

Big Problems for **Small Networks**: Statistical Analysis of Small Networks and Team Performance

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Acknowledgements

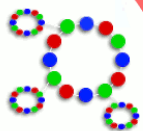


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Network Science of Teams

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Research Problem

Part I: Exponential Random Graph Models for Small Networks

Exponential Random Graph Models (ERGMs)



Figure 1: Friendship network of a UK university faculty. Source: **igraphdata** R package (Csardi, 2015). Figure drawn using the R package **netplot** ((<https://github.com/usccana/netplot>))

How can we explain what we see here?

ERGMs... the *lingua franca* of SNA

- ▶ Seeks to answer the question: What local social structures gave origin to a given observed graph?
- ▶ The model is centered around a vector of **sufficient statistics** $s()$, and is operationalized as:

$$\Pr(\mathbf{G} = \mathbf{g} \mid \theta, \mathbf{X}) = \frac{\exp\{\theta^t s(\mathbf{g}, \mathbf{X})\}}{\kappa(\theta, \mathbf{X})}, \quad \forall \mathbf{g} \in \mathcal{G} \quad (1)$$

Where $\kappa(\theta, \mathbf{X})$ is the normalizing constant and equals $\sum_{\mathbf{g}' \in \mathcal{G}} \exp\{\theta^t s(\mathbf{g}', \mathbf{X})\}$. Figure 2 shows some examples of values in $s()$.

- ▶ In the case of directed networks, \mathcal{G} has $2^{n(n-1)}$ terms.
- ▶ See Wasserman, Pattison, Robins, Snijders, Handcock and others.

Structures




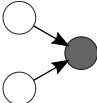

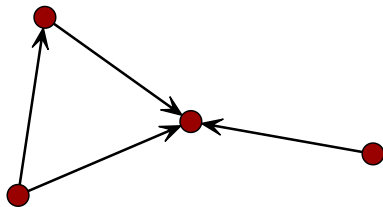
Representation	Description
	Mutual Ties (Reciprocity) $\sum_{i \neq j} y_{ij} y_{ji}$
	Transitive Triad (Balance) $\sum_{i \neq j \neq k} y_{ij} y_{jk} y_{ik}$
	Homophily $\sum_{i \neq j} y_{ij} \mathbf{1}(x_i = x_j)$
	Covariate Effect for Incoming Ties $\sum_{i \neq j} y_{ij} x_j$
	Four Cycle $\sum_{i \neq j \neq k \neq l} y_{ij} y_{jk} y_{kl} y_{li}$

Figure 2: Besides of the common edge count statistic (number of ties in a graph), ERGMs allow measuring other more complex structures that can be captured as sufficient statistics.

Example of model

In this network



We see 4 **edges**, 1 **transitive triad** and **no mutual ties**.

The probability function of this model would be

$$\Pr(\mathbf{G} = \mathbf{g} \mid \theta) = \frac{\exp \{ 4\theta_{edges} + \theta_{ttriads} + 0\theta_{mutual} \}}{\sum_{\mathbf{g}' \in \mathcal{G}} \exp \{ \theta^t s(\mathbf{g}') \}}$$

with $\theta = [\theta_{edges} \quad \theta_{ttriads} \quad \theta_{mutual}]^t$

This model has **MLE parameter estimates** of -0.19 (low density), 0.27 (high chance of ttriads), and -9.75 (low chance of mutuality) for the parameters edges, ttriads, and mutual respectively.

Estimation of ERGMs

- ▶ Calculating of the normalizing constant in (1), $\kappa = \sum_{\mathbf{g}' \in \mathcal{G}} \exp \{ \theta^t s(\mathbf{g}', \mathbf{X}) \}$, makes ERGMs difficult to estimate.
- ▶ For this reason, statistical methods have focused on avoiding the direct calculation of κ ; most modern methods for estimating ERGMs rely on MCMC.

Estimation of ERGMs (cont'd)

Description of the MCMC-MLE algorithm (Geyer and Thompson 1992)

1. Make an initial guess of the model parameters using MPL (maximum pseudo likelihood), $\hat{\theta}_0$
2. While the algorithm doesn't converge, do:
 - a. Generate a large sample of graphs from $\Pr(\mathbf{G} = \mathbf{g} \mid \hat{\theta}_0, \mathbf{X})$ using MCMC
 - b. Use the simulated sample of graphs to approximate the likelihood function. And with this approximation update the parameter θ_0 using a Newton-Raphson step.
 - c. next iteration

Once it converges, the last value of $\hat{\theta}_0$ is the MCMC-MLE estimates. The variance approximation is another story.

- ▶ While significant advances have been made in the area, simulation based models can suffer from **model degeneracy** (Handcock 2003).
- ▶ This occurs when the model parameters *live in a degenerate region*. This impacts directly the MCMC process yielding poor mixing

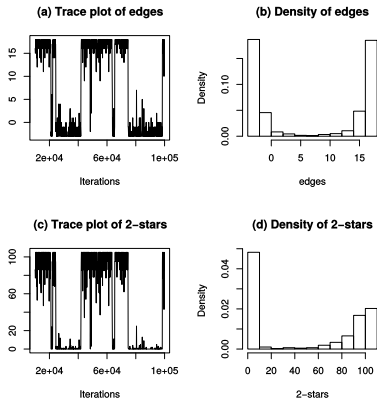
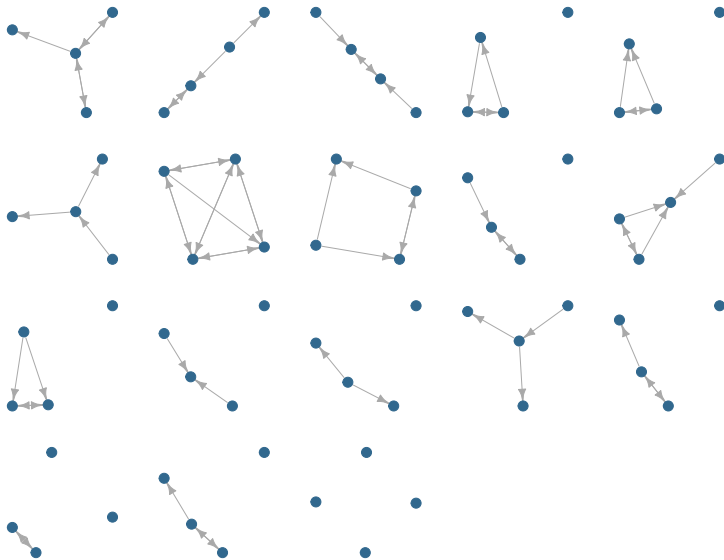


Figure 3: Model degeneracy. Figure 5 in Handcock (2003)

- ▶ Model degeneracy is particularly problematic with small networks. . . (says anyone who has tried to fit one).

ERGMs for Small Networks



ERGMs for Small Networks (cont'd)

- ▶ In the case of small networks (e.g. at most 6 nodes), the calculation of κ becomes computationally feasible.
- ▶ This allows direct calculation of (1), **avoiding the need for simulations** and allowing us to obtain Maximum Likelihood Estimates using *standard* optimizations techniques.
- ▶ In addition, if a common data generating process can be assumed, a pooled version of the ERGMs can be estimated easily:

$$\Pr(\mathbf{G}_1 = \mathbf{g}_1, \dots, \mathbf{G}_p = \mathbf{g}_p \mid \theta, \mathbf{X}_1, \dots, \mathbf{X}_p) = \prod_p \frac{\exp \{ \theta^t s(\mathbf{g}_p, \mathbf{X}_p) \}}{\kappa_p(\theta, \mathbf{X}_p)}$$

How different is this from the “normal” way to fit ERGMs?

Estimation of ERGMitos (cont'd)

Description of the MCMC-MLE algorithm (Geyer and Thompson 1992)

1. Make an initial guess of the model parameters using MPL (maximum pseudo likelihood), $\hat{\theta}_0$
2. While the algorithm doesn't converge, do:
 - a. Simulate B graphs from $\Pr(\mathbf{G} = \mathbf{g} \mid \hat{\theta}_0, \mathbf{X})$ using MCMC
 - b. Use the simulated sequence of graphs to approximate the likelihood function. And with this approximation **update the parameter θ_0 using a Newton-Raphson step.**
 - c. next iteration

By skipping the MCMC part we:

1. are able to get MLE estimates directly,
2. avoiding the degeneracy problem latent in MCMC, and
3. obtain more accurate estimates faster.

We have implemented this and more in the `ergmito` R package

An important pause. . .

ito, ita: From the latin *-īttus*. suffix in spanish used to denote small or affection.
e.g.:

*¡Que lindo ese perr**ito**! / What a beautiful little dog!*

*¿Me darías una tac**ita** de azúcar? / Would you give me a small cup of sugar?*

Special thanks to George Barnett who proposed the name during the 2018 NASN!

Features of `ergmito`

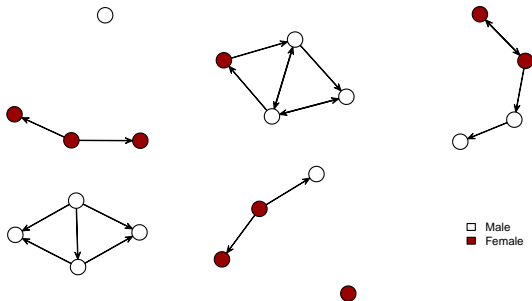
This (`lifecycle` `experimental`) R package has the following features

- ▶ Built on top of **statnet**'s `ergm` R package (Hunter et al. 2008; Handcock et al. 2018).
- ▶ Allows estimating ERGMs for small networks (less than 7 and perhaps 6)¹ via MLE.
- ▶ Implements pooled ERGM models.
- ▶ In the same spirit of the exhaustive enumeration, includes a simulation function for small networks sampling from the true distribution.

¹A directed graph of size 6 has a support set with $2^{6 \times (6-1)} = 1,073,741,824$ elements.

ergmito example

```
library(ergmito)
data(fivenets, package = "ergmito")
```



```
# Looking at one of the five networks  
fivenets[[1]]
```

```
## Network attributes:  
##   vertices = 4  
##   directed = TRUE  
##   hyper = FALSE  
##   loops = FALSE  
##   multiple = FALSE  
##   bipartite = FALSE  
##   total edges= 2  
##     missing edges= 0  
##     non-missing edges= 2  
##  
## Vertex attribute names:  
##   female name  
##  
## No edge attributes
```

So how can we fit this model?

ergmito example (cont'd)

The same as you would do with the `ergm` package:

```
(model1 <- ergmito(fivenets ~ edges + nodematch("female")))
```

```
##  
## ERGMito estimates  
##  
## Coefficients:  
##           edges  nodematch.female  
##          -1.705             1.587
```

	Model 1
edges	-1.70** (0.54)
nodematch.female	1.59* (0.64)
AIC	73.34
BIC	77.53
Log Likelihood	-34.67
# Networks	5
*** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$	

Table 1: Statistical models

Simulation Study

We conducted a simulation study to explore the properties of MLE for small networks (a.k.a. ERGMito). To generate each sample of teams:

1. Draw the **population parameters** from a piecewise Uniform with values in $[-4, -.1] \cup [.1, 4]$
2. We will draw **groups of sizes 3 to 5**. The number of networks per group size are drawn from a Poisson distribution with parameter 10 (hence, an expected size of 30 networks per sample).
3. Use the **drawn parameters** and **group sizes** to **generate random graphs** using an ERGM data generating process.

We simulated 100,000 samples, each one composed of an average of 30 networks.

Simulation Study (cont'd)

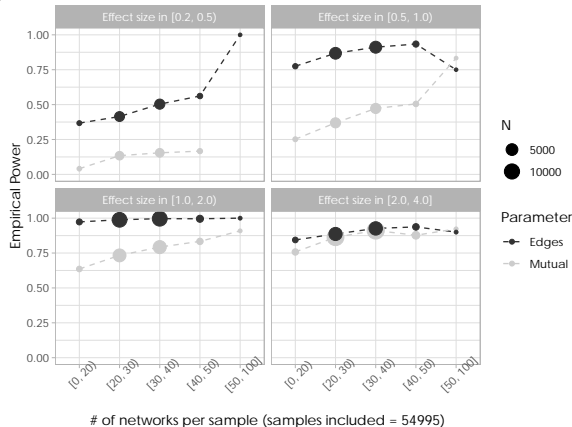


Figure 4: Empirical power of Pooled-ERGM estimates at various levels of effect size. As expected, power increases significantly with sample size (# of networks per sample). Interestingly, the discovery rate of an effect size within [1, 2) is very high even with a sample size of 20-30 networks. More extreme points have higher volatility due to small number of samples included.

Preliminary results

	nodeicov(RME)	nodeocov(Female)	RME+Female
Edges	-1.22*** (0.20)	-0.82*** (0.16)	-1.64*** (0.25)
Transitive Triads	0.23*** (0.05)	0.24*** (0.05)	0.24*** (0.06)
Indeg. RME	0.29*** (0.07)		0.35*** (0.08)
Outdeg. Female		0.45** (0.17)	0.53** (0.18)
AIC	772.98	711.79	693.08
BIC	786.06	724.59	710.14
Log Likelihood	-383.49	-352.90	-342.54
# Networks	41	38	38

*** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$

Table 2: Final set of specifications. The two statistics that showed the most invariant were **Indeg. RME** and **Outdeg. Female**. These two effects can be described as (1) individuals with high levels of RME receive more ties, and (2) female subjects were more likely of seeking advice than male. Other statistics such as GPA, religiousness, age, and ethnicity were not significant.

Part II: Association between group structure and team performance

Testing effects of social network structure on group performance

Two common approaches: Generalized Linear Models (GLMs), or permutation-like tests. Both have limitations:

- ▶ GLMs:
 - ▶ Sample size is problematic: How costly is getting enough teams to run get a desired level of power?
- ▶ Permutation-like tests:
 - ▶ Permutation: shuffle the dependent variable and compute, for example, correlations. **This assumes the outcome is completely detached from the observation.**
 - ▶ Rewiring (see Milo et al. 2004): Sample networks by generating new graphs using a rewiring algorithm, usually preserving some property such as degree sequence—the observed sequence of in/out degree. **This assumes an oversimplifying data generating process.** And worse, in a network of size 4, how many different networks can be observed **holding the degree sequence fixed?**

BTW Degree sequence \mapsto Scale-free networks

“

The **structural diversity of real-world networks** uncovered here presents both a puzzle and an opportunity. The strong focus in the scientific literature on **explaining and exploiting scale-free patterns** has meant **relatively less is known about mechanisms that produce non-scale-free structural patterns**, e.g., those with degree distributions better fitted by a log-normal. Two important directions of future work will be the **development and validation of novel mechanisms for generating more realistic degree structure in networks**, and novel statistical

”

techniques for identifying or untangling them given empirical data

– p. 8, Broido and Clauset (2019)

See Holme (2019) for a recent reference on the Scale-free issue.

An Idea

What about using ERGMs to generate null distributions?

- ▶ Even for small networks, we can generate thousands of unique graphs, so sample size is not a problem.
- ▶ As a difference from permutation tests:
 - ▶ ERGMs provide a more meaningful/realistic (or comprehensive if you like) data generating process
 - ▶ On average, sampled networks will **look like the original** graph (addressing Broido and Clauset (2019)'s comments).

In principle, this would be equivalent to revise a rewiring test. . .

Algorithm

1. Estimate an ERGM (estimates can come from a single graph or pooled estimates). We denote the data-generating-process of this model as $\mathcal{D} : \Theta \times \mathcal{X} \mapsto \mathcal{G}$.
2. Calculate the value $s_0 = s(\mathbf{G}, Y)$.
3. Now, for $b \in \{1, \dots, B\}$ do:
 - 3.1 For each group j in $\{1, \dots, J\}$, draw a new network $\mathbf{g}_j^b \sim \mathcal{D}(\hat{\theta}, X_j)$, this new sequence is denoted \mathbf{G}^b
 - 3.2 Using \mathbf{G}^b and Y , calculate $s_b = s(\mathbf{G}^b, Y)$
 - 3.3 Next b .

This will generate a null distribution for the statistic s , which we can use to compare against the observed statistic, s_0 .

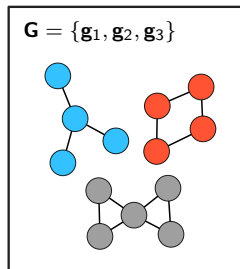
Note An important distinction to make is that structures that gave origin to the graph need not to be relevant for the team's performance per se.

Illustrated example

Suppose that we have a 3 networks of sizes 4, 4, and 5 respectively. The

Step 1:

Fit the ERGMito



Fit the ERGMito,
This will give us $\mathcal{D}(\hat{\theta}, X_j)$

Step 2:

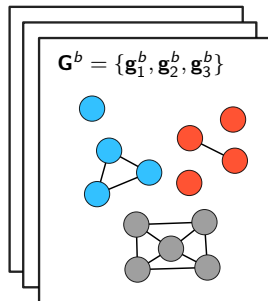
Calculate $s_0 =$

$$s \left(\begin{bmatrix} \text{blue network} \\ \text{red network} \\ \text{gray network} \end{bmatrix}, \begin{bmatrix} y_1 \\ y_2 \\ y_3 \end{bmatrix} \right)$$

Throughout the simulations
the only part that changes is
the networks, not Y

Step 3:

For $b \in 1, \dots, B$ do



3.1) For $j \in \{1, 2, 3\}$ draw a
new network from \mathcal{D}

3.2) Use the new sample to
calculate $s_b = s(\mathbf{G}^b, Y)$

We can use the distribution of the sequence $\{s_1, \dots, s_B\}$ as null to compare against s_0

Part III: An empirical example

Small teams performance

Data

Results

Concluding remarks

Exponential Random Graph Models for Small Networks:

1. Not a new thing, what's new is the tool to do so in a smooth way
2. Still work to do (on the development side of things): Goodness of fit tests, better algorithms for drawing random graphs, bayesian model
3. Can be extended to other types of ERGMs. . . our next target: TERGMs (Separable Exponential Random Graph Models)

Test for Association between graph level outcomes and graph structures:

1. Still need to run simulation studies
2. Key question 1: What happens when the graph structure is in both the ERGM and the stat
3. What about other properties such as type I error?

Thanks!



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Let's chat!

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