

CIS 700: “algorithms for Big Data”

Lecture 5: Dimension Reduction

Slides at <http://grigory.us/big-data-class.html>

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Today

- Dimensionality reduction
 - AMS as dimensionality reduction
 - Johnson-Lindenstrauss transform

L_p -norm Estimation

- Stream: m updates $(x_i, \Delta_i) \in [n] \times \mathbb{R}$ that define vector f where $f_j = \sum_{i:x_i=j} \Delta_i$.
- **Example:** For $n = 4$

$$\langle (1,3), (3, 0.5), (1,2), (2, -2), (2,1), (1, -1), (4,1) \rangle$$
$$f = (4, -1, 0.5, 1)$$

- L_p -norm: $\|f\|_p = (\sum_i |f_i|^p)^{\frac{1}{p}}$

L_p -norm Estimation

- L_p -norm: $\|f\|_p = (\sum_i |f|^p)^{\frac{1}{p}}$
- Two lectures ago:
 - $\|f\|_0 = F_0$ -moment
 - $\|f\|_2^2 = F_2$ -moment (via AMS sketching)
- Space: $O\left(\frac{\log n}{\epsilon^2} \log \frac{1}{\delta}\right)$
- Technique: linear sketches
 - $\|f\|_0$: $\sum_{i \in S} f_i$ for random sets S
 - $\|f\|_2^2$: $\sum_i \sigma_i f_i$ for random signs σ_i

AMS as dimensionality reduction

- Maintain a “linear sketch” vector

$$\mathbf{Z} = (Z_1, \dots, Z_k) = Rf$$

$$Z_i = \sum_{j \in [n]} \sigma_{ij} f_j, \text{ where } \sigma_{ij} \in_R \{-1, 1\}$$

- Estimator \mathbf{Y} for $\|f\|_2^2$:

$$\frac{1}{k} \sum_{i=1}^k Z_i^2 = \frac{\|Rf\|_2^2}{k}$$

- “Dimensionality reduction”: $x \rightarrow Rx$, “heavy” tail

$$\Pr \left[\left| \mathbf{Y} - \|f\|_2^2 \right| \geq c \left(\frac{2}{k} \right)^{\frac{1}{2}} \|f\|_2^2 \right] \leq \frac{1}{c^2}$$

Normal Distribution

- Normal distribution $N(0,1)$
 - Range: $(-\infty, +\infty)$
 - Density: $\mu(x) = (2\pi)^{-\frac{1}{2}} e^{-\frac{x^2}{2}}$
 - Mean = 0, Variance = 1
- Basic facts:
 - If X and Y are independent r.v. with normal distribution then $X + Y$ has normal distribution
 - $Var[cX] = c^2 Var[X]$
 - If X, Y are independent, then $Var[X + Y] = Var[X] + Var[Y]$

Johnson-Lindenstrauss Transform

- Instead of ± 1 let σ_i be i.i.d. random variables from normal distribution $N(0,1)$

$$Z = \sum_i \sigma_i f_i$$

- We still have $\mathbb{E}[Z^2] = \sum_i f_i^2 = \|f\|_2^2$ because:
 - $\mathbb{E}[\sigma_i]\mathbb{E}[\sigma_j] = 0$; $\mathbb{E}[\sigma_i^2] = \text{“variance of } \sigma_i \text{”} = 1$
- Define $\mathbf{Z} = (Z_1, \dots, Z_k)$ and define:

$$\|\mathbf{Z}\|_2^2 = \sum_j Z_j^2 \quad \left(\mathbb{E}[\|\mathbf{Z}\|_2^2] = k\|f\|_2^2 \right)$$

- **JL Lemma:** There exists $C > 0$ s.t. for small enough $\epsilon > 0$:

$$\Pr \left[\left| \|\mathbf{Z}\|_2^2 - k\|f\|_2^2 \right| > \epsilon k\|f\|_2^2 \right] \leq \exp(-C\epsilon^2 k)$$

Proof of JL Lemma

- **JL Lemma:** $\exists C > 0$ s.t. for small enough $\epsilon > 0$:
$$\Pr \left[\left| \|\mathbf{Z}\|_2^2 - k \|f\|_2^2 \right| > \epsilon k \|f\|_2^2 \right] \leq \exp(-C \epsilon^2 k)$$
- Assume $\|f\|_2^2 = 1$.
- We have $\mathbf{Z}_i = \sum_j \sigma_{ij} f_j$ and $\mathbf{Z} = (\mathbf{Z}_1, \dots, \mathbf{Z}_k)$
$$\mathbb{E} \left[\|\mathbf{Z}\|_2^2 \right] = k \|f\|_2^2 = k$$
- **Alternative form of JL Lemma:**
$$\Pr \left[\|\mathbf{Z}\|_2^2 > k(1 + \epsilon)^2 \right] \leq \exp(-\epsilon^2 k + O(k \epsilon^3))$$

Proof of JL Lemma

- Alternative form of JL Lemma:

$$\Pr \left[\|\mathbf{Z}\|_2^2 > k(1 + \epsilon)^2 \right] \leq \exp(-\epsilon^2 k + O(k \epsilon^3))$$

- Let $Y = \|\mathbf{Z}\|_2^2$ and $\alpha = k(1 + \epsilon)^2$
- For every $s > 0$ we have:

$$\Pr[Y > \alpha] = \Pr[e^{sY} > e^{s\alpha}]$$

- By Markov and independence of Z'_i s:

$$\Pr[e^{sY} > e^{s\alpha}] \leq \frac{\mathbb{E}[e^{sY}]}{e^{s\alpha}} = e^{-s\alpha} \mathbb{E} \left[e^{s \sum_i Z_i^2} \right] = e^{-s\alpha} \prod_{i=1}^k \mathbb{E} \left[e^{s Z_i^2} \right]$$

- We have $Z_i \in N(0,1)$, hence:

$$\mathbb{E} \left[e^{s Z_i^2} \right] = (2\pi)^{-\frac{1}{2}} \int_{-\infty}^{\infty} e^{s t^2} e^{-\frac{t^2}{2}} dt = \frac{1}{\sqrt{1 - 2s}}$$

Proof of JL Lemma

- Alternative form of JL Lemma:

$$\Pr \left[\|\mathbf{Z}\|_2^2 > k(1 + \epsilon)^2 \right] \leq \exp(-\epsilon^2 k + O(k \epsilon^3))$$

- For every $\mathbf{s} > 0$ we have:

$$\Pr[\mathbf{Y} > \alpha] \leq e^{-\mathbf{s}\alpha} \prod_{i=1}^k \mathbb{E} \left[e^{\mathbf{s}Z_i^2} \right] = e^{-\mathbf{s}\alpha} (1 - 2\mathbf{s})^{-\frac{k}{2}}$$

- Let $\mathbf{s} = \frac{1}{2} \left(1 - \frac{k}{\alpha} \right)$ and recall that $\alpha = k(1 + \epsilon)^2$
- A calculation finishes the proof:

$$\Pr[\mathbf{Y} > \alpha] \leq \exp(-\epsilon^2 k + O(k \epsilon^3))$$

Johnson-Lindenstrauss Transform

- Single vector: $k = O\left(\frac{\log \frac{1}{\delta}}{\epsilon^2}\right)$
 - Tight: $k = \Omega\left(\frac{\log \frac{1}{\delta}}{\epsilon^2}\right)$ [Woodruff'10]
- n vectors simultaneously: $k = O\left(\frac{\log n \log \frac{1}{\delta}}{\epsilon^2}\right)$
 - Tight: $k = \Omega\left(\frac{\log n \log \frac{1}{\delta}}{\epsilon^2}\right)$ [Molinaro, Woodruff, Y. '13]
- Distances between n vectors = $O(n^2)$ vectors:

$$k = O\left(\frac{\log n \log \frac{1}{\delta}}{\epsilon^2}\right)$$

Random Variables and Norms

- For a random variable \mathbf{X} and $p \geq 1$ let:

$$\|\mathbf{X}\|_p = \mathbb{E}[X^p]^{1/p}$$

Facts:

- For any c : $\|c\mathbf{X}\|_p = c\|\mathbf{X}\|_p$
- $\|\cdot\|_p$ is a norm (Minkowski's inequality)
- $\|\cdot\|_p \leq \|\cdot\|_q$ for $p \leq q$ (Monotonicity of norms)
- Jensen's inequality (used a lot for $F = |x|^p$):
If F is convex then $F(\mathbb{E}[\mathbf{X}]) \leq \mathbb{E}[F(\mathbf{X})]$

Khintchine Inequality

- [Khintchine] For $p \geq 1, x \in \mathbb{R}^n$ and σ_i i.i.d. Rademachers:

$$\left\| \sum_i \sigma_i x_i \right\|_p \leq \sqrt{p} \|x\|_2$$

- For r_i (either σ_i or $g_i \sim N(0,1)$) expand $\mathbb{E}[(\sum_i r_i x_i)^p]$
- All odd powers of r_i are zero
- All even moments for σ_i are 1, and for g_i are ≥ 1
- $\left\| \sum_i \sigma_i x_i \right\|_p \leq \left\| \sum_i g_i x_i \right\|_p$
- $\sum_i g_i x_i \sim N(0, \|x\|_2^2) \Rightarrow \left\| \sum_i g_i x_i \right\|_p \leq \sqrt{p} \|x\|_2$

Symmetrization

- [Symmetrization]: If Z_1, \dots, Z_n are independent and σ_i are i.i.d. Rademachers:

$$\left\| \sum_i Z_i - \mathbb{E} \sum_i Z_i \right\|_p \leq 2 \left\| \sum_i \sigma_i Z_i \right\|_p$$

- Let $Y_1 \dots Y_n$ be independent with the same distribution as Z_i
- $\left\| \sum_i Z_i - \mathbb{E} \sum_i Z_i \right\|_p = \left\| \sum_i Z_i - \mathbb{E}_Y \sum_i Y_i \right\|_p$
 $\leq \left\| \sum_i (Z_i - Y_i) \right\|_p$ (Jensen)
 $= \left\| \sum_i \sigma_i (Z_i - Y_i) \right\|_p$ ($Z_i - Y_i$ are independent and symmetric)
 $\leq 2 \left\| \sum_i \sigma_i Z_i \right\|_p$ (triangle inequality)

Decoupling

- Let x_1, \dots, x_n be independent with mean 0 and x'_1, \dots, x'_n identically distributed as x_i and independent of them. For any a_{ij} and $p \geq 1$:

$$\left\| \sum_{i \neq j} a_{ij} x_i x_j \right\|_p \leq 4 \left\| \sum_{i,j} a_{ij} x_i x'_j \right\|_p$$

- Let η_1, \dots, η_n be i.i.d. Bernoullis (0/1 w.p. 1/2):

$$\begin{aligned} \left\| \sum_{i \neq j} a_{ij} x_i x_j \right\|_p &= 4 \left\| \mathbb{E}_\eta \sum_{i \neq j} a_{ij} x_i x_j \eta_i (1 - \eta_j) \right\|_p \\ &\leq 4 \left\| \sum_{i \neq j} a_{ij} x_i x_j \eta_i (1 - \eta_j) \right\|_p \text{ (Jensen)} \end{aligned}$$

- There exists $\eta' \in \{0,1\}^n$ such that:

$$\left\| \sum_{i \neq j} a_{ij} x_i x_j \eta_i (1 - \eta_j) \right\|_p \leq \left\| \sum_{i \in S} \sum_{j \in \bar{S}} a_{ij} x_i x_j \right\|_p$$

where $S = \{i : \eta' = 1\}$.

Decoupling (continued)

Let x_S be an S -dimensional vector of x_i for $i \in S$.

$$\begin{aligned}
 & \bullet \left\| \sum_{i \in S} \sum_{j \in \bar{S}} a_{ij} x_i x_j \right\|_p = \left\| \sum_{i \in S} \sum_{j \in \bar{S}} a_{ij} x_i x'_j \right\|_p \\
 &= \left\| \mathbb{E}_{x_{\bar{S}}} \mathbb{E}_{x'_S} \sum_{i,j} a_{ij} x_i x'_j \right\|_p \quad (\mathbb{E}[x_i] = \mathbb{E}[x'_i] = 0) \\
 &\leq \left\| \sum_{i,j} a_{ij} x_i x'_j \right\|_p \text{ (Jensen)}
 \end{aligned}$$

• Overall:

$$\left\| \sum_{i \neq j} a_{ij} x_i x_j \right\|_p \leq 4 \left\| \sum_{i,j} a_{ij} x_i x'_j \right\|_p$$

Hanson-Wright Inequality

- For $\sigma_1, \dots, \sigma_n$ independent Rademachers and $A \in \mathbb{R}^{n \times n}$ real and symmetric for all $p \geq 1$:
$$\left| \left| \sigma^T A \sigma - \mathbb{E}[\sigma^T A \sigma] \right| \right|_p \leq \sqrt{p} \|A\|_F + p \|A\|$$
- Recall:

$$- \|A\|_F = \sqrt{\sum_{ij} a_{ij}^2} = \sqrt{\text{Tr}(A^T A)}$$

$$- \|A\| = \sup_{\{v \neq 0\}} \frac{\|Av\|_2}{\|v\|_2}$$

Hanson-Wright Inequality

- For $\sigma_1, \dots, \sigma_n$ independent Rademachers and $A \in \mathbb{R}^{n \times n}$ real and symmetric for all $p \geq 1$:

$$\left| \left| \sigma^T A \sigma - \mathbb{E}[\sigma^T A \sigma] \right| \right|_p \leq \sqrt{p} \|A\|_F + p \|A\|$$

$$\left| \left| \sigma^T A \sigma - \mathbb{E}[\sigma^T A \sigma] \right| \right|_p \leq \left| \left| \sigma^T A \sigma' \right| \right|_p \text{ (decoupling)}$$

$$\leq \sqrt{p} \left| \left| \left| A \sigma \right|_2 \right| \right|_p \text{ (Khintchine)}$$

$$= \sqrt{p} \left| \left| \left| A \sigma \right|_2^2 \right| \right|_{p/2}^{\frac{1}{2}}$$

$$\leq \sqrt{p} \left| \left| \left| A \sigma \right|_2^2 \right| \right|_p^{\frac{1}{2}} \text{ (monotonicity of norms)}$$

Hanson-Wright (continued)

$$\begin{aligned}
 \sqrt{p} \left\| \left\| |A\sigma|_2 \right\| \right\|_p &\leq \dots \leq \sqrt{p} \left\| \left\| |A\sigma|_2^2 \right\| \right\|_p^{\frac{1}{2}} \\
 &\leq \sqrt{p} \left(\mathbb{E} \left[\left\| |A\sigma|_2^2 \right\| \right] + \left\| \left\| |A\sigma|_2^2 \right\| - \mathbb{E} \left[\left\| |A\sigma|_2^2 \right\| \right] \right\|_p \right)^{\frac{1}{2}} \text{ (triangle ineq.)} \\
 &= \sqrt{p} \left(\|A\|_F^2 + \left\| \left\| |A\sigma|_2^2 \right\| - \mathbb{E} \left[\left\| |A\sigma|_2^2 \right\| \right] \right\|_p \right)^{\frac{1}{2}} \\
 &\leq \sqrt{p} \|A\|_F + \sqrt{p} \left\| \left\| |A\sigma|_2^2 \right\| - \mathbb{E} \left[\left\| |A\sigma|_2^2 \right\| \right] \right\|_p^{\frac{1}{2}} \\
 &\preccurlyeq \sqrt{p} \|A\|_F + \sqrt{p} \left| \sigma^T A^T A \sigma' \right|_p^{\frac{1}{2}} \text{ (decoupling)} \\
 &\preccurlyeq \sqrt{p} \|A\|_F + p^{\frac{3}{4}} \left\| \left\| A^T A \sigma \right\|_2 \right\|_p^{1/2} \text{ (Khintchine)} \\
 &\preccurlyeq \sqrt{p} \|A\|_F + p^{\frac{3}{4}} \|A\|^{\frac{1}{2}} \left\| \left\| Ax \right\|_2 \right\|_p^{\frac{1}{2}}
 \end{aligned}$$

Hanson-Wright (continued)

$$\sqrt{p} \left| \left| \left| A\sigma \right|_2 \right|_p \right| \leq \sqrt{p} \|A\|_F + p^{\frac{3}{4}} \|A\|^{\frac{1}{2}} \left| \left| \left| A\sigma \right|_2 \right|_p \right|^{\frac{1}{2}}$$

Let $E = \left| \left| \left| Ax \right|_2 \right|_p \right|^{\frac{1}{2}}$ then $E^2 - Cp^{\frac{1}{4}} \|A\|^{\frac{1}{2}} E - C\|A\|_F \leq 0$

- $E \leq$ larger root of the quadratic equation above
- $E^2 \leq \sqrt{p} \|A\|_F + p \|A\|$
- (Hanson-Wright) For $\sigma_1, \dots, \sigma_n$ independent Rademachers and $A \in \mathbb{R}^{n \times n}$ real and symmetric for all $p \geq 1$:

$$\left| \left| \sigma^T A \sigma - \mathbb{E}[\sigma^T A \sigma] \right|_p \right| \leq \sqrt{p} \|A\|_F + p \|A\|$$

Recap

- For a random variable X and $p \geq 1$ let:

$$\|X\|_p = \mathbb{E}[X^p]^{1/p}$$

- [Khintchine] For $p \geq 1, x \in \mathbb{R}^n$ and σ_i i.i.d. Rademachers:

$$\left\| \sum_i \sigma_i x_i \right\|_p \leq \sqrt{p} \|x\|_2$$

- [Symmetrization]: If Z_1, \dots, Z_n are independent and σ_i are i.i.d. Rademachers:

$$\left\| \sum_i Z_i - \mathbb{E} \sum_i Z_i \right\|_p \leq 2 \left\| \sum_i \sigma_i Z_i \right\|_p$$

- [Hanson-Wright] For $\sigma_1, \dots, \sigma_n$ independent Rademachers and $A \in \mathbb{R}^{n \times n}$ real and symmetric for all $p \geq 1$:

$$\left\| \sigma^T A \sigma - \mathbb{E}[\sigma^T A \sigma] \right\|_p \leq \sqrt{p} \|A\|_F + p \|A\|$$

Bernstein Inequality

- Let X_1, \dots, X_n be indep. r.v's such that $|X_i| \leq K$ almost surely and $\mathbb{E}[\sum_i X_i^2] \leq \sigma^2$. For all $p \geq 1$:

$$\left\| \sum_i X_i - \mathbb{E}[X_i] \right\|_p \leq \sigma \sqrt{p} + Kp$$

$$\left\| \sum_i X_i - \mathbb{E}[X_i] \right\|_p \leq 2 \left\| \sum_i \sigma_i X_i \right\|_p \text{ (symmetrization)}$$

$$\leq \sqrt{p} \left\| \left(\sum_i X_i^2 \right)^{\frac{1}{2}} \right\|_p \text{ (Khintchine)}$$

$$= \sqrt{p} \left\| \sum_i X_i^2 \right\|_{\frac{p}{2}}^{\frac{1}{2}}$$

$$\leq \sigma \sqrt{p} + \sqrt{p} \left\| \sum_i X_i^2 - \mathbb{E}[X_i^2] \right\|_p^{1/2} \text{ (triangle inequality)}$$

Bernstein Inequality (cont.)

$$\begin{aligned}
 \left\| \sum_i X_i - \mathbb{E}[X_i] \right\|_p &\leq \dots \preccurlyeq \sqrt{p} \left\| \left(\sum_i X_i^2 \right)^{\frac{1}{2}} \right\|_p \\
 &\leq \sigma \sqrt{p} + \sqrt{p} \left\| \left\| \sum_i X_i^2 - \mathbb{E}[X_i^2] \right\|_p^{\frac{1}{2}} \right\|_p \\
 &\leq \sigma \sqrt{p} + \sqrt{p} \left\| \left\| \sum_i \sigma_i X_i^2 \right\|_p^{\frac{1}{2}} \right\|_p \quad (\text{symmetrization}) \\
 &\preccurlyeq \sigma \sqrt{p} + p^{\frac{3}{4}} \left\| \left\| \sum_i (X_i^4)^{1/2} \right\|_p^{\frac{1}{2}} \right\|_p \quad (\text{Khintchine}) \\
 &\leq \sigma \sqrt{p} + p^{\frac{3}{4}} \sqrt{K} \left\| \left\| \sum_i (X_i^2)^{1/2} \right\|_p^{\frac{1}{2}} \right\|_p
 \end{aligned}$$

Bernstein Inequality (cont.)

- Let $E = \|(\sum_i X_i^2)^{\frac{1}{2}}\|_p$ then for some $C > 0$:

$$E^2 - Cp^{\frac{1}{4}}\sqrt{K}E - C\sigma \leq 0$$

- $E \geq$ larger root of this quadratic equation
- $E \leq \sigma\sqrt{p} + Kp$
- [Bernstein] Let X_1, \dots, X_n be indep. r.v's such that $|X_i| \leq K$ almost surely and $\mathbb{E}[\sum_i X_i^2] \leq \sigma^2$. For all $p \geq 1$:

$$\left\| \sum_i X_i - \mathbb{E}[X_i] \right\|_p \leq \sigma\sqrt{p} + Kp$$

Sparse Johnson-Lindenstrauss Transform

- Let $\Pi \in \mathbb{R}^{m \times n}$ be a JL-matrix where $m = O\left(\frac{1}{\epsilon^2} \log \frac{1}{\delta}\right)$ which satisfies for $\|x\|_2 = 1$:

$$\Pr_{\Pi} \left[\left| \|\Pi x\|_2^2 - 1 \right| \geq \epsilon \right] \leq \delta$$

- Takes $O\left(m \|x\|_0\right)$ time to compute JL
- Would be $O\left(s \|x\|_0\right)$ time Π only had s non-zero entries per column

Basic Sparse JL Transform

- Pick 2-wise indep. hash function $h : [n] \rightarrow [m]$
- Pick 4-wise indep. hash function $\sigma : [n] \rightarrow \{-1, 1\}$
- For each $i \in [n]$ let $\Pi_{h(i), i} = \sigma(i)$, the rest are 0
- [Thorup, Zhang'12]: This is JL if $m \gtrsim \frac{1}{\epsilon^2 \delta}$
- Best possible since $s = 1$
- Analysis: standard expectation/variance using bounded independence + Chebyshev
- To improve m let's use Hanson-Wright (higher moment than Chebyshev's second)

Sparse JL Transform: Construction

- $\Pi_{r,i} = \eta_{r,i}\sigma_{r,i}/\sqrt{s}$, where η_i are Bernoullis and σ_i are Rademachers
- For all r, i : $\mathbb{E}[\eta_{r,i}] = \frac{s}{m}$
- For all i : $\sum_i \eta_{r,i} = s$ (s non-zeros per column)
- $\eta_{r,i}$ are negatively correlated:

$$\mathbb{E} \left[\prod_{(r,i) \in S} \eta_{r,i} \right] \leq \prod_{(r,i) \in S} \mathbb{E}[\eta_{r,i}] = \left(\frac{s}{m} \right)^{|S|}$$

- Each column chosen uniformly from $\text{Binom}(m, s)$ columns of weight s works here

Sparse JL Transform: Analysis

Thm [KN'14]: If $m = O\left(\frac{1}{\epsilon^2} \log \frac{1}{\delta}\right)$ and $s \approx \epsilon m$:

$$\forall x: \|x\|_2 = 1, \Pr_{\Pi} \left[\left| \|\Pi x\|_2^2 - 1 \right| \geq \epsilon \right] \leq \delta$$

- $Z = \|\Pi x\|_2^2 - 1 =$

$$\frac{1}{s} \sum_{r=1}^m \sum_{i \neq j} \eta_{r,i} \eta_{r,j} \sigma_{r,i} \sigma_{r,j} x_i x_j \equiv \sigma^T A_{x,\eta} \sigma$$
- $A_{x,\eta}$ is a block-diagonal matrix with m blocks where r -th block is $\frac{1}{s} x^{(r)} (x^{(r)})^T$ but with zeros on the diagonal
- $x^{(r)}$ is a vector with entries $x_i^{(r)} = \eta_{r,i} x_i$
- By Hanson-Wright: $\|Z\|_p \leq \left\| \sqrt{p} \|A_{x,\eta}\|_F + p \|A_{x,\eta}\| \right\|_p$

$$\leq \sqrt{p} \| \|A_{x,\eta}\|_F \|_p + p \| \|A_{x,\eta}\| \|_p$$

Sparse JL Transform: Analysis

- (Operator norm) Since $A_{x,\eta}$ is block-diagonal $\|A_{x,\eta}\|$ is the largest norm of any block
- Eigenvalues in the r -th block are at most

$$\frac{2}{s} \max \left(\|x^{(r)}\|_2^2, \|x^{(r)}\|_\infty^2 \right) \leq \frac{2}{s}$$

- $\|A_{x,\eta}\| \leq \frac{2}{s}$

Sparse JL Transform: Analysis

- Define $Q_{i,j} = \sum_{r=1}^m \eta_{r,i} \eta_{r,j}$ so that:

$$\left\| A_{x,\eta} \right\|_F^2 = 1/s^2 \sum_{i \neq j} x_i^2 x_j^2 Q_{i,j}$$

- Lemma:** If $p \approx s^2 m$ then $\forall i, j \left\| Q_{i,j} \right\|_p \leq p$

$$\left\| \left\| A_{x,\eta} \right\|_F \right\|_p = \left\| \left\| A_{x,\eta} \right\|_F^2 \right\|_p^{\frac{1}{2}}$$

$$\leq \left\| \left| \frac{1}{s^2} \sum_{i \neq j} x_i^2 x_j^2 Q_{i,j} \right| \right\|_p^{\frac{1}{2}}$$

$$\leq \frac{1}{s^2} \sum_{i \neq j} x_i^2 x_j^2 \left\| Q_{i,j} \right\|_p \text{ (triangle ineq.)}$$

$$\leq 1/\sqrt{m}$$

Sparse JL Transform: Analysis

- By Markov ($m = O\left(\frac{1}{\epsilon^2} \log 1/\delta\right)$, $s \approx \epsilon m$, $p \approx \frac{s^2}{m}$):

$$\Pr\left[\left|\left|\Pi x\right|_2^2 - 1\right| > \epsilon\right] =$$

$$\Pr\left[\left|\sigma^T A_{x,\eta} \sigma\right|^p > \epsilon^p\right] \leq$$

$$\epsilon^{-p} \mathbb{E}\left[\left|\sigma^T A_{x,\eta} \sigma\right|^p\right] \text{ (Markov)}$$

$$\leq \epsilon^{-p} C^p \left(\frac{\sqrt{p}}{\sqrt{m}} + \frac{p}{s}\right)^p = \epsilon^{-p} C^p \left(\frac{1}{\epsilon} + \frac{1}{\epsilon}\right)^p \leq \delta$$

Sparse JL Transform: Analysis

- **Lemma:** If $p \approx s^2 m$ then $\forall i, j \left\| Q_{i,j} \right\|_p \leq p$
- Suppose $\eta_{a_1,i}, \dots, \eta_{a_s,i}$ are all 1 where $a_1 < \dots < a_s$.
- Note that $Q_{ij} = \sum_{t=1}^s Y_t$ where t is an indicator r.v. for the event $\eta_{a_t,j} = 1$.
- Y_t 's are not indep. but negatively correlated \Rightarrow p-th moment at most p-th moments of i.i.d. Bernoullis with expectation $\frac{s}{m}$ (expand $(\sum_t Y_t)^p$ and compare term by term)
- By Bernstein inequality:

$$\left\| Q_{ij} \right\|_p = \left\| \sum_t Y_t \right\| \leq \sqrt{\frac{s^2}{m}} \sqrt{p} + p \approx p$$