Hidden Markov Model

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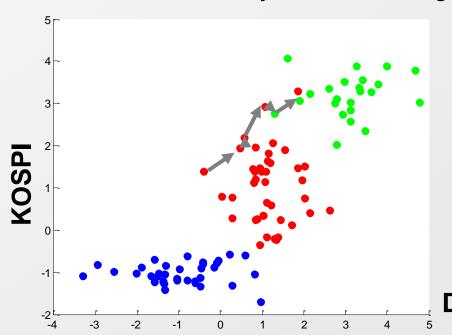
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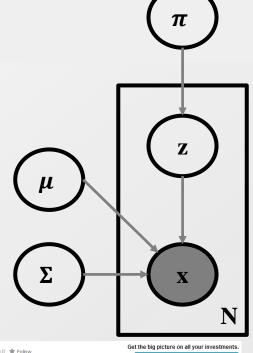
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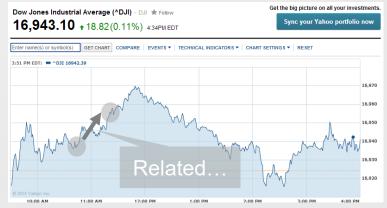
HIDDEN MARKOV MODEL

Times Series Data for GMM

- Imagine the following case
 - Data points on the plane
 - Have a temporal trace of data points
 - Now, any broken assumption in the analysis?
- Any real world applications
 - Many, many, many...
 - Stock market analysis, text mining…







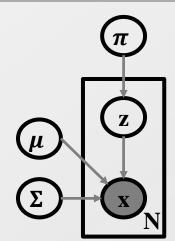
DJIA

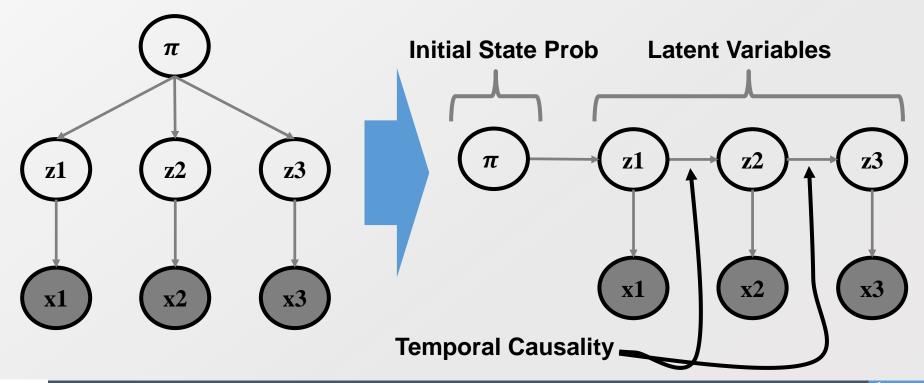
If they make the ballot in November, an array of proposals will be among the first in the nation to ask a state's voters to sharply

What to Model and How to Model



- Previously, all data points are independent trials
 - Now, they are not any further
- Temporal relation: causality from time t to time t+1
- Overall trend: latent state variables

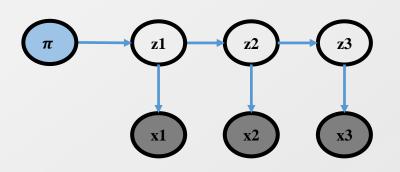


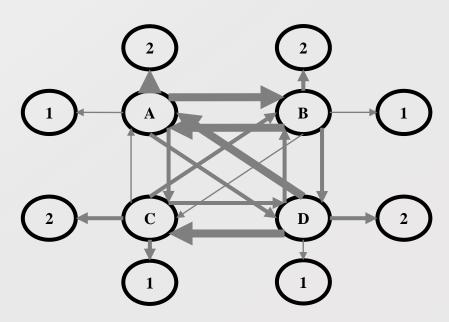


Hidden Markov Model



- Observation, x
 - Can be either discrete or continuous
 - Just a difference in probability distributions
 - Will only handle discrete case in this course
 - $x_1...x_T$: Observation from time 1 to time T
 - $x_i \in \{c_1, ..., c_m\}$: m types of observation values
- Latent state, z
 - Vector variable with K elements
 - Let's say that there are K types of state values corresponding to each element
 - Can be either discrete or continuous
 - If continuous → Kalman filter, and this is a continuous version of HMM
 - Out of scope of this course
- Initial State probabilities
 - $P(z_1) \sim Mult(\pi_1, ..., \pi_k)$
- Transition probabilities
 - $P(z_t|z_{t-1}^i=1) \sim Mult(a_{i,1},...,a_{i,k})$
 - Or, $P(z_t^j = 1 | z_{t-1}^i = 1) = a_{i,j}$
- Emission probabilities
 - $P(x_t|z_t^i = 1) \sim Mult(b_{i,1}, ..., b_{i,m}) \sim f(x_t|\theta_i)$
 - Or, $P(x_t^j = 1 | z_t^i = 1) = b_{i,j}$
- A stochastic generative model





Main Questions on HMM

- Given the topology of the Bayesian network, HMM, or M
- Evaluation question
 - Given **π**, **a**, **b**, X
 - Find P(X|M, π , a, b)
 - How much is X likely to be observed in the trained model?
- Decoding question
 - Given *π*, *a*, *b*, X
 - Find $argmax_z P(Z|X, M, \pi, a, b)$
 - What would be the most probable sequences of latent states?
- Learning question
 - Given X
 - Find $argmax_{\pi, a, b}P(X|M, \pi, a, b)$
 - What would be the underlying parameters of the HMM given the observations?
- Decoding questions and learning questions are very similar to
 - Supervised and unsupervised learning
- Anyhow, we often need to find π , a, b prior to the supervised learning with X

Initial State probabilities $P(z_1) \sim Mult(\pi_1, ..., \pi_k)$

Transition probabilities

$$P(z_{t}|z_{t-1}^{i}=1) \sim Mult(a_{i,1},...,a_{i,k})$$

Or, $P(z_{t}^{j}=1|z_{t-1}^{i}=1)=a_{i,j}$

$$\begin{split} &P(x_t|z_t^i=1) \sim \textit{Mult}(b_{i,1}, \dots, b_{i,m}) \sim f(x_t|\theta_i) \\ &\text{Or, } P\Big(x_t^j=1 \Big| z_t^i=1\Big) = b_{i,j} \end{split}$$

Obtaining π , a, b given X and M

Initial State probabilities

$$P(z_1) \sim Mult(\pi_1, ..., \pi_k)$$

Transition probabilities

$$P(z_t|z_{t-1}^i = 1) \sim Mult(a_{i,1}, ..., a_{i,k})$$

 $Or, P(z_t^j = 1|z_{t-1}^i = 1) = a_{i,j}$

Emission probabilities

$$P(x_t|z_t^i = 1) \sim Mult(b_{i,1}, ..., b_{i,m}) \sim f(x_t|\theta_i)$$

Or, $P(x_t^j = 1|z_t^i = 1) = b_{i,j}$



M_i observations for i-th sequence

- Finding π , a, b from the data in the supervised learning approach requires X as well as Z
- Example scenario
 - Loaded dice and fair dice
 - Two dices yield different probability distributions from one to six
 - · Dealer changes the dice as he wishes
- Probability estimation
 - Use MLE, MAP and counting...
 - Find out
 - Dealer starts with a certain dice type: $P(z_1^L = 1) = 1/2$
 - Dealer switches the dice: $P(z_t^L = 1 | z_{t-1}^L = 1) = 0.7, P(z_t^L = 1 | z_{t-1}^F = 1) = 0.5$
 - Loaded dice: P(X = 1) = P(X = 2) = P(X = 3) = P(X = 4) = P(X = 5) = 1/10, P(X = 6) = 1/2
 - Fair dice: P(X = 1) = P(X = 2) = P(X = 3) = P(X = 4) = P(X = 5) = P(X = 6) = 1/6
- What if the X is continuous? Use a known distribution, and estimate its parameters

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Joint Probability

 Let's assume that we have a training dataset with X and Z

Initial State probabilities

 $P(z_1) \sim Mult(\pi_1, ..., \pi_k)$

Transition probabilities

$$P(z_t|z_{t-1}^i=1) \sim Mult(a_{i,1},...,a_{i,k})$$

Or,
$$P(z_t^j = 1 | z_{t-1}^i = 1) = a_{i,j}$$

Emission probabilities

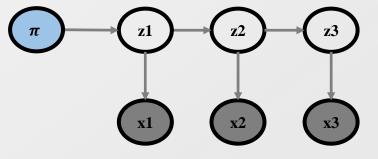
$$P(x_t|z_t^i=1) \sim Mult(b_{i,1},...,b_{i,m}) \sim f(x_t|\theta_i)$$

Or, $P(x_t^j=1|z_t^i=1) = b_{i,j}$



M_i observations for i-th sequence

- Can we compute the joint probability, P(X,Z)
 - Yes. Easily by the virtue of the network structure
- Anyway, a Bayesian network, so...
 - Factorize
 - $P(X,Z) = P(x_1,...,x_t,z_1,...,z_t)$
 - $= P(z_1)P(x_1|z_1)P(z_2|z_1)P(x_2|z_2)P(z_3|z_2)P(x_3|z_3)$
 - Nothing but a combination of initial, transition, and emission probabilities
 - = $\pi_{idx(z_1=1)}b_{idx(x_1=1),idx(z_1=1)}a_{idx(z_1=1),idx(z_2=1)}...$
- Assume that we have 166 as X
 - Let's check Z=LLL and FFF
 - $P(166, LLL) = \frac{1}{2} \times \frac{1}{10} \times \frac{7}{10} \times \frac{1}{2} \times \frac{7}{10} \times \frac{1}{2} = 0.0061$
 - $P(166, FFF) = \frac{1}{2} \times \frac{1}{6} \times \frac{1}{2} \times \frac{1}{6} \times \frac{1}{2} \times \frac{1}{6} = 5.7870e 04$
 - What about FLL, FFL, FLF.....? Exponential combination to check



Marginal Probability

- Eventually, we only want to use X and marginalize Z
 - Just like GMM, $P(X|\theta) = \sum_{Z} P(X, Z|\theta)$
 - In HMM,

$$P(X|\pi, a, b) = \sum_{Z} P(X, Z|\pi, a, b)$$

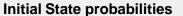
- $P(X) = \sum_{Z} P(X, Z) = \sum_{z_1} ... \sum_{z_t} P(x_1, ..., x_t, z_1, ..., z_t)$
- $= \sum_{z_1} \dots \sum_{z_t} \pi_{z_1} \prod_{t=2}^T \alpha_{z_{t-1}, z_t} \prod_{t=1}^T b_{z_t, x_t}$
 - Many summations yield an exponential number of combinations
- Need to avoid a repetitive computing
 - Compute only necessary terms for a single time
 - Let's work on the formula
 - P(A,B,C)=P(A)P(B|A)P(C|A,B)
 - $P(x_1, ..., x_t, z_t^k = 1) = \sum_{z_{t-1}} P(x_1, ..., x_{t-1}, x_t, z_{t-1}, z_t^k = 1)$
 - = $\sum_{z_{t-1}} P(x_1, ..., x_{t-1}, z_{t-1}) P(z_t^k = 1 | x_1, ..., x_{t-1}, z_{t-1}) P(x_t | z_t^k = 1, x_1, ..., x_{t-1}, z_{t-1})$
 - By the virtue of the structure

•
$$= \sum_{z_{t-1}} P(x_1, ..., x_{t-1}, z_{t-1}) P(z_t^k = 1 | z_{t-1}) P(x_t | z_t^k = 1)$$

•
$$= P(x_t|z_t^k = 1) \sum_{z_{t-1}} P(x_1, ..., x_{t-1}, z_{t-1}) P(z_t^k = 1|z_{t-1})$$

•
$$= b_{z_t^k, x_t} \sum_{z_{t-1}} P(x_1, ..., x_{t-1}, z_{t-1}) a_{z_{t-1}, z_t^k}$$

- Now, we see a repeating structure of terms
- $P(x_1, ..., x_t, z_t^k = 1) = \alpha_t^k = b_{k,x_t} \sum_i \alpha_{t-1}^i a_{i,k}$



$$P(z_1) \sim Mult(\pi_1, ..., \pi_k)$$

Transition probabilities

$$P(z_t|z_{t-1}^i = 1) \sim Mult(a_{i,1}, ..., a_{i,k})$$

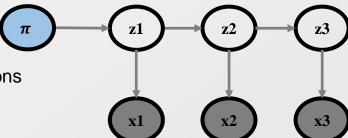
Or, $P(z_t^j = 1|z_{t-1}^i = 1) = a_{i,j}$

$$P(x_t|z_t^t=1) \sim Mult(b_{i,1},...,b_{i,m}) \sim f(x_t|\theta_i)$$

Or,
$$P(x_t^j = 1 | z_t^i = 1) = b_{i,j}$$



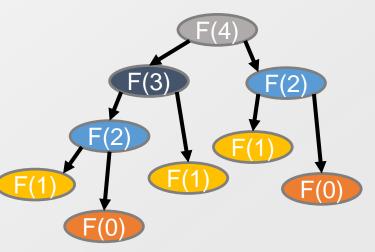
M_i observations for i-th sequence



Detour: Dynamic Programming



- Dynamic programming:
 - A general algorithm design technique for solving problems defined by or formulated as recurrences with overlapping subinstances
 - In this context, Programming == Planning
- Main storyline
 - Setting up a recurrence
 - Relating a solution of a larger instance to solutions of some smaller instances
 - Solve small instances once
 - Record solutions in a table
 - Extract a solution of a larger instance from the table

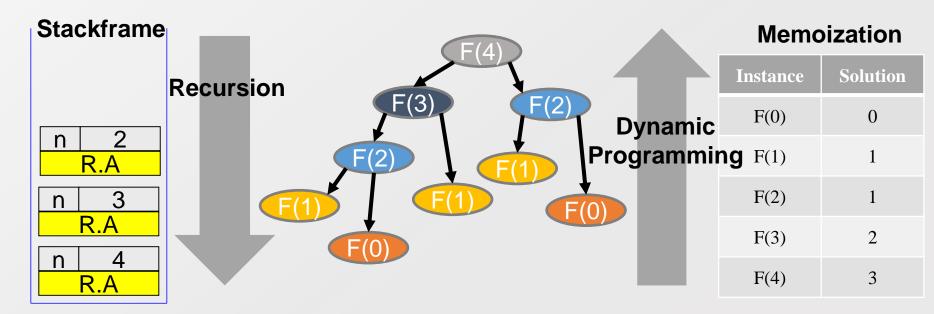


Instance	Solution
F(0)	0
F(1)	1
F(2)	1
F(3)	2
F(4)	?

Detour: Memoization



- Key technique of dynamic programming
 - Simply put
 - Storing the results of previous function calls to reuse the results again in the future
 - More philosophical sense
 - Bottom-up approach for problem-solving
 - Recursion: Top-down of divide and conquer
 - Dynamic programming: Bottom-up of storing and building



Forward Probability Calculation



- Need to know α_t^k
 - Time X States
 - When we know α_t^k with X, then we know the value of P(X)
 - Answering the evaluation question without Z
- ForwardAlgorithm
 - Initialize

$$\bullet \quad \alpha_1^k = b_{k,x_1} \pi_k$$

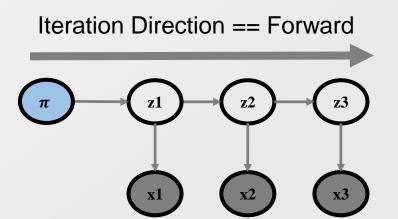
Iterate until time T

•
$$\alpha_t^k = b_{k,x_t} \sum_i \alpha_{t-1}^i a_{i,k}$$

- Return $\sum_i \alpha_T^i$
- Proof of correctness

•
$$\sum_{i} \alpha_{T}^{i} = \sum_{i} P(x_{1}, \dots, x_{T}, z_{T}^{i} = 1) = P(x_{1}, \dots, x_{t})$$

- Where to use the memoization table?
 - α_t^k
- Limitation of the forward probability
 - Only takes the input sequence of X before time t
 - $P(x_1, ..., x_t, z_t^k = 1) = \alpha_t^k \text{ and } t \neq T$
 - Need to see a probability distribution of a latent variable at time t given the whole X
 - Recall the Bayes ball algorithm



Backward Probability Calculation



- We need $P(z_t^k = 1|X)$ instead of $P(x_1, ..., x_t, z_t^k = 1)$
- Let's derive from the joint probability

•
$$P(z_t^k = 1, X) = P(x_1, ..., x_t, z_t^k = 1, x_{t+1}, ..., x_T)$$

• =
$$P(x_1, ..., x_t, z_t^k = 1)P(x_{t+1}, ..., x_T | x_1, ..., x_t, z_t^k = 1)$$

By the virtue of the structure

•
$$= P(x_1, ..., x_t, z_t^k = 1) P(x_{t+1}, ..., x_T | z_t^k = 1)$$

- We already handled $P(x_1, ..., x_t, z_t^k = 1)$
 - $P(x_1, ..., x_t, z_t^k = 1) = \alpha_t^k$
- So, we need to compute $P(x_{t+1}, ..., x_T | z_t^k = 1)$
 - $P(x_{t+1},...,x_T|z_t^k=1)=\beta_t^k$

•
$$P(x_{t+1}, ..., x_T | z_t^k = 1)$$

• =
$$\sum_{z_{t+1}} P(z_{t+1}, x_{t+1}, ..., x_T | z_t^k = 1)$$

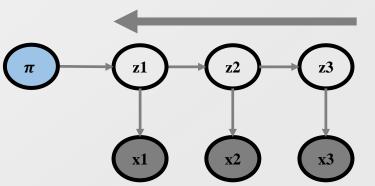
•
$$= \sum_{i} P(z_{t+1}^{i} = 1 | z_{t}^{k} = 1) P(x_{t+1} | z_{t+1}^{i} = 1, z_{t}^{k} = 1) P(x_{t+2}, ..., x_{T} | x_{t+1}, z_{t+1}^{i} = 1, z_{t}^{k} = 1)$$

•
$$= \sum_{i} P(z_{t+1}^{i} = 1 | z_{t}^{k} = 1) P(x_{t+1} | z_{t+1}^{i} = 1) P(x_{t+2}, ..., x_{T} | z_{t+1}^{i} = 1)$$

- = $\sum_{i} a_{k,i} b_{i,x_t} \beta_{t+1}^{l}$
- Again, recursive structure. How to calculate this efficiently?

•
$$P(z_t^k = 1, X) = \alpha_t^k \beta_t^k = (b_{k, x_t} \sum_i \alpha_{t-1}^i a_{i, k}) \times (\sum_i a_{k, i} b_{i, x_t} \beta_{t+1}^i)$$

Iteration Direction == Backward

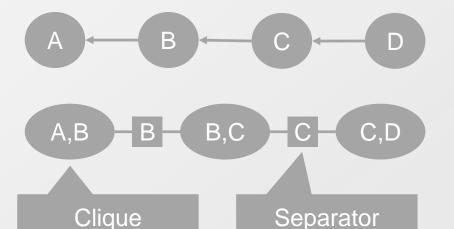


Detour: Potential Functions



- P(A,B,C,D) = P(A|B)P(B|C)P(C|D)P(D)
- Let's define a potential function
 - Potential function:

 a function which is not a probability
 function yet, but once normalized it can
 be a probability distribution function
 - Potential function on nodes
 - $\psi(a,b), \psi(b,c), \psi(c,d)$
 - Potential function on links
 - $\phi(b), \phi(c)$
- How to setup the function?
 - $P(A, B, C, D) = P(U) = \frac{\prod_{N} \psi(N)}{\prod_{L} \phi(L)} = \frac{\psi(a, b)\psi(b, c)\psi(c, d)}{\phi(b)\phi(c)}$
 - $\psi(a,b) = P(A|B), \psi(b,c) = P(B|C), \psi(c,d) = P(C|D)P(D)$
 - $\phi(b) = 1, \phi(c) = 1$
 - $P(A, B, C, D) = P(U) = \frac{\prod_{N} \psi(N)}{\prod_{L} \phi(L)} = \frac{\psi^*(a,b)\psi^*(b,c)\psi^*(c,d)}{\phi^*(b)\phi^*(c)}$
 - $\psi^*(a,b) = P(A,B), \psi^*(b,c) = P(B,C), \psi^*(c,d) = P(C,D)$
 - $\phi^*(b) = P(B), \phi^*(c) = P(C)$



Marginalization is also applicable:

$$\psi(w) = \sum_{v-w} \psi(v)$$

Constructing a potential of a subset (w) of all variables (v)

Detour: Absorption in Clique Graph



- Only applicable to the tree structure of clique graph
- Let's assume

•
$$P(B) = \sum_{A} \psi(A, B)$$

•
$$P(B) = \sum_{C} \psi(B, C)$$

- $P(B) = \phi(B)$
- How to find out the ψ s and the ϕ s?
 - When the ψ s change by the observations: P(A,B) \rightarrow P(A=1,B)
 - A single ψ change can result in the change of multiple ψ s
 - The effect of the observation propagates through the clique graph
 - Belief propagation!
- How to propagate the belief?
 - Absorption (update) rule
 - Assume $\psi^*(A,B),\psi(B,C)$, and $\phi(B)$
 - Define the update rule for separators

- Define the update rule for cliques
 - $\psi^*(B,C)=\psi(B,C)\frac{\phi^*(B)}{\phi(B)}$

Why does this work?

$$\sum_{C} \psi^*(B,C) = \sum_{C} \psi(B,C) \frac{\phi^*(B)}{\phi(B)}$$

$$= \frac{\phi^*(B)}{\phi(B)} \sum_{C} \psi(B,C) = \frac{\phi^*(B)}{\phi(B)} \phi(B) = \sum_{A} \psi^*(A,B)$$

Guarantees the local consistency

→ Global consistency after iterations

Detour: Simple Example of Belief Propagation



- Initialized the potential functions
 - $\psi(a,b) = P(a|b), \psi(b,c) = P(b|c)P(c)$
 - $\phi(b) = 1$
- Example 1. P(b) = ?
 - $\phi^*(b) = \sum_a \psi(a, b) = 1$
 - $\psi^*(b,c) = \psi(b,c) \frac{\phi^*(b)}{\phi(b)} = P(b|c)P(c) = P(b,c)$
 - $\phi^{**}(b) = \sum_{c} \psi(b, c) = \sum_{c} P(b, c) = P(b)$
 - $\psi^*(a,b) = \psi(a,b) \frac{\phi^{**}(b)}{\phi^*(b)} = \frac{P(a|b)P(b)}{1} = P(a,b)$
 - $\phi^{***}(b) = \sum_{a} \psi^{*}(a, b) = P(b)$
- Example 2. P(b|a = 1, c = 1) = ?
 - $\phi^*(b) = \sum_a \psi(a, b) \delta(a = 1) = P(a = 1|b)$
 - $\psi^*(b,c) = \psi(b,c) \frac{\phi^*(b)}{\phi(b)} = P(b|c=1)P(c=1) \frac{P(a=1|b)}{1}$
 - $\phi^{**}(b) = \sum_{c} \psi(b,c) \, \delta(c=1) = P(b|c=1)P(c=1)P(a=1|b)$
 - $\psi^*(a,b) = \psi(a,b) \frac{\phi^{**}(b)}{\phi^*(b)} = P(a=1|b) \frac{P(b|c=1)P(c=1)P(a=1|b)}{P(a=1|b)} = P(b|c=1)P(c=1)P(a=1|b)$
 - $\phi^{***}(b) = \sum_{a} \psi^{*}(a,b) \, \delta(a=1) = P(b|c=1)P(c=1)P(a=1|b)$



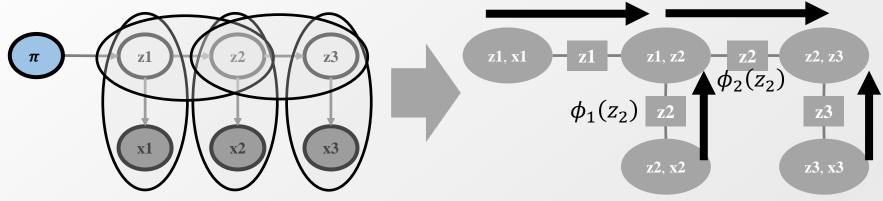
Bayesian Network

A,B B B,C

Clique Graph

Message Passing and Forward-Backward





•
$$P(z_1, z_2, z_3, x_1, x_2, x_3) = \frac{\prod_N \psi(N)}{\prod_L \phi(L)}$$

•
$$= \frac{\psi(z_1, x_1)\psi(z_1, z_2)\psi(z_2, z_3)\psi(z_2, x_2)\psi(z_3, x_3)}{\phi(z_1)\phi_1(z_2)\phi_2(z_2)\phi(z_3)}$$

- - $= P(z_1)P(x_1|z_1)P(z_2|z_1)P(x_2|z_2)P(z_3|z_2)P(x_3|z_3)$

$$P(z_1, z_2, z_3, x_1, x_2, x_3)$$
• = $P(z_1)P(x_1|z_1)P(z_2|z_1)P(x_2|z_2)P(z_2|z_2)P(x_2|z_2)$

Define the update rule for separators

$$\phi^*(B) = \sum_A \psi^*(A, B)$$

Define the update rule for cliques $\psi^*(B,C)=\psi(B,C)\frac{\phi^*(B)}{\phi(B)}$

 $P(z_1) \sim Mult(\pi_1, ..., \pi_k)$

Transition probabilities

$$P(z_t|z_{t-1}^i = 1) \sim Mult(a_{i,1}, ..., a_{i,k})$$

Or,
$$P(z_t^j = 1 | z_{t-1}^i = 1) = a_{i,j}$$

Emission probabilities

$$P(x_t|z_t^i = 1) \sim Mult(b_{i,1}, ..., b_{i,m})$$

Or, $P(x_t^j = 1|z_t^i = 1) = b_{i,i}$

- Initialized the potential functions
 - $\psi(z_1, x_1) = P(z_1)P(x_1|z_1), \psi(z_1, z_2) = P(z_2|z_1), \psi(z_2, z_3) = P(z_3|z_2), \psi(z_2, x_2) = P(x_2|z_2), \psi(z_3, x_3) = P(x_3|z_3)$
 - $\phi(z_1) = \phi_1(z_2) = \phi_2(z_2) = \phi(z_3) = 1$
- Start absorbing and updating
 - $\phi_2^*(z_2) = \sum_{z_1} \psi^*(z_1, z_2) = \sum_{z_1} \psi(z_1, z_2) \phi^*(z_1) \phi_1^*(z_2) = \sum_{z_1} \psi(z_1, z_2) \phi^*(z_1) \phi_1^*(z_2)$
 - Because x_2 is already observed, so the summation on x_2 does not happen, use just fixed x_2
 - $= \sum_{z_1} P(z_2|z_1) \, \phi^*(z_1) P(x_2|z_2) = P(x_2|z_2) \sum_{z_1} P(z_2|z_1) \, \phi^*(z_1) = b_{idx(z_2),x_2} \sum_{i \in z_1} \alpha_{i,z_2}^i$
 - Same as the forward probability calculation
 - This is the upward process, then the downward process is same as the backward probability calculation

Viterbi Decoding

- $P(z_t^k = 1, X) = \alpha_t^k \beta_t^k = (b_{k, x_t} \sum_i \alpha_{t-1}^i a_{i, k}) \times (\sum_i a_{k, i} b_{i, x_t} \beta_{t+1}^i)$
 - This dictates the most probable assignment to a single latent variable, z_t, given the whole observed sequence, X
 - $k_t^* = argmax_k P(z_t^k = 1|X) = argmax_k P(z_t^k = 1, X) = argmax_k \alpha_t^k \beta_t^k$
 - What if we want to have the most probable assignment of Z given X?
 - Exactly the decoding question
 - Different from the most probable assignment of a single latent variable
- Viterbi decoding
 - $k^* = argmax_k P(z^k = 1|X) = argmax_k P(z^k = 1, X)$
 - Need to model the sequence of Z.
 - Let's use the forward approach (Bottom-up)
 - $V_t^k = max_{z_1...z_{t-1}} P(x_1, ..., x_{t-1}, z_1, ..., z_{t-1}, x_t, z_t^k = 1)$
 - Most probable sequence of latent states until t-1 and fixing the state k at time t

•
$$= max_{z_1...z_{t-1}} P(x_t, z_t^k = 1 | x_1, ..., x_{t-1}, z_1, ..., z_{t-1}) P(x_1, ..., x_{t-1}, z_1, ..., z_{t-1})$$

•
$$= max_{z_1...z_{t-1}} P(x_t, z_t^k = 1 | z_{t-1}) P(x_1, ..., x_{t-2}, z_1, ..., z_{t-2}, x_{t-1}, z_{t-1})$$

•
$$= max_{z_{t-1}} P(x_t, z_t^k = 1 | z_{t-1}) max_{z_1...z_{t-2}} P(x_1, ..., x_{t-2}, z_1, ..., z_{t-2}, x_{t-1}, z_{t-1})$$

•
$$= \max_{i \in z_{t-1}} P(x_t, z_t^k = 1 | z_{t-1}^i = 1) V_{t-1}^i = \max_{i \in z_{t-1}} P(x_t | z_t^k = 1) P(z_t^k = 1 | z_{t-1}^i = 1) V_{t-1}^i$$

•
$$= P(x_t|z_t^k = 1)max_{i \in z_{t-1}}P(z_t^k = 1|z_{t-1}^i = 1)V_{t-1}^i = b_{k,idx(x_t)}max_{i \in z_{t-1}}a_{i,k}V_{t-1}^i$$

Keep going until time T

$$P(z_1) \sim Mult(\pi_1, ..., \pi_k)$$

Transition probabilities

$$P(z_t|z_{t-1}^i = 1) \sim Mult(a_{i,1}, ..., a_{i,k})$$

Or, $P(z_t^j = 1|z_{t-1}^i = 1) = a_{i,j}$

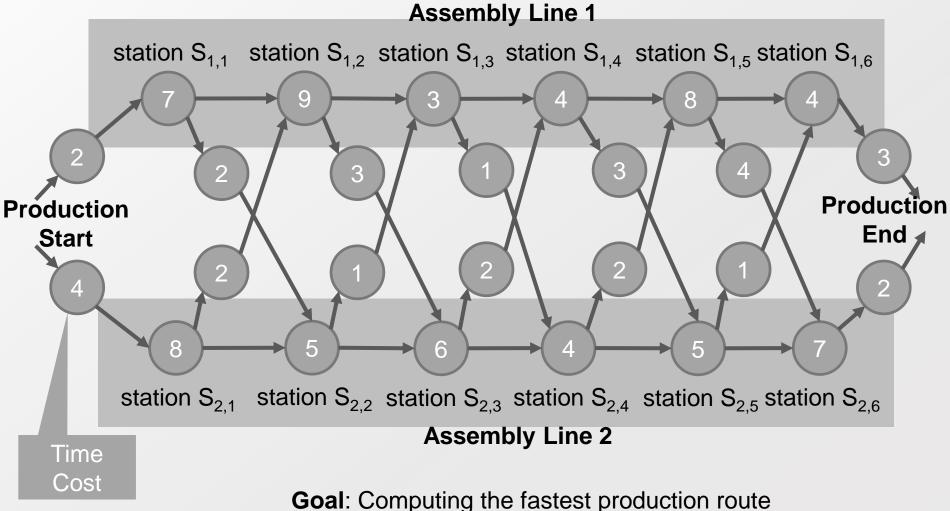
Emission probabilities

$$P(x_t|z_t^i = 1) \sim Mult(b_{i,1}, ..., b_{i,m})$$

Or, $P(x_t^j = 1|z_t^i = 1) = b_{i,i}$

Detour: Assembly Line Scheduling



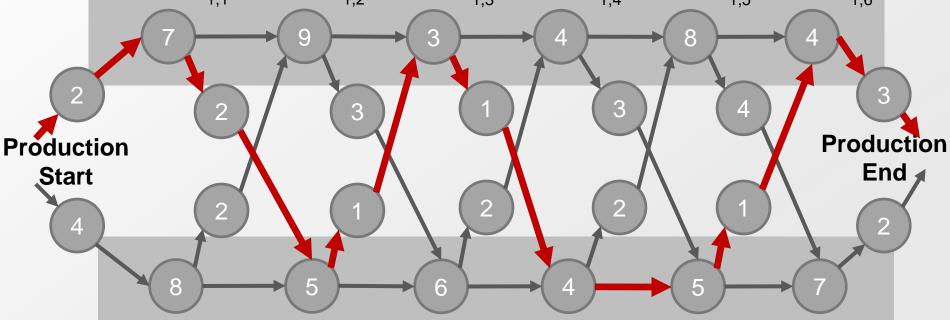


Detour: Tracing Assembly Line Scheduling in DP





station $S_{1,1}$ station $S_{1,2}$ station $S_{1,3}$ station $S_{1,4}$ station $S_{1,5}$ station $S_{1,6}$



station $S_{2,1}$ station $S_{2,2}$ station $S_{2,3}$ station $S_{2,4}$ station $S_{2,5}$ station $S_{2,6}$

Assembly Line 2

Used for retrace purpose

Time	1	2	3	4	5	6
L1	9	18	20	24	32	35
L2	12	16	22	25	30	37

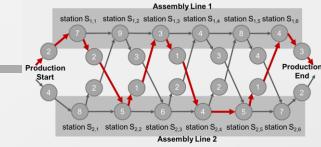
Trace	1	2	3	4	5	6
L1	S	1	2	1	1	2
L2	S	1	2	1	2	2

Viterbi Decoding Algorithm

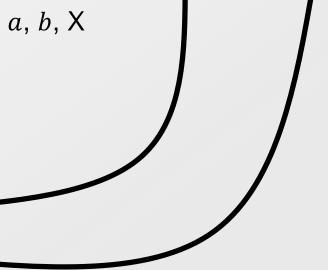
- Need to know V_t^k
 - Time X States
 - Store V_t^kTwo variables to store the trace and the probability up to time t. Two memoization tables
 - Answering the decoding question with π, a, b, X
- ViterbiDecodingAlgorithm
 - Initialize

•
$$V_1^k = b_{k,x_1}\pi_k$$

- Iterate until time T
 - $V_t^k = b_{k,idx(x_t)} max_{i \in Z_{t-1}} a_{i,k} V_{t-1}^l$
 - $trace_t^k = argmax_{i \in Z_{t-1}} a_{i,k} V_{t-1}^i$
- Return $P(X, Z^*) = max_k V_T^k$, $z_T^* = argmax_k V_T^k$, $z_{t-1}^* = trace_t^{z_t^*}$
- Technical difficulties in the implementation
 - Very frequent underflow problems.
 - Turn this into the log domain → from multiplication to summation







Learning Parameters with Only X



- Importance of π , a, b
 - HMM parameters
 - Forward algorithm (evaluation) and Viterbi algorithm (decoding) depends on knowing $\pi,\,a,\,b$
- However, knowing π , a, b assumes that we have observed X and Z
 - But, often Z is hard to observe. Need tagging, annotation, etc
 - Often the latent space is what we want to know, so we can't assume that we know Z
- If we don't know Z, we can assign the most probable Z to X
 - However, this is decoding problem, and this requires knowing π , a, b
- Most likely scenario in the real world
 - You have only X
 - You don't have Z, π , a, b, and you need to find out Z, π , a, b
- Strategy
 - Finding the optimized $\bar{\pi}$, \bar{a} , \bar{b} with X
 - Finding the most probable Z with X, $\bar{\pi}$, \bar{a} , \bar{b}
 - How to find the unknown parameter of the latent distribution without supervision?
- EM algorithm!
 - Iteratively optimizing $\overline{\pi}$, \overline{a} , \overline{b} and Z

Detour: EM Algorithm

$$\begin{split} l(\theta) &= \ln P(X|\theta) = \ln \left\{ \sum_{Z} q(Z) \frac{P(X,Z|\theta)}{q(Z)} \right\} \geq \sum_{Z} q(Z) \ln \frac{P(X,Z|\theta)}{q(Z)} = Q(\theta,q) \\ Q(\theta,q) &= E_{q(Z)} \ln P(X,Z|\theta) + H(q) \\ L(\theta,q) &= \ln P(X|\theta) - \sum_{Z} \{q(Z) \ln \frac{q(Z)}{P(Z|X,\theta)} \} \end{split}$$

- EM algorithm
 - Finds the maximum likelihood solutions for models with latent variables
 - $P(X|\theta) = \sum_{Z} P(X, Z|\theta) \rightarrow \ln P(X|\theta) = \ln \{\sum_{Z} P(X, Z|\theta)\}$
- EM algorithm
 - Initialize θ^0 to an arbitrary point
 - Loop until the likelihood converges
 - Expectation step
 - $q^{t+1}(z) = argmax_q Q(\theta^t, q) = argmax_q L(\theta^t, q) = argmin_q KL(q||P(Z|X, \theta^t))$
 - $\rightarrow q^t(z) = P(Z|X,\theta) \rightarrow \text{Assign Z by } P(Z|X,\theta)$
 - Maximization step
 - $\theta^{t+1} = argmax_{\theta}Q(\theta, q^{t+1}) = argmax_{\theta}L(\theta, q^{t+1})$
 - fixed Z means that there is no unobserved variables
 - → Same optimization of ordinary MLE

EM for HMM

$$\begin{split} &l(\theta) = \ln P(X|\theta) = \ln \left\{ \sum_{Z} q(Z) \frac{P(X,Z|\theta)}{q(Z)} \right\} \\ &\geq \sum_{Z} q(Z) \ln \frac{P(X,Z|\theta)}{q(Z)} = Q(\theta,q) \\ &Q(\theta,q) = E_{q(Z)} lnP(X,Z|\theta) + H(q) \\ &L(\theta,q) = lnP(X|\theta) - \sum_{Z} \{q(Z) \ln \frac{q(Z)}{P(Z|X,\theta)}\} \end{split}$$

Initial State probabilities

Transition probabilities

Emission probabilities

 $P(z_1) \sim Mult(\pi_1, ..., \pi_k)$

 $P(z_t|z_{t-1}^i = 1) \sim Mult(a_{i,1}, ..., a_{i,k})$ $Or, P(z_t^j = 1|z_{t-1}^i = 1) = a_{i,j}$

 $P(x_t|z_t^i = 1) \sim Mult(b_{i,1}, ..., b_{i,m})$ Or, $P(x_t^j = 1|z_t^i = 1) = b_{i,i}$

- EM algorithm for HMM
 - Initialize π^0, a^0, b^0 to an arbitrary point
 - Loop until the likelihood converges
 - Expectation step
 - $q^{t+1}(z) = P(Z|X, \pi^t, a^t, b^t) \rightarrow \text{Assign Z by } P(Z|X, \pi^t, a^t, b^t)$
 - Maximization step
 - π^{t+1} , a^{t+1} , $b^{t+1} = argmax_{\pi, a, b}Q(\pi, a, b, q^{t+1}) = argmax_{\pi, a, b}E_{q^{t+1}(z)}lnP(X, Z|\pi, a, b) + H(q)$
- Assign Z and optimize π , a, b alternatively
 - Coordinated optimization
 - How to optimize? Derivation of EM update formula from HMM?
 - $P(X, Z | \pi, a, b) = \pi_{z_1} \prod_{t=2}^{T} a_{z_{t-1}, z_t} \prod_{t=1}^{T} b_{z_t, x_t}$
 - $\ln P(X, Z | \pi, a, b) = \pi_{z_1} + \sum_{t=2}^{T} \ln a_{z_{t-1}, z_t} + \sum_{t=1}^{T} \ln b_{z_t, x_t}$
 - $E_{q^{t+1}(z)} lnP(X, Z | \pi, a, b) = \sum_{Z} q^{t+1}(z) lnP(X, Z | \pi, a, b)$ = $\sum_{Z} P(Z | X, \pi^{t}, a^{t}, b^{t}) lnP(X, Z | \pi, a, b)$

Derivation of EM Update Formula

- $Q(\pi, a, b, q^{t+1}) = E_{q^{t+1}(z)} ln P(X, Z | \pi, a, b)$
- = $\sum_{Z} q^{t+1}(z) \ln P(X, Z | \pi, a, b)$
- = $\sum_{Z} P(Z|X, \pi^t, a^t, b^t) \ln P(X, Z|\pi, a, b)$
 - $\ln P(X, Z | \pi, a, b)$
 - $= \ln \pi_{z_1} + \sum_{t=2}^{T} \ln a_{z_{t-1}, z_t} + \sum_{t=1}^{T} \ln b_{z_t, x_t}$
- = $\sum_{Z} P(Z|X, \pi^t, a^t, b^t) \{ \ln \pi_{z_1} + \sum_{t=2}^{T} \ln a_{z_{t-1}, z_t} + \sum_{t=1}^{T} \ln b_{z_t, x_t} \}$
- Need to optimize $Q(\pi, a, b, q^{t+1})$ by using π, a, b
 - Remember that π , a, b is actually probabilities. $\sum_i \pi_i = 1$
 - Since there are constraints on π , a, b and Q is smooth \rightarrow Lagrange method!
- $L(\pi, a, b, q^{t+1})$
- = $Q(\pi, a, b, q^{t+1}) \lambda_{\pi} (\sum_{i=1}^{K} \pi_i 1) \sum_{i=1}^{K} \lambda_{a_i} (\sum_{j=1}^{K} a_{i,j} 1) \sum_{i=1}^{K} \lambda_{b_i} (\sum_{j=1}^{M} b_{i,j} 1)$
- Now, a typical optimization with partial derivative
- $\frac{dL(\pi, a, b, q^{t+1})}{d\pi_i} = \frac{d}{d\pi_i} \sum_{Z} P(Z|X, \pi^t, a^t, b^t) \ln \pi_{z_1} \lambda_{\pi} \left(\sum_{i=1}^K \pi_i 1 \right) = 0$
 - Only the terms with $z_1 = i$ survives
- $\frac{\frac{d}{d\pi_{i}} \left\{ \sum_{Z} P(Z|X, \pi^{t}, a^{t}, b^{t}) \ln \pi_{Z_{1}} \lambda_{\pi} \left(\sum_{i=1}^{K} \pi_{i} 1 \right) \right\} = \frac{P\left(z_{1}^{i} = 1 \middle| X, \pi^{t}, a^{t}, b^{t} \right)}{\pi_{i}} \lambda_{\pi} = 0 \Rightarrow \pi_{i} = \frac{P\left(z_{1}^{i} = 1 \middle| X, \pi^{t}, a^{t}, b^{t} \right)}{\pi_{i}}$
- $\frac{d}{d\lambda_{\pi}}L(\pi, a, b, q^{t+1}) = -(\sum_{i=1}^{K} \pi_i 1) = 0 \rightarrow \sum_{i=1}^{K} \pi_i = 1$
- Together, $\pi^{t+1}_{i} = \frac{P(z_1^i = 1 | X, \pi^t, a^t, b^t)}{\sum_{j=1}^K P(z_1^j = 1 | X, \pi^t, a^t, b^t)}$
 - This is an update function for π_i at the M Step

$$a^{t+1}_{i,j} = \frac{\sum_{t=2}^{T} P(z_{t-1}^{i} = 1, z_{t}^{j} = 1 | X, \pi^{t}, a^{t}, b^{t})}{\sum_{t=2}^{T} P(z_{t-1}^{i} = 1 | X, \pi^{t}, a^{t}, b^{t})}$$

$$b^{t+1}_{i,j} = \frac{\sum_{t=1}^{T} P(z_{t}^{i} = 1 | X, \pi^{t}, a^{t}, b^{t}) \delta(idx(x_{t}) = j)}{\sum_{t=1}^{T} P(z_{t}^{i} = 1 | X, \pi^{t}, a^{t}, b^{t})}$$

KAIST

Baum Welch Algorithm



- Answer to the learning question of HMM
- Again, EM for HMM with more details
- EM algorithm for HMM, a.k.a. Baum-Welch, Forward-Backward...
 - Initialize π^0, a^0, b^0 to an arbitrary point
 - Loop until the likelihood converges
 - Expectation step
 - $q^{t+1}(z) = P(Z|X, \pi^t, a^t, b^t) \rightarrow \text{Assign Z by } P(Z|X, \pi^t, a^t, b^t)$
 - Maximization step

•
$$\pi^{t+1}_{i} = \frac{P(z_{1}^{i} = 1 | X, \pi^{t}, a^{t}, b^{t})}{\sum_{j=1}^{K} P(z_{1}^{j} = 1 | X, \pi^{t}, a^{t}, b^{t})}$$

•
$$a^{t+1}_{i,j} = \frac{\sum_{t=2}^{T} P(z_{t-1}^{i} = 1, z_{t}^{j} = 1 | X, \pi^{t}, a^{t}, b^{t})}{\sum_{t=2}^{T} P(z_{t-1}^{i} = 1 | X, \pi^{t}, a^{t}, b^{t})}$$

•
$$b^{t+1}_{i,j} = \frac{\sum_{t=1}^{T} P(z_t^i = 1 | X, \pi^t, a^t, b^t) \delta(idx(x_t) = j)}{\sum_{t=1}^{T} P(z_t^i = 1 | X, \pi^t, a^t, b^t)}$$

$$P(z_t^k = 1, X) = \alpha_t^k \beta_t^k$$

$$\alpha_t^k = b_{k,x_t} \sum_i \alpha_{t-1}^i \alpha_{i,k}$$

$$\beta_t^k = \sum_i \alpha_{k,i} b_{i,x_t} \beta_{t+1}^i$$

$$P(X) = \sum_i \alpha_T^i$$

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Further Readings



Bishop Chapter 13