DSCI561: Regression I

Lecture 2: November 20, 2017

Gabriela Cohen Freue Department of Statistics, UBC

In Lect 1: two-sample t-test vs ANOVA

- In both cases we are interested in studying the expected number of runs per game
 - Quantitative response variable: tax assessment value for a property built in period $i \colon Y_i$
 - Population mean: $\mu_i = E[Y_i]$
- two-sample t-test: compares the means of two populations (groups)

$$H_0: \mu_C = \mu_M$$

- one-way ANOVA: compares the means of K groups
 - 1 factor, K levels)
 - 1 factor: age period; K=3 levels: C, M, O $H_0: \mu_C = \mu_M = \mu_O$
 - K=2: it is equivalent to a two-sample t-test
- ANOVA: Study the effect of one or more qualitative variables (factors) on a quantitative variable (response):
 - Quantitative response: tax property assessment value



```
#t-test vs ANOVA
#responses within each group
tax.M <-dat.small %>% subset(age_factor =="M", select=assessment_k)
tax.C <-dat.small %>% subset(age_factor == "C", select=assessment_k)
t.test(tax.M,tax.C,var.equal=T)
##
   Two Sample t-test
##
##
## data: tax.M and tax.C
## t = -2.2034, df = 18, p-value = 0.04083
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -180.111457 -4.288543
## sample estimates:
                                                          (-2.2034)^2 = 4.855
## mean of x mean of y
      428.8
                521.0
##
#subset of 2 age periods
summary(aov(assessment_k~age_factor,data=subset(dat.small,age_factor %in% c("M","C"))))
              Df Sum Sq Mean Sq F value (>F)
##
## age_factor 1 31878
                          31878
                                  4.855 0.0408 *
## Residuals 18 118188
                         6566
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

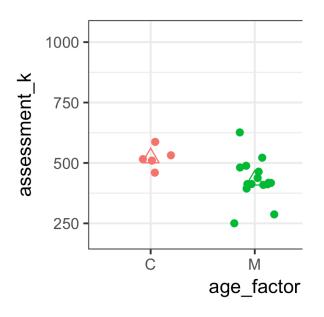
```
#t-test vs ANOVA
#responses within each group
tax.M <-dat.small %>% subset(age_factor =="M", select=assessment_k)
tax.C <-dat.small %>% subset(age_factor == "C", select=assessment_k)
t.test(tax.M,tax.C,var.equal=T)
##
                                                           H_0: \mu_C = \mu_M
   Two Sample t-test
##
## data: tax.M and tax.C
## t = -2.2034, df = 18, p-value = 0.04083
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -180.111457 -4.288543
## sample estimates:
                                                                 same test
## mean of x mean of y
      428.8
                521.0
##
#subset of 2 age periods
summary(aov(assessment_k~age_factor,data=subset(dat.small,age_factor %in% c("M","C"))))
              Df Sum Sq Mean Sq F value Pr(>F)
##
                          31878
                                 4.855 0.0408
## age_factor
              1 31878
                           6566
## Residuals
              18 118188
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

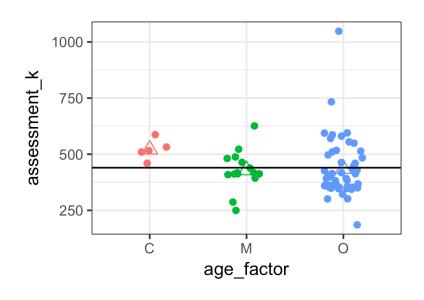
ANOVA as a special case of Regression

Two groups

```
summary(lm(assessment_k~age_factor,data=subset(dat.small,age_factor %in% c("M","C"))))
##
## Call:
## lm(formula = assessment_k ~ age_factor, data = subset(dat.small,
      age_factor %in% c("M", "C")))
##
##
## Residuals:
      Min
               10 Median
##
                               3Q
                                     Max
## -178.80 -17.55 -10.90
                          39.45 197.20
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
                            36.24 14.377 2.62e-11 ***
## (Intercept) 521.00
## age_factorM -92.20
                           41.84 -2.203
                                           0.0408 *
                                                                    same as t-test
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 81.03 on 18 degrees of freedom
## Multiple R-squared: 0.2124. Adjusted R-squared: 0.1687
                                                                  same as ANOVA
## F-statistic: 4.855 on 1 and 18 DF, p-value: 0.04083
```

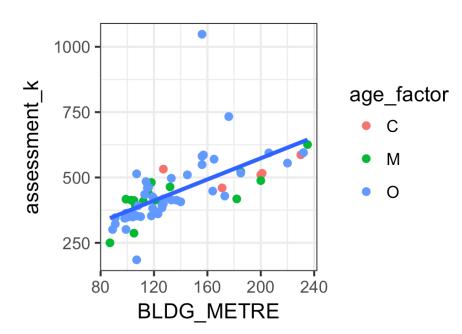
```
#More than 2 groups
                                                       More than 2 groups
#LM with 3 age periods
summary(lm(assessment_k~age_factor,data=dat.small))
##
## Call:
## lm(formula = assessment k ~ age factor, data = dat.small)
##
## Residuals:
##
               10 Median
      Min
                               3Q
                                      Max
## -250.14 -74.89 -16.97
                            51.36 612.86
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                                    9.631 2.87e-14 ***
                521.00
                            54.10
                                                            different from t-test!!
## age_factorM
                -92.20
                            62.46 -1.476
                                             0.145
                -85.86
                            56.74 -1.513
                                             0.135
## age factor0
## ---
## Signif. codes:
                  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 121 on 67 degrees of freedom
## Multiple R-squared: 0.0353, Adjusted R-squared: 0.006498
## F-statistic: 1.226 on 2 and 67 DF, p-value: 0.3001
                                                                        same test
#ANOVA with 3 age periods
summary(aov(assessment_k~age_factor,data=dat.small))
              Df Sum Sq Mean Sq F value Pr(>F)
##
## age factor
                  35867
                          17934
                                  1.226
                                           0.3
## Residuals
              67 980328
                          14632
```





two-samples t-test: 2 groups

one-way ANOVA: more than 2 groups



Linear regression: quantitative and qualitative explanatory variables

In today's lecture

- Comparison with the output of the `lm()` function in R
- Review of linear algebra operations
- Review the mathematical notation of a linear model and its connection with the R-code
- Build the matrix notation of a linear model

Some linear algebra...

Sum of matrices

- Let A and B be $n \times m$ matrices (n rows, m columns)
- $\, {f A+B} \,$ is an $\, n\, {f x} \,$ m matrix with $\it ij$ th element equal to $a_{ij}+b_{ij}$

$$egin{bmatrix} a_{11} & a_{12} \ a_{21} & a_{22} \ dots & dots \ a_{m1} & a_{m2} \end{bmatrix} + egin{bmatrix} b_{11} & b_{12} \ b_{21} & b_{22} \ dots & dots \ b_{m1} & b_{m2} \end{bmatrix} = egin{bmatrix} a_{11} + b_{11} & a_{12} + b_{12} \ a_{21} + b_{21} & a_{22} + b_{22} \ dots & dots \ a_{m1} + b_{m1} & a_{m2} + b_{m2} \end{bmatrix}$$
 $egin{bmatrix} A & B & A+B \end{bmatrix}$

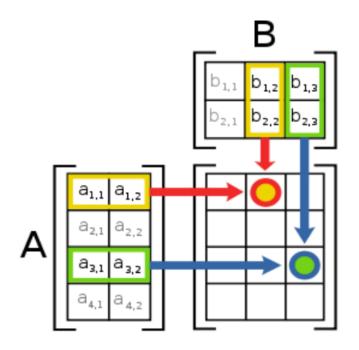
$$\begin{bmatrix} 1 & 0 & 3 & -1 \\ 0 & 1 & 1 & -1 \\ 0 & 4 & 0 & 0 \end{bmatrix} + \begin{bmatrix} 0 & 0 & -1 & 1 \\ 2 & 0 & 0 & 0 \\ 0 & -2 & 0 & 0 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 2 & 0 \\ 2 & 1 & 1 & -1 \\ 0 & 2 & 0 & 0 \end{bmatrix}$$

Multiplication of matrices

- If A is an $n \times m$ matrix, AB is defined only if B has m rows (number of columns in B doesn't matter)
 - **AB** is an $n \times m$ matrix

Dot product

$$a \cdot b = \sum a_i b_i$$



Example

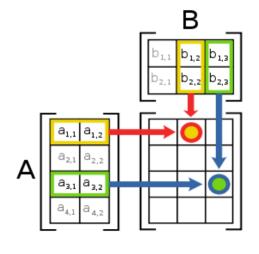
$$\begin{bmatrix} 1 & 2 \\ 1 & 4 \\ 0 & 5 \end{bmatrix} * \begin{bmatrix} -1 & -2 \\ -2 & 1 \end{bmatrix} = \begin{bmatrix} -5 & 0 \\ -9 & 2 \\ -10 & 5 \end{bmatrix}$$
(3x2) (2x2) (3x2)

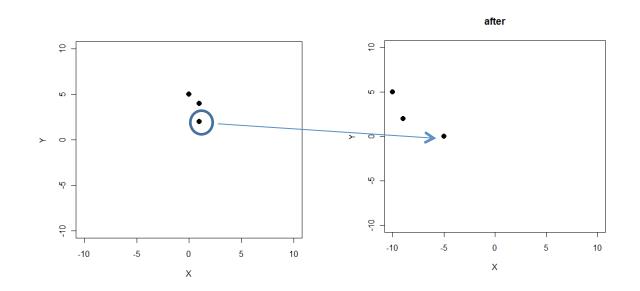
```
#Multiplication
A <- matrix(c(1, 2, 1,4, 0,5),3,2,byrow=T)
B <- matrix(c(-1, -2, -2,1),2,2,byrow=T)

A%*%B
```

```
## [,1] [,2]
## [1,] -5 0
## [2,] -9 2
## [3,] -10 5
```

#Note: A%*%B is not the same as A*B





Matrix operations as transformations in space

- Multiply by a scalar: moving the points further apart in space (or closer together)
- Multiply by another matrix: e.g., rotation, or projection
- Projection in particular is a fundamental operation: often want to project from original space to a reduced space that is "explanatory"

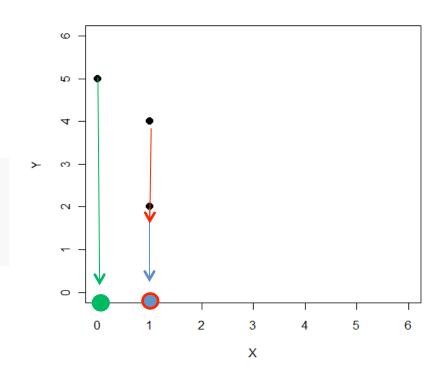
Examples: Projections

$$\begin{bmatrix} 1 & 2 \\ 1 & 4 \\ 0 & 5 \end{bmatrix} \times \begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 0 & 0 \end{bmatrix}$$

#Projection

```
A <- matrix(c(1,2,1,4,0,5),3,2,byrow=T)
Px <- matrix(c(rep(0,3),1),2,2)
A%*%Px
```

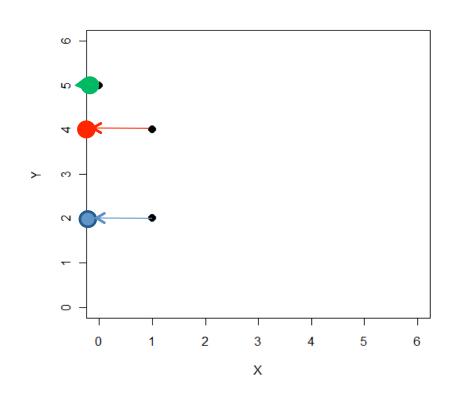
```
## [,1] [,2]
## [1,] 0 2
## [2,] 0 4
## [3,] 0 5
```



Examples: Projections

$$\begin{bmatrix} 1 & 2 \\ 1 & 4 \\ 0 & 5 \end{bmatrix} \times \begin{bmatrix} 0 & 0 \\ 0 & 1 \end{bmatrix} = \begin{bmatrix} 0 & 2 \\ 0 & 4 \\ 0 & 5 \end{bmatrix}$$

```
## [,1] [,2]
## [1,] 1 0
## [2,] 1 0
## [3,] 0 0
```



We can project onto any line we like

Inverse of a matrix

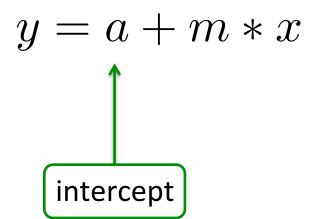
- An $n \times n$ square matrix A is invertible, if there exist an $n \times n$ square matrix B such that $AB=BA=I_n$
- This matrix **B** is called the inverse of **A**: **A**⁻¹

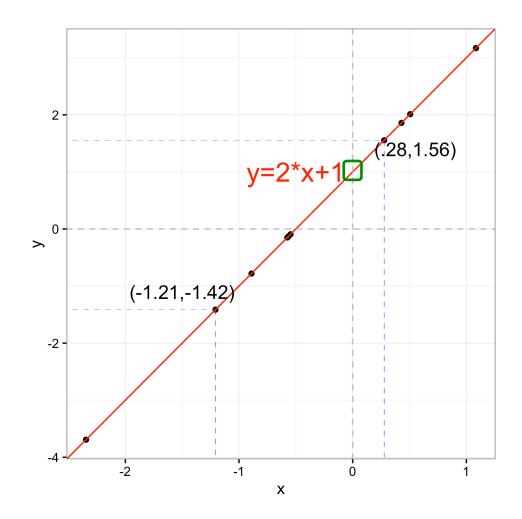
$$\begin{bmatrix} 7 & 2 & 1 \\ 0 & 3 & -1 \\ -3 & 4 & -2 \end{bmatrix}^{-1} = \begin{bmatrix} -2 & 8 & -5 \\ 3 & -11 & 7 \\ 9 & -34 & 21 \end{bmatrix}$$

```
#Inverse
A <- matrix(c(7, 2 , 1,0 , 3 ,-1, -3, 4 , -2),3,3,byrow=T)
solve(A)
```

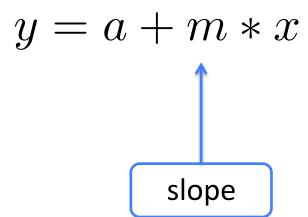
```
## [,1] [,2] [,3]
## [1,] -2 8 -5
## [2,] 3 -11 7
## [3,] 9 -34 21
```

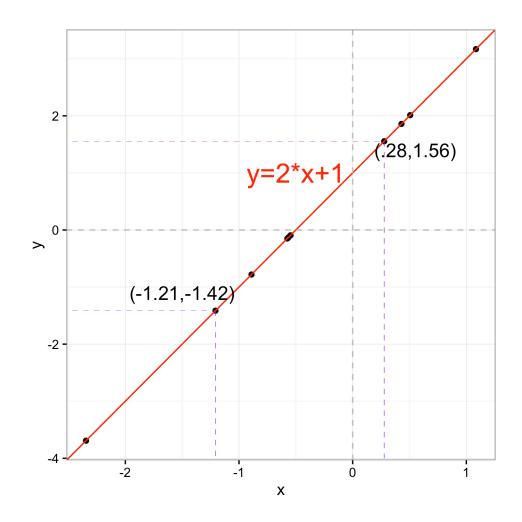
A line



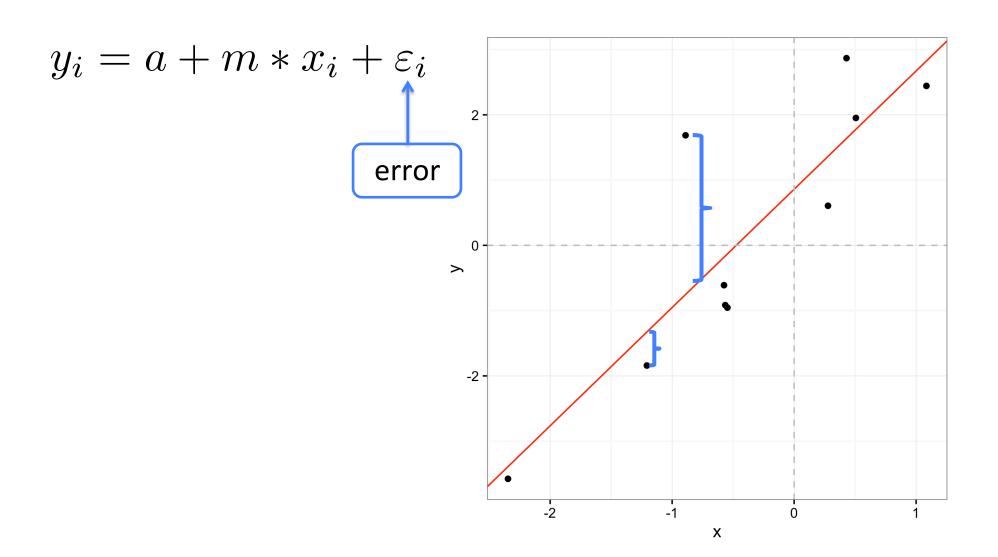


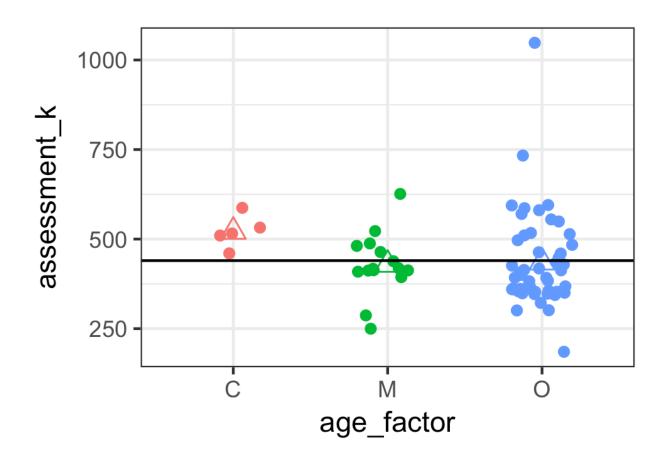
A line





A regression line





Where is the linear regression?

YEAR_BUILT	age_factor	assessment_k	
2013	С		510
2003	С		516
2002	С		460
2002	С		532
2005	С		587
1995	M		481
1989	М		409
1991	M		522
1998	M		413
1989	M		413
1988	M		394
1990	M		418
1998	M		464
1990	M		412
1989	M		488
1990	M		626
1984	M		417
1989	M		250
1997	M		438
1980	М		287

Glance at the data

$$Y_C; Y_{C1}, \dots, Y_{C5}, n_C = 5$$

$$Y_M; Y_{M1}, \ldots, Y_{M15}, n_M = 15$$

$$H_0: \mu_C = \mu_M$$

```
## Call:
## lm(formula = assessment_k ~ age_factor, data = subset(dat.small,
## age_factor %in% c("M", "C")))
```

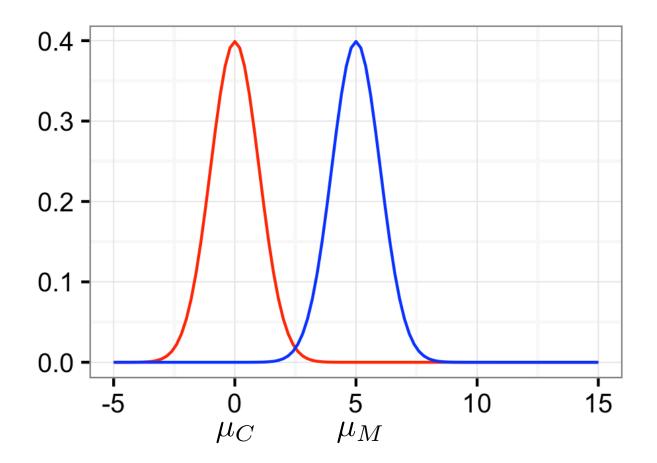
$$Y_{ij} = \mu_j + \varepsilon_{ij}$$
, where $\varepsilon_{ij} \sim F_j$, $E(\varepsilon_{ij}) = 0$

YEAR_BUILT	age_factor	$assessment_\hat{k}$	
2013	С		510
2003	С		516
2002	С		460
2002	С		532
2005	С		587
1995	M		481
1989	M		409
1991	M		522
1998	M		413
1989	M		413
1988	M		394
1990	M		418
1998	M		464
1990	M		412
1989	M		488
1990	M		626
1984	M		417
1989	M		250
1997	M		438
1980	М		287

$$\begin{bmatrix} Y_{11} \\ \vdots \\ Y_{n_1 1} \\ \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \vdots \\ \mu_1 \\ \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \vdots \\ \varepsilon_{n_1 1} \\ \varepsilon_{12} \\ \vdots \\ \vdots \\ \vdots \\ \varepsilon_{n_2 2} \end{bmatrix}$$

$$H_0: \mu_C = \mu_M$$

Change in notation...



$$Y_{ij} = \mu_j + \varepsilon_{ij}, \ \varepsilon_{ij} \sim F_j, \ E[\varepsilon_{ij}] = 0$$
$$Y_M \sim F_M; \ E[Y_M] = \mu_M$$
$$Y_C \sim F_C; \ E[Y_B] = \mu_C$$

 $H_0: \mu_C = \mu_M$

We don't know or observe these curves and parameters

$$Y_{ij} = \mu_j + \varepsilon_{ij}$$
, where $\varepsilon_{ij} \sim F_j$, $E(\varepsilon_{ij}) = 0$

$$\begin{bmatrix} Y_{11} \\ \vdots \\ Y_{n_1 1} \\ \hline Y_{12} \\ \vdots \\ Y_{n_2 2} \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \vdots \\ \mu_1 \\ \hline \mu_2 \\ \vdots \\ \mu_2 \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \vdots \\ \varepsilon_{n_1 1} \\ \varepsilon_{12} \\ \vdots \\ \varepsilon_{n_2 2} \end{bmatrix}$$

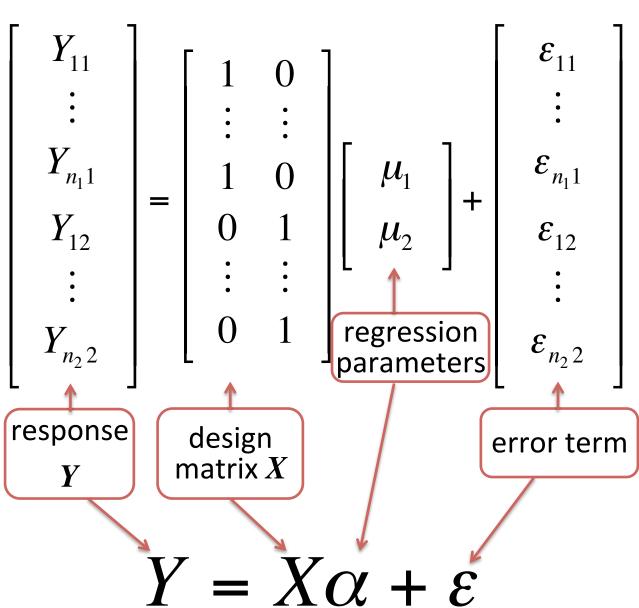
$$Y_{ij} = \mu_j + \varepsilon_{ij}$$
, where $\varepsilon_{ij} \sim F_j$, $E(\varepsilon_{ij}) = 0$

$$\begin{bmatrix} Y_{11} \\ \vdots \\ Y_{n_{1}1} \\ Y_{12} \\ \vdots \\ Y_{n_{2}2} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ \vdots & \vdots \\ 1 & 0 \\ 0 & 1 \\ \vdots & \vdots \\ 0 & 1 \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \vdots \\ \varepsilon_{n_{1}1} \\ \varepsilon_{12} \\ \vdots \\ \varepsilon_{n_{2}2} \end{bmatrix} + \begin{bmatrix} \mu_{1} \\ \vdots \\ \mu_{2} \\ \vdots \\ \mu_{2} \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \vdots \\ \varepsilon_{n_{1}1} \\ \varepsilon_{12} \\ \vdots \\ \varepsilon_{n_{2}2} \end{bmatrix}$$

For example:

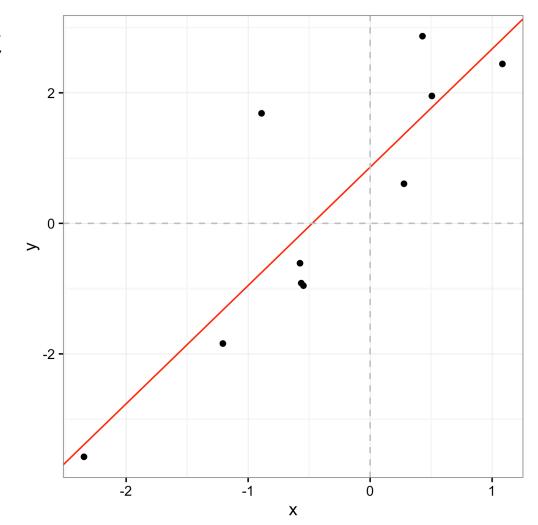
$$Y_{11} = 1 * \mu_1 + 0 * \mu_2 + \varepsilon_{11} = \mu_1 + \varepsilon_{11}$$
$$Y_{n_2 2} = 0 * \mu_1 + 1 * \mu_2 + \varepsilon_{n_2 2} = \mu_2 + \varepsilon_{n_2 2}$$

$$Y_{ij} = \mu_j + \varepsilon_{ij}$$
, where $\varepsilon_{ij} \sim F_j$, $E(\varepsilon_{ij}) = 0$



A regression line

$$y_i = a + m * x_i + \varepsilon_i$$



the column vector of the responses one element per experimental unit

a column vector of the errors

 $Y = X\alpha + \varepsilon$

a (design) matrix that represents covariate info, one row per experimental unit

a column vector of the parameters in the linear model

Generic linear model, using conventional matrix formulation

$$\begin{bmatrix} Y_{11} \\ \vdots \\ Y_{n_1 1} \\ Y_{12} \\ \vdots \\ Y_{n_2 2} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ \vdots & \vdots \\ 1 & 0 \\ 0 & 1 \\ \vdots & \vdots \\ 0 & 1 \end{bmatrix} \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \vdots \\ \varepsilon_{n_1 1} \\ \varepsilon_{12} \\ \vdots \\ \varepsilon_{n_2 2} \end{bmatrix}$$

Two groups

sample mean of C

```
summary(lm(assessment_k~age_factor,data=subset(dat.small,age_factor %in% c("M","C"))))
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                521.00
                            36.24 14.377 2.62e-11 ***
                                                                  NOT the sample
## age_factorM
                -92.20
                            41.84 -2.203
                                            0.0408 *
                                                                     mean of M
                     '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 81.03 on 18 degrees of freedom
## Multiple R-squared: 0.2124, Adjusted R-squared: 0.1687
## F-statistic: 4.855 on 1 and 18 DF, p-value: 0.04083
```

$$\begin{bmatrix} Y_{11} \\ \vdots \\ Y_{n_{1}1} \\ Y_{12} \\ \vdots \\ Y_{n_{2}2} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ \vdots & \vdots \\ 1 & 0 \\ 0 & 1 \\ \vdots & \vdots \\ 0 & 1 \end{bmatrix} \begin{bmatrix} \mu_{1} \\ \mu_{2} \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \vdots \\ \varepsilon_{n_{1}1} \\ \varepsilon_{12} \\ \vdots \\ \varepsilon_{n_{2}2} \end{bmatrix}$$

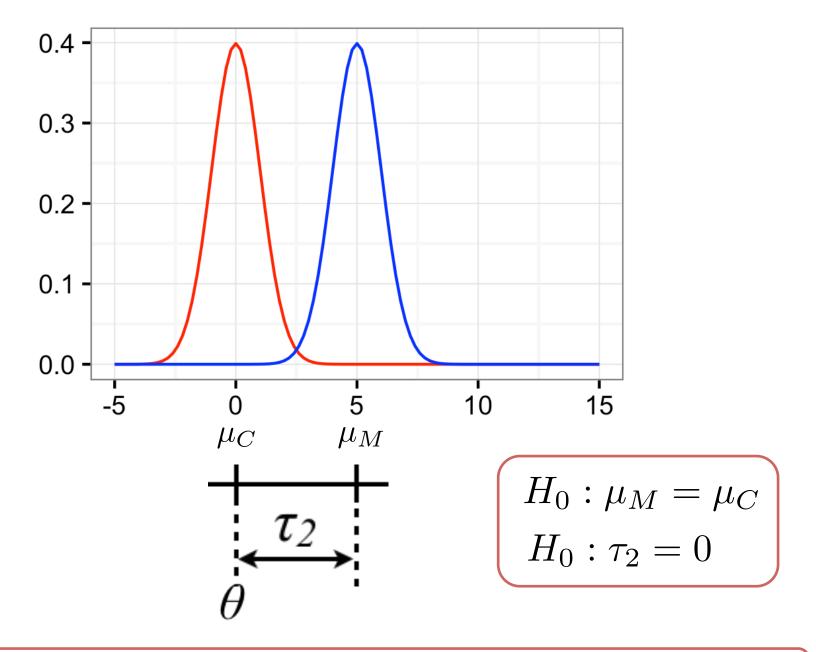
$$Y = X\alpha + \varepsilon$$

This is a one way of writing our problem as a linear regression

"cell means" parametrization

... but there are other ways!!

By default, R does not estimate these parameters



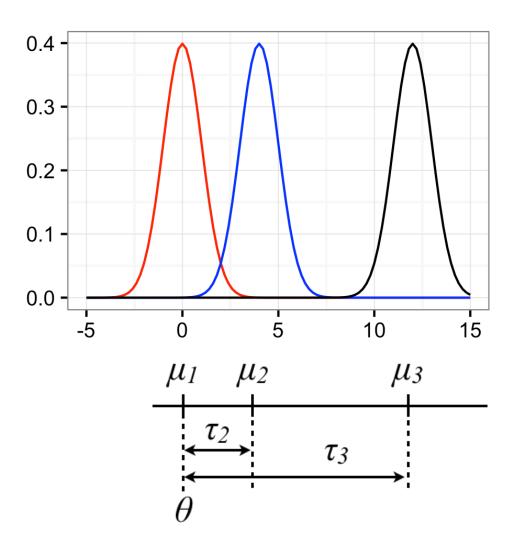
A different parametrization: "reference-treatment effect"

ANOVA-style, "cell means"

ANOVA-style, "ref + tx effects"

$$Y_{ij} = \mu_j + \varepsilon_{ij}$$

$$Y_{ij} = \theta + \tau_j + \varepsilon_{ij}, (\tau_1 = 0)$$



ANOVA-style, "cell means"

ANOVA-style, "ref + tx effects"

$$Y_{ij} = \mu_j + \varepsilon_{ij}$$

$$Y_{ij} = \theta + \tau_j + \varepsilon_{ij}, (\tau_1 = 0)$$

$$\begin{bmatrix} y_{11} \\ y_{21} \\ \vdots \\ y_{n_{3}3} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} \mu_{1} \\ \mu_{2} \\ \mu_{3} \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \varepsilon_{21} \\ \vdots \\ \varepsilon_{n_{3}3} \end{bmatrix} \begin{bmatrix} y_{11} \\ y_{21} \\ \vdots \\ y_{n_{3}3} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ 1 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 1 \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \varepsilon_{21} \\ \vdots \\ \varepsilon_{n_{3}3} \end{bmatrix}$$

The design matrix specifies how the observed data relates to the regression parameters.

ANOVA-style, "cell means"

ANOVA-style, "ref + tx effects"

$$Y_{ij} = \mu_j + \varepsilon_{ij}$$

$$Y_{ij} = \theta + \tau_j + \varepsilon_{ij}, (\tau_1 = 0)$$

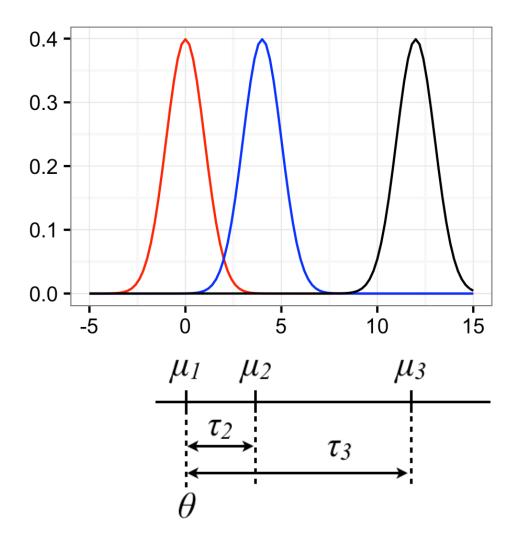
$$Y = X\alpha + \varepsilon$$

$$\begin{bmatrix} y_{11} \\ y_{21} \\ \vdots \\ y_{n_33} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 1 \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \varepsilon_{21} \\ \vdots \\ \varepsilon_{n_33} \end{bmatrix} + \begin{bmatrix} y_{11} \\ y_{21} \\ \vdots \\ y_{n_33} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 1 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 1 \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \varepsilon_{21} \\ \vdots \\ \varepsilon_{n_33} \end{bmatrix}$$

For example:

$$Y_{11} = 1 * \theta + 0 * \tau_2 + 0 * \tau_3 + \varepsilon_{11} = \theta + \varepsilon_{11} \implies E[Y_{11}] = \theta$$

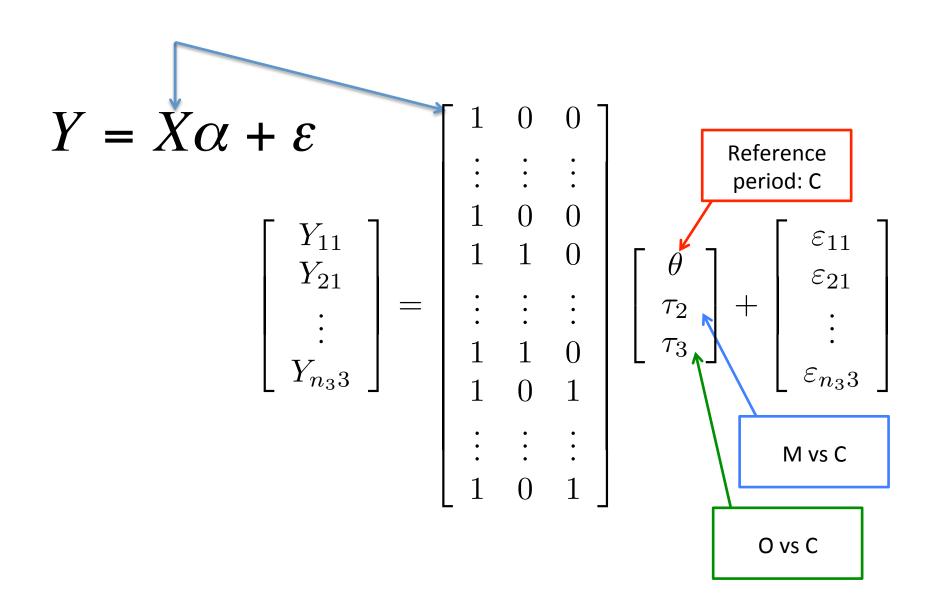
 $Y_{13} = 1 * \theta + 0 * \tau_2 + 1 * \tau_3 + \varepsilon_{13} = \theta + \tau_3 + \varepsilon_{13} \implies E[Y_{13}] = \theta + \tau_3$



$$H_0: \mu_1 = \mu_2 = \mu_3$$

$$H_0: \tau_2 = \tau_3 = 0$$

$$Y_{i1} \sim F_1; \ E[Y_{i1}] = \mu_1$$
 $E[Y_{i1}] = \theta$ reference reference $E[Y_{i1}] = \theta$ $E[Y_{i2}] = \theta + \tau_2 \implies E[Y_{i2}] - E[Y_{i1}] = \tau_2$ $Y_{i3} \sim F_3; \ E[Y_{i3}] = \mu_3$ $E[Y_{i3}] = \theta + \tau_3 \implies E[Y_{i3}] - E[Y_{i1}] = \tau_3$



difference in *population* means