

A Universal Algorithm for Continuum Limit Distributions of Continuous Time Random Walks

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Abstract

We propose a universal “Semi-Markov algorithm” for the computation of probability distributions of Continuous Time Random Walks (CTRWs) and their continuum limits. The algorithm is universal in the following sense: Any CTRW continuum limit can be mapped to a bivariate Langevin equation which tracks the cumulative sum of jumps and waiting times, and given the coefficients of this Langevin equation, the algorithm will compute the desired probability distributions. Besides subdiffusion with space- and time-dependent drift, this approach covers subdiffusion with spatially varying exponent $\beta(x) \in (0, 1)$ and tempering parameter $\theta(x) \geq 0$, and subdiffusion with mixed (distributed) order where the mixture can vary in space. Our Semi-Markov algorithm generalizes the recent “Discrete Time Random Walk” algorithm, and shares the same properties: it is consistent, conserves mass, generates strictly non-negative solutions, and has the same computational complexity. To illustrate applicability, we set up an interface problem, where two subdiffusive media with different anomalous exponents meet, and compute the evolution of probability densities at the interface.

1. Introduction

Subdiffusive transport processes are characterized via a sublinear growth of the mean squared displacement: $\langle X_t \rangle \sim t^\alpha$, where $0 < \alpha < 1$. Such processes are usually modelled either by fractional Brownian motion or Continuous Time Random Walks (CTRWs), depending on whether the auto-correlation of jumps decays slowly or the waiting times between jumps are heavy-tailed with parameter α , modelling traps or dead ends (Henry et al. 2010). The CTRW model has proven to be a particularly useful model, predominantly in biophysics (Metzler and Klafter 2000; Tolić-Nørrelykke et al. 2004; Wong et al. 2004; Banks and Fradin 2005; Santamaria et al. 2006; Höfling, Franosch, and Article 2012; Regner et al. 2013), but also in groundwater hydrology (Berkowitz, Emmanuel, and Scher 2008; Schumer et al. 2003) and econophysics (Scalas 2006).

A modelling framework for the evolution of probability densities of random walks is given by the Fokker–Planck equation (Gardiner 2004):

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$$\frac{\partial P(y, t)}{\partial t} = \mathcal{L}^*(y, t)P(y, t) + \delta_{(0,0)}(y, t), \quad (1)$$

where

$$\mathcal{L}^* g(y, t) = -\frac{\partial}{\partial y}[b(y, t)g(y, t)] + \frac{1}{2} \frac{\partial^2}{\partial y^2}[a(y, t)g(y, t)] \quad (2)$$

is called the Fokker–Planck operator. CTRWs generalize random walks by allowing a larger, heavy-tailed class of waiting times before each jump. This translates into a *memory kernel* $V(y, t)$ acting on the time variable in the equation (Baeumer and Straka 2016):

$$\frac{\partial P(y, t)}{\partial t} = \mathcal{L}^*(y, t) \left[\frac{\partial}{\partial t} \int_0^t P(y, t-s)V(y, s) ds \right] + \delta_{(0,0)}(y, t). \quad (3)$$

The table below gives an overview over frequently studied forms of $V(y, s)$:

	Kernel	Laplace Transform	Reference
no memory	1	λ^{-1}	
subdiffusion	$s^{\alpha-1}/\Gamma(\alpha)$	$\lambda^{-\alpha}$	Sokolov and Klafter (2006)
tempered subdiffusion	unknown	$((\lambda+\theta)^\alpha - \theta^\alpha)^{-1}$	Gajda and Magdziarz (2010)

As the table indicates, most researchers have studied spatially constant memory kernels, without any dependence on the space variable y . This implies a homogeneous distribution of waiting times throughout the entire medium, i.e. that diffusion is equally anomalous everywhere. This assumption is of course too restrictive for some applications in biophysics (Wong et al. 2004; Straka and Fedotov 2015), e.g. when trapping varies due to locally different compositions of the cellular matrix. Moreover, media with two different anomalous exponents exhibit interesting, paradoxical behaviour (Korabel and Barkai 2010; Straka 2018), and have been studied (analytically) in the physics literature (Stickler and Schachinger 2011; Fedotov and Falconer 2012).

Numerous methods for the computation of solutions to *homogeneously anomalous* diffusion have been developed, among them explicit methods (Yuste and Acedo 2005), implicit methods (Langlands and Henry 2005), spectral methods (Li and Xu 2009; Hanert and Piret 2014) and Galerkin methods (Mustapha and McLean 2011). In the domain of *inhomogeneously anomalous* diffusion, several

authors have developed computational methods for *variable order* fractional Fokker–Planck Equations, but only the equation studied by Chen et al. (2010) is consistent with a CTRW scaling limit representation (Straka 2018).

The algorithm we introduce in this paper computes solutions to all Fokker–Planck equations of type (3) with spatially varying memory. Its only requirement is that the coefficients of the underlying bivariate Langevin process (Y_u, Z_u) , which tracks the cumulative sum of jumps resp. waiting times, can be evaluated numerically.

Similarly to the Discrete Time Random Walk (DTRW) method (C. N. Angstmann, Donnelly, Henry, and Nichols 2015; Angstmann et al. 2016), our algorithm calculates the probability distributions of a CTRW whose waiting times are grid-valued, and which approximates the continuum limit process. The advantages of this approach are that mass is necessarily conserved in each timestep; that solutions are guaranteed to be nowhere negative; and that stochastic process convergence implies the consistency of the algorithm. However, we do not rely on discrete Z-transforms, which means that our method remains tractable not just for Shibuya-distributed waiting times.

In the rest of this paper: We give a short account of bivariate Langevin dynamics and their relevance for this article in Section 2. In Section 3 we construct a sequence of DTRWs which converges to a CTRW continuum limit process, represented by a general bivariate Langevin equation (Y_u, Z_u) . The stochastic process convergence guarantees the convergence of probability distributions to the solutions $P(y, t)$ of (3), which translates into the consistency of the algorithm. In Section 4 we calculate the probability distributions of the DTRW. Because of the Semi-Markov property, this can be done via gearalized master equations in a higher dimensional state space which tracks the time since the last jump. In Section 5 we consider an interface problem, where the anomalous exponent interpolates continuously between two different anomalous exponents for each half axis. Section 6 concludes.

2. Stochastic solution to Fokker-Planck equation with memory

The Langevin representation of a stochastic process whose distribution $P(y, t)$ solves a Fokker-Planck equation with memory has been studied in various articles (Weron and Magdziarz 2008; Henry, Langlands, and Straka 2010; Gajda and Magdziarz 2010; Hahn et al. 2011). Recently, a Langevin representation for inhomogeneous anomalous diffusion was given (Straka 2018): Consider the bivariate Langevin process with state space $\mathbb{R} \times [0, \infty)$

$$dY_u = b(Y_u, Z_u) du + \sqrt{a(Y_u, Z_u)} dW_u \quad (4)$$

$$dZ_u = d(Y_u) du + \int_{w>0} w n(dw, du) \quad (5)$$

Here, u is auxiliary time, corresponding to the number of jumps; $b(y, t)$ and $a(y, t)$ are drift and diffusivity coefficients (of units length resp. length² per unit

auxiliary time) appearing in (3); $d(y)$ is a temporal drift coefficient (unit physical time per unit auxiliary time). Finally, $n(dw, du)$ denotes Levy noise that can be spatially varying. Recall that Levy noise has a representation as a Counting Measure, where for any rectangle $R = (u_1, u_2) \times (w_1, w_2) \subset [0, \infty) \times (0, \infty)$ the number of points $n(R)$ in R is Poisson distributed, and independent of any counts in other, disjoint rectangles (Applebaum 2009). The Poisson distribution, and hence the entire Counting Measure, is governed by a unique mean measure $m(dw, du)$ which satisfies $m(R) = \langle n(R) \rangle$. Examples:

- If $m(R) = (u_2 - u_1) \times \int_{w_1}^{w_2} \frac{\beta w^{-1-\beta}}{\Gamma(1-\beta)} dw$, then Z_u has independent and identically distributed increments, i.e. it is a Levy flight.
- Letting $m(R) = (u_2 - u_1) \times \int_{w_1}^{w_2} \frac{\beta w^{-1-\beta} e^{-\theta w}}{\Gamma(1-\beta)} dw$ results in Z_u being a *tempered stable Levy flight* with tempering parameter $\theta \geq 0$.

A dependence of the Levy measure on the position Y_u of the walker can be achieved via letting

$$m(R) = \int_{u_1}^{u_2} \int_{w_1}^{w_2} \nu(w|Y_u) dw du$$

for some *Levy measure* $\nu(dw|y)$, which may vary with y . Recall that a Levy measure is defined by the requirement

$$\int_0^\infty \min\{1, w\} \nu(dw|y) < \infty.$$

For instance, letting the fractional exponent $\beta(y) \in (0, 1]$ depend on space, choosing $\nu(w|y) = \frac{\beta(y) w^{-1-\beta(y)}}{\Gamma(1-\beta(y))}$ results in Z_u having independent increments, which follow the stable distribution with continuously varying exponent $\beta(Y_u)$ (Straka 2018).

It will be convenient to introduce the tail function of the Levy measure

$$\bar{\nu}(w|y) := \int_w^\infty \nu(dw|y), \quad w > 0.$$

and its Laplace transform

$$\hat{\bar{\nu}}(\lambda|y) = \int_0^\infty \bar{\nu}(w|y) e^{-\lambda w} dw.$$

We can then define the *renewal function* $V(y, s)$ via its Laplace transform

$$\hat{V}(y, \lambda) := \int_0^\infty V(y, s) e^{-\lambda s} ds = \frac{1}{\lambda[d(y) + \hat{\bar{\nu}}(\lambda|y)]}$$

As shown by Baeumer and Straka (2016), the Fokker–Planck equation with memory (3) has, under certain continuity conditions on the four coefficient

functions, a unique solution $P(y, t)$. This solution coincides with the probability distribution at time t of the subordinated process

$$X(t) := Y_{E(t)}, \quad E(t) := \inf\{u : Z_u > t\}. \quad (6)$$

$X(t)$ is also called a CTRW limit or the *continuum limit* of the CTRW.

Coefficient representation

We note that the 4-tuple

$$(a(x, t), b(x, t), d(x), \bar{\nu}(w, x)) \quad (7)$$

concisely represents the Langevin process (4)–(5). However, the representation is only unique up to a multiplicative factor: if every element in (7), say, doubled, then the speed of (Y_u, Z_u) is doubled. But, this has no effect on the distribution of the points that are traversed by (Y_u, Z_u) , and hence does not affect the distribution of the trajectories $X(t)$. This remains true even if the speed varies with the location (x, t) of (Y_u, Z_u) .

Assuming that the coefficients are all bounded functions in (x, t) , we hence divide by a large enough number so that $a(x, t) < 1$ for all (x, t) . (At $a(x, t) = 1$, numerical instabilities may occur, which are smoothed out if e.g. $a(x, t) < 0.9$ throughout the domain.) In the derivation of our algorithm, we will transform the tuple (7) as follows: Define $\theta(x)$ via $d(x) = \theta(x)/(1 - \theta(x))$. Then multiply the tuple (7) by $(1 - \theta(x))$, to get the transformed tuple

$$((1 - \theta(x))a(x, t), (1 - \theta(x))b(x, t), \theta(x), (1 - \theta(x))\bar{\nu}(w, y)). \quad (8)$$

Hence if we assume that $d(y)$ is bounded, then we can also assume WLOG that $0 \leq d(y) \leq 1$.

Finally, we add the technical but non-restrictive condition

$$\bar{\nu}(w|y) \leq G(y) \frac{w^{-\beta(y)}}{\Gamma(1 - \beta(y))}, \quad w \downarrow 0, \quad (9)$$

for some bounded function $G(y)$, which prevents the Levy measure from blowing up in regions where $\beta(y) \uparrow 1$, see Lemma 1.

3. Discrete Langevin Dynamics

Let $c > 0$ be a scaling parameter, and define a spatio-temporal grid $\#$ with spacings $\chi = c^{-1/2}$ and $\tau = c^{-1/\alpha_0}$, where $\alpha_0 \in (0, 1)$ is some reference value to be defined later. Assuming for simplicity that space is one-dimensional, the grid is embedded in space-time $\mathbb{R} \times [0, \infty)$. In this section we define for each

$c > 0$ a Langevin process $(Y_u^{(c)}, Z_u^{(c)})$ with state space $\#$ such that as $c \rightarrow \infty$, $(Y_u^{(c)}, Z_u^{(c)})$ converges to (Y_u, Z_u) .

It is clear that $(Y_u^{(c)}, Z_u^{(c)})$ must be a jump process hopping on $\#$. Since Y_u has continuous sample paths, nothing is gained by allowing $Y_u^{(c)}$ to jump to non-neighbouring lattice sites. Also, since Z_u is increasing, $Z_u^{(c)}$ need not jump backwards. It is helpful to view the sequence of grid points traversed by $(Y_u^{(c)}, Z_u^{(c)})$ as locations and times of a walker performing a DTRW (discrete time random walk), with jumps and waiting times given by the increments of $Y_u^{(c)}$ resp. $Z_u^{(c)}$.

3.1. Waiting time distribution

We define the discrete waiting time distribution $\psi^{(c)}(j\tau|x)$ as a mixture of a “local” and a “nonlocal” component:

$$\psi^{(c)}(j\tau|x) := \theta(x)\psi_{\text{loc}}^{(c)}(j\tau|x) + (1 - \theta(x))\psi_{\text{nonloc}}^{(c)}(j\tau|x), \quad j = 1, 2, \dots \quad (10)$$

where, by definition, $0 \leq d(x) \leq 1$. The local part is simply deterministic, with all mass at τ , that is $\psi_{\text{loc}}^{(c)}(\tau|x) = 1$ and $\psi_{\text{loc}}^{(c)}(k\tau|x) = 0$ for $k = 2, 3, \dots$. The nonlocal part is the truncated, normalized and discretized Lévy measure: First, define the function

$$H^{(c)}(w|x) = 1 \wedge \frac{\bar{\nu}(w|x)}{c},$$

where $a \wedge b := \min\{a, b\}$. For convenience, we say that $\bar{\nu}(w|x) = \infty$ if $w \leq 0$. Then, define

$$w_\tau := j\tau, \quad \text{where} \quad j\tau \leq w < (j+1)\tau. \quad (11)$$

Finally, note that $\Psi_{\text{nonloc}}^{(c)}(w|x) := H^{(c)}(w_\tau|x)$ is piecewise constant with jumps in $\tau, 2\tau, \dots$, and decreasing from 1 to 0. We take this function to be the tail function of $\psi_{\text{nonloc}}^{(c)}(w|x)$, that is,

$$\psi_{\text{nonloc}}^{(c)}(j\tau|x) = H^{(c)}((j-1)\tau|x) - H^{(c)}(j\tau|x), \quad k = 1, 2, \dots \quad (12)$$

We then have $\psi^{(c)}(0\tau|y) = 0$, meaning that waiting times are always strictly positive.

3.2. Jump distribution

We assume that the DTRW jumps can have one of the three values $\{-\chi, 0, +\chi\}$, where $\bar{a} = \sup\{a(x, t)\}$ and $\chi = (\bar{a}/c)^{1/2}$. The probabilities to jump left, to “self-jump” (i.e. jump back to the original location), and to jump right, are given by

$$\ell^{(c)}(x, t) = \frac{a(x, t) - \chi b(x, t)}{2\bar{a}}, \quad n(x, t) = 1 - a(x, t)/\bar{a}, \quad r^{(c)}(x, t) = \frac{a(x, t) + \chi b(x, t)}{2\bar{a}}.$$

where x is the location of the walker before the jump, and t is the time at which the jump occurs. In order for r, n and ℓ to be between 0 and 1, we need χ to be small enough so that

$$\chi|b(x, t)| \leq a(x, t), \quad (x, t) \in \mathbb{R} \times [0, \infty).$$

3.3. Convergence

At scale c , the probabilities $\psi^{(c)}(j\tau|y)$ and $\ell^{(c)}(x, t)$, $n(x, t)$ and $r^{(c)}(x, t)$ define a jump kernel on $\#$, which defines the distribution of jump z and waiting time w given the current location of the walker at x at time t :

$$K^{(c)}(z, w|x, t) = \left[r^{(c)}(x, t+w)\delta_{+\chi}(z) + n(x, t+w)\delta_0(z) + \ell^{(c)}(x, t+w)\delta_{-\chi}(z) \right] \psi^{(c)}(w|x). \quad (13)$$

Note that we evaluate the jump probabilities at the end $t+w$ of a waiting time, as is common for CTRWs. Th. 2.1 in (Straka 2018) specifies conditions on $K^{(c)}(z, w|x, t)$ which imply the convergence of $(Y_u^{(c)}, Z_u^{(c)})$ to (Y_u, Z_u) and which we repeat here for convenience:

$$\lim_{\epsilon \downarrow 0} \lim_{c \rightarrow \infty} c \iint_{|z| < \epsilon, 0 < w < \epsilon} z K^{(c)}(z, w|x, s) dz dw = b(x, s) \quad (14)$$

$$\lim_{\epsilon \downarrow 0} \lim_{c \rightarrow \infty} c \iint_{|z| < \epsilon, 0 < w < \epsilon} z^2 K^{(c)}(z, w|x, s) dz dw = a(x, s) \quad (15)$$

$$\lim_{\epsilon \downarrow 0} \lim_{c \rightarrow \infty} c \iint_{|z| < \epsilon, 0 < w < \epsilon} w K^{(c)}(z, w|x, s) dz dw = \theta(x) \quad (16)$$

$$\lim_{c \rightarrow \infty} c \iint_{z \in \mathbb{R}, w \geq 0} g(z, w) K^{(c)}(z, w|x, s) dz dw = \int_{w > 0} g(0, w)(1 - \theta(x)) \nu(w|x) dw \quad (17)$$

for any bounded continuous function $g(z, w)$ which vanishes in a neighbourhood of the origin. We give calculations in the appendix which confirm that the above four conditions indeed hold for $K^{(c)}(z, w|x, t)$.

Remark

The alternative kernel

$$K^{(c)}(z, w|x, t) = \left[r^{(c)}(x, t)\delta_{+\chi}(z) + n(x, t)\delta_0(z) + \ell^{(c)}(x, t)\delta_{-\chi}(z) \right] \psi^{(c)}(w|x). \quad (18)$$

also satisfies (14) – (17). The difference to (13) is that the probabilities $r^{(c)}(x, t)$, $\ell^{(c)}(x, t)$ and $n(x, t)$ are evaluated at the *beginning* of a waiting time, rather than the end. As investigated by C. N. Angstmann, Donnelly, Henry, Langlands, et al. (2015), this difference vanishes in the limit as $c \rightarrow \infty$.

4. Semi-Markov numeric scheme

As described at the beginning of Section 3, the discrete Langevin process $(Y_u^{(c)}, Z_u^{(c)})$ has an embedded DTRW, for which we write $X^{(c)}(t)$. By Theorem 2.2 in Straka (2018), $X^{(c)}(t)$ converges to the CTRW continuum limit process $X(t)$ from (6). For large c , the probability distributions of $X^{(c)}(t)$ may hence be taken as approximations of $P(y, t)$. In this section, we derive master equations for the probability distributions of $X^{(c)}(t)$.

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$$X^{(c)}(t) \text{ converges to the CTRW continuum limit process } X(t) \quad (19)$$

from (6). (Convergence here means weak convergence with respect to the J_1 topology of right-continuous sample paths with left-hand limits, see Whitt (2001).) For large c , the probability distributions of $X^{(c)}(t)$ may hence be taken as approximations of $P(y, t)$. In this section, we derive master equations for the probability distributions of $X^{(c)}(t)$.

4.1. Semi-Markov property

A DTRW starting at x at time t is defined by the jump kernel (13) as follows: first, a waiting time is drawn from the distribution $\psi^{(c)}(w|x)$; then a jump left or right or a self-jump is drawn from the probabilities $\ell^{(c)}(x, t+w)$, $r^{(c)}(x, t+w)$ and $n(x, t+w)$. The Semi-Markov approach embeds $X^{(c)}(t)$ into a Markov process as follows: Define the *age* of a walker as the time that has passed since he last arrived at his current location. In each timestep τ , either the waiting time has not expired yet, in which case no jump occurs and age is increased by τ ; or age is reset to 0 and a jump occurs. Since this recipe determines the future evolution of position and age based on only the current position and age, the process is Markovian, and it is straightforward to derive master equations.

Recall that a waiting time W at a spatial lattice point $i\chi$ is drawn from $\psi^{(c)}(w|i\chi)$ and thus satisfies

$$\mathbf{P}(W > j\tau) = H^{(c)}(j\tau|i\chi) =: h_{i,j}.$$

Conditional on $W > j\tau$, the probability that $W > (j+1)\tau$ is

$$\mathbf{P}(W > (j+1)\tau | W > j\tau) = h_{i,j+1}/h_{i,j}.$$

That is, if at time $k\tau$, position and age are $(x_k, v_k) = (i, j)$, then at time $(k+1)\tau$

- with probability $h_{i,j+1}/h_{i,j}$ we have $(x_{k+1}, v_{k+1}) = (x_k, v_k + 1)$, and
- with probability $1 - h_{i,j+1}/h_{i,j}$ we have $(x_{k+1}, v_{k+1}) = (x_k + \zeta, 0)$,

where $\zeta \in \{-1, 0, +1\}$ with probabilities $\ell^{(c)}(i_\chi, (k+1)\tau)$, $n(i_\chi, (k+1)\tau)$ and $r^{(c)}(i_\chi, (k+1)\tau)$.

The above dynamics uniquely determine the stepwise evolution of (x_k, v_k) . We write $\xi_{i,j}^k = \mathbf{P}(x_k = i, v_k = j)$ for the probability distribution of (i, j) at time k . The master equations for $\xi_{i,j}^k$ then read:

$$\xi_{i,j}^{k+1} = \frac{h_{i,j}}{h_{i,j-1}} \xi_{i,j-1}^k, \quad 1 \leq j < J-1, \quad (20)$$

$$\xi_{i,0}^{k+1} = \sum_{j=0}^J \left(1 - \frac{h_{i,j+1}}{h_{i,j}}\right) (\ell_{i+1}^k \xi_{i+1,j}^k + r_{i-1}^k \xi_{i-1,j}^k + n_{i,j}^k \xi_{i,j}^k) \quad (21)$$

The line (20) states that for a walker to have age $j \geq 1$, it must have had age $j-1$ in the previous time step, and not jumped. The line (21) states that for a walker to have age $j=0$, it must have jumped to its location i in the previous time step, from a neighbouring lattice site or from i itself. The probability mass of all walkers jumping from site i during time step $k \rightarrow k+1$ is $\sum_{j=0}^J (1 - h_{i,j+1}/h_{i,j}) \xi_{i,j}^k$, which is redistributed according to the probabilities $r_{i,j}^{k+1}$, $\ell_{i,j}^{k+1}$ and $c_{i,j}^{k+1}$. This interpretation shows that (20)–(21) **conserve probability mass**.

Iterating the equation pair (20)–(21) from some initial condition computes the evolution of the joint probability distribution of position and age. The marginal distribution of the position is calculated simply via

$$\mathbf{P}(X_t^{(c)} = i_\chi) =: \rho_i^k = \sum_{j=0}^J \xi_{i,j}^k, \quad k = \lfloor t/\tau \rfloor.$$

Here we note that $X^{(c)}(t) = X^{(c)}(t_\tau)$, where t_τ is the left-nearest lattice point defined exactly as w_τ in (11).

4.2. Boundary conditions

In practice, one can only allocate a finite number J of points to the lattice of ages. If we cannot allocate $\lfloor T/\tau \rfloor$ lattice points, where T is the largest time of interest, then it is possible that the age of walkers may reach the end of the lattice. In this case, and if the walker does not jump in the next time step, we do not increase its age any further, until it eventually does jump:

$$\xi_{i,J}^{k+1} = \frac{h_{i,J}}{h_{i,J-1}} \xi_{i,J-1}^k + \frac{h_{i,J+1}}{h_{i,J}} \xi_{i,J}^k.$$

Finally, assuming that the spatial coordinates of the lattice go from $-I$ to I , we implement Neumann boundary conditions by placing a walker back on the boundary whenever it would otherwise have jumped off the lattice, that is:

$$\ell_{-I}^k = 0, \quad n_{-I}^k = \ell(-I\chi, k\tau) + n(-I\chi, k\tau), \quad r_{-I}^k = r(-I\chi, k\tau), \quad (22)$$

$$\ell_I^k = \ell(I\chi, k\tau), \quad n_I^k = n(I\chi, k\tau) + r(I\chi, k\tau), \quad r_I^k = 0 \quad (23)$$

4.3. Properties of the algorithm

Positivity

From (20)–(21), it is evident that the $\xi_{i,j}^k$ are necessarily non-negative, and hence the solution ρ_i^k cannot be negative.

Consistency of the algorithm

Due to the convergence (19), we have

$$\sum_{i=-I}^I f(i\chi) \rho_i^{\lfloor t/\tau \rfloor} = \langle f(X_t^{(c)}) \rangle \longrightarrow \langle f(X_t) \rangle \quad \text{as } c \rightarrow \infty, \quad (24)$$

for all bounded continuous real-valued f defined on \mathbb{R} . Assuming that the distribution of X_t has a probability density, we may replace f by an indicator function of an interval (a, b) , and (24) reads

$$\sum_{a < i\chi < b} \rho_i^{\lfloor t/\tau \rfloor} \longrightarrow \mathbf{P}(a < X_t < b) \quad \text{as } c \rightarrow \infty. \quad (25)$$

Equivalence with DTRW approach

The Discrete Time Random Walk algorithm by C. N. Angstmann, Donnelly, Henry, and Nichols (2015) assumes discrete waiting times with the Sibuya distribution, whose tail function $\Psi(n)$ has the asymptotics $\Psi(n) \sim n^{-\beta}$. In (21), see that we have $\xi_{i,j}^k = \xi_{i,0}^{k-j} h_{i,j}$, by telescoping (20) and $h_{i,0} = 1$. Hence (21) rewrites to

$$\xi_{i,0}^{k+1} = \sum_{j=0}^J (h_{i,j} - h_{i,j+1}) (\ell_{i+1}^{k+1} \xi_{i+1,0}^{k-j} + r_{i-1}^{k+1} \xi_{i-1,0}^{k-j} + c_{i,j}^{k+1} \xi_{i,0}^{k-j}),$$

assuming that $h_{i,j}$ is constant in i (homogeneous waiting times). Since $h_{i,j} - h_{i,j+1}$ is the probability of a waiting time being $j+1$, one sees the equivalence of methods by comparing with Equation (16) in C. N. Angstmann, Donnelly, Henry, and Nichols (2015), if we choose $h_{i,j} = \Psi(j)$.

5. Examples

Within our unifying semi-Markov framework, we may approximate probability distributions of a great variety of CTRW limits. We study several examples.

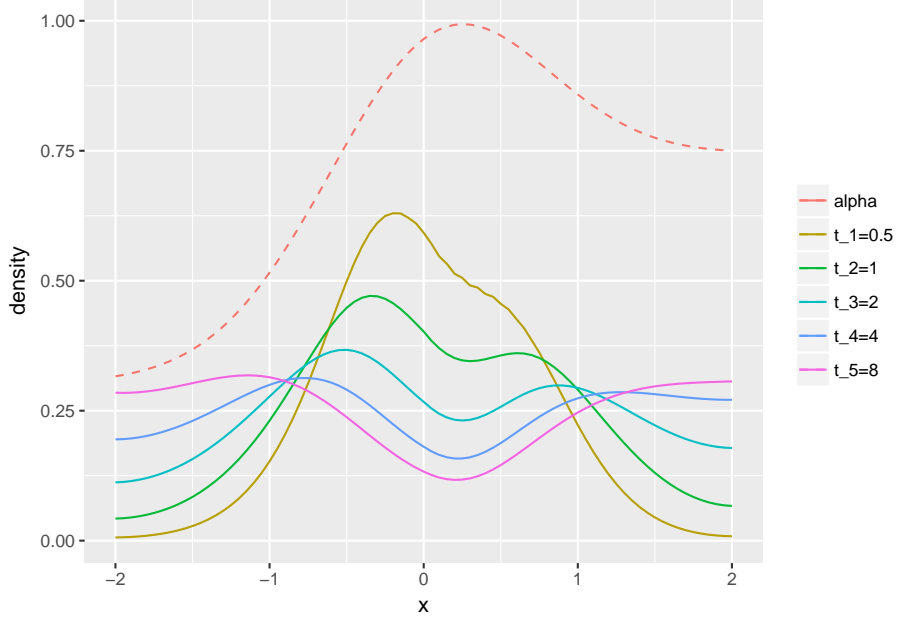


Figure 1: Continuous interface problem. Coefficients are as given in the text. Other numeric parameters are $c = 100$, $\chi = 1/10$ and $\tau = 1/100$.

5.1. Continuous interface problem

Korabel and Barkai (2010) have studied a one-dimensional subdiffusive lattice with exponent $\beta = 0.3$ for $x < 0$ and $\beta = 0.75$ for $x > 0$, where at the interface ($x = 0$) the waiting time is exponentially distributed. Even if particles are biased to jump to the right at $x = 0$ and thus the net drift becomes positive, in the long-time limit *all particles end up in the left half*.

Here we consider a continuous medium that mimics this setup with the coefficients $(\bar{\nu}(w|x), d(x), a(x, t), b(x))$ chosen as follows:

$$\bar{\nu}(w|x) = \frac{w^{-\alpha(x)}}{\Gamma(1 - \alpha(x))} \text{ where } \alpha(x) = 0.45e^{-x^2} + 0.3 + 0.45/(1 + e^{-2x}),$$

$$d(x) \equiv 0, \quad a(x, t) \equiv 1, \quad b(x, t) = 0.1 * \phi(x|0, 0.2)$$

where $\phi(x|\mu, \sigma)$ denotes the probability density of the Gaussian distribution with mean μ and standard deviation σ . Note that $\alpha(x, t)$ is chosen so that it approaches 0.3 for large negative x , 0.75 for large positive x and remains just under 1 near $x = 0$.

Figure 1 shows the evolution of the density $P(y, t)$ with a delta function initial condition. At small times we observe two peaks reflecting the trapping that occurs either side of the interface. For late times we can see the aggregation of all particles towards the left hand side ($y < 0$) where trapping is stronger.

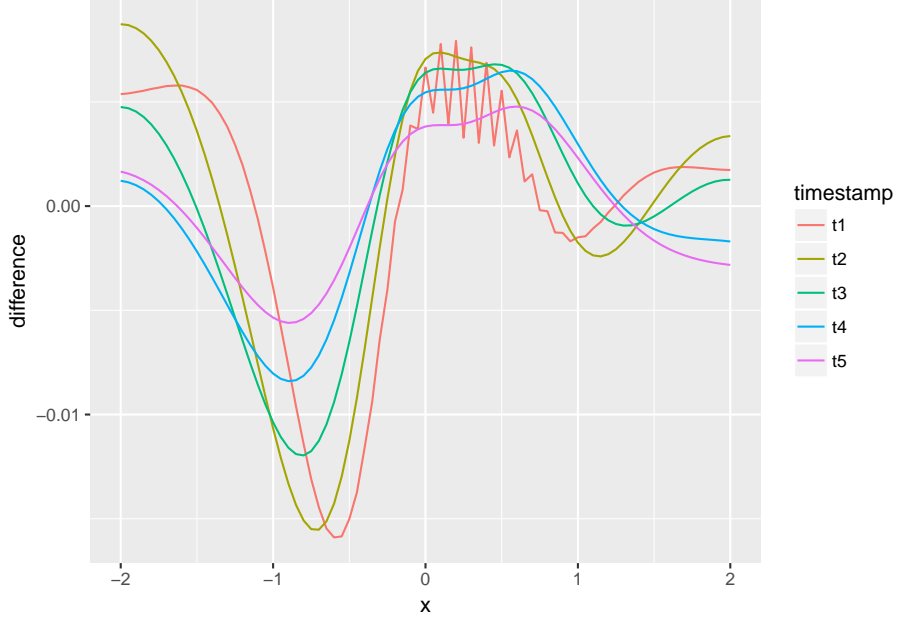


Figure 2: Absolute difference between densities, calculated at corresponding timestamps, for two choices of time scale $T_0 = 1$ and $T_0 = 2$.

Straka (2018) shows that changing time units from $T_0 = 1$ to $T_0 = 2$ results in the the updated diffusivity and drift coefficients

$$a_{\beta(y)}(x, t) = \frac{a(x, t)}{T_0^{-\alpha(x)}}, \quad b_{\beta(y)}(x, t) = \frac{b(x, t)}{T_0^{-\alpha(x)}},$$

leading to spatially inhomogeneous temporal scaling. We confirm this by computing probability densities for the parameter tuple $(\bar{\nu}(w|x), d(x), a_{\alpha(x)}(x, t), b_{\alpha(x)}(x))$, at the timestamps multiplied by $T_0 = 2$, and plotting the differences (Figure 2). Relative differences are near 1%, and remain stable after 8 units of time, indicating that indeed the same densities are calculated in both cases.

5.2. Temporal drift $d(x)$

CTRW limits with positive temporal drift $d(x)$ as per representation (4)–(5) have been studied by Straka (2011): In the case where Z_u is a β stable Lévy flight, Z_u grows superlinearly at the rate $u^{1/\beta}$ both in the short time limit $t \downarrow 0$ and the long time limit $t \uparrow \infty$. Accordingly, the inverse stable subordinator $E(t)$ in (6) grows as $\propto t^\beta$, also both in the short time and long time limit. Adding a drift to Z_u , e.g. $d(x) \equiv d > 0$, means that Z_u now grows linearly $\propto du$ at short times. Accordingly, its inverse $E(t)$ also grows linearly as $\propto t/d$ at short times. The growth behaviour at late times of Z_u and $E(t)$ remains dominated by large

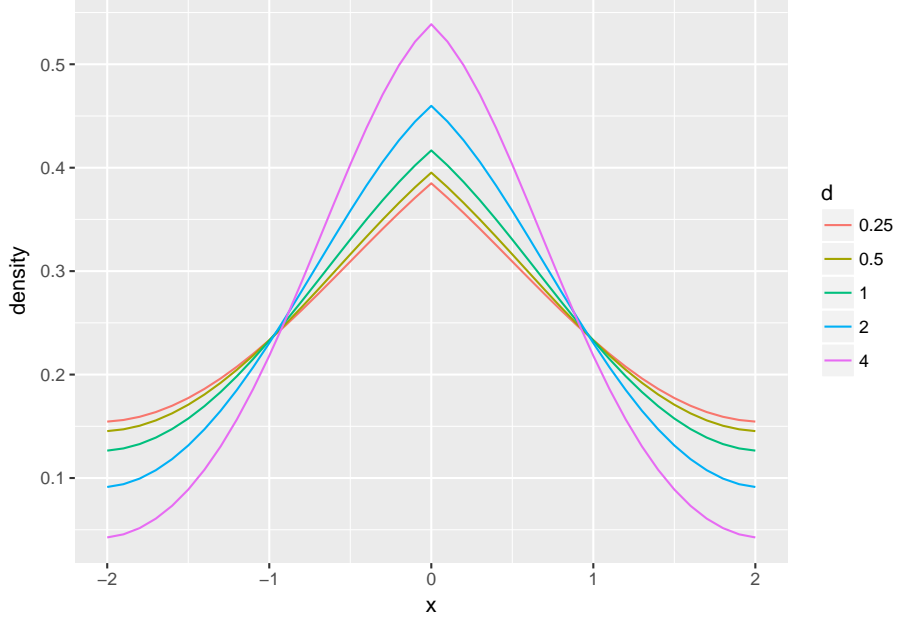


Figure 3: Increasing temporal drift $d(x)$ decreases the speed of the diffusion and increases resemblance to a standard Gaussian process. Parameters: $a(x, t) = 0.8, b(x, t) = 0, \bar{\nu}(w|x) = w^{-0.7}/\Gamma(1 - 0.7), c = 100, \tau = 1/100, \chi = 1/10$.

jumps resp. long rests, and remains $\propto u^{1/\beta}$ resp. $\propto t^\beta$. Hence the addition of the drift $d > 0$ means that the slope of $E(t)$ is no longer infinite, and thus the speed of $E(t)$ is tempered at very short times. Figure 3 illustrates the effect of increasing the temporal drift. As can be seen, the jump component of Z_u becomes less pronounced as the temporal drift increases, increasing resemblance to a Gaussian process. Figure 4 shows anomalous diffusion with exponent 0.7 with spatially varying temporal drift $d(x)$. Particles accumulate in patches of low mobility, corresponding to high $d(x)$.

5.3. Variably distributed fractional order

Anomalous diffusion with *distributed order* assumes a mixing probability distribution $p(\beta) d\beta$ on the interval $(0, 1]$, and the

For constant a, b and d with $d \geq 0$ and $\nu(\omega|y) \equiv \nu(\omega)$, it can be shown that [equation 1.3] can be written as a special case of a Distributed order FFPE with a distribution over the orders of β and 1 [equation (3.8) in JphysA paper]. This leads to a mixture of orders with the smaller one β dominating the long-time behaviour. The behaviour of such a process is similar to that of a CTRW with two different underlying waiting time densities $\psi_1(t)$ and $\psi_2(t)$ [SandeV paper].

Here we are now able to consider a qualitatively similar mixed type fractional diffusion by defining $\bar{\nu}(\omega|y) = p(y)\bar{\nu}_{\beta_1}(\omega) + (1 - p(y))\bar{\nu}_{\beta_2}(\omega)$ where $y, p(y) \in [0, 1]$. The weight function $p(y)$ allows us to vary the bifractional mixture in space.

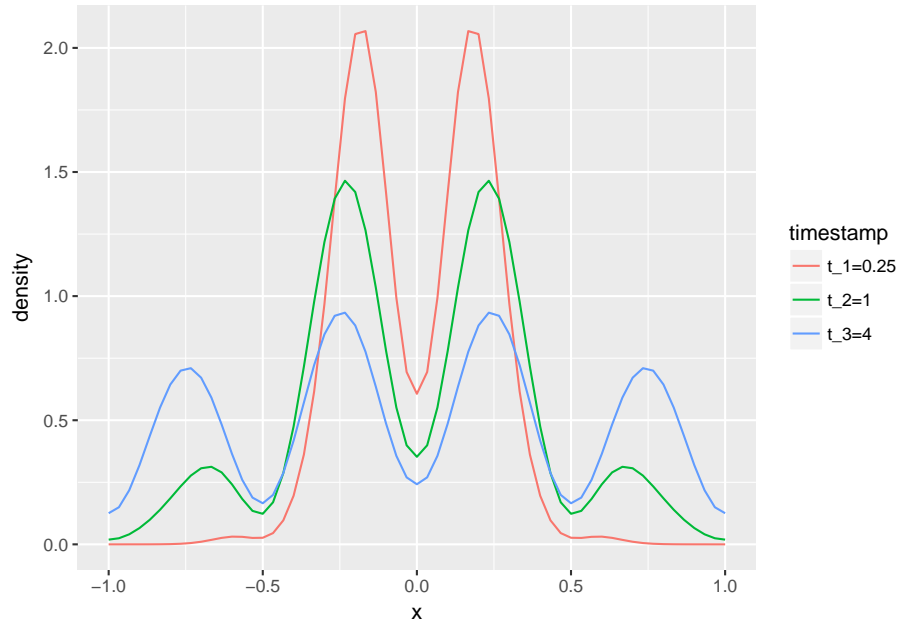


Figure 4: A system with a spatially varying temporal drift $d(x) = 10 * \sin(2\pi x)$. Particles accumulate in the slow patches where $d(x)$ is high, while trapping is homogeneous in space. Other parameters: $a(x, t) = 0.8, b(x, t) = 0, \bar{v}(w|x) = w^{-0.7}/\Gamma(1-0.7), c = 900, \tau = 1/900, \chi = 1/30$.

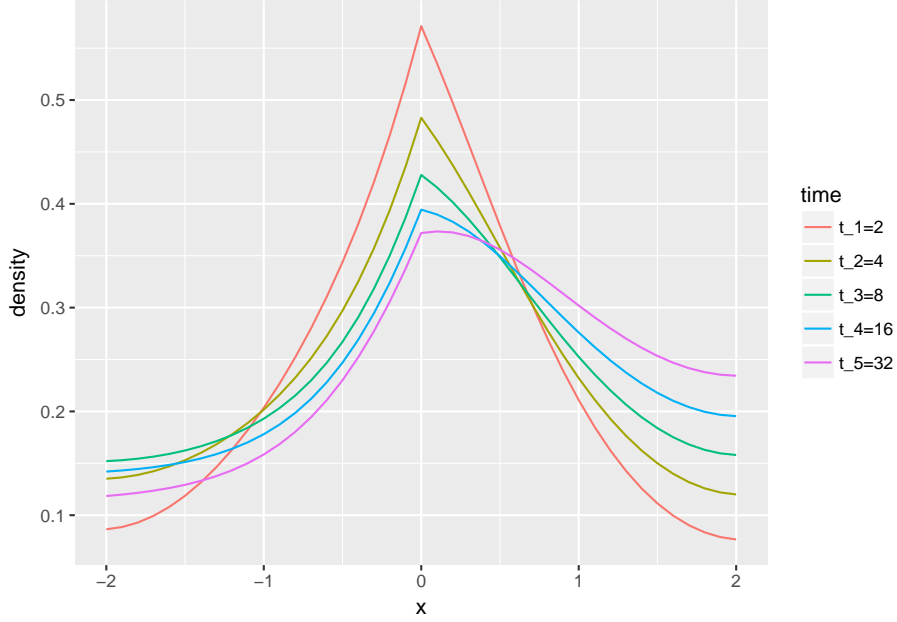


Figure 5: Distributed order anomalous diffusion with spatially varying mixture. The mixture has two modes, $\beta_1 = 0.7$ (left) and $\beta_2 = 0.3$ (right). The mixing is according to $(p(x), 1 - p(x))$, where $p(x) = 1/(1 + \exp(-2x))$.

Figure [fig:varyingmixture] shows the resulting PDF $P(y, t)$ when the weight function follows a logistic function $p(y) = 1/(1 + e^{-10y+5})$ on $y \in [0, 1]$ and initial condition is $P(dy, 0) = \delta_0(dy)$. Note that as $t \rightarrow \infty$, $P(y, t) \rightarrow \delta(y - 1)$ (see [Fedotov/Falconer]), hence [fig:varyingmixture] represents the intermediate evolution of the density, the speed of which is controlled by parameters β_1 , β_2 and $p(y)$.

5.4. Inverse tempered stable subordinator

A. Checking conditions (14) – (17)

The following lemma pertains to the calculations in the waiting times of (14) – (17):

Lemma 1. *Under condition (9), the waiting time distribution (10) satisfies, as*

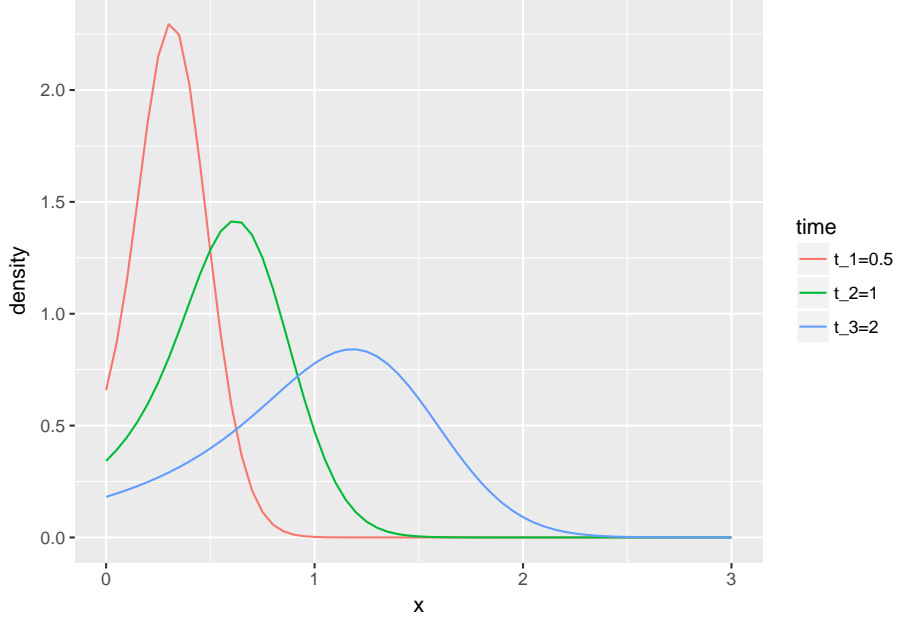


Figure 6: Probability densities of the tempered inverse stable subordinator.

$c \rightarrow \infty$,

$$\int f(w)\psi^{(c)}(w|y) dw = \sum_{j=1}^{\infty} f(j\tau)\psi^{(c)}(j\tau|y) \rightarrow f(0), \quad (26)$$

$$c \int g(w)\psi^{(c)}(w|y) dw = c \sum_{j\tau>0} g(j\tau)\psi^{(c)}(j\tau|y) \rightarrow \int g(w)\nu(w|y) dw, \quad (27)$$

$$c \int_0^{\varepsilon} w\psi^{(c)}(w|y) dw = c \sum_{0<j\tau\leq\varepsilon} j\tau\psi^{(c)}(j\tau|y) \rightarrow d(y) + \mathcal{O}\left(\frac{\varepsilon^{1-\beta(y)}}{\Gamma(1-\beta(y))}\right), \quad \varepsilon > 0. \quad (28)$$

for any bounded continuous f and g , where g vanishes in a neighbourhood of 0.

Proof. (26) holds since $\psi^{(c)}(w|y)$ is a probability distribution on the positive numbers with tail function

$$\Psi^{(c)}(w|x) = d(x)\mathbf{1}\{w \leq \tau\} + (1 - d(x))H^{(c)}(w_\tau|x)$$

which for all $w > 0$ satisfies $\Psi^{(c)}(w|y) \rightarrow 0$ as $c \rightarrow \infty$ (recall that $\tau = 1/c \downarrow 0$). For (27), we first note that

$$c\Psi^{(c)}(w|y) = cd(y)\mathbf{1}(w \leq \tau) + [c(1 - d(y))] \wedge \bar{\nu}(w_\tau) \rightarrow \bar{\nu}(w), \quad c \rightarrow \infty, \quad (29)$$

for every $w > 0$. Assume that g is differentiable, and let $\varepsilon > 0$ be small enough so that $g(\varepsilon) = 0$. Using (Lebesgue-Stieltjes) integration by parts, we may calculate

$$\begin{aligned} c \int_0^\infty g(w) \psi^{(c)}(w|y) dw &= c \int_\varepsilon^\infty g(w) \psi^{(c)}(w|y) dw = c \int_\varepsilon^\infty g'(w) \Psi^{(c)}(w|y) dw \\ &\rightarrow \int_\varepsilon^\infty g'(w) \bar{\nu}(w|y) dw = \int_\varepsilon^\infty g(w) \nu(w|y) dw = \int_0^\infty g(w) \nu(w|y) dw. \end{aligned}$$

But bounded continuous functions can be approximated by differentiable functions with arbitrary accuracy, so (27) follows.

Finally, for (28), we consider the local and nonlocal parts $\psi_{\text{loc}}^{(c)}(w|x)$ and $\psi_{\text{nonloc}}^{(c)}(w|x)$ separately. For the local part, we have

$$c \int_0^\varepsilon w \psi_{\text{loc}}^{(c)}(w|x) dw = c\tau \rightarrow 1.$$

For the nonlocal part, we use Lebesgue-Stieltjes integration by parts:

$$\begin{aligned} \int_0^\varepsilon w \psi_{\text{nonloc}}^{(c)}(w|x) dw &= \int_0^\varepsilon w \left(-d\Psi_{\text{nonloc}}^{(c)}(w|x) \right) \\ &= - \left[w \Psi_{\text{nonloc}}^{(c)}(w|x) \right]_0^\varepsilon + \int_0^\varepsilon \Psi_{\text{nonloc}}^{(c)}(w|x) dw \\ &= -\varepsilon \Psi_{\text{nonloc}}^{(c)}(\varepsilon|x) + \int_0^\varepsilon \Psi_{\text{nonloc}}^{(c)}(w|x) dw \end{aligned}$$

Multiplying with c and letting $c \rightarrow \infty$, the right hand side converges to

$$-\frac{\varepsilon \bar{\nu}(\varepsilon|x)}{1-d(x)} + \int_0^\varepsilon \frac{\bar{\nu}(w|x)}{1-d(x)} dw$$

where both terms are of order $\mathcal{O}(\varepsilon^{1-\beta(x)}/\Gamma(1-\beta(x)))$ according to the technical assumption (9). (28) now follows from the definition (10) of $\psi^{(c)}(w|x)$. \square

The final lemma pertains to the jump distributions in (14)–(17):

Lemma 2. *The jump probabilities $\ell^{(c)}(x, t)$, $r^{(c)}(x, t)$ and $n(x, t)$ satisfy*

$$c[-\chi \ell^{(c)}(x, t+w) + \chi r^{(c)}(x, t+w)] = b(x, t+w) \quad (30)$$

$$c\chi^2[\ell^{(c)}(x, t+w) + r^{(c)}(x, t+w)] = a(x, t+w) \quad (31)$$

$$\int_{\mathbb{R}} f(z) \left[r^{(c)}(x, t) \delta_\chi(z) + n(x, t) \delta_0(z) + \ell^{(c)}(x, t) \delta_{-\chi}(z) \right] dz \rightarrow f(0) \quad (32)$$

as $c \rightarrow \infty$ for all bounded continuous f .

Proof. This follows easily from the definitions of the jump probabilities. \square

Finally, to see that (14)–(15) hold, use (30)–(31) and (26). To see (16), use (28) and let $\varepsilon \downarrow 0$; and finally, to see (17), use (27) and (32).

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