

— Answers for homework 4

$$\begin{aligned} 3. \Pr(B) &= \Pr(B|A) \times \Pr(A) + \Pr(B|A^c) \times \Pr(A^c) \\ &= \Pr(B|A_1) \times \Pr(A_1) + \Pr(B|A_2) \times \Pr(A_2) + \Pr(B|A_3) \times \Pr(A_3) \\ &= 0.001 \times 0.99 + 0.90 \times 0.001 + 0.90 \times 0.009 = 0.00999 \end{aligned}$$

$$\begin{aligned} 4. a \Pr &= 1 - \left[ \binom{15}{0} (0.2)^0 (0.8)^{15} + \binom{15}{1} (0.2)^1 (0.8)^{14} + \binom{15}{2} (0.2)^2 (0.8)^{13} + \binom{15}{3} (0.2)^3 (0.8)^{12} \right. \\ &\quad \left. + \binom{15}{4} (0.2)^4 (0.8)^{11} + \binom{15}{5} (0.2)^5 (0.8)^{10} \right] = 0.061 \end{aligned}$$

$$b \quad E(X) = np = 15 \times 0.2 = 3$$

$$\begin{aligned} 5. a \Pr(X < x_0) &= \Pr\left(\frac{X - \mu}{\sigma} < \frac{x_0 - \mu}{\sigma}\right) = \Phi\left(\frac{200 - 219}{50}\right) = \Phi(-0.38) \\ &= 0.352 \end{aligned}$$

$$\begin{aligned} b \Pr(X > x_0) &= 1 - \Pr(X < x_0) = 1 - \Pr\left(\frac{X - \mu}{\sigma} < \frac{x_0 - \mu}{\sigma}\right) = 1 - \Phi\left(\frac{250 - 219}{50}\right) \\ &= 1 - \Phi(0.62) = 0.2676 \end{aligned}$$

$$\begin{aligned} c \Pr(x_1 < X < x_2) &= \Pr(X < x_2) - \Pr(X < x_1) = \Pr\left(\frac{X - \mu}{\sigma} < \frac{x_2 - \mu}{\sigma}\right) - \Pr\left(\frac{X - \mu}{\sigma} < \frac{x_1 - \mu}{\sigma}\right) \\ &= \Phi\left(\frac{250 - 219}{50}\right) - \Phi\left(\frac{200 - 219}{50}\right) = \Phi(0.62) - \Phi(-0.38) = 0.3804 \end{aligned}$$



## Answers for homework 5

1. a  $\text{sd}(\bar{X}_1)$   
 standard error =  $\sigma_1/\sqrt{n_1} \approx 0.5/\sqrt{40} = 0.0791$   
 $\text{sd}(\bar{X}_2) = \sigma_2/\sqrt{n_2} \approx 0.4/\sqrt{32} = 0.0707$   
 b If central limit theorem applies,  $\bar{X} \sim N(\mu, \sigma^2/n)$   
 which is  $\bar{X}_1 \sim N(1.35, 0.00625)$ ,  $\bar{X}_2 \sim N(0.92, 0.005)$

2. variable	Baseline	Follow up
Serum creatinine	$CI = (\bar{X} - t_{n-1, 1-\alpha/2} S/\sqrt{n}, \bar{X} + t_{n-1, 1-\alpha/2} S/\sqrt{n})$ $= (0.97 - t_{101, 0.975} \times \frac{0.22}{\sqrt{102}}, 0.97 + t_{101, 0.975} \times \frac{0.22}{\sqrt{102}})$ $= (0.97 - 1.984 \times \frac{0.22}{\sqrt{102}}, 0.97 + 1.984 \times \frac{0.22}{\sqrt{102}})$ $= (0.9268, 1.0132)$	$CI = (0.95436, 1.0456)$
Serum potassium	$CI = (4.3042, 4.5557)$	$CI = (4.3195, 4.6605)$
Serum phosphate	$CI = (1.5876, 1.7723)$	$CI = (1.4739, 1.6661)$
PAIs	$CI = (33.341, 39.6588)$	$CI = (19.9581, 26.5819)$

3. a. point estimator  $\hat{p} = \frac{7}{525} = 13.33\%$

b.  $\text{sd}(\hat{p}) = \sqrt{\frac{pq}{n}} = \sqrt{\frac{0.1333 \times (1-0.1333)}{525}} = 0.01483$

$CI = (\hat{p} - z_{1-\alpha/2} \text{sd}(\hat{p}), \hat{p} + z_{1-\alpha/2} \text{sd}(\hat{p})) = (0.1043, 0.1623)$



# Answers for homework 6

1. a)  $H_0: \mu = 1.2$   $H_1: \mu \neq 1.2$

b/ Type I error = rejection of  $H_0$  that is actually true

c/ Type II error = fail to reject  $H_0$  that is actually wrong

d/  $t = \frac{\bar{X} - \mu}{s/\sqrt{n}} = \frac{1.05 - 1.2}{0.5/\sqrt{15}} = -1.1619$

$P\text{value} = P(T \leq -1.1619 \text{ or } T \geq 1.1619) = 0.1324 \times 2 = 0.2648 > \alpha$

Fail to reject  $H_0$ .

2. a/  $H_0: \mu = 10$   $H_1: \mu \neq 10$

b/ Type I error = rejection of  $H_0$   $\mu = 9.1$  that is actually true

c/ Type II error = fail to reject  $H_0$   $\mu = 9.1$  that is actually wrong

d/  $t = \frac{\bar{X} - \mu}{s/\sqrt{n}} = \frac{9.1 - 10}{2.3/\sqrt{100}} = -3.9130$

$P\text{value} = P(T \leq -3.9130 \text{ or } T \geq 3.9130) = 2 \times 0.0001 = 0.0002 < \alpha$

Reject  $H_0$ , which indicates  $\mu$  is significantly different from 10

e/  $CI = (\bar{X} - t_{n-1, \frac{\alpha}{2}} \frac{s}{\sqrt{n}}, \bar{X} + t_{n-1, \frac{\alpha}{2}} \frac{s}{\sqrt{n}})$   
 $= (8.6437, 9.5563)$

3. a/ paired data

b/  $H_0: \mu_D = 0$   $H_1: \mu_D \neq 0$

c/ For patient  $i$ ,  $D_i = Y_i - X_i$ .  $\bar{D} = \frac{\sum_{i=1}^{14} D_i}{14} = -0.3629$   $t = \frac{-0.3629 - 0}{0.4059/\sqrt{14}} = -3.3453$

$P\text{value} = P(T \leq -3.3453 \text{ or } T \geq 3.3453) = 2 \times 0.0026 = 0.0052 < \alpha$

Reject  $H_0$ .

d/  $\mu_D$  is significantly different from 0, which means two treatments are significantly different.

$$4. \quad H_0 = \mu_1 = \mu_2$$

$$H_1 = \mu_1 \neq \mu_2$$

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{s_p^2 \left( \frac{1}{n_1} + \frac{1}{n_2} \right)}}$$

$$= \frac{(54.8 - 69.5) - 0}{\sqrt{s_p^2 \left( \frac{1}{156} + \frac{1}{148} \right)}}$$

$$= -4.0687$$

$$s_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{(n_1 - 1) + (n_2 - 1)}$$

$$= \frac{(156 - 1) \times 28.1^2 + (148 - 1) \times 34.7^2}{(156 - 1) + (148 - 1)}$$

$$= 991.3602$$

$$P_{\text{value}} = P(T < -4.0687 \text{ or } T > 4.0687) < 0.0001 < \alpha$$

The dietary fat intake between two groups of men are significantly different.



# Answers for homework 8

1.a.  $H_0: \mu_1 = \mu_2 = \mu_3$

$H_1: \mu_i \neq \mu_j$  for some  $i, j$

b. Source of variation	SS	df	MS	F value	p
between group	SST = 180.067	2	MST = 90.033	4.662	0.018
within group	SSE = 521.459	27	MSE = 19.313		
Total	Tot. 526	29			

c. Since  $p\text{-value} < 0.05$ , reject  $H_0$ . There is evidence that at least  $\mu_j$  differs from the rest.

2.

a.  $\hat{p}_1 = \frac{10}{10+90} = 0.1$      $\hat{p}_2 = \frac{14}{14+86} = 0.14$      $\hat{p}_3 = \frac{19}{19+81} = 0.19$

b.  $H_0: p_1 = p_2 = p_3$      $H_1: p_i \neq p_j$  for some  $i, j$

$\hat{p} = \frac{10+14+19}{100+100+100} = 0.143$

	# of Tumors	
Rat group	one or more	none
Control	Observed = 10 expected = $0.143 \times 100 = 14.3$	Observed = 90 expected = 85.7
Low dose	observed = 14 expected = 14.3	observed = 86 expected = 85.7
High dose	observed = 19 expected = 14.3	observed = 81 expected = 85.7

$$\chi^2 = \sum \frac{(O-E)^2}{E} = \frac{(10-14.3)^2}{14.3} + \frac{(14-14.3)^2}{14.3} + \frac{(19-14.3)^2}{14.3} + \frac{(90-85.7)^2}{85.7} + \frac{(86-85.7)^2}{85.7} + \frac{(81-85.7)^2}{85.7}$$

$$= 3.3186$$

$P\text{-value} = 0.81 > \alpha$ , Fail to reject  $H_0$

# Answers für homework 9.

1. a.  $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$

$$\hat{\beta}_1 = \frac{L_{xy}}{L_{xx}} = \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2} = \frac{\sum X_i Y_i - \frac{\sum X_i \sum Y_i}{n}}{\sum X_i^2 - \frac{(\sum X_i)^2}{n}} = \frac{423643.3 - \frac{23670 \times 214.9}{12}}{46689410 - \frac{(23670)^2}{12}} = -0.7372$$

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X} = \frac{214.9}{12} + 0.7372 \times \frac{23670}{12} = 1471.965$$

b.  $H_0: \beta_1 = 0$

$H_1: \beta_1 \neq 0$

$$t = \frac{\hat{\beta}_1 - 0}{S_{\hat{\beta}_1}}$$

$$S_{y.x} = \sqrt{\frac{\sum (Y_i - \hat{Y}_i)^2}{n-2}} = 0.5733$$

$$S_{\hat{\beta}_1} = \frac{S_{y.x}}{\sqrt{\sum (X_i - \bar{X})^2}} = 0.03122$$

$$= \frac{-0.7372}{0.03122}$$

$$= -23.54$$

$$\Pr(>|t|) < 0.0001 < \alpha. \text{ Reject } H_0$$

c.  $\hat{Y} = -0.7372 \times 1989 + 1471.965 = 5.6742$

d. It will be inaccurate to apply this model to make predictions or forecasts outside of relevant range

2. a.  $H_0: \mu_y = 0$

$H_1: \mu_y \neq 0$

$$t = \frac{\bar{Y} - \mu_y}{S_y / \sqrt{n}} = \frac{\frac{290}{16} - 0}{15.10106 / \sqrt{16}} = 4.82992$$

$$P\text{-value} = 0.00022 < \alpha$$

reject  $H_0$

b, d, see attached plots.

c.  $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$

$$\hat{\beta}_1 = \frac{\sum X_i Y_i - \frac{\sum X_i \sum Y_i}{n}}{\sum X_i^2 - \frac{(\sum X_i)^2}{n}} = 0.2866$$

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X} = -33.8507$$



$$e. H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

$$t = \frac{\hat{\beta}_1 - 0}{s_{\hat{\beta}_1}} = \frac{0.2866}{0.0780} = 3.672$$

$$\Pr(>|t|) = 0.00251 < \alpha. \text{ Reject } H_0$$

