

## 1 Basic Properties

1.  $E(X) = \sum xp(x)$
2.  $Var(X) = \sum (x - \mu)^2 f(x)$
3. X is around  $E(X)$ , give or take  $SD(X)$
4.  $E(aX + bY) = aE(X) + bE(Y)$
5.  $Var(aX + bY) = a^2 Var(X) + b^2 Var(Y)$
6.  $Var(X) = E(X^2) - [E(X)]^2$
7.  $Cov(X_1, X_2) = E(X_1 X_2) - E(X_1)E(X_2)$
8.  $P(AB) = P(A)P(B)$  if A and B independent
9. RV is centered when  $E(X) = 0$ , and any RV can be centered via  $Y = X - E(X)$ , with SD and variance unaffected
10. In  $X = \mu + \epsilon$ ,  $\mu$  is the unknown constant of interest, and  $\epsilon$  represents random measurement error.
11. if  $X, Y$  are independent:
  - (a)  $M_{X+Y}(t) = M_X(t)M_Y(t)$
  - (b)  $E(XY) = E(X)E(Y)$ , converse is true if  $X$  and  $Y$  are bivariate normal, extends to multivariate normal

## 2 Approximations

### 2.1 Law of Large Numbers

Let  $X_1, X_2, \dots, X_n$  be IID, with expectation  $\mu$  and variance  $\sigma^2$ .  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow[n]{\infty} \mu$ . Let  $x_1, x_2, \dots, x_n$  be realisations of the random variable  $X_1, X_2, \dots, X_n$ , then  $\bar{x}_n = \frac{1}{n} \sum_{i=1}^n x_n \xrightarrow[n]{\infty} \mu$

### 2.2 Central Limit Theorem

Let  $S_n = \sum_{i=1}^n X_i$  where  $X_1, X_2, \dots, X_n$  IID.  $\frac{S_n - n\mu}{\sqrt{n}\sigma} \xrightarrow[n]{\infty} \mathcal{N}(0, 1)$

## 3 Distributions

### 3.1 Normal $X \sim \mathcal{N}(\mu, \sigma^2)$

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right), -\infty < x < \infty$$

1. When  $\mu = 0$ ,  $f(x)$  is an even function, and  $E(X^k) = 0$  where  $k$  is odd
2.  $Y = \frac{X - E(X)}{SD(X)}$  is the standard normal

### 3.2 Gamma $\Gamma$

$$g(t) = \frac{\lambda^\alpha}{\Gamma(\alpha)} t^{\alpha-1} e^{-\lambda t}, t \geq 0$$

### 3.3 $\chi^2$ Distribution

Let  $\mathcal{Z} \sim \mathcal{N}(0, 1)$ ,  $\mathcal{U} = \mathcal{Z}^2$  has a  $\chi^2$  distribution with 1 d.f.

$$f_{\mathcal{U}}(u) = \frac{1}{\sqrt{2\pi}} u^{-\frac{1}{2}} e^{-\frac{u}{2}}, u \geq 0$$

$$\chi_1^2 \sim \Gamma(\alpha = \frac{1}{2}, \lambda = \frac{1}{2})$$

Let  $U_1, U_2, \dots, U_n$  be  $\chi_1^2$  IID, then  $V = \sum_{i=1}^n U_i$  is  $\chi_n^2$  with  $n$  degree freedom,  $V \sim \Gamma(\alpha = \frac{n}{2}, \lambda = \frac{1}{2})$   
 $E(\chi_n^2) = n, Var(\chi_n^2) = 2n$   
 $M(t) = (1 - 2t)^{-\frac{n}{2}}$

### 3.4 t-distribution

Let  $\mathcal{Z} \sim \mathcal{N}(0, 1)$ ,  $\mathcal{U}_n \sim \chi_n^2$  be independent,  $t_n = \frac{\mathcal{Z}}{\sqrt{\mathcal{U}_n/n}}$  has a t-distribution with  $n$  d.f.

$$f(t) = \frac{\Gamma((n+1)/2)}{\sqrt{n\pi}\Gamma(n/2)} \left(1 + \frac{t^2}{n}\right)^{-\frac{n+1}{2}}$$

1.  $t$  is symmetric about 0
2.  $t_n \xrightarrow[n]{\infty} \mathcal{Z}$

### 3.5 F-distribution

Let  $U \sim \chi_m^2, V \sim \chi_n^2$  be independent,  $W = \frac{U/m}{V/n}$  has an F distribution with  $(m, n)$  d.f.

If  $X \sim t_n$ ,  $X^2 = \frac{\mathcal{Z}/1}{\mathcal{U}_n/n}$  is an F distribution with  $(1, n)$  d.f, with  $w \geq 0$ :

$$f(w) = \frac{\Gamma((n+1)/2)}{\Gamma(n/2)\Gamma(n/2)} \frac{m}{n} \frac{w^{\frac{m}{2}-1}}{w^{\frac{m}{2}-1}} \left(1 + \frac{m}{n}w\right)^{-\frac{m+n}{2}}$$

For  $n > 2$ ,  $E(W) = \frac{n}{n-2}$

## 4 Sampling

Let  $X_1, X_2, \dots, X_n$  be IID  $\mathcal{N}(\mu, \sigma^2)$ .

sample mean,  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$

sample variance,  $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

### 4.1 Properties of $\bar{X}$ and $S^2$

1.  $\bar{X}$  and  $S^2$  are independent
2.  $\bar{X} \sim \mathcal{N}(\mu, \frac{\sigma^2}{n})$
3.  $\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$
4.  $\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{n-1}$

### 4.2 Survey Sampling

In population of size  $N$ , we are interested in a variable  $x$ . The  $i$ th individual has fixed value  $x_i$ .

mean of population =  $\mu = \frac{1}{N} \sum_{i=1}^N x_i$

total of population =  $\tau = \sum_{i=1}^N x_i = \mu N$

SD of population =  $\sigma$

$$\sigma^2 = \sum_{i=1}^N (x_i - \mu)^2 \frac{1}{N} \sum_{i=1}^N x_i^2 - \mu^2$$

### 4.2.1 Dichotomous case

Population are members with value 0 or 1. Let  $p$  be the proportion of members with value 1.  $\mu = p, \sigma^2 = p(1 - p)$

### 4.3 Simple Random Sampling (SRS)

Assume  $n$  random draws are made without replacement. (Not SRS, will be corrected for later).

### 4.3.1 Lemma A

The draws  $X_i$  have the same distribution, and denote  $\xi_1, \xi_2, \dots, \xi_n$  as values assumed by the population, and let the number of members with value  $\xi_j$  be  $n_j$

$$P(X_i = \xi_j) = \frac{n_j}{N}$$

$$E(X_i) = \mu, Var(x_i) = \sigma^2$$

### 4.3.2 Lemma B

For  $i \neq j$ ,  $Cov(X_i, X_j) = -\frac{\sigma^2}{N-1}$

We use sample mean  $\bar{X}$  to estimate  $\mu$ :

$E(\bar{X}) = \mu$  from Lemma A, and

$Var(\bar{X}) = \frac{\sigma^2}{n} \left(\frac{N-n}{N-1}\right)$  from Lemma B, where  $\frac{N-n}{N-1}$  is the finite population correction factor.

In 0-1 population, let  $\hat{p}$  be proportion of 1s in the sample:

$$E(\hat{p}) = p, SD(\hat{p}) = \sqrt{\frac{p(1-p)}{n} \frac{N-n}{N-1}}$$

### 4.3.3 Estimation Problem

Let  $X_1, X_2, \dots, X_n$  be random draws with replacement. Then  $\bar{X}$  is an estimator of  $\mu$ . and the observed value of  $\bar{X}$ ,  $\bar{x}$  is an estimate of  $\mu$ .

### 4.3.4 Standard Error (SE)

Since  $E(\bar{X}) = \mu$ , the estimator is unbiased.

The error in a particular estimate  $\bar{X}$  is unknown, but on average its size is about  $SD(\bar{x}) = \frac{\sigma}{\sqrt{n}}$

Standard error of an  $\bar{X}$  is defined to be  $SD(\bar{X})$

An unbiased estimator for  $\sigma^2$  is  $s^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

param	est	SE	Est.	SE
$\mu$	$\bar{X}$	$\frac{\sigma}{\sqrt{n}}$	$\bar{x}$	$\frac{s}{\sqrt{n}}$
$p$	$\hat{p}$	$\sqrt{\frac{p(1-p)}{n}}$	$\hat{p}$	$\sqrt{\frac{\hat{p}(1-\hat{p})}{n-1}}$

### 4.3.5 Without Replacement

SE is multiplied by  $\frac{N-n}{N-1}$ , because  $s^2$  is biased for  $\sigma^2$ :  $E(\frac{N-1}{N}s^2) = \sigma^2$ , but  $N$  is normally large.

### 4.3.6 Confidence Interval

An approximate  $1 - \alpha$  CI for  $\mu$  is  $(\bar{x} - z_{\alpha/2} \frac{s}{\sqrt{n}}, \bar{x} + z_{\alpha/2} \frac{s}{\sqrt{n}})$

### 4.4 Measurement Error

Let  $x_1, x_2, \dots, x_n$  be independent measurements of unknown constant  $\mu$ .  $X_i = \mu + \epsilon_i$ . The errors are IID with expectation 0, and variance  $\sigma^2$ .  $x_i = \mu + e_i$ , where  $x_i$  and  $e_i$  are realisations of the RV. Then  $\bar{x}$  is an estimate of  $\mu$ , with SE  $\frac{\sigma}{\sqrt{n}}$ .

### 4.4.1 Biased Measurements

Let  $X = \mu + \epsilon$ , where  $E(\epsilon) = 0$ ,  $Var(\epsilon) = \sigma^2$ . Suppose  $\bar{X}$  is used to measure an unknown constant  $a$ ,  $a \neq \mu$ .  $X = a + (\mu - a) + \epsilon$ , where  $\mu - a$  is the bias.

Mean square error (MSE) is  $E((X - a)^2) = \sigma^2 + (\mu - a)^2$

with  $n$  IID measurements,  $\bar{x} = \mu + \bar{\epsilon}$

$$E((x - a)^2) = \frac{\sigma^2}{n} + (\mu - a)^2$$

MSE = SE<sup>2</sup> + bias<sup>2</sup>, hence  $\sqrt{\text{MSE}}$  is a good measure of the accuracy of the estimate  $\bar{x}$  of  $a$ .