

CSCI B609:

“Foundations of Data Science”

Lecture 8: Faster Power Method and Applications of SVD

Slides at <http://grigory.us/data-science-class.html>

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Faster Power Method

- PM drawback: $A^T A$ is dense even for sparse A
- Pick random Gaussian \mathbf{x} and compute $B^{\mathbf{k}} \mathbf{x}$
- $\mathbf{x} = \sum_{i=1}^{\mathbf{d}} c_i \mathbf{v}_i$ (augment \mathbf{v}_i 's to o.n.b. if $r < \mathbf{d}$)
- $B^{\mathbf{k}} \mathbf{x} \approx (\sigma_1^{2\mathbf{k}} \mathbf{v}_1 \mathbf{v}_1^T) (\sum_{i=1}^{\mathbf{d}} c_i \mathbf{v}_i) = \sigma_1^{2\mathbf{k}} c_1 \mathbf{v}_1$
 $B^{\mathbf{k}} \mathbf{x} = (A^T A)(A^T A) \dots (A^T A) \mathbf{x}$
- **Theorem:** If \mathbf{x} is unit $\mathbb{R}^{\mathbf{d}}$ -vector, $|\mathbf{x}^T \mathbf{v}_1| \geq \delta$:
 - V = subspace spanned by \mathbf{v}_i 's for $\sigma_j \geq (1 - \epsilon)\sigma_1$
 - \mathbf{w} = unit vector after $\mathbf{k} = \frac{1}{2\epsilon} \ln \left(\frac{1}{\epsilon\delta} \right)$ iterations of PM

$\Rightarrow \mathbf{w}$ has a component at most ϵ orthogonal to V

Faster Power Method: Analysis

- $A = \sum_{i=1}^r \sigma_i \mathbf{u}_i \mathbf{v}_i^T$ and $\mathbf{x} = \sum_{i=1}^d c_i \mathbf{v}_i$
- $B^k \mathbf{x} = \sum_{i=1}^d \sigma_i^{2k} \mathbf{v}_i \mathbf{v}_i^T \sum_{j=1}^d c_j \mathbf{v}_j = \sum_{i=1}^d \sigma_i^{2k} c_i \mathbf{v}_i$

$$\|B^k \mathbf{x}\|_2^2 = \left\| \sum_{i=1}^d \sigma_i^{2k} c_i \mathbf{v}_i \right\|_2^2 = \sum_{i=1}^d \sigma_i^{4k} c_i^2 \geq \sigma_1^{4k} c_1^2 \geq \sigma_1^{4k} \delta^2$$

- (Squared) component orthogonal to V is

$$\sum_{i=m+1}^d \sigma_i^{4k} c_i^2 \leq (1 - \epsilon)^{4k} \sigma_1^{4k} \sum_{i=m+1}^d c_i^2 \leq (1 - \epsilon)^{4k} \sigma_1^{4k}$$

- Component of $\mathbf{w} \perp V \leq (1 - \epsilon)^{2k} / \delta \leq \epsilon$

Choice of \mathbf{x}

- \mathbf{y} random spherical Gaussian with unit variance
- $\mathbf{x} = \frac{\mathbf{y}}{\|\mathbf{y}\|_2}$:
$$\Pr \left[|\mathbf{x}^T \mathbf{v}| \leq \frac{1}{20\sqrt{d}} \right] \leq \frac{1}{10} + 3e^{-d/64}$$
- $\Pr \left[\|\mathbf{y}\|_2 \geq 2\sqrt{d} \right] \leq 3e^{-d/64}$ (Gaussian Annulus)
- $\mathbf{y}^T \mathbf{v} \sim N(0,1) \Rightarrow \Pr \left[\left| \mathbf{y}^T \mathbf{v} \right| \leq \frac{1}{10} \right] \leq \frac{1}{10}$
- Can set $\delta = \frac{1}{20\sqrt{d}}$ in the “faster power method”

Singular Vectors and Eigenvectors

- Right singular vectors are eigenvectors of $A^T A$
- σ_i^2 are eigenvalues of $A^T A$
- Left singular vectors are eigenvectors of AA^T
- $A^T A$ satisfies $\forall \mathbf{x}: \mathbf{x}^T B \mathbf{x} \geq 0$
 - $B = \sum_i \sigma_i^2 \mathbf{v}_i \mathbf{v}_i^T$
 - $\forall \mathbf{x}: \mathbf{x}^T \mathbf{v}_i \mathbf{v}_i^T \mathbf{x} = (\mathbf{x}^T \mathbf{v}_i)^2 \geq 0$
 - Such matrices are called positive semi-definite
- Any p.s.d matrix can be decomposed as $A^T A$

Application of SVD: Centering Data

- Minimize sum of squared distances from A_i to S_k
- **SVD**: best fitting S_k if data is centered
- What if not?
- **Thm.** S_k that minimizes squared distance goes through centroid of the point set:

$$\frac{1}{n} \sum A_i$$

- Will only prove for $k = 1$, analogous proof for arbitrary k (see textbook)

Application of SVD: Centering Data

- **Thm.** Line that minimizes squared distance goes through the centroid
- Line: $\ell = \mathbf{a} + \lambda \mathbf{v}$; distance $\text{dist}(\mathbf{A}_i, \ell)$
- $\|\mathbf{A}_i - \mathbf{a}\|_2^2 = \text{dist}(\mathbf{A}_i, \ell)^2 + \langle \mathbf{v}, \mathbf{A}_i \rangle^2$
- Center so that $\sum_{i=1}^n \mathbf{A}_i = \mathbf{0}$ by subtracting the centroid

$$\begin{aligned} \sum_i^n \text{dist}(\mathbf{A}_i, \ell)^2 &= \sum_{i=1}^n (\|\mathbf{A}_i - \mathbf{a}\|_2^2 - \langle \mathbf{v}, \mathbf{A}_i \rangle^2) \\ &= \sum_{i=1}^n (\|\mathbf{A}_i\|_2^2 + \|\mathbf{a}\|_2^2 - 2\langle \mathbf{A}_i, \mathbf{a} \rangle - \langle \mathbf{v}, \mathbf{A}_i \rangle^2) \\ &= \sum_{i=1}^n \|\mathbf{A}_i\|_2^2 + n\|\mathbf{a}\|_2^2 - 2\langle \sum_{i=1}^n \mathbf{A}_i, \mathbf{a} \rangle - \sum_{i=1}^n \langle \mathbf{v}, \mathbf{A}_i \rangle^2 \\ &= \sum_{i=1}^n \|\mathbf{A}_i\|_2^2 + n\|\mathbf{a}\|_2^2 - \sum_{i=1}^n \langle \mathbf{v}, \mathbf{A}_i \rangle^2 \end{aligned}$$

- Minimized when $\mathbf{a} = \mathbf{0}$

Principal Component Analysis

- $n \times d$ matrix: customers \times movies preference
- $n = \# \text{customers}$, $d = \# \text{movies}$
- A_{ij} = how much customer i likes movie j
- Assumption: A_{ij} can be described with k factors
 - Customers and movies: vectors in \mathbf{u}_i and $\mathbf{v}_j \in \mathbb{R}^k$
 - $A_{ij} = \langle \mathbf{u}_i, \mathbf{v}_j \rangle$
- Solution: $A \approx U V^T$

$\begin{pmatrix} \vdots \\ \vdots \\ \vdots \end{pmatrix}$

customers

$\begin{pmatrix} \vdots \\ \vdots \\ \vdots \end{pmatrix}$

factors

$\begin{pmatrix} \vdots \\ \vdots \\ \vdots \end{pmatrix}$

movies

$$\begin{matrix} & & \text{factors} \\ \text{customers} & \begin{pmatrix} A \end{pmatrix} & = & \begin{pmatrix} U \end{pmatrix} \begin{pmatrix} \text{movies} \\ V \end{pmatrix} \end{matrix}$$

Separating mixture of k Gaussians

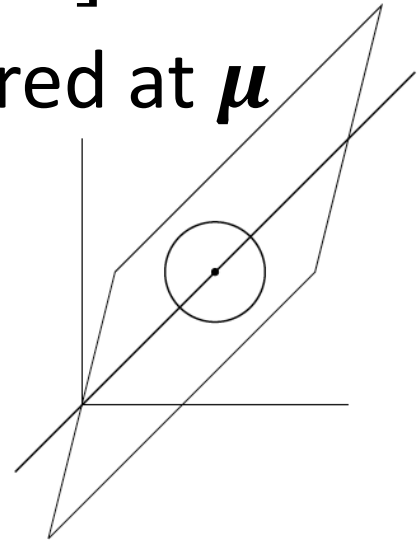
- **Sample origin problem:**
 - Given samples from k **well-separated** spherical Gaussians
 - **Q:** Did they come from the same Gaussian?
- δ = distance between centers
- For two Gaussians naïve separation requires
$$\delta > \omega(d^{1/4})$$
- **Thm.** $\delta = \Omega(k^{\frac{1}{4}})$ suffices
- **Idea:**
 - Project on a k -dimensional subspace through centers
 - **Key fact:** This subspace can be found via SVD
 - Apply naïve algorithm

Separating mixture of k Gaussians

- **Easy fact:** Projection preserves the property of being a unit-variance spherical Gaussian
- **Def.** If p is a probability distribution, **best fit line** $\{c\mathbf{v}, c \in \mathbb{R}\}$ is:

$$\mathbf{v} = \operatorname{argmax}_{|\mathbf{v}|=1} \mathbb{E}_{\mathbf{x} \sim p} \left[(\mathbf{v}^T \mathbf{x})^2 \right]$$

- **Thm:** Best fit line for a Gaussian centered at $\boldsymbol{\mu}$ passes through $\boldsymbol{\mu}$ and the origin



Best fit line for a Gaussian

- **Thm:** Best fit line for a Gaussian centered at μ passes through μ and the origin

$$\begin{aligned}\mathbb{E}_{x \sim p} \left[(v^T x)^2 \right] &= \mathbb{E}_{x \sim p} \left[(v^T (x - \mu) + v^T \mu)^2 \right] \\&= \mathbb{E}_{x \sim p} \left[v^T (x - \mu)^2 + 2(v^T \mu) v^T (x - \mu) + (v^T \mu)^2 \right] \\&= \mathbb{E}_{x \sim p} [v^T (x - \mu)^2] + 2(v^T \mu) \mathbb{E}_{x \sim p} [v^T (x - \mu)] + (v^T \mu)^2 \\&= \mathbb{E}_{x \sim p} [v^T (x - \mu)^2] + (v^T \mu)^2 \\&= \sigma^2 + (v^T \mu)^2\end{aligned}$$

- Where we used:

$$- \mathbb{E}_{x \sim p} [v^T (x - \mu)] = 0$$

$$- \mathbb{E}_{x \sim p} [v^T (x - \mu)^2] = \sigma^2$$

- Best fit line maximizes $(v^T \mu)^2$