

# Improved Runtime Bounds for the Univariate Marginal Distribution Algorithm via Anti-Concentration

Phan Trung Hai Nguyen & Per Kristian Lehre {PXN683, P.K.Lehre}@cs.bham.ac.uk
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School of Computer Science University of Birmingham Birmingham B15 2TT United Kingdom

### **Outlines**

### Background

- Estimation of Distribution Algorithms (EDAs)
- Univariate Marginal Distribution Algorithm (UMDA)

#### Useful tools

- Level-based theorem
- Anti-concentration bound
- Feige's inequality

#### Our result

 $\bullet$  Improved upper bound on runtime of  $\operatorname{UMDA}$  on  $\operatorname{OneMax}$ 

#### Conclusion

# **Estimation of Distribution Algorithms (EDAs)**

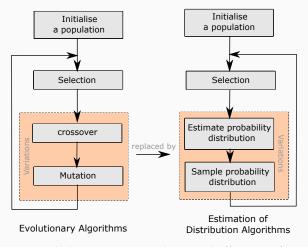


Figure 1: Comparison between  $\mathrm{EAs}$  and  $\mathrm{EDAs}$  (Shakya 2006)

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# Univariate Marginal Distribution Algorithm (UMDA)

Solution representation:  $(X_1,\ldots,X_n)\in\{0,1\}^n$ .

Probabilistic model at generation t:  $\mathcal{M}_t := (p_t^1, p_t^2, \dots, p_t^n)$ .

Marginal probabilities:  $p_t^i := \Pr(X_i = 1)$ .

#### Algorithm 1: UMDA

#### begin

initial model  $\mathcal{M}_0 = (1/2, 1/2, \dots, 1/2)$ 

#### repeat

Sample  $P_t$  of  $\lambda$  individuals using joint probability

$$\Pr(X_1, X_2, \dots, X_n) = \prod_{i=1}^n \Pr(X_i).$$

Select the  $\boldsymbol{\mu}$  best individuals according to fitness. Update the probabilistic model

$$p_{t+1}^i := \frac{1}{\mu} \sum_{i=1}^{\mu} X_i^{[j]} \in \left[\frac{1}{n}, 1 - \frac{1}{n}\right]$$

where  $X_i^{[j]}$  is the *i*-th bit of the *j*-th individual. **until** *termination condition fulfilled* 

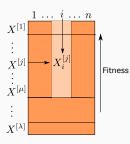


Figure 2: UMDA illustration

## **Previous Work**

	(1+1) EA	UMDA	
LeadingOnes	$\Theta(n^2)$	$\mathcal{O}(n\lambda\log\lambda+n^2)$	(Dang & Lehre 2015) <sup>1</sup>
BVLEADINGONES	$\Theta(n \log n)$	$\infty$ w.o.p.	(Chen et al. 2010)
		$\mathcal{O}(n\lambda)$	(Chen et al. $2010)^2$
SubString	$2^{\Omega(n)}$ w.o.p.	$\mathcal{O}(n\lambda)$	(Chen et al. 2009) <sup>4</sup>
ONEMAX	$\Theta(n \log n)$	$\mathcal{O}(n\lambda\log\lambda)$	(Dang & Lehre 2015) <sup>3</sup>
		$\Omega(\mu\sqrt{n} + n\log n)$	(Krejca & Witt 2017) <sup>3</sup>

 $<sup>\</sup>begin{array}{l}
1 \lambda = \Omega(\log n) \\
^2 \lambda = \omega(n^2) \\
^3 \lambda = (1 + \Theta(1))\mu
\end{array}$ 

## **Open Question**

### Runtime analysis of UMDA on ONEMAX (latest results):

- Upper bound:  $\mathcal{O}(n \log n \log \log n)$  by Dang & Lehre (Dang & Lehre 2015).
- Lower bound:  $\Omega(\mu\sqrt{n} + n\log n)$  by Krejca & Witt (Krejca & Witt 2017).

The upper and lower bounds are still different by  $\Theta(\log \log n)$ .

Open Question: Could this gap be closed?

#### UMDA on ONEMAX

#### Theorem

The expected optimisation time of  $\operatorname{UMDA}$  with

- $c \log n \le \mu \le c' \sqrt{n}$  for some constants c, c' > 0,
- $\lambda \ge a\mu$  for some constant  $a \ge 13e$

on ONEMAX is

$$\mathcal{O}(n\lambda)$$
.

Intuition: UMDA can optimise ONEMAX within  $\mathcal{O}(n)$  generations.

## UMDA on ONEMAX (n = 1000)

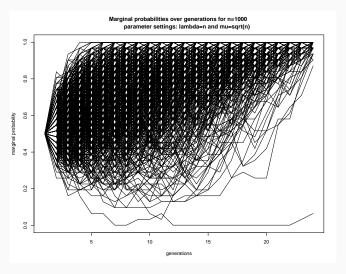
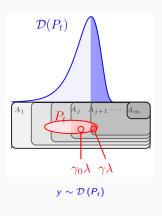


Figure 3: Marginal probabilities of UMDA on ONEMAX for n=1000

### Level-Based Theorem<sup>4</sup>

Provide **upper bounds** on the expected runtime of some population-based algorithms on a wide range of optimisation problems.



• (G1) 
$$j \in [m-1]$$
, if  $|P_t \cap A_{>i}| \ge \gamma_0 \lambda$  then

$$\Pr\left(\mathbf{y}\in A_{\geq j+1}\right)\geq z_{j}.$$

• (G2) similar to (G1) and  $|P_t \cap A_{\geq j+1}| \geq \gamma \lambda$  then

$$\Pr\left(y \in A_{\geq j+1}\right) \geq (1+\delta)\gamma.$$

• (G3) and the population size  $\lambda \in \mathbb{N}$  satisfies

$$\lambda \geq \left(rac{4}{\gamma_0\delta^2}
ight) \ln \left(rac{128m}{z_*\delta^2}
ight)$$

where  $z_* := \min_{j \in [m-1]} \{z_j\}$ , then expected runtime

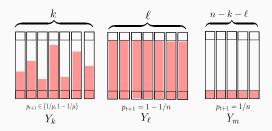
$$\mathbb{E}\left[T\right] \leq \left(\frac{8}{\delta^2}\right) \sum_{i=1}^{m-1} \left[\lambda \ln \left(\frac{6\delta\lambda}{4 + z_j\delta\lambda}\right) + \frac{1}{z_j}\right].$$

<sup>&</sup>lt;sup>4</sup>See Corus et al. (2016) for more details on the theorem.

## Proof idea (UMDA with margins)

- Level definition:  $A_j := \{x \in \{0,1\}^n \mid \text{OneMax}(x) = j 1\}.$
- Let  $A_i$  denote the current level.
- ullet We choose  $\gamma_0:=\mu/\lambda\Rightarrow$  all  $\mu$  best individuals have  $\geq j-1$  ones.
- Upgrade  $\Rightarrow$  Sample an offspring with  $\geq j$  ones, i.e.  $\Pr(Y \geq j)$ .
- Verify condition (G1):  $Pr(Y \ge j) \ge z_j$ .
- Verify condition (G2):  $Pr(Y \ge j) \ge (1 + \delta)\gamma$ .

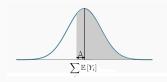
# Verify condition (G2)



The probability of sampling j one-bits:

$$\begin{split} \Pr(Y \geq j) \geq & \Pr(Y_k + Y_m \geq j - \ell) \cdot \Pr(Y_\ell = \ell) \\ \geq & \Pr(Z > j - \ell - 1) \cdot \left(1 - \frac{1}{n}\right)^{\ell} \qquad (Z := Y_k + Y_m) \\ \geq & \Pr\left(Z > \mathbb{E}\left[Z\right] - \frac{\gamma}{\gamma_0}\right) \cdot \frac{1}{e} \qquad \left(\mathbb{E}\left[Z\right] \geq j - \ell - 1 + \frac{\gamma}{\gamma_0}\right) \\ \geq & \dots \\ \geq & \dots \\ \geq & \dots \\ \geq & (1 + \delta)\gamma. \end{split}$$

# Feige's Inequality



### Theorem (Feige (2004))

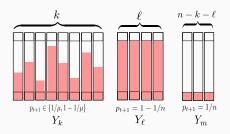
Given *n* independent r.v.  $Y_1, \ldots, Y_n \in [0,1]$ , then for all  $\Delta > 0$ 

$$\mathsf{Pr}\left(\sum_{i=1}^n Y_i > \sum_{i=1}^n \mathbb{E}\left[Y_i\right] - \Delta\right) \geq \min\bigg\{\frac{1}{13}, \frac{\Delta}{1+\Delta}\bigg\}.$$

The probability of sampling j one-bits now becomes

$$\begin{split} \Pr\left(Y \geq j\right) &\geq \Pr\left(Z > \mathbb{E}\left[Z\right] - \frac{\gamma}{\gamma_0}\right) \cdot \frac{1}{e} \\ &\geq \min\left\{\frac{1}{13}, \frac{\gamma/\gamma_0}{1 + \gamma/\gamma_0}\right\} \cdot \frac{1}{e} \\ &\geq \frac{\gamma}{13\gamma_0} \cdot \frac{1}{e} \geq (1 + \delta)\gamma. \end{split}$$

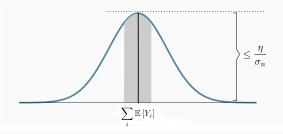
## Verify (G1): k is sufficiently large



The upgrade probability in this case:

$$\begin{split} \Pr(Y_k \geq j - \ell) &= \Pr(Y_k \geq j - \ell - 1) - \Pr(Y_k = j - \ell - 1) \\ &= \Pr(Y_k \geq \mathbb{E}[Y_k]) - \Pr(Y_k = j - \ell - 1) \\ \geq & \dots \\ > & \dots \end{split}$$

### **Anti-Concentration Bound**



### Theorem (Baillon et al. (2016))

Given n independent Bernoulli r.v.  $Y_1, \ldots, Y_n$  with success probabilities  $p_1, \ldots, p_n$ . For all n, y and  $p_i$ .

$$\Pr\left(\sum_{i=1}^n Y_i = y\right) \le \frac{\eta}{\sigma_n}$$

where  $\sigma_n^2 := \sum_{i=1}^n p_i (1-p_i)$  and  $\eta \sim 0.4688$  is an absolute constant.

# Verify (G1): k is sufficiently large

#### Theorem (Jogdeo & Samuels 1968)

Let  $Y_1,Y_2,\ldots,Y_n$  be n independent Bernoulli random variables. Let  $Y:=\sum_{i=1}^n Y_i$  be the sum of these random variables. If the expectation of Y is an integer, then

$$\Pr(Y \geq \mathbb{E}[Y]) \geq \frac{1}{2}.$$

Now the upgrade probability now becomes

$$\begin{split} \Pr\left(Y_k \geq j - \ell\right) &= \Pr\left(Y_k \geq j - \ell - 1\right) - \Pr\left(Y_k = j - \ell - 1\right) \\ &= \Pr\left(Y_k \geq \mathbb{E}\left[Y_k\right]\right) - \Pr(Y_k = j - \ell - 1) \\ &\geq \frac{1}{2} - \frac{\eta}{\sqrt{\text{Var}\left[Y_k\right]}} \\ &\geq \Omega(1). \end{split}$$

### Remaining cases

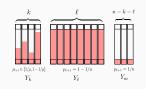
**CASE 2:** small k and large  $\ell$  (so is j)

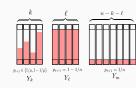
- Being very close to optimal level.
- If  $j \ge n + 1 \frac{n}{\mu}$ , then according to Dang & Lehre (2015)

$$\Pr(Y \ge j) = \Omega\left(\frac{1}{\mu}\right) = \Omega\left(\frac{n-j+1}{n}\right).$$

**CASE 3:** Small k and not close to optimal level (i.e. not too large j).

$$\begin{split} \Pr\left(Y \geq j\right) & \geq \Pr\left(Y_k \geq j - \ell - 1\right) \cdot \Pr\left(Y_\ell = \ell\right) \cdot \Pr\left(Y_m \geq 1\right) \\ & \geq \frac{1}{2} \cdot \frac{1}{e} \cdot \frac{n - k - \ell}{n} \\ & \geq \Omega\left(\frac{n - j + 1}{n}\right). \end{split}$$





# Condition (G3) and Expected runtime

Combining all three cases

$$z_j := \min \left\{ \Omega(1), \Omega\left(\frac{n-j+1}{n}\right) \right\} = \Omega\left(\frac{n-j+1}{n}\right).$$

(G3) satisfied iff

$$\lambda \geq \left(\frac{4}{\gamma_0 \delta^2}\right) \ln \left(\frac{128m}{z_* \delta^2}\right) = \Omega(\log n).$$

The expected optimisation time is

$$\begin{split} \mathcal{O}\left(\lambda\sum_{j=1}^{n}\ln\left(\frac{1}{z_{j}}\right)+\sum_{j=1}^{n}\frac{1}{z_{j}}\right) &= \mathcal{O}\left(\lambda\sum_{j=1}^{n}\ln\left(\frac{n}{n-j+1}\right)+\sum_{j=1}^{n}\frac{n}{n-j+1}\right) \\ &= \mathcal{O}\left(\lambda\ln\prod_{j=1}^{n}\frac{n}{n-j+1}+n\sum_{k=1}^{n}\frac{1}{k}\right) \\ &= \mathcal{O}\left(\lambda\ln\frac{n^{n}}{n!}+n\sum_{k=1}^{n}\frac{1}{k}\right) \quad \text{(Stirling's approximation}^{5}) \\ &= \mathcal{O}(n\lambda)+\mathcal{O}(n\log n) \\ &= \mathcal{O}(n\lambda). \end{split}$$

<sup>&</sup>lt;sup>4</sup>Stirling's approximation:  $n! \propto n^{n+0.5}e^{-n}$ 

### **Summary**

- The upper bound  $\mathcal{O}(n\lambda)$  holds for  $a\log(n) \leq \mu \leq a'\sqrt{n}$  where a,a' are some positive constants.
- The result finally closes the gap  $\Theta(\log \log n)$  between the first upper bound  $\mathcal{O}(n \log n \log \log n)$  (Dang & Lehre 2015) and recently discovered lower bound  $\Omega(\mu\sqrt{n} + n \log n)$  (Krejca & Witt 2017).
- Anti-concentration may be applied to analyse runtime of other algorithms.

Witt (2017) independently obtained the same upper bound  $\mathcal{O}(n\lambda)$  when  $\lambda=(1+\Theta(1))\mu$  and  $\mu\geq c\log(n)$ . Our result relaxes the proportional relationship between  $\lambda$  and  $\mu$  but convers smaller range of  $\mu$ .

### Thank you

Improved Runtime Bound for the
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