# Distribution

#### 1.1 $\Gamma$ distribution

A random variable X is said to have a  $\Gamma$  distribution with parameters  $\alpha$ ,  $\beta$  if its probability density function is given by

$$f(x;\alpha,\beta) = \frac{x^{\alpha-1}e^{-\frac{x}{\beta}}}{\beta^{\alpha}\Gamma(\alpha)}, \quad \alpha,\beta \ge 0, \quad x \ge 0.$$
 (1.1)

The quantity  $\Gamma(\alpha)$  is known as the  $\Gamma$  function and it is equal to:

$$\Gamma(\alpha) = \int_0^\infty x^{\alpha - 1} e^{-x} dx \tag{1.2}$$

 $\alpha$  is *shape*,  $\beta$  is called *scale*, and  $\theta = \frac{1}{\beta}$  is called *rate*.

Some useful results:

$$\mathbb{E}[X] = \alpha \beta, \quad \mathbb{V}[X] = \alpha \beta^2, \quad \Gamma(\frac{1}{2}) = \sqrt{\pi}, \quad \Gamma(n) = (n-1)!.$$

If we set  $\alpha = 1$  and  $\beta = \frac{1}{\lambda}$ , we get  $f(x) = \lambda e^{-\lambda x}$ . We see that the exponential distribution is a special case of the  $\Gamma$  distribution.

#### 1.1.1 Moment Generating function

Moment generating function of  $X \sim \Gamma(\alpha, \beta)$  is

$$M_X(t) = (1 - \beta t)^{-\alpha}$$
 (1.3)

**Proof**:

$$M_X(t) = \mathbb{E}[e^{tX}] = \int_0^\infty e^{tx} \frac{x^{\alpha - 1} e^{-\frac{x}{\beta}}}{\beta^{\alpha} \Gamma(\alpha)} dx = \frac{1}{\beta^{\alpha} \Gamma(\alpha)} \int_0^\infty x^{\alpha - 1} e^{-x(\frac{1 - \beta t}{\beta})} dx \tag{1.4}$$

Let  $y = x(\frac{1-\beta t}{\beta})$ , then  $x = (\frac{\beta}{1-\beta t})y$ , and  $dx = (\frac{\beta}{1-\beta t})dy$ . Substitute these in the expression above

$$M_X(t) = \frac{1}{\beta^{\alpha} \Gamma(\alpha)} \int_0^{\infty} \left(\frac{\beta}{1 - \beta t}\right)^{\alpha - 1} y^{\alpha - 1} e^{-y} \frac{\beta}{1 - \beta t} dy \tag{1.5}$$

$$\beta^{\alpha}\Gamma(\alpha) J_0 \qquad (1 - \beta t) \qquad 1 - \beta t$$

$$= \frac{1}{\beta^{\alpha}\Gamma(\alpha)} \left(\frac{\beta}{1 - \beta t}\right)^{\alpha} \int_0^{\infty} y^{\alpha - 1} e^{-y} dy \qquad (1.6)$$

$$= (1 - \beta t)^{-\alpha} \qquad (1.7)$$

$$= (1 - \beta t)^{-\alpha} \tag{1.7}$$

## $\chi^2$ distribution

Let  $Z_1, Z_2, \dots, Z_k$  be independent random variables with  $Z_i \sim \mathcal{N}(0, 1)$  (iid), then

$$Z = Z_1^2 + Z_2^2 + \dots + Z_k^2 = \sum_{i=1}^k Z_i^2 \sim \chi_k^2$$
 (1.8)

 $\chi^2$  is a class of distribution indeXed by its degree of freedom, like the t-distribution. In fact,  $\chi^2$  has a

If  $X_1, X_2, \ldots, X_n$  are independent random variables with  $X_i \sim \mathcal{N}(\mu, \sigma)$ , then

$$X = \sum_{i=1}^{n} \left(\frac{X_i - \mu}{\sigma}\right)^2 \sim \chi_n^2 \tag{1.9}$$

Let  $X_1 \sim \chi_n^2$  and  $X_2 \sim \chi_m^2$ . If  $X_1$  and  $X_2$  are independent, then

$$X_1 + X_2 \sim \chi_{n+m}^2. \tag{1.10}$$

Let  $X_1, X_2, \ldots, X_n$  be independent random variables with  $X_i \sim \mathcal{N}(\mu, \sigma)$ . Define the sample variance

$$S^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2}$$
(1.11)

Then

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2. \tag{1.12}$$

### shape of $\chi^2$ distribution

Figure 1.1:  $\chi^2$  with different df