CTMC – Time spent in a state

- For a time homogeneous CTMC, we have a transition probability matrix with entries $p_{ij}(t)$, i.e., $P(t) = [[p_{ij}(t)]]$.
- ► Let $Y_t = \inf\{s > 0 : X(t+s) \neq X(t)\}$
- $\succ Y_t$ is the remaining time that the process spends in whichever state it is in, at time t.
- Intuitively, the time spent in a state should depend only on what state it is in, and not on the previous state.

Theorem

$$P(Y_t > u | X(t) = i) := \bar{G}_i(u) = e^{-a_i u}$$

for all $i \in S$ and $t \ge 0$, $u \ge 0$ and for some real number $a_i \in [0, \infty]$.

Proof 1

- $\bar{G}_i(u+v) = P(X(s)=i, s \in [t, t+u+v]|X(t)=i)$
- $\bar{G}_i(u+v) = P(X(s) = i, s \in [t+u, t+u+v]; X(p) = i, p \in [t, t+u]|X(t) = i)$
- ightharpoonup P(AB|C) = P(A|BC)P(B|C)
- ▶ Due to Markov property we have P(AB|C) = P(A|B)P(B|C)
- $P(X(s) = i, s \in [t + u, t + u + v]|X(p) = i, p \in [t, t + u]) =$
- $P(X(s) = i, s \in [t + u, t + u + v]|X(t + u) = i) = \bar{G}_i(v)$
- $P(X(p) = i, p \in [t, t + u]|X(t = i)) = \bar{G}_i(u)$
- $ightharpoonup ar{G}_i(u+v) = ar{G}_i(u)ar{G}_i(v)$
- Only CCDF function which satisfies this equation is the exponential distribution. This requires a proof. We will skip this part.

Simpler Proof

- Let τ_i denote the time the CTMC spends in state i before moving out. Suppose the CTMC is in state i at time 0.
- $\qquad \qquad \textbf{What is } P(\tau_i > s + t | \tau_i > s)?$
- Note that X(s) = i and therefore from the Markov property,

$$P(\tau_{i} > s + t | \tau_{i} > s) = P(X(u) = i, u \in [s, s + t] | X(t) = i, t \in [0, s])$$

$$= P(X(u) = i, u \in [s, s + t] | X(s) = i)$$

$$= P((X(u) = i, u \in [0, t] | X(0) = i)$$

$$= P(\tau_{i} > t).$$

Since $P(\tau_i > s + t | \tau_i > s) = P(\tau_i > t)$, this implies the distribution has memoriless property and must be exponential.

Finite dimensional distributions

- ▶ Consider a DTMC $\{X_n, n \ge 0\}$ with tpm denoted by P.
- \triangleright We assume M states and X_0 denotes the initial state.
- You can start in any starting state or may pick your starting state randomly.
- Let $\bar{\mu} = (\mu_1, \dots, \mu_M)$ denote the initial distribution.
- How does one obtain the finite dimensional distribution $P(X_0 = x_0, X_1 = x_1, ..., X_k = x_k)$?

Finite dimensional distributions

- ▶ Consider a CTMC $\{X_t, t \ge 0\}$ with t-time pm given by P(t).
- \triangleright We assume M states and X_0 denotes the initial state.
- Let $\bar{\mu} = (\mu_1, \dots, \mu_M)$ denote the initial distribution.
- How does one obtain the finite dimensional distribution $P(X_0 = x_0, X_{t_1} = x_1, ..., X_{t_k} = x_k)$?

Chapman Kolmogorov Equations for DTMC

- $ightharpoonup P = [[p_{ij}]]$ denotes the one step transition probability matrix.
- Let $P^{(n)}$ denote the n-step transition probability matrix.
- ► CK equation tells us that $P^{(n+l)} = P^{(n)}P^{(l)}$.
- $p_{ij}^{(n+l)} = P(X_{n+l} = j | X_0 = i) = \sum_k P(X_{n+l} = j, X_n = k | X_0 = i)$
- $p_{ij}^{(n+1)} = \sum_{k} P(X_{n+1} = j | X_n = k, X_0 = i) P(X_n = k | X_0 = i)$
- $p_{ij}^{(n+l)} = \sum_{k} P(X_{n+l} = j | X_n = k) P(X_n = k | X_0 = i)$
- $p_{ij}^{(n+l)} = \sum_{k} p_{ik}^{(n)} p_{kj}^{(l)} = [P^{(n)} P^{(l)}]_{ij}$
- At which step did we use time homogeneity and the Markov property?

n step transition probabilities

- $ightharpoonup P = [[p_{ij}]]$ denotes the one step transition probability matrix.
- Let $P^{(n)}$ denote the n-step transition probability matrix.
- From the CK equation we know that $P^{(n+l)} = P^{(n)}P^{(l)}$.
- ▶ It is easy to see that $P^{(n)} = P^{(n-1)}P$.
- For an M state DTMC, $p_{ij}^{(2)} = \sum_{k=1}^{M} p_{ik} p_{kj}$.
- This implies that that the n-step transition probability matrix can be obtained as $P^{(n)} = P^n$
- ▶ Given X_0 and P, you can generate n-step probabilities or $P_{X_0}(X_n)$

Chapman Kolmogorov Equations for CTMC

- Let P(t) denote the t-time transition probability matrix.
- ▶ CK equation for a CTMC is P(t+I) = P(t)P(I).
- $P_{ij}(t+1) = P(X(t+1) = j|X(0) = i)$
- $ightharpoonup = \sum_{k} P(X(t+1) = j, X(t) = k | X(0) = i)$
- $= \sum_{k} P(X(t+1) = j | X(t) = k, X(0) = i) P(X(t) = k | X(0) = i)$
- $= \sum_{k} P(X(t+1) = j | X(t) = k) P(X(n) = k | X(0) = i)$
- $ho_{ij}(t+I) = \sum_{k} p_{ik}(t) p_{kj}(I) = [P(t)P(I)]_{ij}$

What generates a CTMC ?

- P(t + I) = P(t)P(I).
- \triangleright In DTMC, we could use P to generate the chain on Matlab.
- ▶ What about CTMC ? Can we use P(t)?
- ▶ What is $\lim_{h\to 0} P(h)$?
- ▶ What is $\frac{dP(h)}{dh}$ evaluated at h = 0?

What generates a CTMC ?

- Lets look at $\frac{dP(h)}{dh}|_{h=0} = \lim_{h\to 0} \frac{P(h)-P(0)}{h} = \lim_{h\to 0} \frac{P(h)-I}{h}$.
- ▶ Define $Q := \lim_{h \to 0} \frac{P(h) I}{h}$
- Does it always exist ? Yes! (Proposition 2.2 and 2.4 (Anderson))
- ▶ Q has terms of the form q_{ii} and q_{ij} for $i, j \in \{1, 2, ..., M\}$.
- $ightharpoonup q_{ii}=rac{dp_{ii}(h)}{dh}|_{h=0}.$ Similarly $q_{ij}=rac{dp_{ij}(h)}{dh}|_{h=0}$

What generates a CTMC?

Theorem

Let P(t) be a transition function. Then the generator matrix $Q = \lim_{h \to 0} \frac{P(h)-l}{h}$ exists.

Theorem

$$P(Y_t > u | X(t) = i) = e^{-a_i u}$$
 where $a_i > 0$.

Theorem

(Proposition 2.8 Anderson)

$$P(Y_t > u | X(t) = i) := e^{q_{ii}u}, i.e., q_{ii} = -a_i.$$

 $P(X(t + Y_t) = j | X(t) = i) = \frac{q_{ij}}{|q_{ii}|}.$

Q generates the CTMC

- \triangleright Cannot generate CTMC directly from P(t).
- From P(t), obtain Q using $Q = \frac{dP(h)}{dh}|_{h=0}$
- ▶ Consider Y_t when X(t) = i.
- Now use the following theorem for generating the CTMC on a computer

Theorem

(Proposition 2.8 Anderson: we won't see proof) $P(Y_t > u | X(t) = i) := e^{q_{ii}u}$, i.e., $q_{ii} = -a_i$. $P(X(t + Y_t) = j | X(t) = i) = \frac{q_{ij}}{|q_{ii}|}$.