Topics to be included in this tutorial are proportion estimator and ratio estimator for simple random sampling.

## 1. Proportion estimator

The sample proportion is basically a special case of sample mean when the variable of interest  $y_i \in \{0,i\}$ . Therefore, all the theory and results of sample mean under SRS still hold for the sample proportion. Besides proportion  $\overline{Y} = \frac{1}{N} \sum_{i=1}^{N} y_i$ , sometimes we may also be interested in the number of units in the population with certain property, i.e.,  $T = \sum_{i=1}^{N} y_i$ . The estimators and Standard errors are given in the table below.

- parameter of interest: 
$$\ddot{\gamma} = \frac{1}{N} \sum_{i=1}^{N} y_i$$

- estimator:  $\hat{p} = \frac{1}{n} \sum_{i=1}^{N} y_i$ 

- variance of the estimator:  $Var(\hat{p}) = (1 - \frac{n}{N}) \frac{S^2}{n}$ ,  $S^2 = \frac{1}{N-1} \sum_{i=1}^{N} (y_i - \tilde{r})^2$ 

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Example. Suppose that the governor wants to estimate the total number of people infected by a certain disease in a certain area. To this end, he takes an SRS of size 1000 and finds that 120 among them are infected. Try to give a point estimate as well as a 95% CI. for the total number of infected people in this area. (It's known that the population in this area is 10°).

Solution. Let  $y_i = \begin{cases} 1 & \text{if unit i is infected}, \\ 0 & \text{otherwise} \end{cases}$  (i=1,..., N=10<sup>5</sup>), then we estimate  $T = \sum_{j=1}^{N} y_i$  by  $\hat{T} = N \hat{p} = N \cdot \frac{1}{n} \sum_{i \in N} y_i = 10^5 \times \frac{1.2 \times 10^4}{10^3} = 1.2 \times 10^4$ . The 95% C.I. is given by  $\hat{T} \pm Z_{0.005} \times N \sqrt{(1-\frac{N}{N}) \frac{\hat{p}(1-\hat{p})}{n-1}} = 12000 \pm 1.96 \times 10^5 \sqrt{0.99 \times \frac{0.12 \times 0.98}{999}} \approx [9995, 14805]$ .

## 2. Ratio estimator

In class, we have learned that the ratio estimator can be used to improve our estimation for  $\bar{\gamma}$  if the variable of interest y: and the auxiliary variable x: are highly correlated. In this case,  $\hat{\Gamma}_r = \bar{\chi} \cdot \frac{\bar{y}}{\bar{\chi}}$  with  $\sec(\hat{\bar{Y}}_r) = \sqrt{(1-\frac{n}{N})} \cdot \frac{\hat{S}_r^2}{n}$ ,  $\hat{S}_r^2 = \frac{1}{N-1} \cdot \frac{1}{12N} (y_1 - Bx_1)^{\frac{1}{N}}$ . Sometimes in practice, it is the ratio itself  $R = \frac{\sum_{i=1}^{N} y_i^2 x_i}{\sum_{i=1}^{N} x_i}$  of interest and we can use the ratio estimator  $\hat{R} = \frac{\bar{y}}{\bar{y}}$  to estimate R. The standard error of  $\hat{R}$  is given by  $\sec(\hat{R}) = \frac{1}{N} \sqrt{(1-\frac{n}{N}) \cdot \frac{\hat{S}_r^2}{n}}$ . As usual, a  $(1-\alpha)-c.i.$  for R is  $\hat{R} \pm Z_N \cdot se(\hat{R})$ .

Example. Suppose the population consists of agricultural fields of different sizes. Let  $Y_i = bushels$  of grain harvested in field i

Xi = acreage of field i

We want to estimate the average yield per acre, i.e.  $R = \frac{T}{X}$ . Suppose we have selected an SRS of size 6 and the numbers are recorded as below.

Then we can estimate the average yield per acre by  $\hat{R} = \sqrt[3]{x} = \frac{12\pi}{16\pi} \frac{1}{16\pi} \frac{1}{16\pi} = 79.18$ and a 95% CI is given by  $\hat{R} \pm \frac{2}{3} \frac{1}{2} \cdot \frac{1}{3} \cdot \frac{1$