

STAT 403 Tutorial - Week 6

For this week's tutorial, I'm going to give a numerical example for the randomized block design. Then I'll take a review of what we've learned so far.

Example. Four different washing solutions are being compared to study their effectiveness in retarding bacteria growth. Suppose that only four trials can be run on a single day and observations (amount of bacteria after applying each washing solution) are taken for three days, as shown below.

	solution 1	solution 2	solution 3	solution 4	average
day 1	33 (Y_{11})	64 (Y_{21})	32 (Y_{31})	64 (Y_{41})	48.25 ($\bar{Y}_{..1}$)
day 2	15 (Y_{12})	26 (Y_{22})	17 (Y_{32})	25 (Y_{42})	20.75 ($\bar{Y}_{..2}$)
day 3	23 (Y_{13})	38 (Y_{23})	16 (Y_{33})	37 (Y_{43})	28.50 ($\bar{Y}_{..3}$)
average	23.67 ($\bar{Y}_{.1}$)	42.67 ($\bar{Y}_{.2}$)	21.67 ($\bar{Y}_{.3}$)	42 ($\bar{Y}_{.4}$)	32.5 ($\bar{Y}_{..}$)

(i) Suppose that each day is a block. - Assess whether four solutions are the same.

(ii) Answer the same question as in (i) if days are not blocks.

Solution:

(i). Denote the observation of solution i on day j by Y_{ij} ($i=1,2,3,4$).

Then we have model $Y_{ij} = \mu + \alpha_i + \beta_j + \epsilon_{ij}$, $\epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$.

Where α_i 's represent treatment effects and β_j 's represent block effects.

The test of hypothesis $H_0: \alpha_1 = \alpha_2 = \alpha_3 = \alpha_4$ v.s. $H_a: \alpha_i$'s not same

is based on

$$F \triangleq \frac{SStreat/(t-1)}{SSerr/(N-b-t-1)} \stackrel{H_0}{\sim} F_{t-1, N-b-t-1}$$

Note here $N = bt = 3 \times 4 = 12$, $b = 3$, $t = 4$.

$$SS_{\text{trt}} = 3 [(\bar{Y}_{1.} - \bar{Y}_{..})^2 + (\bar{Y}_{2.} - \bar{Y}_{..})^2 + (\bar{Y}_{3.} - \bar{Y}_{..})^2 + (\bar{Y}_{4.} - \bar{Y}_{..})^2] = 1167$$

$$SS_{\text{blk}} = 4 [(\bar{Y}_{.1} - \bar{Y}_{..})^2 + (\bar{Y}_{.2} - \bar{Y}_{..})^2 + (\bar{Y}_{.3} - \bar{Y}_{..})^2] = 1608.5$$

$$SS_{\text{total}} = \sum_{i=1}^4 \sum_{j=1}^3 (Y_{ij} - \bar{Y}_{..})^2 = 3043$$

$$SS_{\text{err}} = SS_{\text{total}} - SS_{\text{trt}} - SS_{\text{blk}} = \sum_{i=1}^4 \sum_{j=1}^3 (Y_{ij} - \bar{Y}_{i.} - \bar{Y}_{.j} + \bar{Y}_{..})^2 = 267.5$$

$$\Rightarrow F_{\text{obs}} = \frac{1167/3}{267.5/6} = 8.73$$

$$p\text{-value} = P(F \geq F_{\text{obs}}) = P(F \geq 8.73) = 0.013 < 0.05$$

Therefore we reject the null hypothesis ^{at 5% level} that the 4 solutions have same effects.

ANOVA table:

Source	SS	df
treatment	1167	t-1=3
block	1608.5	b-1=2
error	267.5	(b-1)(t-1)=6
total	3043	N-1=11

(ii). If days were not used as block factors, then the model

becomes $Y_{ij} = \mu + \alpha_i + \epsilon_{ij}$, $\epsilon_{ij} \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$, which is essentially

the model for complete randomized design. Compare the calculations

We'll see

Recall that the testing statistic becomes $F' = \frac{SS_{\text{trt}}/(t-1)}{SS_{\text{err}}/(N-t)} \sim F_{t-1, N-t}$

$$SS'_{\text{trt}} = SS_{\text{trt}} = 1167$$

$$SS'_{\text{tot}} = SS_{\text{tot}} = 3043$$

$$SS'_{\text{err}} = SS'_{\text{tot}} - SS'_{\text{trt}} = 1876$$

$$\Rightarrow F'_{\text{obs}} = \frac{1167/4-1}{1876/12-4} = 1.66$$

$$p\text{-value} = P(F' \geq F'_{\text{obs}}) = 0.252 > 0.05$$

\Rightarrow do not reject at 5% level

ANOVA table:

Source	SS	df
treatment	1167	3
error	1876	8
total	3043	11

Review of STAT 403 from WEEK 1 to WEEK 5.

- One-sample problem : $X_1, \dots, X_n \sim N(\mu, \sigma^2)$ $\left\{ \begin{array}{l} H_0: \mu = 0 \text{ v.s. } H_a: \mu \neq 0, (\mu \geq 0) \\ \text{C.I. for } \mu \end{array} \right.$

- Two-sample problem :

- non-paired $X_1, \dots, X_n \sim N(\mu_1, \sigma^2)$
 $Y_1, \dots, Y_m \sim N(\mu_2, \sigma^2)$

- ① $H_0: \mu_1 = \mu_2$ v.s. $H_a: \mu_1 \neq \mu_2$ $\left\{ \begin{array}{l} t \text{ test} \\ \text{randomization test.} \end{array} \right.$
- ② C.I. for $\mu_1 - \mu_2$

- paired version : $X_1 - Y_1, \dots, X_n - Y_n \sim N(\mu_1 - \mu_2, \sigma^2)$

- ① $H_0: \mu_1 = \mu_2$ v.s. $H_a: \mu_1 \neq \mu_2$ $\left\{ \begin{array}{l} \text{paired } t \text{ test} \\ \text{paired randomization test} \end{array} \right.$
- ② C.I. for $\mu_1 - \mu_2$.

- Multiple-sample (treatments) problem.

- complete randomized designs. $Y_{ij} = \mu_i + \varepsilon_{ij}$, $\varepsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$ $\left(\begin{array}{l} i=1, \dots, t \\ j=1, \dots, n_i \end{array} \right)$

- ① Assess $\mu_1 = \mu_2 = \dots = \mu_t$: F-test & ANOVA table.

- ② Contrasts & their inferences $\theta = c_1\mu_1 + \dots + c_t\mu_t$ ($c_1 + \dots + c_t = 0$).

- randomized block designs. $Y_{ij} = \mu + \alpha_i + \beta_j + \varepsilon_{ij}$, $\varepsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$ $\left(\begin{array}{l} i=1, \dots, t \\ j=1, \dots, b \end{array} \right)$.

- ① F-test & ANOVA table.