

Two-party differential privacy and deterministic extraction from Santha-Vazirani sources

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Plan

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Differential privacy in client-server setting

For strings $x, y \in \{0, 1\}^n$, let $|x - y|_H$ denote Hamming distance. A mechanism M on $\{0, 1\}^n$ is a family of probability distributions $\{\mu_x : x \in \{0, 1\}^n\}$ on \mathbb{R} .

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Definition (Differential privacy)

The mechanism is ϵ -differentially private if for any x and y such that $|x - y|_H = 1$ and any measurable subset $S \subset \mathbb{R}$ we have

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Two-party differential privacy

- $VIEW_P^A(x, y) = (T, R_{AB}, R_A)$ random variable, where the probability space is public and private randomness of both parties.
- For each x , $VIEW_P^A(x, \cdot)$ is a mechanism over the y 's.
- $VIEW_P^B(\cdot, y)$ is defined similarly.

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Definition (Differential privacy for two-party protocols)

Protocol $P(x, y)$ has ϵ -differential privacy if the mechanism $VIEW_P^A(x, \cdot)$ is ϵ -differentially private for all values of x and same holds for $VIEW_P^B(\cdot, y)$ and all values of y .

Hamming distance (independent setting)

Question

Suppose Bob knows that Alice's x comes from *uniform* distribution X , independent of Bob's distribution Y . How can he approximate $|x - y|_H$ up to *an expected additive error $O(\sqrt{n})$* without any communication?

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Answer

- Just say $n/2$.
- W.l.o.g. assume that $y = 0^n$, then correct answer is $|x|_H$.
- $|x|_H$ is distributed by $B(n, 1/2)$.
- Using Hoeffding's inequality: $\Pr[||x|_H - n/2| > c\sqrt{n}] < 2e^{-c^2}$.

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Exercise: How to do better than this?

Hint: Use randomized response.

Santha-Vazirani sources

Definition (α -unpredictable bit source)

For $\alpha \in [0, 1]$, random variable $X = (X_1, \dots, X_n)$ taking values in $\{0, 1\}^n$ is an **α -unpredictable bit source** if for every $i \in [n]$, and every $x_1, \dots, x_{i-1} \in \{0, 1\}^{i-1}$, we have

$$\alpha \leq \frac{\Pr[X_i = 0 | X_1 = x_1, \dots, X_{i-1} = x_{i-1}]}{\Pr[X_i = 1 | X_1 = x_1, \dots, X_{i-1} = x_{i-1}]} \leq 1/\alpha$$

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Properties

- No string has probability mass greater than $1/(1 + \alpha)^n$
- Min-entropy ($\min_x \log_2(1/\Pr[X = x])$), is at least βn , where $\beta = \log_2(1 + \alpha) \geq \alpha$.

Santha-Vazirani sources

Definition (Strongly α -unpredictable bit source)

For $\alpha \in [0, 1]$, a random variable $X = (X_1, \dots, X_n)$ taking values in $\{0, 1\}^n$ is a **strongly α -unpredictable bit source** if for every $i \in [n]$, and every $x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n \in \{0, 1\}^{n-1}$, we have

$$\alpha \leq \frac{\Pr[X_i = 0 | X_1 = x_1, \dots, X_{i-1} = x_{i-1}, X_{i+1} = x_{i+1}, \dots, X_n = x_n]}{\Pr[X_i = 1 | X_1 = x_1, \dots, X_{i-1} = x_{i-1}, X_{i+1} = x_{i+1}, \dots, X_n = x_n]} \leq 1/\alpha$$

Two-party DP and Santha-Vazirani sources

Lemma (Two-party DP and Santha-Vazirani sources)

- Let $P(x, y)$ be a ϵ -differentially private randomized protocol with inputs $x, y \in \{0, 1\}^n$.
- Let X and Y be independent random variables uniformly distributed in $\{0, 1\}^n$.
- Let random variable $T(X, Y)$ denote the transcript on input (X, Y) .

Then for every $t \in \text{Supp}(T)$, the random variables $X|_{T=t}$ and $Y|_{T=t}$ are independent strongly $e^{-\epsilon}$ -unpredictable bit sources.

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Proof

① Independence

Proof by induction on the number of rounds.

② Strong $e^{-\epsilon}$ unpredictability (next slide)

Proof of strong $e^{-\epsilon}$ unpredictability

Using Bayes' Rule and the uniformity of X :

$$\begin{aligned}
 & \frac{\Pr[X_i = 0 | X_1 = x_1, \dots, X_{i-1} = x_{i-1}, X_{i+1} = x_{i+1}, \dots, X_n = x_n, T = t]}{\Pr[X_i = 1 | X_1 = x_1, \dots, X_{i-1} = x_{i-1}, X_{i+1} = x_{i+1}, \dots, X_n = x_n, T = t]} = \\
 &= \frac{\Pr[T = t | X_1 = x_1, \dots, X_{i-1} = x_{i-1}, X_i = 0, X_{i+1} = x_{i+1}, \dots, X_n = x_n]}{\Pr[T = t | X_1 = x_1, \dots, X_{i-1} = x_{i-1}, X_i = 1, X_{i+1} = x_{i+1}, \dots, X_n = x_n]} = \\
 &= \frac{\Pr[T(x_1 \cdots x_{i-1} 0 x_{i+1} \cdots x_n, Y) = t]}{\Pr[T(x_1 \cdots x_{i-1} 1 x_{i+1} \cdots x_n, Y) = t]}
 \end{aligned}$$

By ϵ -differential privacy the latter ratio is between $e^{-\epsilon}$ and e^{ϵ} .

Deterministic extraction from Santha-Vazirani sources

- Vazirani [Vaz87]: Inner product modulo 2 extracts an almost-uniform bit from two independent unpredictable sources
- Not possible for one source (no function can be more than α -unpredictable ([SV86])
Exercise: Prove this.
- Generalization[MMP+10]: Inner product modulo m extracts an almost-uniform element of \mathbb{Z}_m , if n is at least roughly m^2

δ -closeness

Definition (Statistical distance and δ -closeness)

For random variables X and X' taking values in Ω , we say that X and X' are δ -close if the statistical distance between their distributions is at most δ , i.e.,

$$\|X - X'\|_{SD} := \frac{1}{2} \sum_{x \in \Omega} |Pr[X = x] - Pr[X' = x]| \leq \delta$$

Deterministic extraction from Santha-Vazirani sources

Theorem (Randomness extraction)

There is a constant c such that if:

- ① X is an α -unpredictable bit source on $\{0, 1\}^n$,
- ② Y is a source on $\{0, 1\}^n$ with min-entropy at least βn ,
- ③ Y is independent from X ,
- ④ $Z = \langle X, Y \rangle \bmod m$ for some $m \in \mathbb{N}$,

then for every $\delta \in [0, 1]$, such that

$$n \geq c \cdot \frac{m^2}{\alpha\beta} \cdot \log\left(\frac{m}{\beta}\right) \cdot \log\left(\frac{m}{\delta}\right) \approx m^2 \log^2 m$$

the random variable (Y, Z) is δ -close to (Y, U) where U is uniform on \mathbb{Z}_m and independent of Y .

Bounding the magnitude of Fourier coefficients of Z

Lemma (Bounding Fourier coefficients)

Let Z be a random variable taking values in \mathbb{Z}_m .

Then the statistical distance between Z and the uniform distribution on \mathbb{Z}_m is at most

$$\frac{1}{2} \sqrt{\sum_{\omega \neq 1} |E[\omega^Z]|^2} = \frac{1}{2} \sqrt{\sum_{k=1}^{m-1} \left| \sum_{\ell=0}^{m-1} \Pr[Z = \ell] e^{-\frac{2\pi i}{m} k\ell} \right|^2},$$

where the sum is over all complex m 'th roots of unity ω other than 1.

Bounding the magnitude of Fourier coefficients of Z

Proof

Let $p_Z(\cdot), p_U(\cdot)$ be probability masses of Z and U .

$$\|Z - U\|_{SD} = \frac{1}{2} \|p_Z - p_U\|_1 \leq \frac{\sqrt{m}}{2} \|p_Z - p_U\|_2$$

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$$\|Z - U\|_{SD} = \frac{1}{2} \|p_Z - p_U\|_1 \leq \frac{\sqrt{m}}{2} \|p_Z - p_U\|_2$$

By Parseval's theorem ($\hat{p}_X(k)$ is k -th Fourier coefficient of DFT of X):

$$\frac{\sqrt{m}}{2} \|p_Z - p_U\|_2 = \frac{1}{2} \sqrt{\sum_{k=0}^{m-1} |\hat{p}_Z(k) - \hat{p}_U(k)|^2}$$

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Substituting Fourier coefficients $\hat{p}_Z(\cdot)$ and $\hat{p}_U(\cdot)$, the claim follows, i.e.

$$\|Z - U\|_{SD} \leq \frac{1}{2} \sqrt{\sum_{\omega \neq 1} |E[\omega^Z]|^2} = \frac{1}{2} \sqrt{\sum_{k=1}^{m-1} \left| \sum_{\ell=0}^{m-1} \Pr[Z = n] e^{-\frac{2\pi i}{m} k \ell} \right|^2}.$$

Proof of randomness extraction theorem

Proof (Randomness extraction theorem)

- X is α -unpredictable source on $\{0,1\}^n$, Y is βn -source on $\{0,1\}^n$.
- For every $\omega \neq 1$, let $BAD = \bigcup_{\omega} BAD_{\omega}$, where

$$BAD_{\omega} = \left\{ y \in \{0,1\}^n : \left| E \left[\omega^{\langle X, y \rangle} \right] \right| > \frac{\delta}{\sqrt{m}} \right\}.$$

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$$\begin{aligned} \|Z - U\|_{SD} &= (\|Z - U\|_{SD} | y \notin BAD) Pr(y \notin BAD) \\ &\quad + (\|Z - U\|_{SD} | y \in BAD) Pr(y \in BAD) \leq \\ &\leq (\|Z - U\|_{SD} | y \notin BAD) + Pr(y \in BAD). \end{aligned}$$

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- Using **Bounding Fourier coefficients** Lemma for every $y \notin BAD$, the statistical distance between $Z|_{Y=y} = \langle X, y \rangle \bmod m$ and the uniform distribution on \mathbb{Z}_m is at most $(1/2)\sqrt{(m-1) \cdot (\delta/\sqrt{m})^2} \leq \delta/2$.

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$$\begin{aligned} |BAD_\omega| &\leq \frac{\sum_{y \in \mathbb{Z}_2^n} |E[\omega^{\langle X, y \rangle}]|^{2t}}{(\delta/\sqrt{m})^{2t}} \leq \frac{\sum_{y \in \mathbb{Z}_m^n} |E[\omega^{\langle X, y \rangle}]|^{2t}}{(\delta/\sqrt{m})^{2t}} \leq \\ &\leq \frac{[1 + m \cdot \exp(-\Omega(\alpha t/m^2))]^n}{(\delta^2/m)^t} \leq \frac{2^{\beta n/2}}{(\delta^2/m)^t} \end{aligned}$$

for $t = \lceil c_0 \cdot (m^2/\alpha) \cdot \log(m/\beta) \rceil$ for a sufficiently large constant c_0 .

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for $t = \lceil c_0 \cdot (m^2/\alpha) \cdot \log(m/\beta) \rceil$ for a sufficiently large constant c_0 . So if $n \geq (2/\beta) \cdot (t \cdot \log(m/\delta^2) + \log(2m/\delta))$ (holds by hypothesis):

$$Pr[Y \in BAD_\omega] \leq 2^{-\beta n} \cdot |BAD_\omega| \leq \frac{2^{-\beta n/2}}{(\delta^2/m)^t} \leq \frac{\delta}{2m}.$$

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Theorem (Lower bound for inner product)

Let $P(x, y)$ be a randomized protocol with ϵ -differential privacy for inputs $x, y \in \{0, 1\}^n$, and let $\delta > 0$. Then with probability at least $1 - \delta$ over $x, y \leftarrow \{0, 1\}^n$ and the coin tosses of P , party B 's output differs from $\langle x, y \rangle$ by at least

$$\Delta = \Omega\left(\frac{\sqrt{n}}{\log n} \cdot \frac{\delta}{e^\epsilon}\right).$$

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Similar result for Hamming distance is implied, because

$$\langle x, y \rangle = |x|_H + |y|_H - |x - y|_H$$

Limitations of the extractor technique

Definition (Sensitivity)

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Theorem (Limitation of extractor technique)

Let $f: \{0, 1\}^n \times \{0, 1\}^n \rightarrow \mathbb{R}$ be a sensitivity-1 function. Then for any distribution μ such that for any input y , the conditional distribution $\mu(X|Y=y)$ is a product distribution $\prod_{i=1}^n \mu_i(X_i|Y=y)$, there is a function $g(y)$ such that $\Pr_{(x,y) \sim \mu}[|g(y) - f(x,y)| > t] \leq 2 \exp(-t^2/2n)$.

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Proof

For any $h: \{0, 1\}^n \leftarrow \mathbb{R}$ of sensitivity 1, and any product distribution ν on X ,

$$\Pr[|h(x) - E_{x \sim \nu}[h(x)]| > t] \leq 2 \exp(-t^2/2n).$$

Applying to $f(X, y)$ and $g(y) = E_{x \in \mu(X|Y=y)}[f(x, y)]$, the result follows.

References

- [MMP+10] Andrew McGregor, Ilya Mironov, Toniann Pitassi, Omer Reingold, Kunal Talwar, Salil Vadhan, "The Limits of Two-Party Differential Privacy"
(<http://research.microsoft.com/pubs/137029/2dplimits.pdf>)
- [Vaz87] U. V. Vazirni, "Strong communication complexity or generating quasirandom sequences from two communicating semirandom sources"
- [SV86] Miklos Santha, Umesh V. Vazirani, "Generating quasi-random sequences from semi-random sources"
- Ryan O'Donnell, "Analysis of Boolean functions"
(<http://www.cs.cmu.edu/~odonnell/boolean-analysis/>).
- Devdatt P. Dubhashi, Alessandro Pancones, "Concentration of Measure for the Analysis of Randomized Algorithms"