

Solution of GATE-ST 2023 Q18

SUJAL GUPTA - EE22BTECH11052

Suppose that X has the probability density function

$$f(x) = \begin{cases} \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} & \lambda > 0 \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

where $\alpha > 0$ and $\lambda > 0$. Which one of the following statements is NOT true?

- 1) $E(X)$ exists for all $\alpha > 0$ and $\lambda > 0$
- 2) Variance of X exists for all $\alpha > 0$ and $\lambda > 0$
- 3) $E(\frac{1}{X})$ exists for all $\alpha > 0$ and $\lambda > 0$
- 4) $E(\ln(1 + X))$ exists for all $\alpha > 0$ and $\lambda > 0$

Solution:

$$E(X) = \int_{-\infty}^{\infty} x p_X(x) dx \quad (2)$$

$$= \int_0^{\infty} x \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} dx \quad (3)$$

$$= \frac{\lambda^\alpha}{\Gamma(\alpha)} \int_0^{\infty} x^\alpha e^{-\lambda x} dx \quad (4)$$

(5)

since we know that

$$\int_0^{\infty} x^{\alpha-1} e^{-\lambda x} dx = \frac{\Gamma(\alpha)}{\lambda^\alpha} \quad \text{for } \lambda > 0 \quad (6)$$

$$E(X) = \frac{\lambda^\alpha}{\Gamma(\alpha)} \frac{\Gamma(\alpha+1)}{\lambda^{\alpha+1}} \quad (7)$$

Using the relation

$$\Gamma(x+1) = \Gamma(x)x \quad (8)$$

$$E(X) = \frac{\alpha}{\lambda} \quad (9)$$

$$Var(X) = E(X^2) - E(X)^2 \quad (10)$$

(11)

$$E(X^2) = \int_0^{\infty} x^2 \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} dx \quad (12)$$

$$= \int_0^{\infty} \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{(\alpha+2)-1} e^{-\lambda x} dx \quad (13)$$

$$= \int_0^{\infty} \frac{1}{\lambda^2} \frac{\lambda^{\alpha+2}}{\Gamma(\alpha)} x^{(\alpha+2)-1} e^{-\lambda x} dx \quad (14)$$

(15)

$$E(X^2) = \int_0^{\infty} \frac{\alpha(\alpha+1)}{\lambda^2} \frac{\lambda^{\alpha+2}}{\Gamma(\alpha+2)} x^{(\alpha+2)-1} e^{-\lambda x} dx \quad (16)$$

using the density of the gamma distribution, we get

$$E(X^2) = \frac{\alpha(\alpha+1)}{\lambda^2} \quad (17)$$

$$Var(X) = \frac{\alpha^2 + \alpha}{\lambda^2} - \frac{\alpha^2}{\lambda} \quad (18)$$

$$= \frac{\alpha}{\lambda^2} \quad (19)$$

Hence option A and B are true. For option C

$$E\left(\frac{1}{X}\right) = \int_0^{\infty} \frac{1}{x} \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} dx \quad (20)$$

$$= \frac{\lambda^\alpha}{\Gamma(\alpha)} \int_0^{\infty} x^{\alpha-2} e^{-\lambda x} dx \quad (21)$$

For this, $\alpha > 1$ is a must condition. Hence C is not a correct option. Hence C is the answer.