# Identification of cell signaling pathways based on biochemical reaction kinetics repositories

Gustavo Estrela

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Instituto de Matemática e Estatística Centro de Toxinas, Resposta-imune e Sinalização Celular (CeTICS) Laboratório Especial de Ciclo Celular, Instituto Butantan

# Introduction

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Understanding the functioning of cell signaling is important in many biological areas.

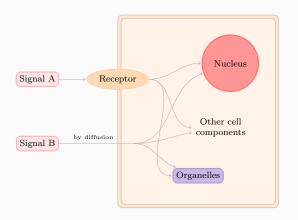
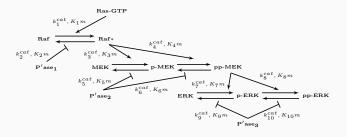


Figura 1: A general cell signaling mechanism.

#### **Cell Signaling Pathways**

A cell signaling network can be characterized by a sequence of chemical reactions that allows the presence of a signal to modify the state or behaviour of a cell.



**Figura 2:** An example of a signaling pathway.

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Using biochemical and enzymatic kinetics, we can write equations that represent the rate of change of concentration for a chemical species.

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Repeating this procedure for all reactions of a pathway allows us to derive a system of ordinary differential equations that can model the signaling pathway.

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As the input, a description of a biological experiment and a set of experimental measurements are given. A possible output to the problem is composed by:

- a model composed by a set of chemical reactions that are relevant for the biological experiment;
- information about the reaction rate constants of the model.

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Hence, it is desirable to construct a method that can systematically modify these models and choose the one that better represents the experiment.

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On her work, the problem of identification of cell signaling pathways is treated as a feature selection problem.

#### **Feature Selection Problem**

The feature selection problem is a combinatorial optimization problem:

Given a set of features S and a cost function c, find subset  $X \in \mathcal{P}(S)$ , with minimum cost c(X).

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#### Feature Selection for Identification of Signaling Pathways

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The methodology proposed by Wu defines the set of features as a set of chemical reactions that can be added to a starting model. This set of chemical reactions is fetched from KEGG and stored in a database of interactions.

#### Wu's Search Algorithm for Feature Selection

The search algorithm used by Wu is the Sequential Forward Selection (SFS).

#### Wu's Cost Function for Feature Selection

Wu defines the cost function as the minimum distance between experimental and simulated data.

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Wu defines the cost function as the minimum distance between experimental and simulated data. The minimum distance is found using a Simmulated Annealing that traverses the parameter space.

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- the database of interactions used could be more nearly complete;
- the search algorithm could also consider removing interactions;
- the cost function could implement a proper penalization of models;

#### What we Propose on this Project

We propose to create a methodology that uses a feature selection approach for identification of signaling pathways, tackling the difficulties encountered by Wu.

To get a more nearly complete database of interactions, we should fetch information from KEGG and other databases,

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To create new search algorithms,

To create new search algorithms, we intend to use more general algorithms that can also remove interactions.

To define new cost functions,

To define new cost functions, we intend to use Bayesian approaches of model selection that allow us to create estimates of probabilities such as p(M|D) or p(D|M).

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- Create search algorithms for the feature selection problem.
- Test the methodology on known signaling pathways.
- Apply the methodology on a real case.

# **Fundamental Concepts**

Kinetics Modeling of Chemical Reactions

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- second order interaction kinetics;
- Michaelis-Menten enzymatic kinetics.

### Kinetic Modeling of First Order Iteration

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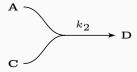
$$A \xrightarrow{k_1} B$$

has rate of:

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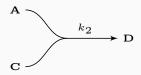
### **Kinetic Modeling of Second Order Iteration**

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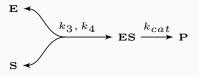


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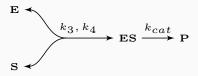
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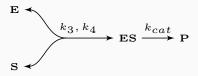
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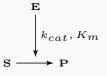
An enzymatic reaction:



Can be divided in two first order reactions plus a second order reaction. However, with the appropriate assumptions, it is possible to use a Michaelis-Menten simplification of this reaction.

#### Michaelis-Menten Kinetics

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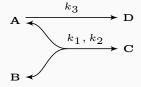
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$$\mathbf{S} \xrightarrow{\mathbf{E}} k_{cat}, K_m$$

and it has rate of:

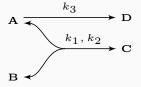
$$k_{cat} \frac{[E][S]}{K_M + [S]}.$$

Suppose we want to model the kinetics of A on these reactions:



This system can be divided in three reactions:

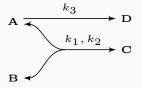
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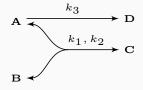
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#### Kinetics of a System of Reactions

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Then, the differential equation that models the concentration change of A is:

$$\frac{d[A]}{dt} = -k_1[A][B] + k_2[C] - k_3[D].$$

Bayesian Methods for Biochemical Model Selection

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- Approximate Bayesian Computation;
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For both methods, we resort to Metropolis-Hastings algorithm to generate samples of distributions.

With Metropolis-Hastings, we can generate a sample of a distribution  $p(\lambda)$  doing the following:

1. Choose some  $\lambda_0$  for which  $p(\lambda_0) > 0$ , and set t = 1;

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- 5. Increase *t* by one and repeat from Step 2 if not reached iteration limit.

# Model Selection

Ranking with Marginal Likelihood Estimation

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$$p(D|M,\theta) = p_{\mathcal{N}_{(\vec{0},\Sigma)}}(\phi(M,\theta) - D).$$

Where  $\phi(M, \theta)$  is the simulated observation.

#### Marginal Likelihood of Data

We can marginalize the likelihood to obtain:

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Calculating this integral is hard, therefore we resort to estimating another integral.

#### Power-posterior distributions

We define a power-posterior distribution as:

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$$p_{\beta=1}(\theta) = \frac{p(D,\theta|M)}{\int_{\Theta} p(D,\theta|M)d\theta} = \frac{p(\theta|D,M)p(D|M)}{p(D|M)} = p(\theta|D,M).$$

# The Thermodynamic Integral

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Using power-posteriors distributions, it is possible to show that

$$\ln p(D|M) = \int_0^1 \mathbb{E}_{p_{\beta}(\theta)}[\ln p(D|\theta, M)]d\beta.$$

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$$\sum_{t=0}^{T-1} (\beta_{t+1} - \beta_t) \frac{\mathbb{E}_{p_{\beta_{t+1}}(\theta)}[\log p(D|M, \theta)] + \mathbb{E}_{p_{\beta_t}(\theta)}[\log p(D|M, \theta)]}{2}$$

To produce the estimates of

$$\mathbb{E}_{p_{\beta_t}(\theta)}[\log p(D|M,\theta)]$$
 for  $t \in \{0,\ldots,T\}$ 

we need to produce samples of the power-posteriors  $p_{\beta_t}(\theta)$ .

The sampling of the power-posteriors are generated using Metropolis-Hastings algorithms in three steps.

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## **Sampling from Power-posteriors**

On the first step, the jump distribution has a diagonal covariance matrix. This matrix is updated according to the rate of acceptance of parameters.

- if the acceptance rate is high, then increase the variance of the jump;
- if the acceptance rate is low, then decrease the variance of the jump.

On the second and third step, the covariance matrix of the jump distributed is estimated with the current sample of the posterior.

## Sampling from the Power-posteriors

On the third step a Populational Monte Carlo Markov Chain is performed. This algorithm allows us to mix samples from different temperatures.

Ranking with Approximate Bayesian Computation

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- 4. Repeat until some iteration limit.

The result of the algorithm is a sample of the distribution

$$p(\theta, M|d(\phi(M, \theta), D) < \epsilon).$$

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We can use the accepted parameters of tolerance  $\epsilon$  and model M to estimate

$$p(M|d(\phi(M,\theta))<\epsilon).$$

**Experiments on Model Selection** 

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We tested the two methods of model ranking using the software:

- SigNetMS: an implementation of the Marginal Likelihood method created in this project.
- ABC-SysBio: an implementation of ABC-SMC.

We ran two experiments based on the same procedure:

• Create 4 candidate models.

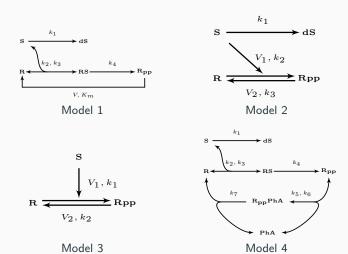
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- Add Gaussian noise to the simulations. Repeat two more times to generate three observations of the system.
- Neglect chosen parameter values and define prior distributions for every parameter.
- Rank the four models.

This experiment is originally from Vyshemirsky and Girolami (2007), in which they present results of Annealing Melting Integration.

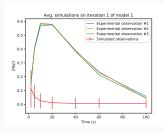


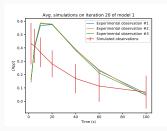
The model used to create the observations was Model 1.

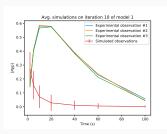
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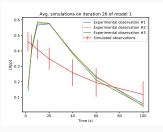
The priors distribution used for all parameters is Gamma(1,3).

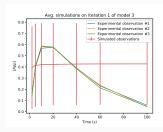
The ABC-SysBio software returned the following ranking of models:

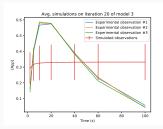


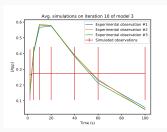


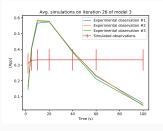












The ranking returned by SigNetMS on the first experiment is:

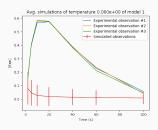
$$1 > 2 > 4 > 3$$
;

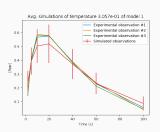
The ranking returned by SigNetMS on the first experiment is:

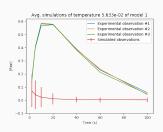
$$1 > 2 > 4 > 3$$
;

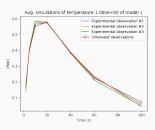
which is very similar to the ranking presented originally by Vyshemirsky and Girolami (2007):

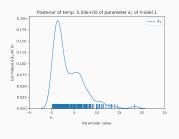
$$1 > 4 > 2 > 3$$
.

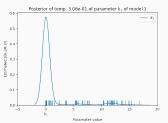


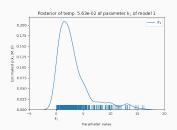


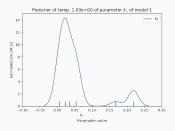




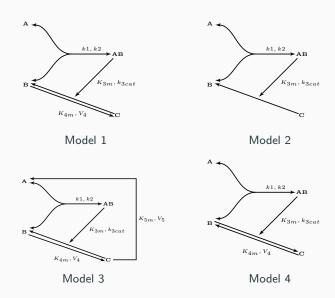








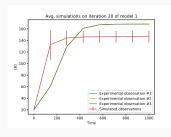
This experiment is very similar to the later and it was designed by us.

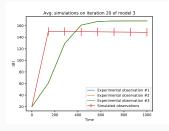


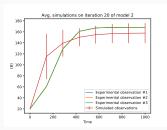
We used the following prior distributions for model parameters:

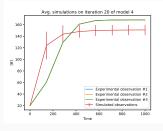
| Parameter         | Models     | Prior          |
|-------------------|------------|----------------|
| k1                | 1, 2, 3, 4 | Gamma(1, 0.01) |
| k2                | 1, 2, 3, 4 | Gamma(2, 0.5)  |
| k <sub>3cat</sub> | 1, 2, 3, 4 | Gamma(4,1)     |
| $K_{3m}$          | 1, 2, 3, 4 | Gamma(2, 1500) |
| $V_4$             | 1, 3, 4    | Gamma(2,1)     |
| $K_{4m}$          | 1, 3, 4    | Gamma(2, 100)  |
| $V_5$             | 3          | Gamma(2, 0.4)  |
| K <sub>5m</sub>   | 3          | Gamma(2, 100)  |

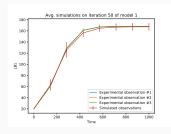
The ABC-SysBio software returned the following ranking of models:

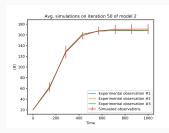


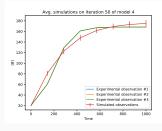


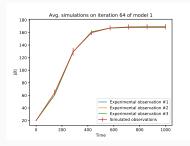


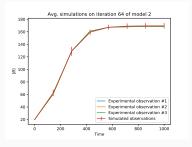


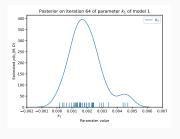


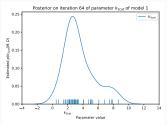


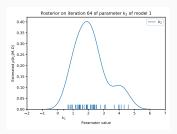


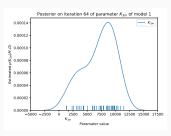












 ${\sf SigNetMS}\ returned\ the\ following\ ranking\ of\ models:$ 

a

# Next Steps