

# Bayesian Learning Lab 1 Report

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## QUESTION 1

### Part A)

Given - Let  $y_1, \dots, y_n | \theta \sim \text{Bern}(\theta)$ , and assume that you have obtained a sample with  $f = 35$  failures in  $n = 78$  trials. Assume a  $\text{Beta}(\alpha_0, \beta_0)$  prior for  $\theta$  and let  $\alpha_0 = \beta_0 = 7$ .

Draw 10000 random values ( $n\text{Draws} = 10000$ ) from the posterior  $\theta | y \sim \text{Beta}(\alpha_0 + s, \beta_0 + f)$ , where  $y = (y_1, \dots, y_n)$ , and verify graphically that the posterior mean  $E[\theta | y]$  and standard deviation  $\text{SD}[\theta | y]$  converges to the true values as the number of random draws grows large.

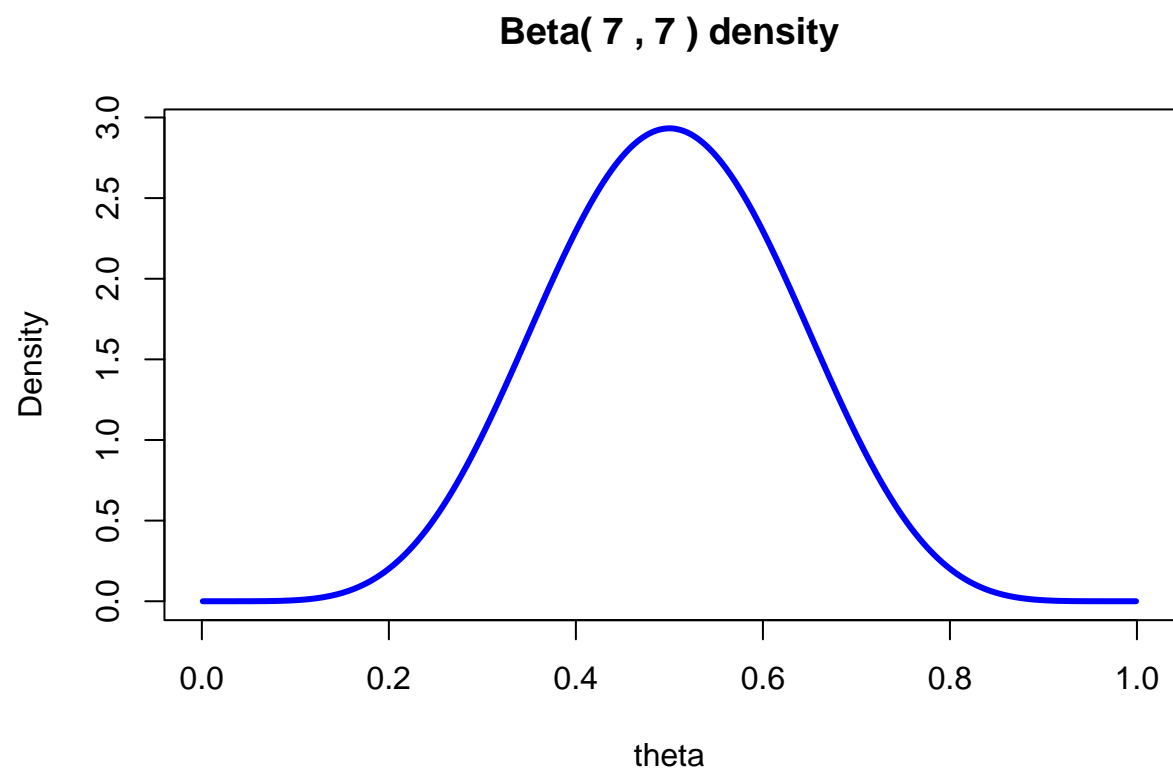
```
### Variables ###
f = 35 # Number of failures
n = 78 # Total trials
a0 = 7 # Prior alpha
b0 = 7 # Prior beta
s = n - f # Number of successes
#HELPER FUNCTIONS WHICH WILL BE USED REPETITIVELY
meanBeta <- function(a, b) {
  return(a / (a + b))
}

varBeta <- function(a, b) {
  return((a * b) / (((a + b)^2) * (a + b + 1)))
}

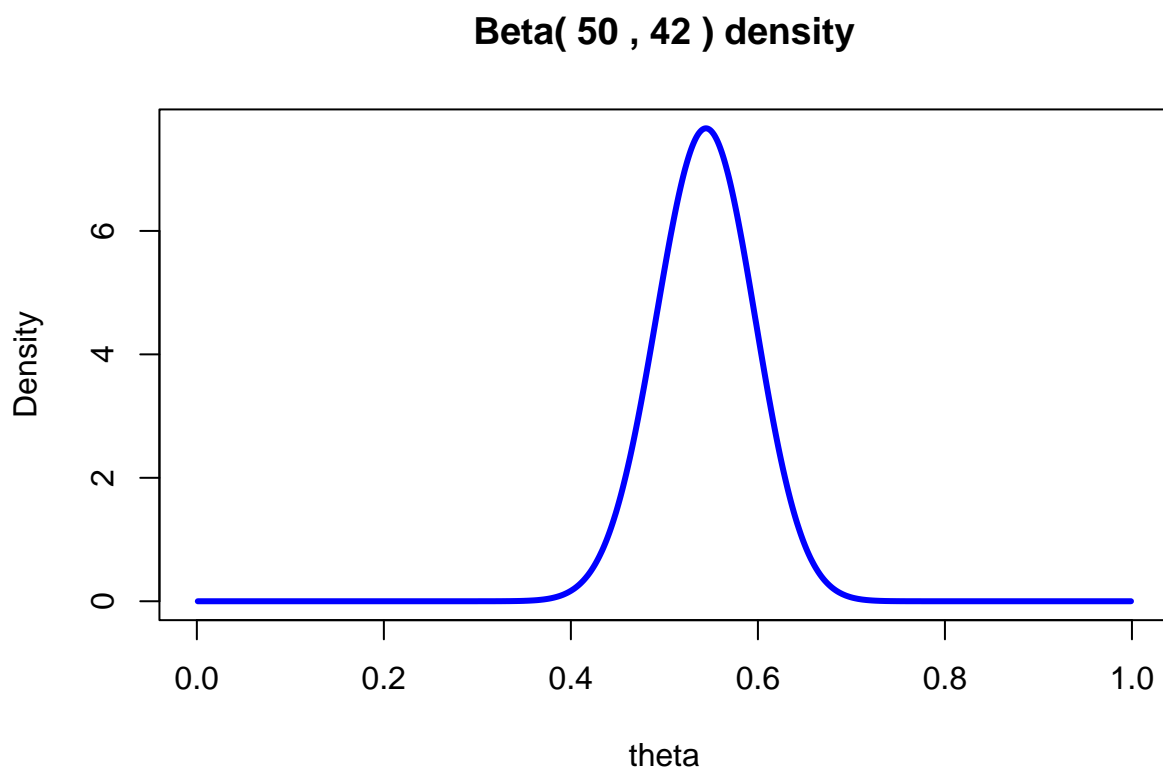
BetaPlot <- function(a, b) {
  xGrid <- seq(0.001, 0.999, by = 0.001)
  prior = dbeta(xGrid, a, b)
  maxDensity <- max(prior)
  plot(xGrid, prior, type = 'l', lwd = 3, col = "blue", xlim = c(0,1), ylim = c(0, maxDensity),
       xlab = "theta", ylab = 'Density', main = paste('Beta(', a, ',', b, ') density'))
}

#Draw 10,000 samples, show posterior mean and SD converge correctly.
# & Visualize convergence of Beta posterior estimates to theoretical true values.

# Prior and Posterior Plots
BetaPlot(a0, b0) # Prior density
```



```
BetaPlot(a0 + s, b0 + f) # Posterior density
```



```
# Posterior Mean and Variance
meanPost = meanBeta(a0 + s, b0 + f)
varPost = varBeta(a0 + s, b0 + f)
stdPost = sqrt(varPost)

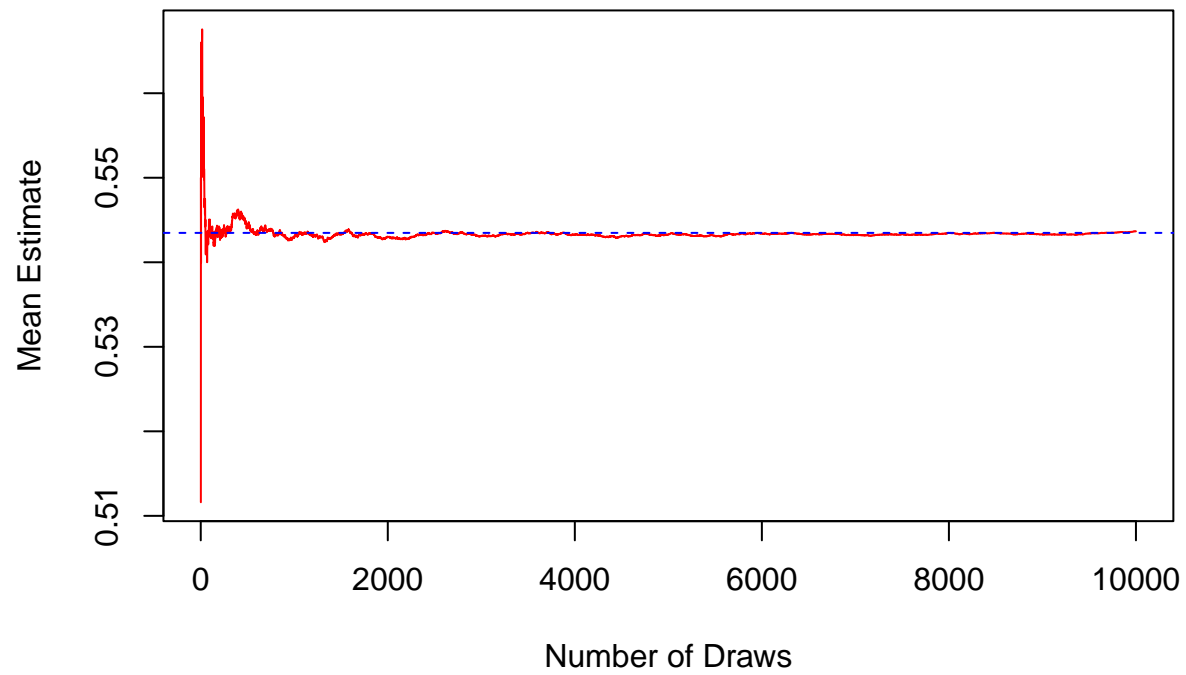
# Sampling from Posterior
Ndraws = 10000
samples = rbeta(Ndraws, shape1 = a0 + s, shape2 = b0 + f)

# Convergence of Mean and Std Dev
mean_vals = rep(0, Ndraws)
std_vals = rep(0, Ndraws)

for (i in 1:Ndraws) {
  mean_vals[i] = mean(samples[1:i])
  std_vals[i] = sd(samples[1:i])
}

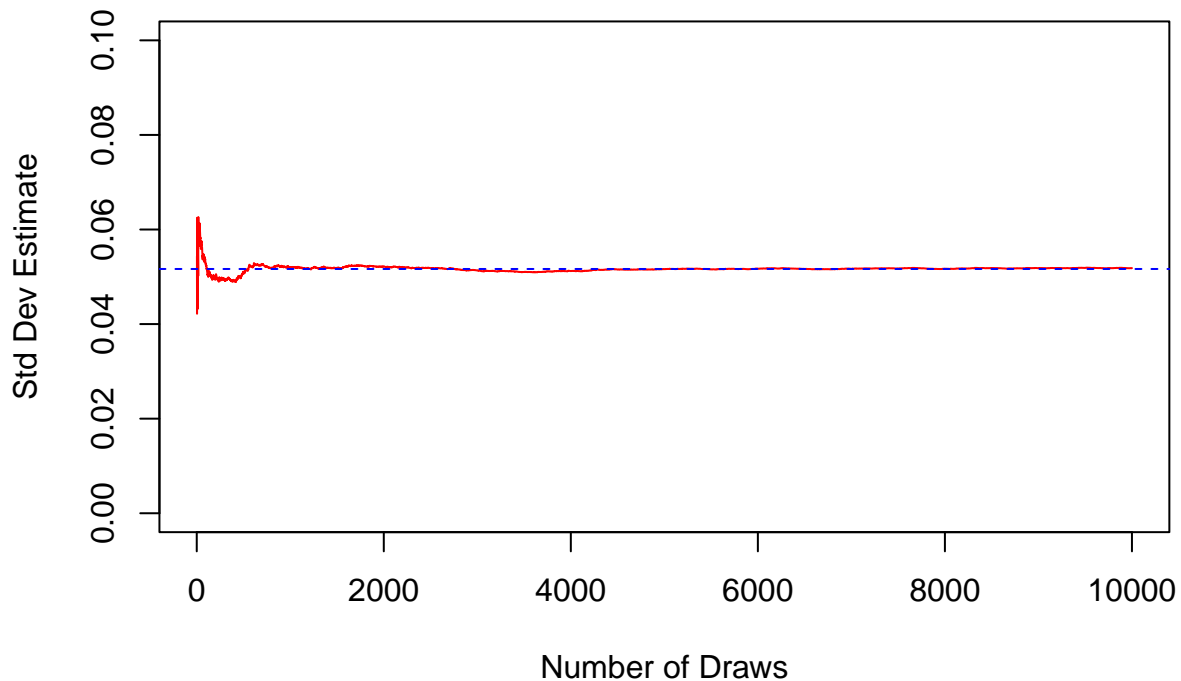
plot(mean_vals, type = "l", col = "red", main = "Convergence of Posterior Mean",
      xlab = "Number of Draws", ylab = "Mean Estimate")
abline(h = meanPost, col = "blue", lty = 2)
```

## Convergence of Posterior Mean



```
plot(std_vals, type = "l", col = "red", main = "Convergence of Posterior Std Dev",  
     xlab = "Number of Draws", ylab = "Std Dev Estimate", ylim = c(0, 0.1))  
abline(h = stdPost, col = "blue", lty = 2)
```

## Convergence of Posterior Std Dev



### Part B)

Draw 10000 random values from the posterior to compute the posterior probability  $\Pr(\theta > 0.5 | y)$  and compare with the exact value from the Beta posterior.

```
# Part (b) - Compute Probability  $P(\theta > 0.5 | y)$ 
threshold = 0.5
true_probability = 1 - pbeta(threshold, shape1 = a0 + s, shape2 = b0 + f)
simulated_probability = mean(samples > threshold)

print(paste("True Probability: ", round(true_probability, 4)))
```

```
## [1] "True Probability: 0.7991"
```

```
print(paste("Simulated Probability: ", round(simulated_probability, 4)))
```

```
## [1] "Simulated Probability: 0.7986"
```

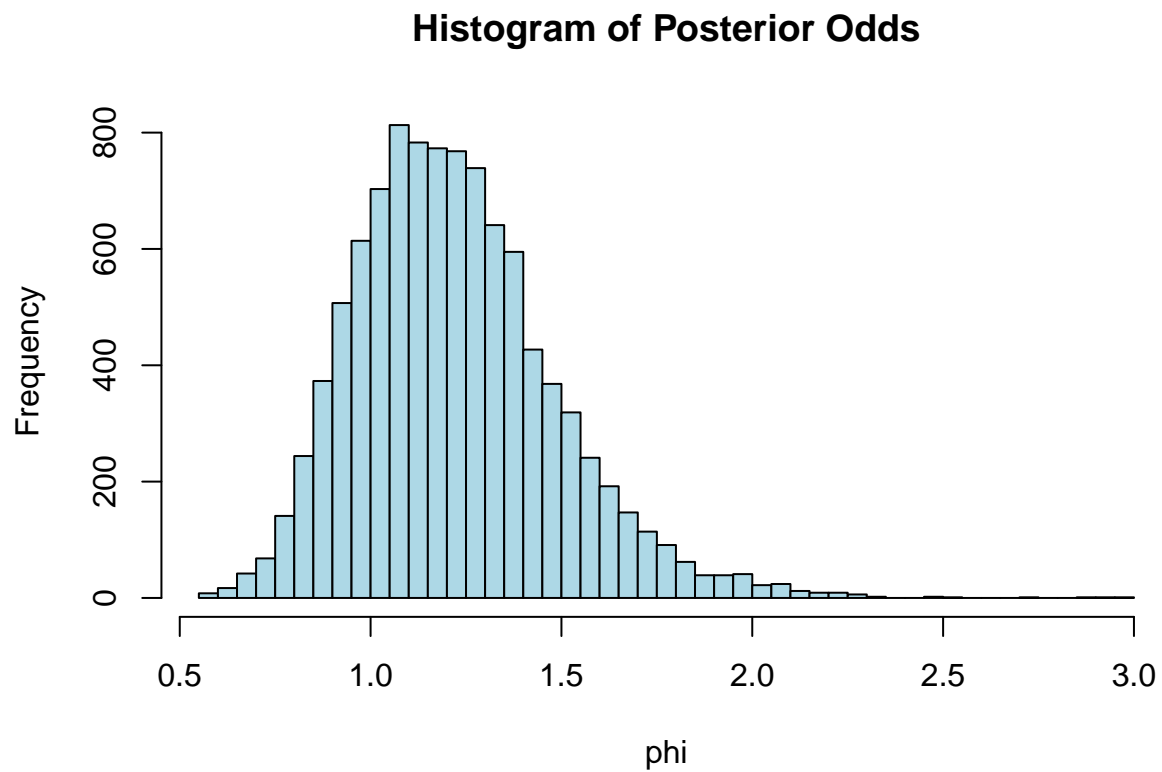
### Part C)

Draw 10000 random values from the posterior of the odds  $\Phi =$

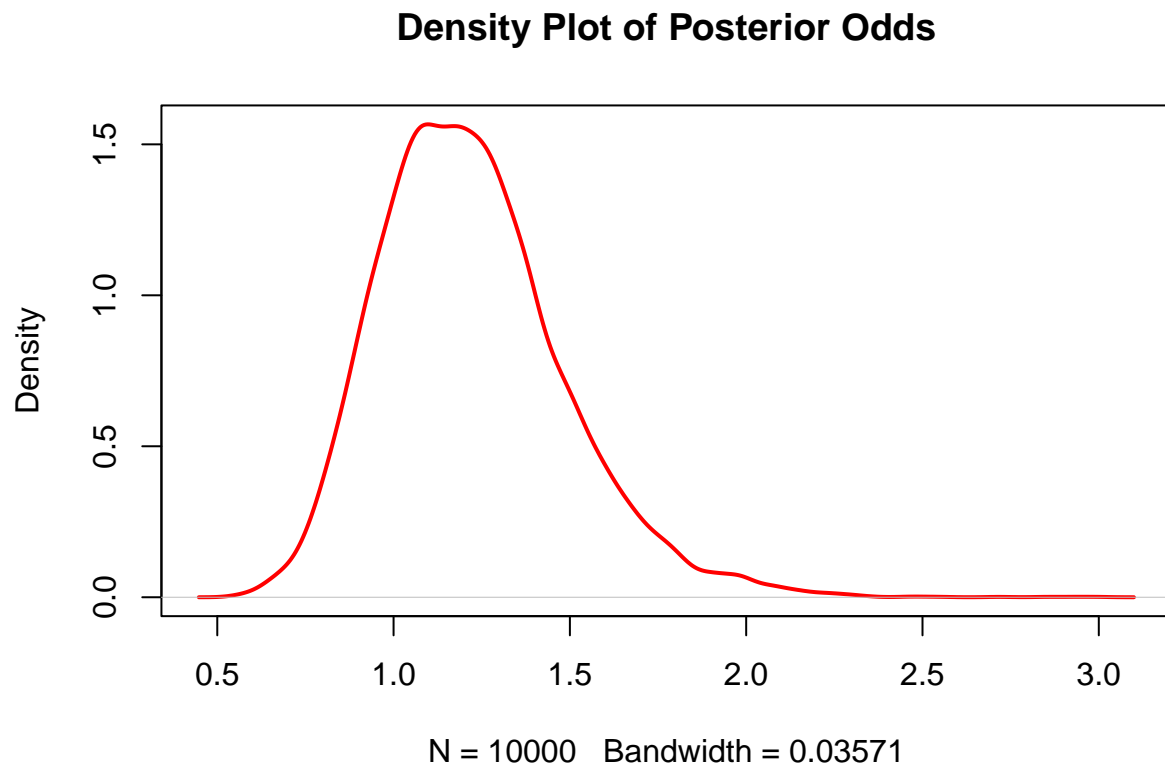
$$\theta / (1 - \theta)$$

by using the previous random draws from the Beta posterior for  $\theta$  and plot the posterior distribution of  $\theta$ .

```
# Part (c) - Posterior Distribution of Odds phi
phi = samples / (1 - samples)
hist(phi, breaks = 50, main = "Histogram of Posterior Odds", xlab = "phi", col = "lightblue")
```



```
plot(density(phi), main = "Density Plot of Posterior Odds", col = "red", lwd = 2)
```



## QUESTION 2

Draw 10000 random values from the posterior of  $\sigma^2$  by assuming  $\mu = 3.65$  and plot the posterior distribution.

### Part A)

```
# Load required libraries  
library(dplyr)
```

```
##  
## Attaching package: 'dplyr'  
  
## The following objects are masked from 'package:stats':  
##  
##   filter, lag  
  
## The following objects are masked from 'package:base':  
##  
##   intersect, setdiff, setequal, union
```

```

library(bayestestR) # Needed for HPDI calculation

## Warning: package 'bayestestR' was built under R version 4.4.3

# Given income data
income <- c(22, 33, 31, 49, 65, 78, 17, 24)

# (a) Compute  $\tau^2$  and draw 10,000 samples from the posterior of  $\sigma^2$ 
mu <- 3.65 # Known mean  $\mu$ 
Ndraw <- 10000 # Number of posterior samples
n <- length(income) # Sample size

# Function to compute  $\tau^2 = (1/n) * \sum (\log(y_i) - \mu)^2$ 
taoFunc <- function(y) {
  return(sum((log(y) - mu)^2) / length(y))
}
tao <- taoFunc(income)

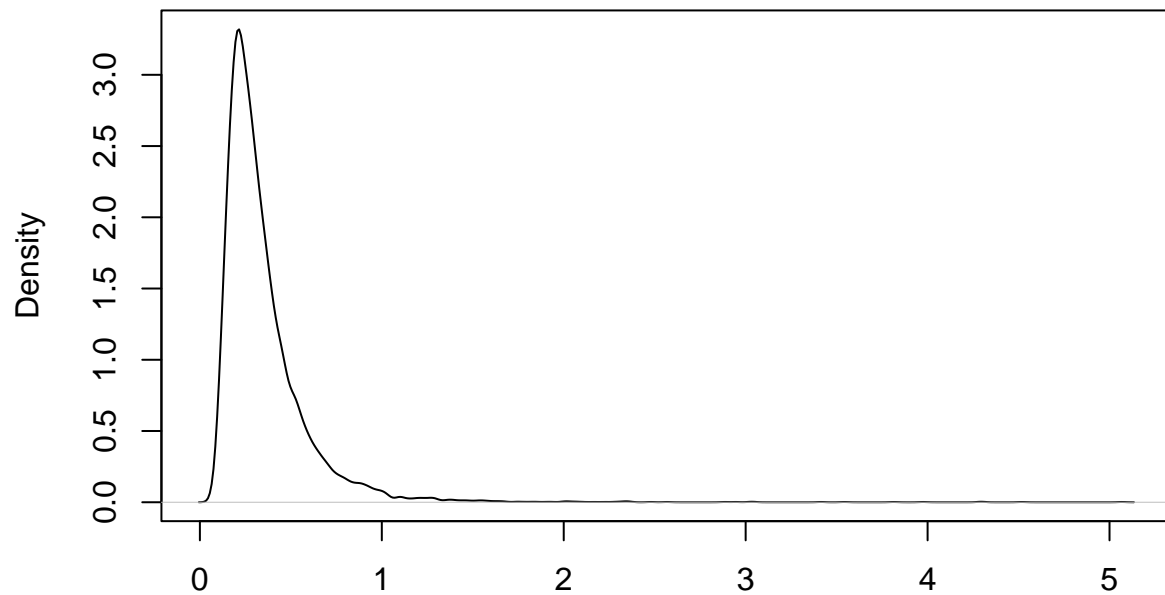
# Draw from chi-squared distribution and compute  $\sigma^2$  posterior samples
set.seed(12345)
chi_vals <- rchisq(Ndraw, n)
sigmaSquared <- (n * tao) / chi_vals # Posterior samples of  $\sigma^2$ 

# Plot posterior distribution of  $\sigma^2$ 
plot(density(sigmaSquared), main = expression(paste("Posterior Distribution of ", sigma^2)))

```



## Posterior Distribution of $\sigma^2$



N = 10000 Bandwidth = 0.02223

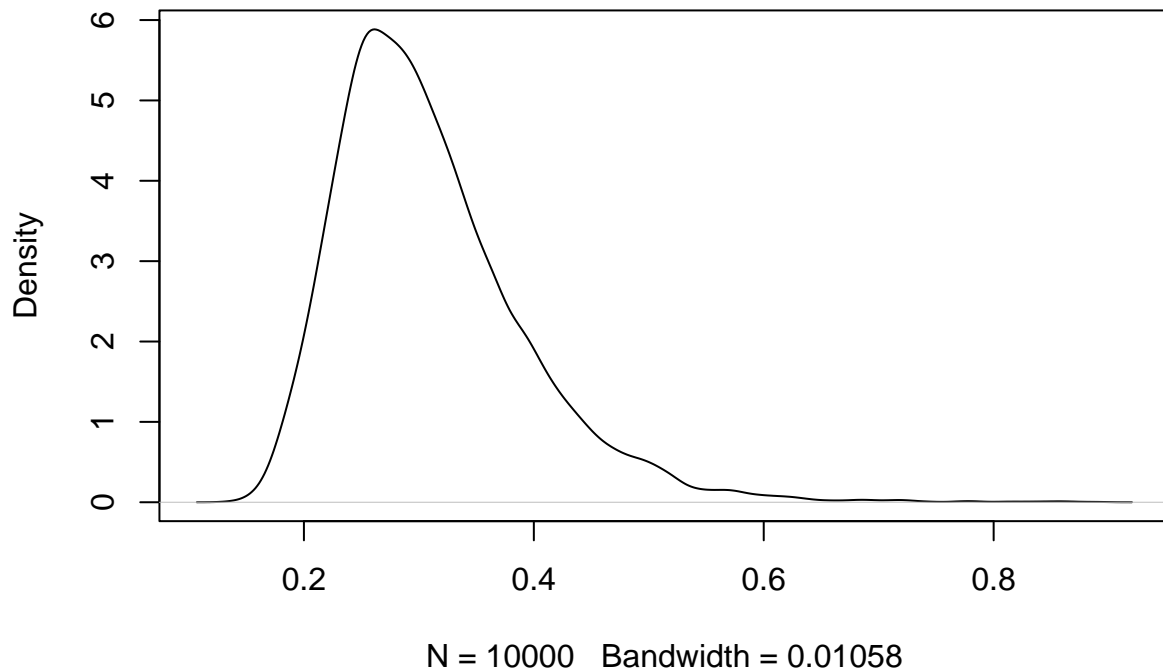
### Part B)

Compute the posterior distribution of Gini coefficient  $G = 2 * \Phi(\sigma / \sqrt{2}) - 1$ , where  $\Phi$  is the CDF of standard normal

```
sigma <- sqrt(sigmaSquared) # Get standard deviation from variance
G <- 2 * pnorm(sigma / sqrt(2)) - 1 # Compute Gini for each draw

# Plot the posterior distribution of Gini coefficient
plot(density(G), main = "Posterior Distribution of Gini Coefficient")
```

## Posterior Distribution of Gini Coefficient



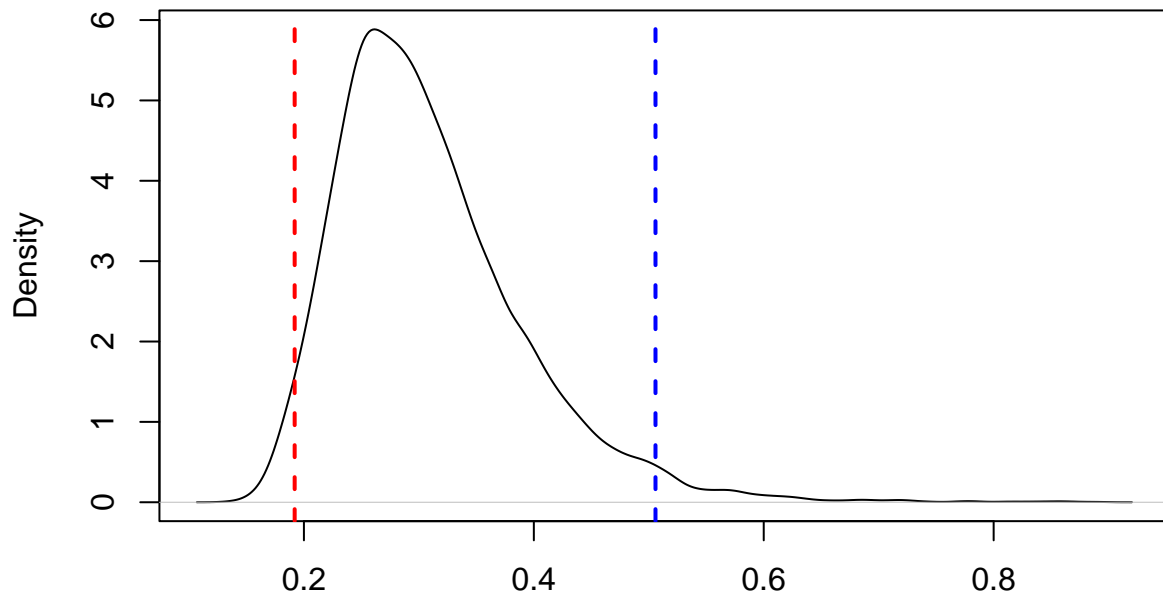
### Part C)

Compute a 95% equal tail credible interval for G.

```
# (c) Compute 95% Equal-Tail Credible Interval for Gini
sorted_G <- sort(G)
lower_ET <- sorted_G[round(Ndraw * 0.025)]
upper_ET <- sorted_G[round(Ndraw * 0.975)]
credible_interval_ET <- c(lower_ET, upper_ET)

# Plot the credible interval on Gini density
plot(density(G), main = "95% Equal-Tail Credible Interval for G")
abline(v = lower_ET, col = "red", lty = 2, lwd = 2)
abline(v = upper_ET, col = "blue", lty = 2, lwd = 2)
```

## 95% Equal-Tail Credible Interval for G



N = 10000 Bandwidth = 0.01058

```
# Print Equal-Tail Credible Interval
cat("95% Equal-Tail Credible Interval for Gini:\n")
```

```
## 95% Equal-Tail Credible Interval for Gini:
```

```
print(credible_interval_ET)
```

```
## [1] 0.1919565 0.5058270
```

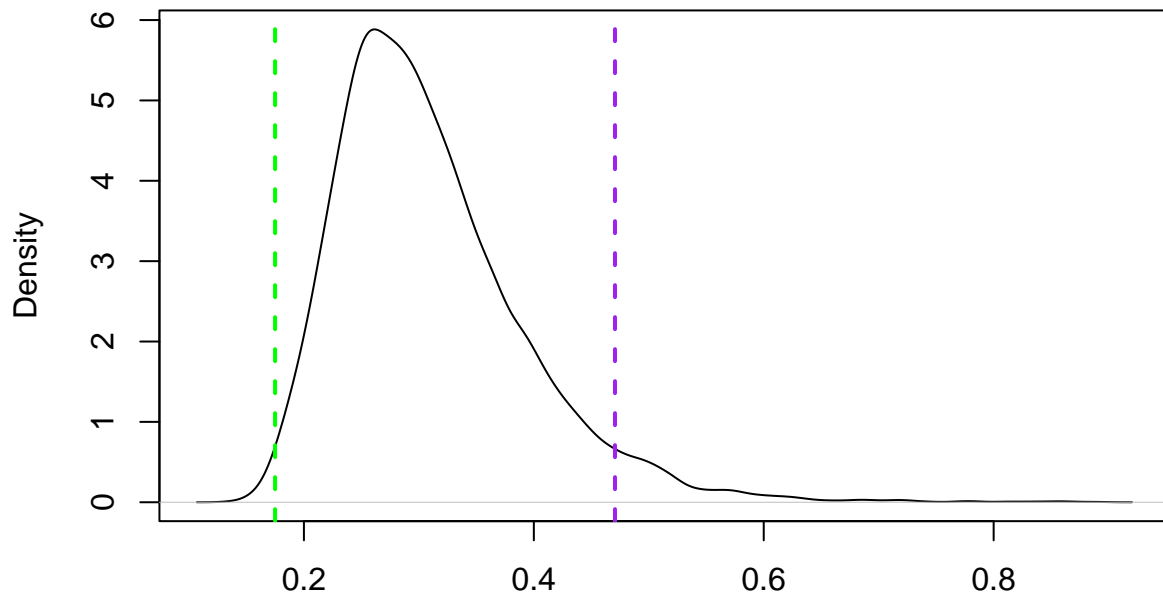
### Part D)

Compute a 95% Highest Posterior Density Interval (HPDI) for G. Also, compare the two intervals in (c) and (d).

```
# (d) Compute 95% Highest Posterior Density Interval (HPDI) using bayestestR
hpdInterval <- hdi(G, ci = 0.95)
```

```
# Plot HPDI on top of Gini density
plot(density(G), main = "95% HPDI for Gini Coefficient")
abline(v = hpdInterval$CI_low, col = "green", lty = 2, lwd = 2)
abline(v = hpdInterval$CI_high, col = "purple", lty = 2, lwd = 2)
```

## 95% HPDI for Gini Coefficient



N = 10000 Bandwidth = 0.01058

```
# Print HPDI Interval
cat("95% Highest Posterior Density Interval (HPDI) for Gini:\n")
```

```
## 95% Highest Posterior Density Interval (HPDI) for Gini:
```

```
print(hpdi_interval)
```

```
## 95% HDI: [0.17, 0.47]
```

```
# Comparison of intervals
cat("\nComparison:\n")
```

```
##
## Comparison:
```

```
cat("Equal-Tail Interval: ", round(credible_interval_ET[1], 4), "-", round(credible_interval_ET[2], 4),
```

```
## Equal-Tail Interval: 0.192 - 0.5058
```

```
cat("HPDI Interval      : ", round(hpdi_interval$CI_low, 4), "-", round(hpdi_interval$CI_high, 4), "\n")
```

```
## HPDI Interval      : 0.1749 - 0.4706
```

## QUESTION 3

### Part A)

Derive the expression for what the posterior pdf  $p(\lambda|y, \sigma)$  is proportional to. Then, plot the posterior distribution of the average goals rate parameter  $\lambda$  over a fine grid of  $\lambda$  values.

```
##### Question 3 #####

### Part A
# Data: Number of goals per match
y <- c(0, 2, 5, 5, 7, 1, 4)

# Parameters
sigma <- 5 # scale of half-normal prior
lambda_grid <- seq(0, 15, length.out = 1000) # Grid of lambda values

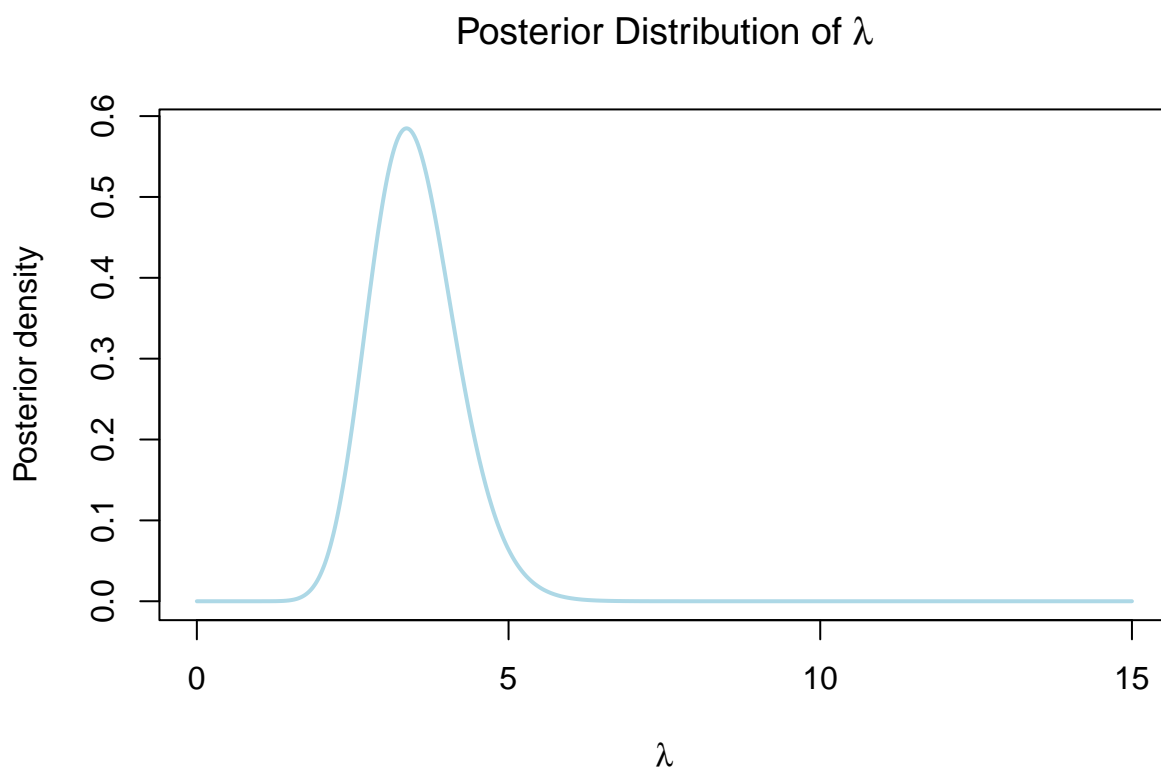
# Prior: PDF of Half-normal distribution
prior <- function(lambda, sigma) {
  sqrt(2) / (sigma * sqrt(pi)) * exp(-lambda^2 / (2 * sigma^2))
}

# Likelihood: Product of Poisson
likelihood <- function(lambda, y) {
  sapply(lambda, function(l) prod(dpois(y, lambda = l)))
}

# Derivation of unnormalized posterior over the grid
unnormalized_posterior <- likelihood(lambda_grid, y) * prior(lambda_grid, sigma)

# Normalize the posterior
posterior <- unnormalized_posterior / sum(unnormalized_posterior * diff(lambda_grid)[1])
# Here, In normalization we multiply by 'diff(lambda_grid)[1]' because it is the width
# between grid points. Instead of integrating, we approximate the integral as a sum over
# small intervals.

# Plot posterior
plot(lambda_grid, posterior, type = "l", lwd = 2, col = "lightblue",
      xlab = expression(lambda), ylab = "Posterior density",
      main = expression(paste("Posterior Distribution of ", lambda)))
```



#### Part B)

The posterior mode of  $\lambda$  from the information above is -

```
### Part B
# Find posterior mode (MAP estimate)
posterior_mode <- lambda_grid[which.max(posterior)]
cat("Posterior mode of lambda:", posterior_mode, "\n")
```

```
## Posterior mode of lambda: 3.363363
```