# CMS SPREAD OPTION PRICING WITH NORMAL AND SHIFTED LOGNORMAL DYNAMICS

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First Version October 29, 2015 - This Version October 29, 2015

ABSTRACT. We extend the cms spread option formula in [1], 13.16.2 to the case of shifted lognormal or normal swap rate dynamics.

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## 1. Original Lognormal Model

The original model in [1], 13.16.2 for cms spread option pricing is as follows. If  $S_i$ , i = 1, 2 are the underlying swap rates, t is the fixing time, T the payment time of the coupon with payoff  $(\phi \in \{-1, +1\})$ 

(1.1) 
$$\Pi(t) = \max(\phi(aS_1(t) + bS_2(t) - K), 0)$$

then under the T-forward measure we assume

$$dS_1 = \mu_1 S_1 dt + \sigma_1 S_1 dZ_1$$

$$dS_2 = \mu_2 S_2 dt + \sigma_2 S_2 dZ_2$$

$$(1.4) dZ_1 dZ_2 = \rho dt$$

[1] gives a formula for the price  $P(0,T)E^T(\Pi(t))$  involving a one dimensional integral as formula (13.34). The drifts  $\mu_i$  are implied from exogeneously given convexity adjustments for  $S_i$ , e.g. computed in a replication model on the respective underlying smiles).

Note that the formulation is [1] is a little more restrictive, because there it is assumed t = T. However, provided that the swap rate adjustments implying  $\mu_i$ , i = 1, 2 are computed w.r.T. the same fixing time t and payment time T > t, one can just replace the pricing expectation  $E^t$  by  $E^T$  without changing anything (inside the expectation a factor P(t,T)/P(t,T) = 1 occurs, which doesn't change anything).

Date: October 29, 2015.

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# 2. Shifted Lognormal extension

We look at the following straightforward extension of the model, introducing shifts for the underlying rates

$$(2.1) dS_1 = \mu_1(S_1 + d_1)dt + \sigma_1(S_1 + d_1)dZ_1$$

$$(2.2) dS_2 = \mu_2(S_2 + d_2)dt + \sigma_2(S_2 + d_2)dZ_2$$

$$(2.3) dZ_1 dZ_2 = \rho dt$$

with  $d_1, d_2 \geq 0$ , possibly different. Writing

$$(2.4) X_1 = (S_1 + d_1)$$

$$(2.5) X_2 = (S_2 + d_2)$$

$$(2.6) L = K + ad_1 + bd_2$$

we see that we can apply the original solution for 1.2 with underlyings  $X_1, X_2$  and the payoff written as  $(aX_1 + bX_2 - L)^+$ . Note that the computation of the drifts  $\mu_i$  change to

(2.7) 
$$\mu_i = \frac{\log((S_i(0) + d_i + c_i)/(S_i(0) + d_i))}{t} = \frac{\log((X_i(0) + c_i)/(X_i(0)))}{t}$$

for i = 1, 2 accordingly with  $c_i$  denoting the convexity adjustment applicable to rate  $S_i$ .

# 3. Normal extension

The normal flavour of the original model reads

$$dS_1 = \mu_1 dt + \sigma_1 dZ_1$$

$$dS_2 = \mu_2 dt + \sigma_2 dZ_2$$

$$(3.3) dZ_1 dZ_2 = \rho dt$$

with drifts now given by

with  $c_i$  again denoting the exogeneously given convexity adjustment for rate  $S_i$ , i = 1, 2.

The option price  $\nu$  is given by

(3.5) 
$$\nu = P(0,T)E^T \left( (\phi(aS_1(t) + bS_2(t) - K))^+ \right)$$

The expectation can, more explicity, written as

(3.6) 
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (\phi(as_1 + bs_2 - K))^+ p(s_1, s_2) ds_1 ds_2$$

with p denoting a bivariate normal distribution with expectation  $\mu$  and covariance matrix  $\Sigma$ 

(3.7) 
$$\mu = \begin{pmatrix} \mu_1 t \\ \mu_2 t \end{pmatrix}, \Sigma = \begin{pmatrix} \sigma_1^2 t & \rho \sigma_1 \sigma_2 t \\ \rho \sigma_1 \sigma_2 t & \sigma_2^2 t \end{pmatrix}$$

It is well known that the distribution of  $S_1(t)$  conditional on  $S_2(t) = s_2$  is given by

(3.8) 
$$S_1(t)|\{S_2(t) = s_2\} \sim \mathcal{N}\left(\mu_1 t + \frac{\rho \sigma_1}{\sigma_2}(s_2 - \mu_2 t), \sigma_1^2 t (1 - \rho^2)\right)$$

We denote the density of this distribution by  $p_{s_2}$  We continue to compute the inner integral in 3.6 for a fixed  $S_2(t) = s_2$ , i.e.

(3.9) 
$$\int_{-\infty}^{\infty} (\phi(as_1 + bs_2 - K))^+ p_{s_2}(s_1) ds_1$$

We have the following

**Lemma 1.** For  $\alpha, \beta \in \mathbb{R}$  we have

(3.10) 
$$\frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} (\alpha x + \beta)^{+} e^{-x^{2}/2} dx = \phi \frac{\alpha}{\sqrt{2\pi}} e^{-\beta^{2}/(2\alpha^{2})} + \beta \phi (1 - N(-\beta/\alpha))$$

where  $\phi = \text{sign}(\alpha) \in \{+1, -1\}$  and N denotes the cumulative normal distribution function. For  $\alpha = 0$  the right hand side simplifies to  $\beta^+$ .

*Proof.* Let  $\alpha > 0$ . The proof for  $\alpha < 0$  is similar and the case  $\alpha = 0$  is obvious from the following steps as well. Obviously  $\alpha x + \beta > 0$  iff  $x > -\beta/\alpha$ , so the integral can be written as

(3.11) 
$$\frac{1}{\sqrt{2\pi}} \int_{-\beta/\alpha}^{\infty} (\alpha x + \beta) e^{-x^2/2} dx$$

Furthermore,

(3.12) 
$$\int_{-\beta/\alpha}^{\infty} x e^{-x^2/2} dx = \left[ -e^{-x^2/2} \right]_{-\beta/\alpha}^{\infty} = e^{-\beta^2/(2\alpha^2)}$$

and

(3.13) 
$$\int_{-\beta/\alpha}^{\infty} e^{-x^2/2} dx = 1 - N(-\beta/\alpha)$$

which proves the identity.

THe integral 3.9 is

(3.14) 
$$\frac{1}{\sqrt{2\pi}}v\int_{-\infty}^{\infty}(\phi(as_1+bs_2-K))^+e^{-\frac{(s_1-\mu)^2}{2v^2}}ds_1$$

with  $\mu = \mu_1 t + \rho \sigma_1 / \sigma_2 (s_2 - \mu_2 t)$  and  $v^2 = \sigma_1^2 t (1 - \rho^2)$ . Substituting  $x = (s_1 - \mu)/v$  this becomes

(3.15) 
$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} (\phi a(\mu + vx) + bs_2 - K)^+ e^{-x^2/2} dx$$

Setting  $\alpha = \phi av$  and  $\beta = \phi a\mu + bs_2 - K$  and  $\psi = \text{sign}(\alpha)$  we can write the npv (using lemma 1)  $\nu$  as

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(3.16) 
$$P(0,T) \frac{1}{\sigma_2 \sqrt{2\pi t}} \int_{-\infty}^{\infty} \psi \left[ \frac{\alpha}{\sqrt{2\pi}} e^{-\beta^2/(2\alpha^2)} + \beta (1 - N(-\beta/\alpha)) \right] e^{-\frac{(s_2 - \mu_2 t)^2}{2\sigma_2^2 t}} ds_2$$
 with

$$(3.17) \alpha = \phi a \sigma_1 \sqrt{t(1-\rho^2)}$$

(3.18) 
$$\beta = \phi a(\mu_1 t + \rho \sigma_1 / \sigma_2 (s_2 - \mu_2 t)) + b s_2 - K$$

$$(3.19) \psi = \operatorname{sign}(\alpha)$$

A final substitution  $s = (s_2 - \mu_2 t)/(\sigma_2 \sqrt{t})$  yields

$$(3.20) \quad \nu = P(0,T) \frac{1}{\sigma_2 \sqrt{2\pi t}} \int_{-\infty}^{\infty} \psi \left[ \frac{\alpha}{\sqrt{2\pi}} e^{-\beta^2/(2\alpha^2)} + \beta (1 - N(-\beta/\alpha)) \right] e^{-s^2/2} ds$$
 with

$$(3.21) \alpha = \phi a \sigma_1 \sqrt{t(1-\rho^2)}$$

(3.22) 
$$\beta = \phi a(\mu_1 t + \rho \sigma_1 / \sigma_2^2 s \sqrt{t}) + b(s \sigma_2 \sqrt{t} + \mu_2 t) - K$$

$$(3.23) \psi = \operatorname{sign}(\alpha)$$

## References

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