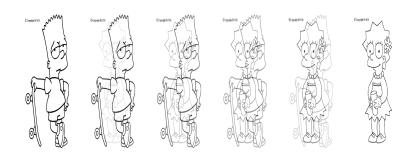
Blind Source Separation

Václav Šmídl

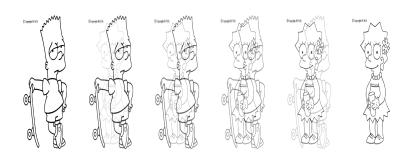
April 4, 2018

Blind source separation – image sequence



2 sources

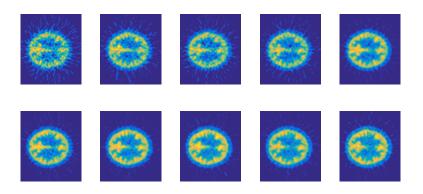
Blind source separation – image sequence



- ▶ 2 sources
- weights:

| | weights. | | | | | | | |
|---|----------|-----|-----|-----|-----|---|--|--|
| | 1 | 8.0 | 0.6 | 0.4 | 0.2 | 0 | | |
| ĺ | 0 | 0.2 | 0.4 | 0.6 | 8.0 | 1 | | |

Medical imaging



- number of sources?
- source images
- ▶ time activity of the source



Mathematical description

Linear model of $p \times n$ matrix

$$d_i = Ax_i + e_i, i = 1..n$$
 $D = AX^T + E,$

where

A is the $p \times r$ matrix of source images, $r < \min(n, p)$

X is the $n \times r$ matrix of time activities,

E is the $p \times n$ noise matrix (Gaussian)

Both images and time activities are unknown!

Issue: ambiguity

$$AX^{T} = ATT^{-1}X^{T}$$
$$(AT)(T^{-1}X^{T}) = \overline{AX}^{T}$$



Principal component analysis

Consider n, p-dimensional vectors d_i , $i = 1, \ldots n$, and their covariance matrix

$$S = \frac{1}{n} \sum (d_i - \overline{d})(d_i - \overline{d})^T.$$

Then r-dimensional vectors x_i ,

$$x_i = U(:, 1:r)^T d_i, S = U \Lambda U^T,$$

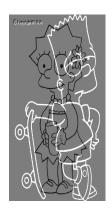
has maximum variance from all possible projections to r dimensions. U are eigenvectors of S sorted with decreasing eigenvalue.

Matrix SVD approach:

$$D \stackrel{\mathsf{svd}}{=} USV,$$
$$= AX^{7}$$

Application to sequences of images





- Popular few years ago for speed of computation,
- Requires to find the rotation matrix T,
- ► Can we do better?
 - Independent Component Analysis (higher order moments)?
 - Structural priors



Probabilistic PCA [Tipping, Bishop, 1999]

Consider model:

$$p(d_i|A, x_i, \sigma) = \mathcal{N}(Ax_i, \sigma I_p),$$

$$p(x_i) = \mathcal{N}(0, I_r),$$

Marginalization over x_i yields

$$p(d_i|W,\sigma) = \mathcal{N}(0,C),$$

 $C = W^T W + \sigma I_p,$

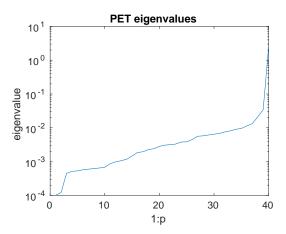
with maximum likelihood

$$\hat{W} = U_{1:r} (\Lambda_{1:r} - \sigma I)^{\frac{1}{2}} T,$$

$$\hat{\sigma} = \frac{1}{d-r} \sum_{i=r+1}^{N} \lambda_i,$$

where $S = U\Lambda U^T$ is eigen-decomposition of S.

Image Sequence (PET)



- ▶ no plateau
- ► Bayesian solution?

Toy matrix decomposition

Consider 1×1 matrix d, decomposed

$$p(d|a,x) = \mathcal{N}(ax, r_e),$$

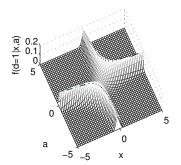
Find a, x.

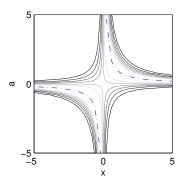
Toy matrix decomposition

Consider 1×1 matrix d, decomposed

$$p(d|a,x) = \mathcal{N}(ax, r_e),$$

Find a, x.





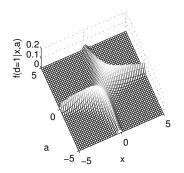
Toy matrix decomposition

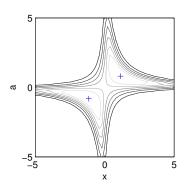
Consider 1×1 matrix d, decomposed

$$p(d|a,x) = \mathcal{N}(ax, r_e),$$

$$p(x) = \mathcal{N}(0, r_x),$$

$$p(a) = \mathcal{N}(0, r_a)$$





Toy maximum likelihood

Joint distribution:

$$\log p(d, a, x) \propto -\frac{1}{2r_e}(d - ax)^2 - \frac{1}{2r_a}a^2 - \frac{1}{2r_x}x^2$$

Find

$$\hat{x}, \hat{a} = \arg\max_{a,x} (\log p(d, a, x))$$

For $d < \frac{r_e}{\sqrt{r_a r_x}}$,

$$\hat{x}=0, \hat{a}=0,$$

For $d \geq \frac{r_e}{\sqrt{r_a r_x}}$,

$$\hat{x} = \pm \left(d\sqrt{\frac{r_x}{r_a}} - \frac{r_e}{r_a} \right)^{\frac{1}{2}}, \qquad \hat{a} = \pm \left(d\sqrt{\frac{r_a}{r_x}} - \frac{r_e}{r_x} \right)^{\frac{1}{2}}.$$

Note that the product of the maxima is

$$\hat{a}\hat{x}=d-\frac{r_e}{\sqrt{r_ar_x}}.$$



Marginal likelihood (PPCA)

Joint distribution:

$$\log p(d, a, x) \propto -\frac{1}{2r_e}(d - ax)^2 - \frac{1}{2r_a}a^2 - \frac{1}{2r_x}x^2$$

Marginal

$$\begin{split} p(a|d) &\propto \int p(d,a,x) dx \\ &\propto \exp(-\frac{1}{2}d^2(a^2r_x+r_e)^{-1})\sqrt{r_er_x}(a^2r_x+r_e)^{-\frac{1}{2}}, \end{split}$$

with maximum:

$$\hat{a} = egin{cases} rac{\sqrt{d^2 - r_e}}{\sqrt{r_x}} & ext{if } d^2 > r_e, \ 0 & ext{otherwise}, \end{cases}$$

Variational Bayes

Joint distribution:

$$\log p(d, a, x) \propto -\frac{1}{2r_e}(d - ax)^2 - \frac{1}{2r_a}a^2 - \frac{1}{2r_x}x^2$$

Factor q(a|d)

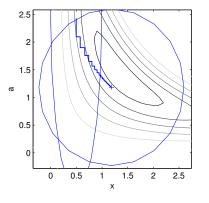
$$\begin{split} p(a|d) &\propto \exp\left(\mathsf{E}_{\mathsf{X}}\left[-\frac{1}{2r_{e}}(d-a\mathsf{X})^{2} - \frac{1}{2r_{a}}a^{2}\right]\right) \\ &= \mathcal{N}(\hat{\mathsf{a}},\sigma_{\mathsf{a}}), \\ \hat{\mathsf{a}} &= \sigma_{\mathsf{a}}d\left\langle \mathsf{X}\right\rangle r_{e}^{-1}, \quad \sigma_{\mathsf{a}}^{-1} = \left\langle a^{2}\right\rangle r_{e}^{-1} + r_{\mathsf{a}}^{-1} \end{split}$$

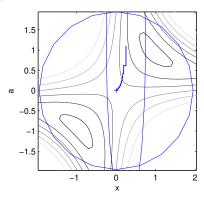
Factor q(x|d)

$$\begin{split} p(a|d) &\propto \exp\left(\mathsf{E}_{x}\left[-\frac{1}{2r_{e}}(d-ax)^{2} - \frac{1}{2r_{a}}a^{2}\right]\right) \\ &= \mathcal{N}(\hat{a},\sigma_{a}), \\ \hat{a} &= \sigma_{a}d\left\langle x\right\rangle r_{e}^{-1}, \quad \sigma_{a}^{-1} = \left\langle a^{2}\right\rangle r_{e}^{-1} + r_{a}^{-1} \end{split}$$

Convergence of VB

- 1. compute \hat{a}, σ_a , and $\langle a \rangle = \hat{a}, \langle a^2 \rangle = \hat{a}^2 + \sigma_a$,
- 2. compute \hat{x}, σ_x , and $\langle x \rangle = \hat{x}$, $\langle x^2 \rangle = \hat{x}^2 + \sigma_x$,





Positive support

What if we are interested only in the positive solution?

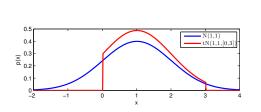
Positive support

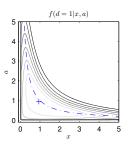
What if we are interested only in the positive solution?

$$p(d|a,x) = \mathcal{N}(ax, r_e),$$

$$p(x) = t\mathcal{N}(0, r_x, \langle 0, \infty \rangle) \propto \mathcal{N}(0, r_x)\chi(x > 0),$$

$$p(a) = t\mathcal{N}(0, r_a, \langle 0, \infty \rangle) \propto \mathcal{N}(0, r_a)\chi(a > 0),$$





Variational Bayes with positive support

Joint distribution:

$$\log p(d, a, x) \propto -\frac{1}{2r_e}(d - ax)^2 - \frac{1}{2r_a}a^2 - \frac{1}{2r_x}x^2 + \log \chi_x + \log \chi_a$$

Factor q(a|d)

$$p(a|d) \propto \exp\left(\mathsf{E}_{x}\left[-\frac{1}{2r_{e}}(d-ax)^{2} - \frac{1}{2r_{a}}a^{2} + \log\chi_{x}\right]\right)$$

$$= t\mathcal{N}(\hat{a}, \sigma_{a}, \langle 0, \infty \rangle),$$

$$\hat{a} = \sigma_{a}d\langle x \rangle r_{e}^{-1}, \quad \sigma_{a}^{-1} = \langle x^{2} \rangle r_{e}^{-1} + r_{a}^{-1}$$

Factor q(x|d)

$$\begin{split} p(x|d) &\propto \exp\left(\mathsf{E}_x \left[-\frac{1}{2r_e} (d-ax)^2 - \frac{1}{2r_a} a^2 + \log \chi_a \right] \right) \\ &= t \mathcal{N}(\hat{a}, \sigma_x, \langle 0, \infty \rangle), \\ \hat{x} &= \sigma_x d \left\langle a \right\rangle r_e^{-1}, \quad \sigma_x^{-1} = \left\langle a^2 \right\rangle r_e^{-1} + r_x^{-1} \end{split}$$

Convergence of VB with positive support

1. compute \hat{a}, σ_a , and

$$\langle a \rangle = \hat{a} + \frac{\phi(\alpha) - \phi(\beta)}{Z_a},$$
$$\langle a^2 \rangle = \sigma_a^2 \left\{ 1 + \frac{\alpha \phi(\alpha) - \beta \phi(\beta)}{Z_a} \right\}$$

where
$$\alpha = \frac{-\hat{a}}{\sqrt{\sigma_a}}$$
, $\beta = \infty$, $\phi(x) = \frac{1}{\sqrt{2\pi}} \exp(-x^2/s)$, $Z = \Phi(\beta) - \Phi(\alpha)$, $\Phi(x) = \frac{1}{2}(1 + \text{erf}(x))$.

2. compute \hat{x}, σ_x , and

$$\langle x \rangle = \hat{x} + \frac{\phi(\alpha) - \phi(\beta)}{Z_x},$$
$$\langle x^2 \rangle = \sigma_x^2 \left\{ 1 + \frac{\alpha\phi(\alpha) - \beta\phi(\beta)}{Z_x} \right\}$$

Zero solution is hard to reach. (ARD)



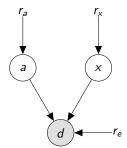
Decomposition with ARD

Original model:

$$p(d|a,x) = N(ax, r_e),$$

$$p(a) = N(0, r_a),$$

$$p(x) = N(0, r_x),$$



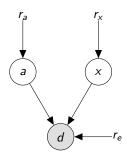
Decomposition with ARD

Original model:

$$p(d|a,x) = N(ax, r_e),$$

$$p(a) = N(0, r_a),$$

$$p(x) = N(0, r_x),$$



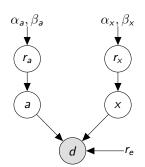
Unknown hyper-parameters:

$$p(a|r_a) = N(0, r_a),$$

$$p(x|r_e) = N(0, r_x),$$

$$p(r_a) = \Gamma(\alpha_a, \beta_a),$$

$$p(r_x) = \Gamma(\alpha_x, \beta_x),$$



Variational PCA [Bishop 200]

Consider model

$$p(d_i|A, x_i, \sigma) = \mathcal{N}(Ax_i, \omega I_p),$$

$$p(x_i) = \mathcal{N}(0, I_r),$$

$$p(\underline{a}_j) = \mathcal{N}(0, I_r),$$

Find $p(A, X, \sigma)$.

Variational PCA [Bishop 200]

Consider model

$$p(d_i|A, x_i, \sigma) = \mathcal{N}(Ax_i, \omega I_p),$$

$$p(x_i) = \mathcal{N}(0, I_r),$$

$$p(\underline{a}_j) = \mathcal{N}(0, I_r),$$

Find $p(A, X, \sigma)$.

Joint likelihood:

$$\log p(D, A, X) \propto -\frac{1}{2}\omega \sum_{i} (d_i - Ax_i)^T (d_i - Ax_i) - \frac{1}{2}\sum_{j} \underline{a}_j^T \underline{a}_j - \frac{1}{2}\sum_{i} x_i^T x_i,$$

Variational PCA [Bishop 200]

Consider model

$$p(d_i|A, x_i, \sigma) = \mathcal{N}(Ax_i, \omega I_p),$$

$$p(x_i) = \mathcal{N}(0, I_r),$$

$$p(\underline{a}_j) = \mathcal{N}(0, I_r),$$

Find $p(A, X, \sigma)$.

Joint likelihood:

$$\log p(D, A, X) \propto -\frac{1}{2}\omega \sum_{i} (d_i - Ax_i)^T (d_i - Ax_i) - \frac{1}{2} \sum_{i} \underline{a}_j^T \underline{a}_j - \frac{1}{2} \sum_{i} x_i^T x_i,$$

Factor $q(x_i)$

$$\begin{split} q(\mathbf{x}_i) &\propto \exp\left(-\frac{1}{2}\omega(d_i - A\mathbf{x}_i)^T(d_i - A\mathbf{x}_i) - \frac{1}{2}\sum_i \mathbf{x}_i^T\mathbf{x}_i,\right) \\ &= \mathcal{N}(\hat{\mathbf{x}}_i, \mathbf{\Sigma}_{\mathbf{x}}), \\ \hat{\mathbf{x}}_i &= \langle \omega \rangle \, \mathbf{\Sigma}_{\mathbf{x}} \, \langle A \rangle \, d_i, \quad \mathbf{\Sigma}_{\mathbf{x}}^{-1} &= \langle \omega \rangle \, \langle A^T A \rangle + I \end{split}$$

Extensions: Multivariate PCA

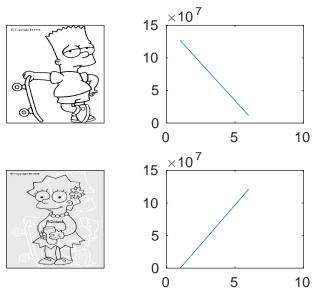
1. Positive support

$$\begin{split} & p(x_i) = t \mathcal{N}(0, I_r, \langle 0, \infty \rangle), \\ & p(\underline{a}_j) = t \mathcal{N}(0, I_r, \langle 0, \infty \rangle), \end{split}$$

moments

$$[\langle x \rangle, \mathsf{diag} \, \big\langle x^\mathsf{T} x \big\rangle] {=} \mathsf{momtrun_low} \big(\hat{x}, \mathsf{sqrt} \big(\mathsf{diag}(\Sigma_x) \big) \big)$$

Non-negative Matrix Factorization (NMF)



Extensions: Multivariate PCA

1. Positive support

$$p(x_i) = t\mathcal{N}(0, I_r, \langle 0, \infty \rangle),$$

$$p(\underline{a}_j) = t\mathcal{N}(0, I_r, \langle 0, \infty \rangle),$$

moments

$$[\langle x \rangle, \operatorname{diag} \langle x^T x \rangle] = \operatorname{momtrun_low}(\hat{x}, \operatorname{sqrt}(\operatorname{diag}(\Sigma_x)))$$

2. Automatic Relevance Determination (#of factors):

$$p(\underline{a}_j) = \mathcal{N}(0, \operatorname{diag}(\alpha)),$$

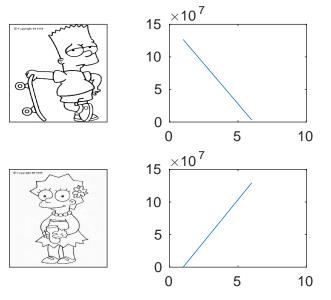
3. Automatic Relevance Determination (#pixels):

$$p(\underline{a}_j) = \mathcal{N}(0, \operatorname{diag}(\alpha_j)),$$

4. many more



Sparse Non-negative Matrix Factorization (SNMF)



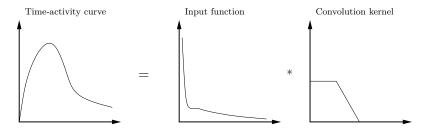
Assignment

| Toy Problem | points | |
|-------------|--------|--|
| Laplace | 5 | |
| VB | 10 | |
| Gibbs | 10 | |
| | | |

| Matrix factorization | points |
|--|--------|
| Bayesian PCA | 20 |
| + positive support | +10 |
| $+$ ARD, $x_{i,j} \sim \mathcal{N}(0, \gamma_{i,j})$ | +10 |
| ICA (other decomp) | 10 |

BSS with deconvolution

Time-activity curves for brain imaging are results of convolution.



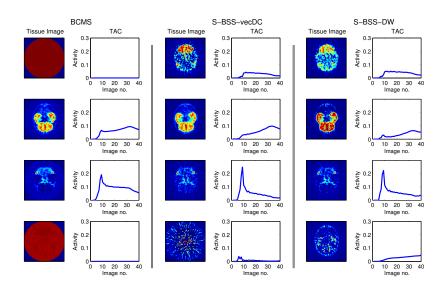
Model of the curve

$$x = b * w = Bw, \qquad B = \begin{pmatrix} b_1 & 0 & 0 & 0 \\ b_2 & b_1 & 0 & 0 \\ \dots & b_2 & b_1 & 0 \\ b_n & \dots & b_2 & b_1 \end{pmatrix}$$

Unknows are b and w. Kernel is sparse.



BSS with deconvolution





(a) out-of-focus blur

Mathematical model:

$$d = Ax + e$$

where

d is the observed (blured) image,

x is the original (true) image,

A is the convolution matrix,

e is the measurement (model) error

Find: x, A, var(e)

unknown

Mathematical model:

$$d = Ax + e$$

where

d is the observed (blured) image,

x is the original (true) image,

A is the convolution matrix,

e is the measurement (model) error

Find: x, A, var(e)

unknown

unknown kernel



Mathematical model:

$$d = Ax + e$$

where

d is the observed (blured) image,

x is the original (true) image,

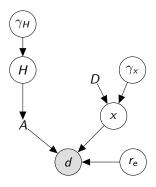
A is the convolution matrix,

e is the measurement (model) error

unknown unknown kernel unknown variance

Find: x, A, var(e)

Model



- number of unknowns > 3× higher than number of observations
- ▶ ARD coefficients γ_H , γ_x , r_e ,
- approximations of the covariance matrices by diagonal

Results:



(a) out-of-focus blur



(b) blind deconvolution