

STAT 547 HW1 - Zhiling GU

Q1

```
#####
##### Q1 #####
#####

# (a)
# Input a vector, return a fd object
nobs <- nrow(pinch)
x <- ((1:nobs)-1)*2
nrep <- ncol(pinch)

bspline4 <- function(y) {
  basis <- create.bspline.basis(rangeval =c(min(x), max(x)), nbasis = 15, norder = 4)
  fdobj <- smooth.basis (x, y, fdParobj=basis)
}

# Store the fd objects in fd_list
fd_list <- list()
# Color the curves gradiantly
col <- sapply(1:nrep, function(i){rgb (1,0.8*i/nrep, 0.8*i/nrep, alpha =0.7)})

par(mar = c(5, 4, 4, 4) + 0.3) # Leave space for z axis
plot(1, type="n", xlab="", ylab="", xlim=range(x), ylim=range(pinch))
for (i in 1:nrep){
  y <- pinch[,i]
  fd_i <- bspline4(y)
  fd_list <- append(fd_list, fd_i)
  lines(fd_i,lwd=2,col = col[i])
}

# (b)
y.mean <- apply(pinch, 1, mean)
# y.mean <- colMeans(pinch)
y.sd <- apply(pinch , 1, sd)
mean_col <- 'black'
sd_col <- 'green'

lines(x,y.mean, bg = mean_col, col = mean_col, cex =0.6, axes = FALSE, lwd=3)

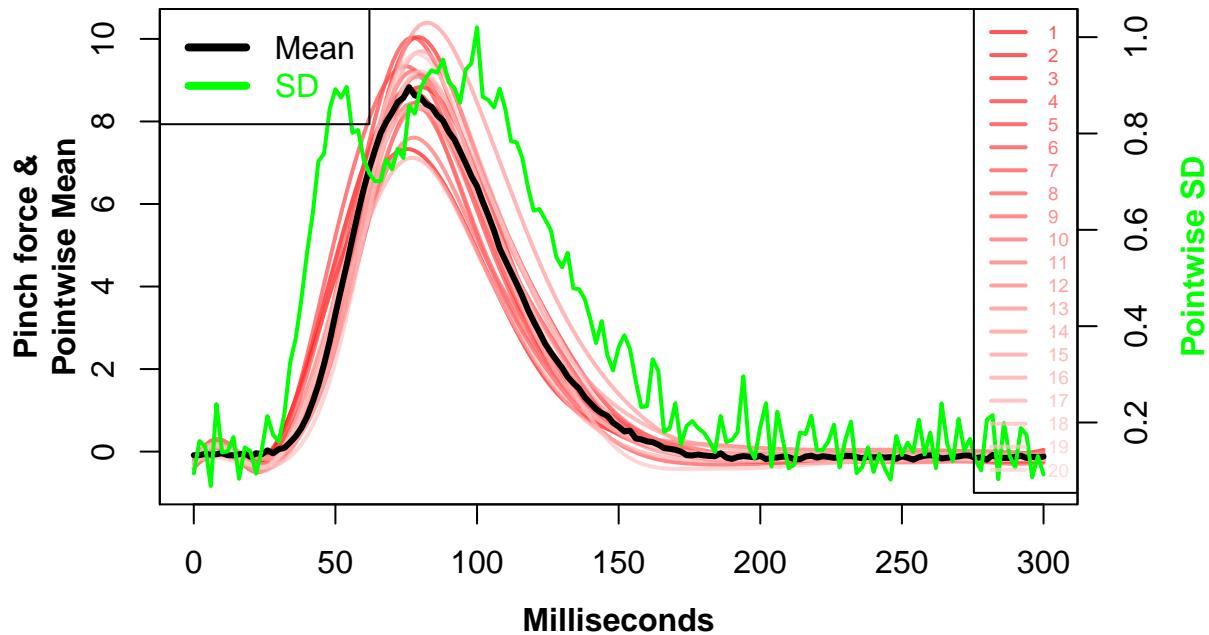
## Warning in plot.xy(xy.coords(x, y), type = type, ...): "axes" is not a
## graphical parameter
par(new = TRUE)
plot(x, y.sd, type = "l", bg = sd_col, col = sd_col, cex =0.5, axes = FALSE, xlab = '', ylab = '',xlim =
axis(side = 4, at = pretty(range(y.sd)))
axis(side = 2, at = pretty(range(y.mean)))
```

```

# axis(side = 1, at = pretty(range(x)))
mtext("Pinch force &", side=2, line=3, col = 'black', font =2 )
mtext("Pointwise Mean", side=2, line=2, col = mean_col, font =2)
mtext("Pointwise SD", side=4, line=2.5, col = sd_col, font =2)
mtext("Milliseconds", side=1, line=2.5, col = 'black', font =2)
legend('topright', legend = c(sapply ((1:ncol(pinch)), function(x){ paste("",x)})), col = c(col), text.col =c(mean_col, sd_col), lty =1, lwd=4)
legend("topleft", legend =c("Mean","SD"),
      col=c(mean_col, sd_col),text.col =c(mean_col, sd_col), lty =1, lwd=4)
title(main = "Pinch Force through time for 20 replications")

```

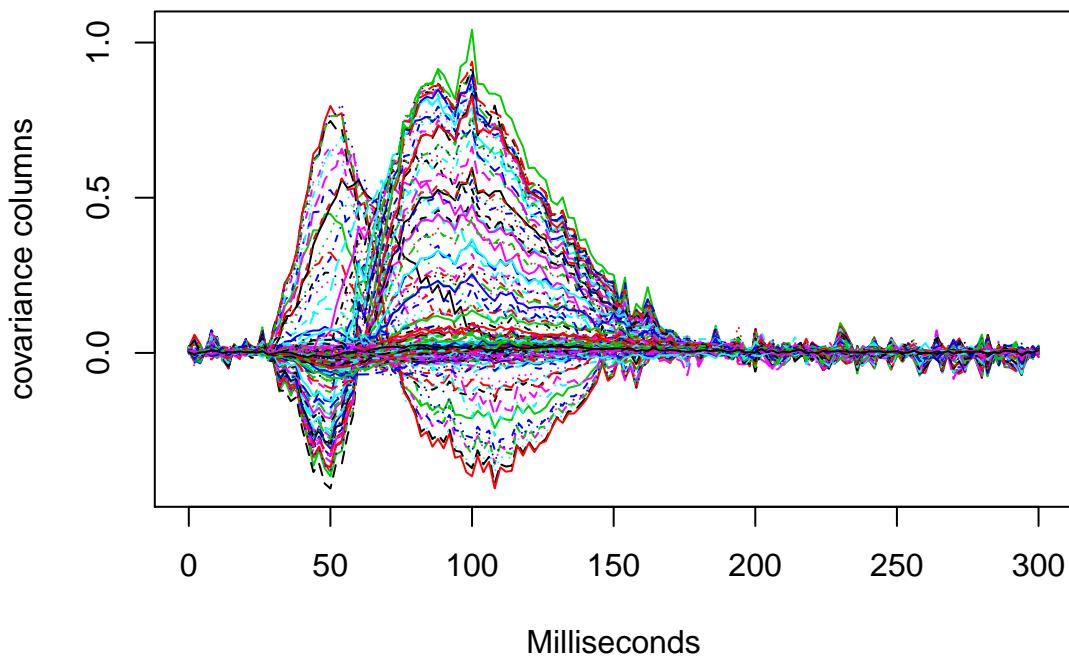
Pinch Force through time for 20 replications



```

# (c)
t.pinch <- t(pinch) # each row is an observation
df <- data.frame(t.pinch)
colnames(df) <- x # change the column units to milliseconds
Ghat <- cov(df)
matplot(x = x, y = Ghat, type='l', xlab="Milliseconds", ylab= "covariance columns")

```

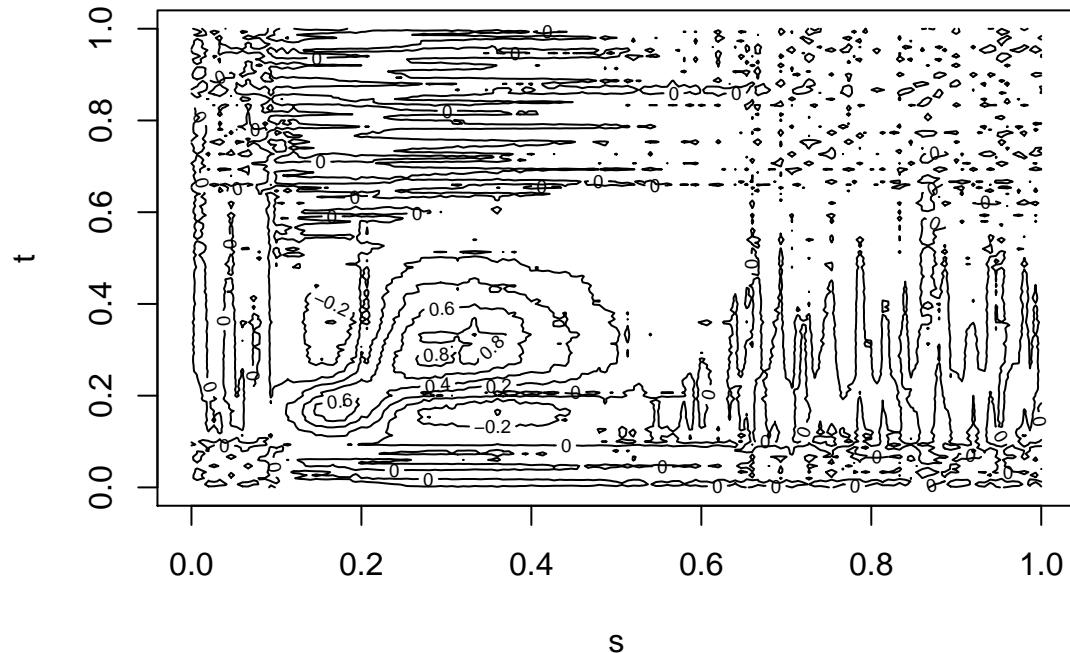


```

rho <- seq(0, 1, length.out=ncol(df))
persp3d(rho, rho, Ghat, col='white', xlab='s', ylab='t' , zlab = 'Sample Cov',main= "Perspective of Sample Covariance")
contour(rho, rho, Ghat, xlab='s', ylab='t', main= "Contour Plot of Sample Covariance")

```

Contour Plot of Sample Covariance



```

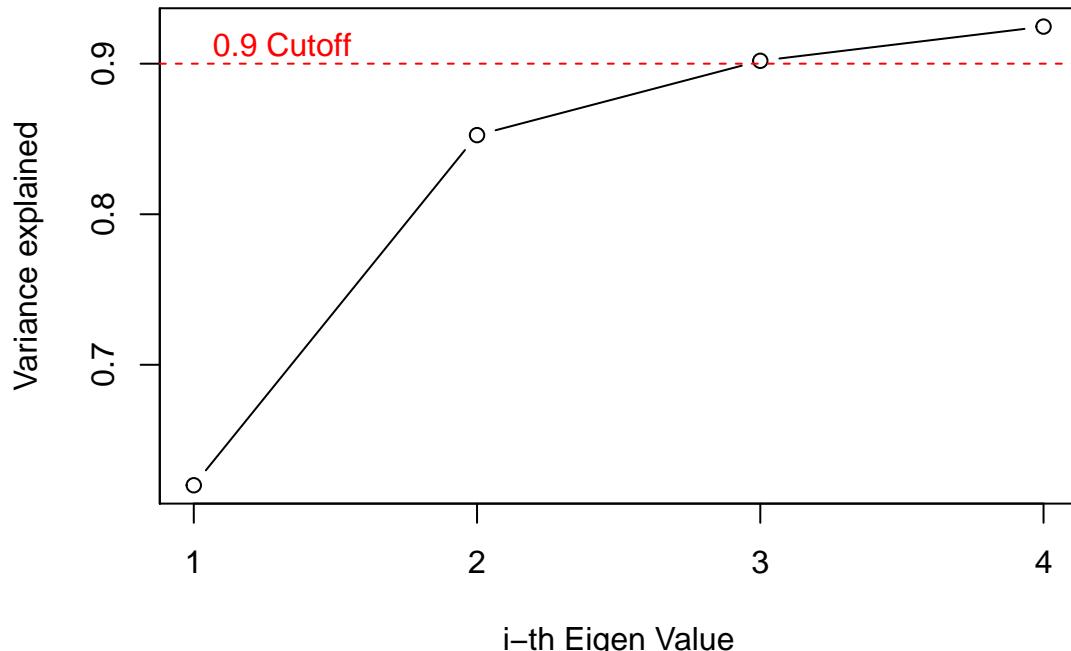
# (d)
Sig <- Ghat
eig <- eigen(Sig)
FVE <- cumsum(eig$values) / sum(eig$values)
plot(FVE[1:4], type='b', axes = FALSE, ann = FALSE) # scree plot

```

```

axis(side =1 , at = 1:4)
axis(side =2 , at = pretty(range(FVE)))
title (xlab = 'i-th Eigen Value')
title (ylab = 'Variance explained')
box()
abline(0.9,0,lwd =1, col = 'red', lty=2)
text(c(0.91,1),labels = "0.9 Cutoff", col = 'red', pos = 4)

```



```

for (i in 1:length(FVE)){
  if(FVE[i]>=0.9){
    flag = i
    break
  }
}
cat(paste('The', flag, "-th eigen value explains 90% of variations" ))

## The 3 -th eigen value explains 90% of variations

```

Q2

```

#####
##### Q2 #####
#####

d <- read.csv("./DataSets/Dow_companies_data.csv")
data = d[,2:31]

# (a)
xom <- d$XOM
head(xom)

## [1] 85.75 85.59 85.99 84.99 85.52 85.19

```

```

r <- sapply (xom, function(x){(x-xom[1])/xom[1]})*100
cat("Exxon-Mobil stock price increase in 2013 has increased by ", r[length(r)], " percent in return.")

## Exxon-Mobil stock price increase in 2013 has increased by 17.1895 percent in return.

# (b)
# Create a 252 by 30 matrix cr that contains the cumulative returns on all stocks in DJIA.
# Plot all the cumulative return functions in one plot. Add the mean and median functions in color.

cr <- apply(d[,2:ncol(d)], 2, function(x){(x - x[1])/x[1]}) *100
rownames(cr) <- 1:252 # d[1:252,1]
colnames(cr) <- 1:30 # colnames(d)[2:31]

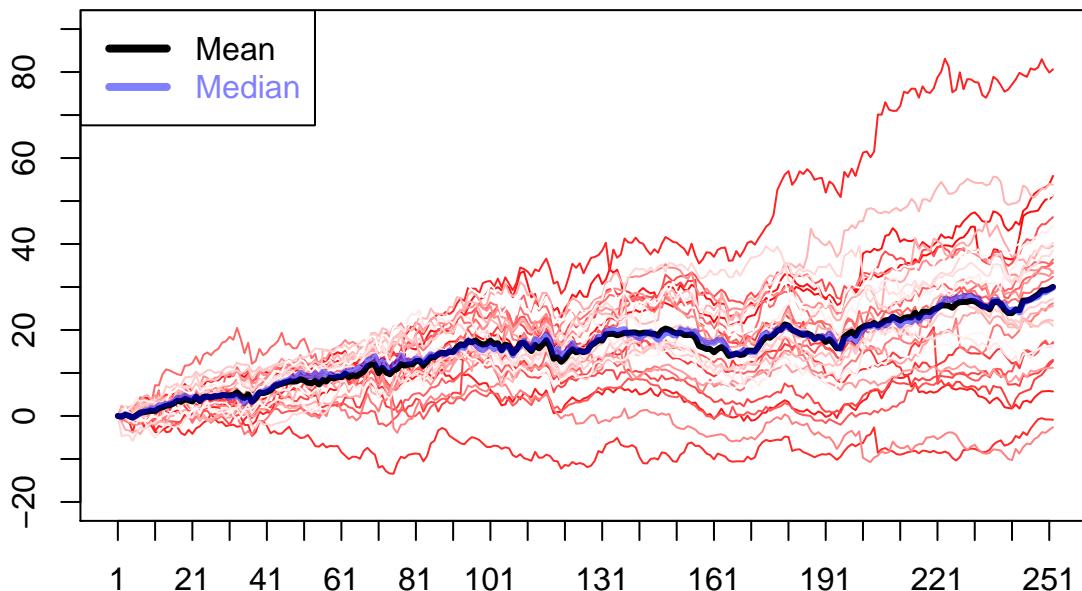
median <- apply (cr,1, median); mean_col = 'black'
mean <- apply(cr,1, mean); median_col = rgb(0,0,1,alpha =0.5)

plot(1, type="n", xlab="", ylab="", axes = FALSE, xlim = c(1,252), ylim =c(-20,90))
for (i in 1:ncol(cr)){
  lines(rownames(cr), cr[,i], col = rgb(1, 1*i/ncol(cr),1*i/ncol(cr)))
}
lines(rownames(cr), mean, col = mean_col,lwd =3)
lines(rownames(cr), median, col = median_col,lwd =3,lty =1)

axis(side =1, at = seq(1,252,by=10))
axis(side =2 , at = seq(-20,90,by=10))
box()
legend("topleft", legend =c("Mean","Median"),
       col=c(mean_col, median_col),text.col =c(mean_col, median_col), lty =1, lwd=4)
title(main = "Cumulative return plot for 30 stocks")

```

Cumulative return plot for 30 stocks



```

# (c)
prob <- c(0.9, 0.6, 0.5, 0.3)
prob_col <- sapply(prob, function(x) {rgb(1,x,x )})

```

```

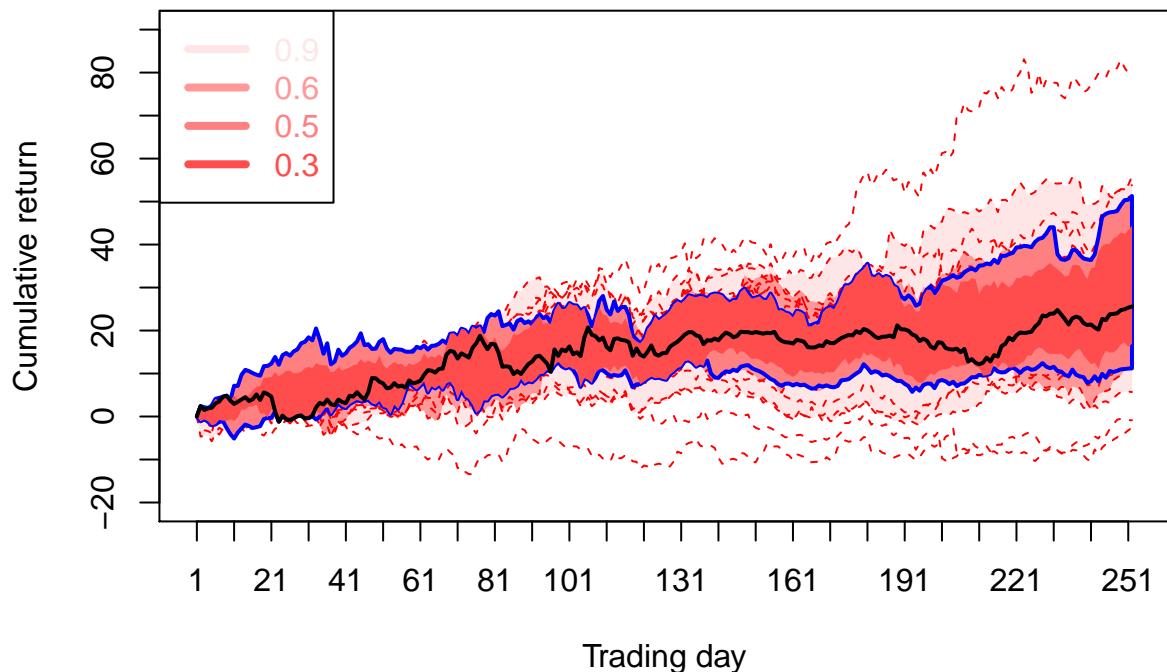
fbplot(fit = cr, ylim = c(-20,90), xlim =c(1,252),xlab="Trading day", ylab = "Cumulative return",prob=prob_col)

## $depth
##      1       2       3       4       5       6       7
## 0.4152915 0.3129926 0.2976852 0.1558498 0.1287926 0.4631477 0.4010787
##      8       9      10      11      12      13      14
## 0.3526204 0.3906062 0.2521096 0.4612320 0.4046730 0.4735837 0.3979771
##     15      16      17      18      19      20      21
## 0.2310368 0.3479132 0.4568167 0.4265850 0.4872856 0.3478768 0.3127919
##     22      23      24      25      26      27      28
## 0.4678366 0.4624726 0.4250160 0.3453772 0.4552842 0.4319308 0.4645525
##     29      30
## 0.4375319 0.3458881
##
## $outpoint
## [1] 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23
## [24] 24 25 26 27 28 29 30
##
## $medcurve
## 19
## 19

axis(side =1, at = seq(1,252, by =10))
axis(side =2 , at = seq(-20,90,by=10))
title(main = "Functional Boxplot for 30 stocks")
box()
legend("topleft", legend =c(0.9, 0.6, 0.5, 0.3),
       col=prob_col,text.col = prob_col, lty =1, lwd=4)

```

Functional Boxplot for 30 stocks



(Q3) : W.L.O.G Let $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d$ be the eigen values of G , G be the covariance matrix of r.v. X . Let u_1, \dots, u_d be the corresponding eigen vectors. Also let u_1, \dots, u_d be normalized vectors.

1° Recall eigen-decomposition:

since $G u_i = \lambda_i u_i$, $i=1, \dots, d$, we have

$$G(u_1 \dots u_d) = (u_1 u_2 \dots u_d) \begin{pmatrix} \lambda_1 & 0 \\ 0 & \ddots & 0 \\ 0 & & \lambda_d \end{pmatrix}$$

$$\text{Denote } Q = (u_1 \dots u_d), \Lambda = \begin{pmatrix} \lambda_1 & 0 \\ 0 & \ddots & 0 \\ 0 & & \lambda_d \end{pmatrix}$$

$$GQ = Q\Lambda \Rightarrow G = Q\Lambda Q^{-1}$$

Lemma For any symmetric matrix, eigen vectors must be orthogonal

$$\lambda_i \langle u_i, u_j \rangle = \langle \lambda_i u_i, u_j \rangle = \langle Gu_i, u_j \rangle$$

$$\begin{aligned} &= \langle u_i, G^T u_j \rangle \stackrel{\text{Sym.}}{=} \langle u_i, Gu_j \rangle \\ &= \langle u_i, \lambda_j u_j \rangle = \lambda_j \langle u_i, u_j \rangle \end{aligned}$$

For $\lambda_i \neq \lambda_j$, equality must hold due to $\langle u_i, u_j \rangle = 0 \neq \#$

use eigen decomposition and the lemma above, we have

$$QQ^T = I = Q^T Q \quad (Q^T = Q^{-1})$$

$$GQ^T = G = Q\Lambda Q^T$$

$$\text{Note } \text{Var}(l_i^T X) = E(l_i^T (X - \mu)(X - \mu)^T l_i)$$

$$= l_i^T E((X - \mu)(X - \mu)^T) l_i$$

$$= l_i^T G l_i$$

$$= l_i^T Q \Lambda Q^T l_i$$

$$= (l_i^T u_1 \dots l_i^T u_d) \Lambda \begin{pmatrix} u_1^T l_i \\ \vdots \\ u_d^T l_i \end{pmatrix}$$

$$= \sum_{i=1}^d l_i^T u_i \lambda_i u_i^T l_i$$

Case I: when $l_i = u_i$, $\text{Var}(l_i^T X) = \underbrace{(l_i^T u_i)}_{=1} \lambda_i (u_i^T l_i) \overset{!}{=} \lambda_i \leq \lambda_i$

Case II: when $l_i = \sum_{i=1}^d c_i u_i$, $\text{Var}(l_i^T X) = \sum_{i=1}^d c_i^2 \lambda_i \leq \sum_{i=1}^d c_i^2 \lambda_1 = \lambda_1$

Therefore

$$v_1 = \underset{l_1}{\operatorname{argmax}} \text{Var}(l_1^T X) = u_1 \#$$

$$\|l_1\| = 1$$

$$2^\circ v_k = \underset{l_k}{\operatorname{argmax}} \text{Var}(l_k^T X)$$

$$\|l_k\| = 1$$

$$l_k^T l_j = 0, j = 1, \dots, k-1$$

Case I: when $l_k = u_i$, $\text{Var}(l_k^T X) = \lambda_i$, where $i \geq k \leftarrow \boxed{l_k^T l_j = 0, j = 1, \dots, k-1}$

Case II: when $l_k = \sum_{i=1}^d c_i u_i$, $\text{Var}(l_k^T X) = \sum_{i=1}^d c_i^2 \lambda_i = \sum_{i=k}^d c_i^2 \lambda_i \leq \sum_{i=k}^d c_i^2 \lambda_k = \lambda_k$
 $\boxed{\text{with } c_1 = \dots = c_{k-1} = 0}$

Therefore

$$v_k = \underset{\|l_k\|=1}{\operatorname{argmax}} \text{Var}(l_k^T X) = \lambda_k \#$$

$$\|l_k\| = 1$$

$$l_k^T l_j = 0, j = 1, \dots, k-1$$

$$\|l_k\| = 1$$

$$(Q4.) \quad a) \quad E(\xi_k) = v_k^T (E(X-\mu)) = 0 \quad \#$$

$$\begin{aligned} b) \quad \text{Var}(\xi_k) &= E(v_k^T (X-\mu)(X-\mu)^T v_k) \\ &= v_k^T E((X-\mu)(X-\mu)^T) v_k \\ &= v_k^T (G v_k) \\ &= v_k^T (\lambda_k v_k) = \lambda_k \quad \# \end{aligned}$$

$$c) \quad \text{Cov}(\xi_j, \xi_k) = E(v_j^T (X-\mu)(X-\mu)^T v_k)$$

$$\begin{aligned} &= v_j^T G v_k \\ &= v_j^T \lambda_k v_k \\ &= \lambda_k v_j^T v_k \\ &= \lambda_k \xi_{jk} \quad \# \quad \leftarrow v_j^T v_k = \begin{cases} 0 & j \neq k \\ 1 & j=k \end{cases} \text{ by orthonormal.} \end{aligned}$$

$$d) \quad \text{Corr}(X_j, \xi_k) = \text{Cov}(X_j, \xi_k) / \sqrt{\sigma_{jj}} \sqrt{\lambda_k}$$

$$= \text{Cov}(X_j, v_k^T (X-\mu)) / \sqrt{\sigma_{jj}} \sqrt{\lambda_k}$$

$$= \text{Cov}(X_j, \sum_{i=1}^K v_{ki} (X_i - \mu_i)) / \sqrt{\sigma_{jj}} \sqrt{\lambda_k}$$

$$= \sum_{i=1}^K v_{ki} \text{Cov}(X_j, X_i) / \sqrt{\sigma_{jj}} \sqrt{\lambda_k}$$

$$= \sum_{i=1}^K v_{ki} \lambda_k \xi_{ij} / \sqrt{\sigma_{jj}} \sqrt{\lambda_k}$$

$$= v_{kj} \lambda_k / \sqrt{\sigma_{jj}} \sqrt{\lambda_k}$$

$$= v_{kj} \sqrt{\lambda_k} / \sqrt{\sigma_{jj}} \quad \#$$

Q5 (EX 10.1 K&R 2017) Prove prop 10.1.1.

$$a) \|ax\| = \sqrt{\langle ax, ax \rangle} = \sqrt{a^2 \langle x, x \rangle} = |a| \sqrt{\langle x, x \rangle} = |a| \|x\| \quad \#$$

$$b) |\langle x, y \rangle| \leq \sqrt{\langle x, x \rangle} \sqrt{\langle y, y \rangle}$$

(Pf by projecting x on \vec{y} and \vec{y}^\perp)

$$\text{Consider } z = x - \frac{\langle x, y \rangle}{\langle y, y \rangle} \cdot y$$

$$\begin{aligned} \langle z, y \rangle &= \langle x, y \rangle - \left\langle \frac{\langle x, y \rangle}{\langle y, y \rangle} \cdot y, y \right\rangle \\ &= \langle x, y \rangle - \frac{\langle x, y \rangle}{\langle y, y \rangle} \cdot \langle y, y \rangle \\ &= 0 \end{aligned} \quad \rightarrow \frac{\langle x, y \rangle}{\langle y, y \rangle} \text{ constant}$$

therefore z is perpendicular to y

Apply pythagorean theorem to $x = z + \frac{\langle x, y \rangle}{\langle y, y \rangle} \cdot y$

$$\text{we have } \|x\|^2 = \|z\|^2 + \left(\frac{\langle x, y \rangle}{\langle y, y \rangle} \right)^2 \|y\|^2$$

$$= \|z\|^2 + \frac{\langle x, y \rangle^2}{(\|y\|^2)^2} \|y\|^2$$

$$\geq 0 + \frac{\langle x, y \rangle^2}{\|y\|^2}$$

$$\begin{aligned} |\langle x, y \rangle| &\leq \sqrt{\|x\|^2} \sqrt{\|y\|^2} \\ &= \sqrt{\langle x, x \rangle} \sqrt{\langle y, y \rangle} \quad \# \end{aligned}$$

$$\begin{aligned} c) \|x+y\| &= \sqrt{\langle x+y, x+y \rangle} = \sqrt{\langle x, x \rangle + \langle y, y \rangle + 2 \langle x, y \rangle} \\ &\stackrel{b)}{\leq} \sqrt{\|x\|^2 + \|y\|^2 + 2 \|x\| \|y\|} \\ &= \|x\| + \|y\| \quad \# \end{aligned}$$

$$d) \quad ① \quad d(x, y) = \|x - y\| = \sqrt{\langle x - y, x - y \rangle} = \sqrt{\langle y - x, y - x \rangle}$$

$$= \|y - x\| = d(y - x) \geq 0, \text{ equal when } x = y$$

$$② \quad d(x, x) = \|x - x\| = \|0\| \stackrel{a)}{=} 0$$

$$③ \quad d(x, y) = \|x - y\| = \|(x - z) + (z - y)\|$$

$$\stackrel{b)}{\leq} \|x - z\| + \|z - y\|$$

$$= d(x, z) + d(z, y) \#$$

Q6 (EX 10.3 K&M 2017) Verify ℓ^2 defined in 10.1.1 is Hilbert

Recall the definitions

A complete innerproduct space is Hilbert Space

A complete space : A Cauchy sequence has limit

A Cauchy sequence : $d(x_m, x_n) \rightarrow 0$ as $m, n \rightarrow \infty$

ℓ^2 : a collection of sequences $x = \{x_1, x_2, \dots\}$ s.t. $\sum_{i=1}^{\infty} |x_i|^2 < \infty$

$$\langle x, y \rangle = \sum_{i=1}^{\infty} x_i \bar{y}_i$$

$$x+y = \{x_1+y_1, x_2+y_2, \dots\}$$

$$ax = \{ax_1, ax_2, \dots\}$$

$$\text{Hint: } \left| \sum_{i=1}^{\infty} x_i \bar{y}_i \right| \leq \left(\sum_{i=1}^{\infty} |x_i|^2 \right)^{1/2} \left(\sum_{i=1}^{\infty} |y_i|^2 \right)^{1/2}$$

$$|\langle x, y \rangle| \leq \sqrt{\langle x, x \rangle} \sqrt{\langle y, y \rangle}$$

[Construct $e_k(j) = \delta_{kj}$ such that $\langle x^m, e_k \rangle = \sum_{j=1}^{\infty} x_j^m e_k(j) = x_k^m$

Consider a Cauchy sequence x^i , $i=1, \dots$, $x^i \triangleq \{x_1^i, x_2^i, \dots\}$

$$\text{then } |x^m(k) - x^n(k)| = |\langle x^m - x^n, e_k \rangle|$$

$$\leq \|x^m - x^n\|^2 \|e_k\|^2 = \|x^m - x^n\|^2 \text{ holds for } j=1, 2, \dots$$

$$= d(x^m, x^n) \rightarrow 0 \text{ as } n, m \rightarrow \infty$$

Therefore $\{x^n(k)\}_{n \in \mathbb{N}}$ is Cauchy for each k , therefore converges.

Let $\tilde{x} = (\tilde{x}(k))_{k \in \mathbb{N}}$ be the limit of $\{x^n(k)\}_{n \in \mathbb{N}}$

only need to show $\tilde{x} \in \ell^2$.

$$\sum_{i=1}^{\infty} |\tilde{x}(k)|^2 = \sum_{i=1}^{\infty} \left| \lim_{n \rightarrow \infty} x^n(k) \right|^2 \xrightarrow{\text{Convergence uniform over } k} \lim_{n \rightarrow \infty} \sum_{i=1}^{\infty} |x^n(k)|^2 = \lim_{n \rightarrow \infty} \|x^n\|^2 < \infty \text{ by def.}$$

Therefore $\tilde{x} \in \ell^2$ #

Q7. Let $X(t)$, $t \in [0,1]$ be a stochastic process for which the sample paths lie in $L^2[0,1]$. Show that the solution to following problem coincides with the projections in the functional PCA.

$$\arg \min_{\{e_1, \dots, e_K, K \geq 1\}} \mathbb{E} \|X - \sum_{k=1}^K \langle X, e_k \rangle e_k\|^2$$

$$\begin{aligned} \mathbb{E} \|X - \sum_{k=1}^K \langle X, e_k \rangle e_k\|^2 &= \mathbb{E} \langle X - \sum_{k=1}^K \langle X, e_k \rangle e_k, X - \sum_{k=1}^K \langle X, e_k \rangle e_k \rangle \\ &= \mathbb{E} (\langle X, X \rangle - 2 \langle X, \sum_{k=1}^K \langle X, e_k \rangle e_k \rangle \\ &\quad + \langle \sum_{k=1}^K \langle X, e_k \rangle e_k, \sum_{k=1}^K \langle X, e_k \rangle e_k \rangle) \\ &= \mathbb{E} (\langle X, X \rangle - 2 \sum_{k=1}^K \langle X, e_k \rangle^2 \\ &\quad + \sum_{i=1}^K \sum_{j=1}^K \langle \langle X, e_i \rangle e_i, \langle X, e_j \rangle e_j \rangle) \\ &= \mathbb{E} (\langle X, X \rangle - 2 \sum_{k=1}^K \langle X, e_k \rangle^2 \\ &\quad + \sum_{i=1}^K \sum_{j=1}^K \langle X, e_i \rangle \cdot \langle X, e_j \rangle \cdot \langle e_i, e_j \rangle) \\ &= \mathbb{E} (\langle X, X \rangle - 2 \sum_{k=1}^K \langle X, e_k \rangle^2 \\ &\quad + \sum_{k=1}^K \langle X, e_k \rangle^2 \cdot 1) \\ &= \mathbb{E} (\langle X, X \rangle - \sum_{k=1}^K \langle X, e_k \rangle^2) \end{aligned}$$

$$\arg \min_{\{e_k\}} \mathbb{E} \| \langle X, X \rangle - \sum_{k=1}^K \langle X, e_k \rangle e_k \|^2$$

$$= \arg \min_{\{e_k\}} \mathbb{E} (\langle X, X \rangle - \sum_{k=1}^K \langle X, e_k \rangle^2) \stackrel{\uparrow}{=} \arg \max_{\{e_k\}} \mathbb{E} (\sum_{k=1}^K \langle X, e_k \rangle^2) \#$$

does not depend on $\{e_k\}$