_2})$ via equation 9, and incorporate it into the cost of traversing segment $(\mathbf{s_1}, \mathbf{s_2})$ using equation 8. Finally, a graph search algorithm (A*) uses the estimated collision cost along each roadmap edge and returns the minimum cost path.

VI. Experiments: Adapted Sampling and Planning

In our preliminary experiments, we measure the effects of adapted sampling strategies on the overall performance of the PRM algorithm. We consider the motion planning problem for a simple circular robot with arbitrary start and goal locations within a 2D plane. The performance is measured by repeatedly simulating a robot trajectory in maps sampled from the distribution shown in figure 2 and recording the number of collisions (example shown in figure 5).

We examine planner robustness under 3 sampling methods in conventional and adapted form, for a total of 6 variants. We use the Nearest Point method to determine whether a sampled pose results in a collision for the adapted sampling variants. The local planner uses the straight-line distance and a maximum of 12 nearest neighbors ($k = 12$). In the query stage, we always use the MCC planner in order to focus on the impact of the sampling method. The resulting path is smoothed by rerunning the *MCC* planner on the fully-connected set of path nodes, including start and goal.

[5]For computational efficiency, we do not perform the Monte Carlo trial if the obstacle and the path edge are statistically far from each other and just set $p_{seg\_coll,i}(\mathbf{s_1}, \mathbf{s_2}) = 0$

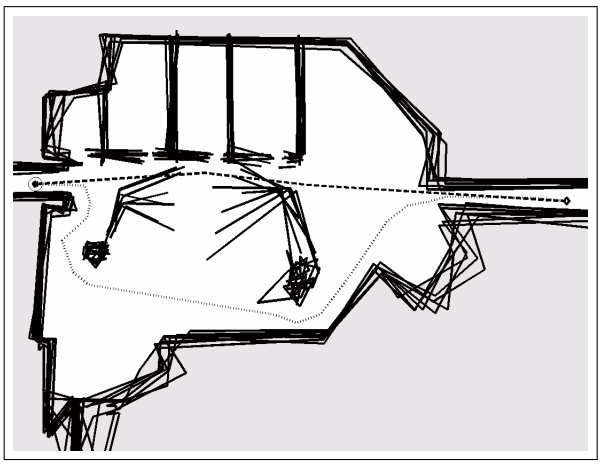

Fig. 5. Map of the MIT Stata Robotic Lab area, trajectories with obstacles perturbed, standard A* planner (dashed), minimum-collision-cost A* planner (dotted)

Plots in figure 6 show that adapted sampling improves the performance of the planner resulting in higher quality paths (fewer collisions) while keeping the number of nodes relatively small. Particularly, in the case of sample numbers less than 200, the adapted sampling methods do substantially better, resulting in collision rates that the standard sampling needs more than 1000 points to obtain. This is because the adapted sampling methods bias the sampling into regions of higher certainty, encouraging pathways through those regions with fewer samples overall.

The adapted sampling methods add a small amount of computational overhead to the standard sampler, because samples are rejected based on their collision probability and more loop iterations are required to collect samples. We found that the time to sample was anywhere from one to three times as long as the standard sampling method. However, the sampling stage of the PRM algorithm is a very small fraction (generally less than 5%) of the overall time to plan and the planner benefits from a reduced sample set to plan with.

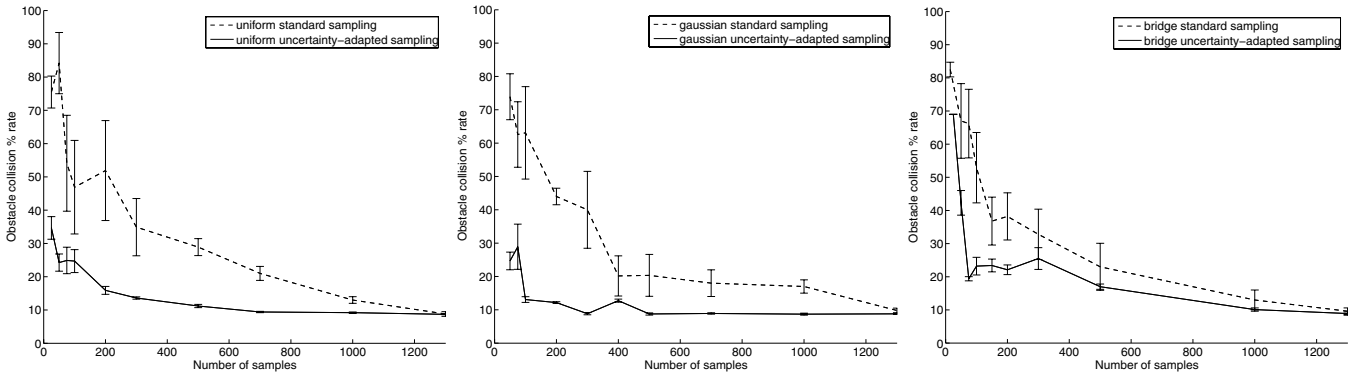

Fig. 6. The PRM performance in terms of collision rates using original and adapted sampling as a function of the sample number. The minimum-collision-cost planner was used in both cases in order to focus on the influence of the sampling strategy. We note that collision rates when sample numbers are small (under 100) are substantially lower when using our adapted sampling techniques.

A. Modelling Orientation

Figure 7 shows planning for rectangular holonomic robots of varying dimensions. For these rectangular robots, we used basic (top) and uncertainty-adapted uniform sampling (bottom) to sample robot poses (x, y, θ) for a total of 3 dimensions. In both cases, as the length of the robot increases, the path accommodates by moving the robot away from the obstacle before it can turn. In figure 7(top) the differences in paths are due purely to robot geometry since conventional planner is used. In figure 7(bottom) paths change due to both geometry and uncertainty. The trajectories obtained with *MCC* planner are further from the uncertain obstacle. Simulations by perturbing the obstacles show that the collision rates decrease from an average of 87% for the conventional planner to 35% for the *MCC* planner.

B. Modelling Environmental Dynamics

Figure 8 shows results obtained when planning a path for a circular B21 robot using a conventional, minimum-distance planner (the solid trajectory) and the probabilistic, minimum-collision-cost (*MCC*) planner (dashed) employing segment collision cost in equation 8. The conventional planner returns a path that is $41.2m$ long where some path segments have $p_{seg\_coll,total}$ estimated by *MCC* to be as high as 0.84. The *MCC* planner selects the longer $(51.0m)$ path, but with smaller (less than 0.40 for a given segment) collision likelihood.

VII. Applicability to Multidimensional Problems

We have shown a method for incorporating uncertainty into the PRM and demonstrated the concepts on 2D world examples, including robots with rotational dependencies. Our extension to model uncertainty can be applied to problems with higher-dimensional configuration, such as multiple dof robots in 3D. There, sampling can be decomposed for each dof of the robot. For example, each limb of a humanoid robot can be treated independently and the overall collision probability estimated by combining the probability of collision for the individual components. It has been shown experimentally [11] that when problem dimensionality increases, randomized sampling and planning methods fare better than deterministic methods with respect to computational complexity. Our extension is at worst polynomial in the number of degrees of freedom of the robot. In contrast, the complexity

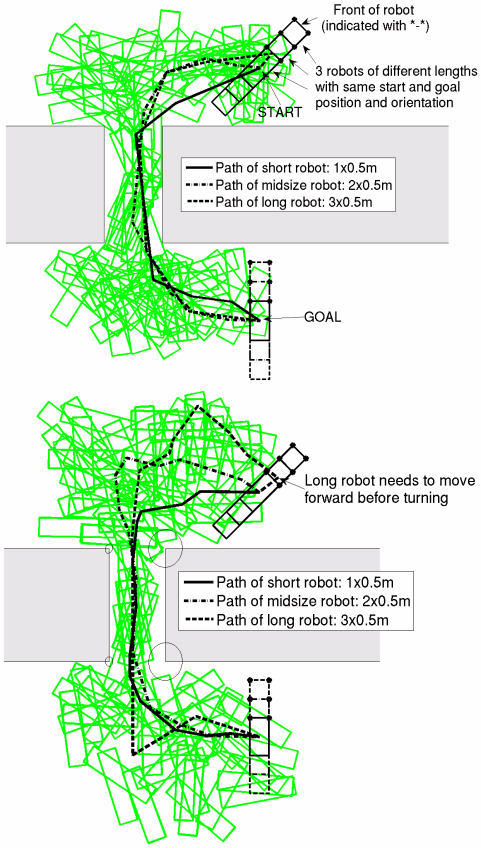

Fig. 7. Rectangular robot sampling of 3 dimensions: position and orientation, and planning; top: standard PRM planning without uncertainty, bottom: sampling and planning incorporating uncertainty. Also note how the robot trajectories change with increasing robot length.

of occupancy grids [18] increases as $O(n^k)$ where k is the number of dimensions, and n is inversely proportional to graph resolution.

We must point out that estimating collision probabilities in a 3D workspace (as distinct from the configuration space) becomes difficult since the closest point belongs to an obstacle plane, and can no longer be approximated by a linear combination of the endpoints. To remedy this problem, a new uncertainty model is needed. Our future work will focus on modelling uncertainty in higher dimensions and methods to incorporate stochastic world information into collision proba-

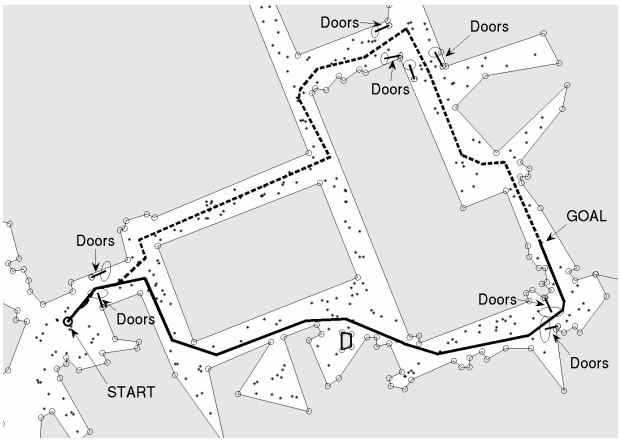

Fig. 8. Map of the 3rd floor MIT Stata Center, the minimum-collision-cost planner (dashed) selects the longer trajectory for the B21 robot to reach the goal based on the uncertain position of the doors along its route, while standard A* planner (solid) selects the shortest path trajectory.

bilities and path costs.

VIII. Related Work

Most randomized planning algorithms are not explicitly robust to model errors. Some sampling methods such as Medial Axis PRM (MAPRM) [19] have attempted to generate trajectories that are robust to model errors since samples on the medial axis of the plane maximize their distance from obstacles. However, such conservative sampling methods do not incorporate uncertainty into the cost function and cannot bias samples to be closer to more certain obstacles [19].

Both RRTs [2] and the Probabilistic Roadmap of Trees method (PRT) [20] use randomized sampling and cost functions over samples or trees. Our modified sampling and planning framework could be incorporated into both the local planner of the RRT and the global planning mechanism of the PRM. In particular, the RRT uses the random tree algorithm as a subroutine in PRM, where the nodes in the PRM roadmap are trees. Since multiple RRTs are grown in parallel we could achieve substantial gains in planning speed by simultaneously exploring and evaluating multiple regions of a map for trajectory quality.

Leven and Hutchinson in [21] address the problem of dynamic environments with a variant on PRM, where sampled nodes are updated according to information about the changing state of the world. Unlike our work, they do not compute feasibility of collisionless travel through different regions of space and their motion plan does not incorporate uncertainty. Similarly, Berg and Overmars in [22] use the PRM to plan in dynamic environments by first generating a global path assuming a static world, and then using local planners to deal with moving obstacles. Their motion plan selects a global trajectory assuming that the world is certain.

IX. Conclusion

We have demonstrated that conventional motion planning algorithms can be extended to allow robust motion planning when the true state of the world is not known exactly. We proposed incorporating uncertainty into the PRM sampling and planning and adapted three popular random sampling techniques to focus samples on regions of higher certainty. In the planning stage, we modelled the cost of potential collisions in travelling through uncertain regions of the configuration space. The experiments showed that using stochastic rejection of samples biased the path into regions of more certainty, resulting in more robust paths with small sample numbers.

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