

Chapter 8. Inferences on a Population Mean

8.1 Confidence Intervals

8.2 Hypothesis Testing

8.1 Confidence Intervals

8.1.1 Confidence Interval Construction

- Confidence Intervals
 - A *confidence interval* for an unknown parameter ϑ is an interval that contains a set of plausible values of the parameter.
 - It is associated with a *confidence level* $1 - \alpha$, which measures the probability that the confidence interval actually **contains the unknown parameter value**.
 - Confidence levels of 90%, 95%, and 99% are typically used.

8.1.1 Confidence Interval Construction

- Inferences on a Population Mean
 - Inference methods on a population mean based upon the *t-procedure* are appropriate for large sample sizes $n \geq 30$ and also for small sample sizes as long as the data can reasonably be taken to be approximately normally distributed.
 - Nonparametric techniques (Chapter 15) can be employed for small sample sizes with data that are clearly not normally distributed.

- Two-Sided t -Interval

- A confidence interval with confidence level $1 - \alpha$ for a population mean μ based upon a sample of n continuous data observations with a sample mean \bar{x} and a sample standard deviation s is

$$\left(\bar{x} - \frac{t_{\alpha/2, n-1} S}{\sqrt{n}}, \bar{x} + \frac{t_{\alpha/2, n-1} S}{\sqrt{n}} \right)$$

- The interval is known as a *two-sided t -interval*.

- Technically speaking, $\frac{\sqrt{n}(\bar{X}-\mu)}{S}$ has a t -distribution only when the random sample is from a Normal distribution.
- Nevertheless, **the central limit theorem ensures** that the distribution of \bar{X} is **approximately normal** for reasonably large sample sizes, and in such cases it is sensible to construct t -intervals **regardless of the actual distribution of the data** observations.

Example 14 : Metal Cylinder Production (p.340)

- Data : 60 metal cylinder diameters (page 271, Figure 6.5).

- Summary statistics:

$n = 60$ Median = 50.01 Max. = 50.36
 $\bar{x} = 49.999$ Upper quartile = 50.07 Min. = 49.74
 $s = 0.134$ Lower quartile = 49.91

- Critical points:

Sample size $n = 60$	
Confidence level 90%:	$t_{0.05,59} = 1.671$
Confidence level 95%:	$t_{0.025,59} = 2.001$
Confidence level 99%:	$t_{0.005,59} = 2.662$

- Confidence interval with confidence level 90%:

$$(49.999 - 1.671 \times \frac{0.134}{\sqrt{60}}, 49.999 + 1.671 \times \frac{0.134}{\sqrt{60}}) = (49.970, 50.028)$$
- Confidence interval with confidence level 95%:

$$(49.999 - 2.001 \times \frac{0.134}{\sqrt{60}}, 49.999 + 2.001 \times \frac{0.134}{\sqrt{60}}) = (49.964, 50.033)$$
- Confidence interval with confidence level 99%:

$$(49.999 - 2.662 \times \frac{0.134}{\sqrt{60}}, 49.999 + 2.662 \times \frac{0.134}{\sqrt{60}}) = (49.953, 50.045)$$

- Conclusion with confidence interval:

With over 99% certainty, the average cylinder diameter lies within 0.05 mm of 50.00mm, that is, within the interval (49.95, 50.05).

- Comment:

It is important to remember that this confidence interval is for the *mean* cylinder diameter, and not for the actual diameter of a randomly selected cylinder.

8.1.2 Effect of the Sample Size on Confidence Intervals

- Interval length(L)

$$L = 2 \times \frac{t_{\alpha/2, n-1} S}{\sqrt{n}}$$

- If a confidence interval with a length no large than L_0 is required, then the desired sample size n must satisfy

$$2 \times \frac{t_{\alpha/2, n-1} S}{\sqrt{n}} \leq L_0.$$

8.1.4 Simulation Experiment

- In practice, **an experimenter observes just one data set**, and it has a probability of 0.95 of providing a 95% confidence interval that does indeed straddle the true value μ .

8.1.5 One-Sided Confidence Intervals

- **One-Sided t -Interval:** One-sided confidence intervals with confidence levels $1-\alpha$ for a population mean μ based on a sample of n continuous data observations with a sample mean \bar{x} and a sample standard deviation s are

$$(-\infty, \bar{x} + \frac{t_{\alpha, n-1}S}{\sqrt{n}})$$

which provides an upper bound on the population mean μ ,
and

$$(\bar{x} - \frac{t_{\alpha, n-1}S}{\sqrt{n}}, \infty)$$

which provides a lower bound on the population mean μ .

Python codes

- `import numpy as np`
- `import pandas as pd`
- `from scipy import stats`
- `import statsmodels.stats.weightstats as sms`
- `data = pd.read_csv('E:/data/taxi.txt', sep='\t', index_col=0)`
 # The data is of tire life times in kilometers.
- `print(data)` # output → next sheet
- `print(data['BrandA'])` # output → next sheet
- `print(data.describe())` # produces basic statistics. → next sheet
- `ds1=sms.DescrStatsW(data['BrandA'])`
- `print("confidence interval=(%.4f,%.4f)" %ds1.tconfint_mean(0.05, 'two-sided'))`
 confidence interval=(31776.7759,42723.2241)

Python outputs

`print(data) →`

	BrandA	BrandB
Taxi		
1	34400	36700
2	45500	46800
3	36700	37700
4	32000	31100
5	48400	47800
6	32800	36400
7	38100	38900
8	30100	31500

`print(data['BrandA']) →`

Taxi

1	34400
2	45500
3	36700
4	32000
5	48400
6	32800
7	38100
8	30100

`print(data.describe()) →`

	BrandA	BrandB
count	8.000000	8.000000
mean	37250.000000	38362.500000
std	6546.754921	6181.062669
min	30100.000000	31100.000000
25%	32600.000000	35175.000000
50%	35550.000000	37200.000000
75%	39950.000000	40875.000000
max	48400.000000	47800.000000

8.1.6 z-Intervals

- Two-Sided z-Interval

If an experimenter wishes to construct a confidence interval for a population mean μ based on a sample of size n with a sample mean \bar{X} and using an assumed **known** value for the population standard deviation σ , then the appropriate confidence interval is

$$(\bar{X} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}}, \bar{X} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}})$$

which is known as a ***two-sided z-interval***.

- One-sided z-intervals are constructed analogously to the one-sided t-intervals with the z-quantile and σ replacing the t-quantile and s.

8.2 Hypothesis Testing

8.2.1 Hypotheses

- Hypothesis Tests of a Population Mean

- A *null hypothesis* H_0 for a population mean μ is a statement that designates possible values for the population mean.
- It is associated with an *alternative hypothesis* H_A , which is the “opposite” of the null hypothesis.
- A *two-sided* set of hypotheses is

$$H_0 : \mu = \mu_0 \quad \text{versus} \quad H_A : \mu \neq \mu_0$$

for a specified value μ_0 of μ .

- A *one-sided* set of hypotheses is either

$$H_0 : \mu \leq \mu_0 \quad \text{versus} \quad H_A : \mu > \mu_0$$

or

$$H_0 : \mu \geq \mu_0 \quad \text{versus} \quad H_A : \mu < \mu_0$$

Example 14 : Metal Cylinder Production

- The machine that produces metal cylinders is **set to make cylinders with a diameter of 50 mm.**
- The two-sided hypotheses of interest are

$$H_0 : \mu = 50 \quad \text{versus} \quad H_A : \mu \neq 50$$

where the null hypothesis states that the machine is calibrated correctly.

Example 48 : Car Fuel Efficiency

- A manufacturer claim : its cars achieve an average of **at least 35 miles per gallon** in highway driving.
- The one-sided hypotheses of interest are

$$H_0 : \mu \geq 35 \quad \text{versus} \quad H_A : \mu < 35$$

- The null hypothesis states that the manufacturer's claim regarding the fuel efficiency of its cars is correct.

8.2.2 Interpretation of p -values

- Types of error
 - Type I error: An error committed by rejecting the null hypothesis when it is true.
 - Type II error: An error committed by accepting the null hypothesis when it is false.
- Significance level
 - is specified as the upper bound of the probability of type I error.

- p -value of a test
 - Definition: The p -value of a test is the **probability of obtaining a given data set or worse when the null hypothesis is true**.
 - A data set can be used to measure the plausibility of null hypothesis H_0 through the construction of a p -value.
 - The smaller the p -value, the less plausible is the null hypothesis.

8.2.3 Calculation of p-values

Example 14 : Metal Cylinder Production

- For testing $H_0: \mu = \mu_0$ vs $H_A: \mu \neq \mu_0$
- The data set of metal cylinder diameters:
- $n = 60, \bar{x} = 49.99856, s = 0.1334$
- $\mu_0 = 50.0 \rightarrow t = \frac{49.99856 - 50.0}{0.1334/\sqrt{60}} = -0.0836$
- $p\text{-value} = 2 \times P(T \geq 0.0836)$ where $T \sim t_{59}$.
- $p\text{-value} = 2 \times 0.467 = 0.934$

P-value for two-sided t-test

- Consider testing

$$H_0: \mu = \mu_0 \text{ vs } H_A: \mu \neq \mu_0$$

- p - value = $2 \times P(T \geq |t|)$

where $T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$.

P-value for one-sided t-test

- Consider testing

$$H_0: \mu \leq \mu_0 \text{ vs } H_A: \mu > \mu_0$$

Then

$$p - \text{value} = P(T \geq t)$$

- Consider testing

$$H_0: \mu \geq \mu_0 \text{ vs } H_A: \mu < \mu_0$$

Then

$$p - \text{value} = P(T \leq t)$$

Making conclusions using p-values

- Rejection of the Null Hypothesis

If a p -value is smaller than the significance level, then the hypothesis H_0 is **rejected** in favor of the alternative hypothesis H_A .

- Acceptance of the Null Hypothesis

A p -value **larger than 0.10** is generally taken to indicate that the null hypothesis H_0 is a plausible statement. The null hypothesis H_0 is therefore **accepted**.

However, this does **not mean** that the null hypothesis H_0 has been **proven to be true**.

Python codes for one-sample tests concerning a mean

- `import numpy as np`
- `import pandas as pd`
- `import statsmodels.stats.weightstats as sms`
- `data = pd.read_csv("taxi.txt",sep='\t',index_col=0)`
- `dat=data/1000`
- `print(dat['BrandA'])`
- `print(dat.describe())`
- `dat_A= dat['BrandA']`
- `ds=sms.DescrStatsW(dat_A)`
- `print("One Sample Two-sided t-test")`
- `print("alternative hypothesis: true mean is not equal to 40")`
- `print("t, p-value, df: %.4f %.4f %.1f" %ds.ttest_mean(40, 'two-sided'))`
- `print("mean: %.4f" %np.mean(dat_A))`

Python output

		BrandA	BrandB
<i>Taxi</i>	Count	8.000000	8.000000
1 34.4	mean	37.250000	38.362500
2 45.5	std	6.546755	6.181063
3 36.7	min	30.100000	31.100000
4 32.0	25%	32.600000	35.175000
5 48.4	50%	35.550000	37.200000
6 32.8	75%	39.950000	40.875000
7 38.1	max	48.400000	47.800000
8 30.1			

[Output]

One Sample Two-sided t-test

alternative hypothesis: true mean is not equal to 40

t, p-value, df: -1.1881 0.2735 7.0

mean: 37.2500

8.2.4 Significance Levels

- Significance Level of a Hypothesis Test
 - A hypothesis test with a *significance level* of *size* α
rejects the null hypothesis H_0 if a p -value *smaller* than α is obtained
and
accepts the null hypothesis H_0 if a p -value *larger* than α is obtained.

Two-Sided Hypothesis Test for a Population Mean with sig. level α

- A size α test for the two-sided hypotheses

$$H_0 : \mu = \mu_0 \text{ vs } H_A : \mu \neq \mu_0$$

rejects the null hypothesis H_0 if the **test statistic** $|t|$ falls in the **rejection region**,

$$\{t: |t| > t_{\alpha/2, n-1}\}$$

and accepts the null hypothesis H_0 if the **test statistic** $|t|$ falls in the **acceptance region**,

$$A = \{t: |t| \leq t_{\alpha/2, n-1}\}$$

- Recall that $T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$

One-Sided Hypothesis Test for a Population Mean with sig. level α

- A size α test for the one-sided hypotheses

$$H_0 : \mu \leq \mu_0 \text{ vs } H_A : \mu > \mu_0$$

rejects the null hypothesis H_0 if the **test statistic t** falls in the **rejection region**,

$$\{t: t > t_{\alpha, n-1}\}$$

and accepts the null hypothesis H_0 if the **test statistic t** falls in the **acceptance region**,

$$A = \{t: t \leq t_{\alpha, n-1}\}$$

- A size α test for the one-sided hypotheses

$$H_0 : \mu \geq \mu_0 \text{ vs } H_A : \mu < \mu_0$$

rejects the null hypothesis H_0 if the **test statistic** t falls in the **rejection region**,

$$\{t: t < -t_{\alpha, n-1}\}$$

and accepts the null hypothesis H_0 if the **test statistic** t falls in the **acceptance region**,

$$A = \{t: t \geq -t_{\alpha, n-1}\}$$

- **Power of a hypothesis Test**

- The *power* of a hypothesis test

- power = $1 - P(\text{Type II error} \mid H_A)$

- which is the probability that the null hypothesis is rejected when it is false.

8.2.5 z -Tests

- Two-Sided z -test

The p -value for the two-sided hypothesis testing problem

$$H_0: \mu = \mu_0 \text{ versus } H_A: \mu \neq \mu_0$$

based upon a data set of size n from $N(\mu, \sigma^2)$ with σ known.

The p -value is given by

$$p\text{-value} = 2 \times P(Z > |z|)$$

$$\text{where } Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}.$$

- As for the two-sided test,
a size α test rejects the null hypothesis H_0 if the *test statistic* z falls in the *rejection region*,

$$\{z: |z| > z_{\alpha/2}\},$$

and accepts the null hypothesis H_0 if the *test statistic* z falls in the *acceptance region*,

$$A = \{z: |z| \leq z_{\alpha/2}\}.$$

- The only difference between t-test and z-test is that the t -statistic is used for the t-test while z -statistic is used for the z-test instead.

Computation of the power of a hypothesis test

- Consider testing $H_0: \mu = \mu_0$ vs $H_A: \mu \neq \mu_0$ with significance level α . Assume that we have a random sample of size n from $N(\mu, \sigma^2)$. For computational convenience, assume σ is known.
- Power of test when $\mu = \mu^* > \mu_0$, $\beta(\mu^*)$.
$$\begin{aligned}\beta(\mu^*) &= 1 - P_{\mu=\mu^*}(\text{Accept } H_0) \\ &= 1 - P_{\mu=\mu^*}(|Z| \leq z_{\alpha/2}) \\ &= 1 - P_{\mu=\mu^*}\left(\frac{|\bar{X} - \mu_0|}{\sigma/\sqrt{n}} \leq z_{\alpha/2}\right) = 1 - P_{\mu=\mu^*}\left(-z_{\alpha/2} \leq \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \leq z_{\alpha/2}\right)\end{aligned}$$

- (Continued)

$$\begin{aligned}
& 1 - P_{\mu=\mu^*}(-Z_{\alpha/2} \leq \frac{\bar{X}-\mu_0}{\sigma/\sqrt{n}} \leq Z_{\alpha/2}) \\
&= 1 - P_{\mu=\mu^*}(-Z_{\alpha/2} - \frac{\mu^*-\mu_0}{\sigma/\sqrt{n}} \leq \frac{\bar{X}-\mu^*}{\sigma/\sqrt{n}} \leq Z_{\alpha/2} - \frac{\mu^*-\mu_0}{\sigma/\sqrt{n}}) \\
&= 1 - \Phi\left(Z_{\alpha/2} - \frac{\mu^*-\mu_0}{\sigma/\sqrt{n}}\right) + \Phi\left(-Z_{\alpha/2} - \frac{\mu^*-\mu_0}{\sigma/\sqrt{n}}\right) \\
&= \beta(\mu^*)
\end{aligned}$$

Determination of sample size in hypotheses testing

- Find n for which $\beta(\mu^*) = \beta^*$ with $\mu^* > \mu_0$.

$$\beta(\mu^*) = 1 - \Phi\left(z_{\alpha/2} - \frac{\mu^* - \mu_0}{\sigma/\sqrt{n}}\right) + \Phi\left(-z_{\alpha/2} - \frac{\mu^* - \mu_0}{\sigma/\sqrt{n}}\right)$$

$$\approx 1 - \Phi\left(z_{\alpha/2} - \frac{\mu^* - \mu_0}{\sigma/\sqrt{n}}\right)$$

$$z_{\alpha/2} - \frac{\mu^* - \mu_0}{\sigma/\sqrt{n}} \approx z_{\beta^*}$$

$$\text{So, } \sqrt{n} \approx \sigma \frac{z_{\alpha/2} - z_{\beta^*}}{\mu^* - \mu_0}$$

Chapter summary

8.1 Confidence Intervals for Mean

t-intervals

z-intervals

8.2 Hypothesis Testing about Mean

p-value

significance level

acceptance region

power of test