

## Chapter 3: Facts about Common Statistical Models \*

### 3 Facts about common statistical models

#### 3.1 Bayes Models

- **Probability Model on Data** We have distributions  $\mathcal{P} \equiv \{P_\theta : \theta \in \Theta\}$  for  $X$  on  $(\mathcal{X}, \mathcal{B})$ , where  $\mathcal{P} \ll \mu$  ( $\sigma$ -finite measure) and R-N derivatives

$$\frac{dP_\theta}{d\mu}(x) = f_\theta(x)$$

- **Prior on Parameter** We now add an assumption of a distribution  $G$  on  $(\Theta, \mathcal{C})$  with  $G \ll \nu$  ( $\sigma$ -finite measure) and R-N derivative

$$\frac{dG}{d\nu}(\theta) = g(\theta)$$

- **Joint Distribution** for  $(X, \theta)$ : Here we consider  $f_\theta(x)$  as a function of both  $x$  and  $\theta$  (i.e., measurable in  $(x, \theta)$ ). If  $f_\theta(x)$  is  $\mathcal{B} \times \mathcal{C}$ -measurable, then there exists a joint probability distribution for  $(X, \theta)$  on  $(\mathcal{X} \times \Theta, \mathcal{B} \times \mathcal{C})$  defined, for  $A \in \mathcal{B} \times \mathcal{C}$ , by

$$\pi^{X, \theta}(A) \equiv P((X, \theta) \in A) = \int_A f_\theta(x) d(\mu \times G)(x, \theta) = \int f_\theta(x) g(\theta) d(\mu \times \nu)(x, \theta)$$

where

$$\frac{d\pi^{X, \theta}}{d(\mu \times G)} \equiv f_\theta(x), \quad \frac{d\pi^{X, \theta}}{d(\mu \times \nu)} \equiv f_\theta(x) g(\theta)$$

- **Marginal Distributions**

– for  $X$  ( $B \in \mathcal{B}$ )

$$\begin{aligned} \pi^X(B) &\equiv P(X \in B) = \pi^{X, \theta}(B \times \Theta) = \int_{B \times \Theta} f_\theta(x) d(\mu \times G)(x, \theta) \stackrel{\text{Fubini}}{=} \int_B \left[ \int_\Theta f_\theta(x) dG(\theta) \right] d\mu(x) \\ &= \int_B \left[ \int_\Theta f_\theta(x) g(\theta) d\nu \right] d\mu(x) \\ 0 &\leq \frac{d\pi^X(x)}{d\mu} = \int_\Theta f_\theta(x) dG(\theta) = \int_\Theta f_\theta(x) g(\theta) \end{aligned}$$

– for  $\theta$  ( $C \in \mathcal{C}$ )

$$\pi^\theta(C) \equiv P(\theta \in C) = \pi^{X, \theta}(\mathcal{X} \times C) = \int_{\mathcal{X} \times C} f_\theta(x) d(\mu \times G)(x, \theta) = \int_C \left[ \int_{\mathcal{X}} f_\theta(x) d\mu(x) \right] dG(\theta) = G(C)$$

Marginal distribution of  $\theta$  is prior distribution  $G$ .

- **Conditional distributions**

– for  $X \mid \theta$

$$\pi^{X|\theta}(B \mid \theta) \equiv P_{X|\theta}(X \in B \mid \theta) = \int_B f_\theta(x) d\mu(x) = P_\theta(B), \quad B \in \mathcal{B}$$

$$\frac{d\pi^{X|\theta}(x)}{d\mu} = \frac{dP_\theta(x)}{d\mu} = f_\theta$$

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– for  $\theta \mid X$

$$\pi^{\theta \mid X}(C \mid x) \equiv P_{\theta \mid X}(\theta \in C \mid X = x) = \int_C \left[ \frac{f_{\theta}(x)}{\int_{\Theta} f_{\theta}(x) dG(\theta)} \right] dG(\theta) = \int_C \frac{f_{\theta}(x)g(\theta)}{\int_{\Theta} f_{\theta}(x)g(\theta) d\nu(\theta)} d\nu(\theta)$$

$$\frac{d\pi^{\theta \mid X}(\theta)}{dG} = \frac{f_{\theta}(x)g(\theta)}{\int_{\Theta} f_{\theta}(x)g(\theta) d\nu(\theta)}, \quad \frac{d\pi^{\theta \mid X}(\theta)}{d\nu} = \frac{f_{\theta}(x)g(\theta)}{\int_{\Theta} f_{\theta}(x)g(\theta) d\nu(\theta)}, G \ll \nu$$

Note priors does not necessarily have a density, i.e.  $G$  is not necessarily dominated by some  $\nu$ . But you can always write the density of posterior with respect to  $G$ .

### 3.2 Exponential Family of Distributions

- **Exponential family:** Definition 16 :  $\mathcal{P} \equiv \{P_{\theta} : \theta \in \Theta\} \ll \mu$  ( $\sigma$ -finite measure) is an exponential family if, for some  $h(x) \geq 0$ , it holds that

$$f_{\theta}(x) \equiv \frac{dP_{\theta}}{d\mu}(x) = \exp \left( \alpha(\theta) + \sum_{i=1}^k \eta_i(\theta) T_i(x) \right) h(x), \quad x \in \mathcal{X}$$

for any  $\theta \in \Theta$

- **Identifiable:** Definition 17 : A family of distributions,  $\mathcal{P} \equiv \{P_{\theta} : \theta \in \Theta\}$  is identifiable if  $\theta_1 \neq \theta_2$  implies  $P_{\theta_1} \neq P_{\theta_2}$
- **Natural parameter space:** Let  $\boldsymbol{\eta} = (\eta_1, \dots, \eta_k) \in \mathbb{R}^k$  and let

$$\Gamma \equiv \left\{ \boldsymbol{\eta} \in \mathbb{R}^k : \int_{\mathcal{X}} h(x) \exp \left( \sum_{i=1}^k \eta_i T_i(x) \right) d\mu(x) < \infty \right\}$$

note: this kernel with linear combinations of  $T_i(X)$  using real numbers  $\eta_i, i = 1, \dots, k$ .

Define a new distributional family

$$\mathcal{P}^* \equiv \left\{ P_{\boldsymbol{\eta}} \text{ has R-N derivative as } f_{\boldsymbol{\eta}}(x) \equiv \frac{dP_{\boldsymbol{\eta}}}{d\mu}(x) = K(\boldsymbol{\eta}) h(x) \exp \left( \sum_{i=1}^k \eta_i T_i(x) \right) : \boldsymbol{\eta} \in \Gamma \right\}$$

- $\mathcal{P} \subset \mathcal{P}^*$
- $\Gamma$  is called the natural parameter space for  $\mathcal{P}^*$  and  $\Gamma$  is a convex subset of  $\mathbb{R}^k$
- If  $\Gamma$  lies in a subspace of dimension less than  $k$ , then  $f_{\boldsymbol{\eta}}(x)$  ( and  $f_{\theta}(x)$ ) can be re-written in a form involving fewer than  $k$  statistics  $T_i(x)$ . (We'll assume  $\Gamma$  to be fully  $k$ -dimensional.)
- $\mathcal{P}$  may be a proper subset of  $\mathcal{P}^*$  or

$$\Gamma_{\theta} \equiv \left\{ (\eta_1(\theta), \eta_2(\theta), \dots, \eta_k(\theta)) \in \mathbb{R}^k : \theta \in \Theta \right\}$$

can be a proper subset of  $\Gamma$ .