



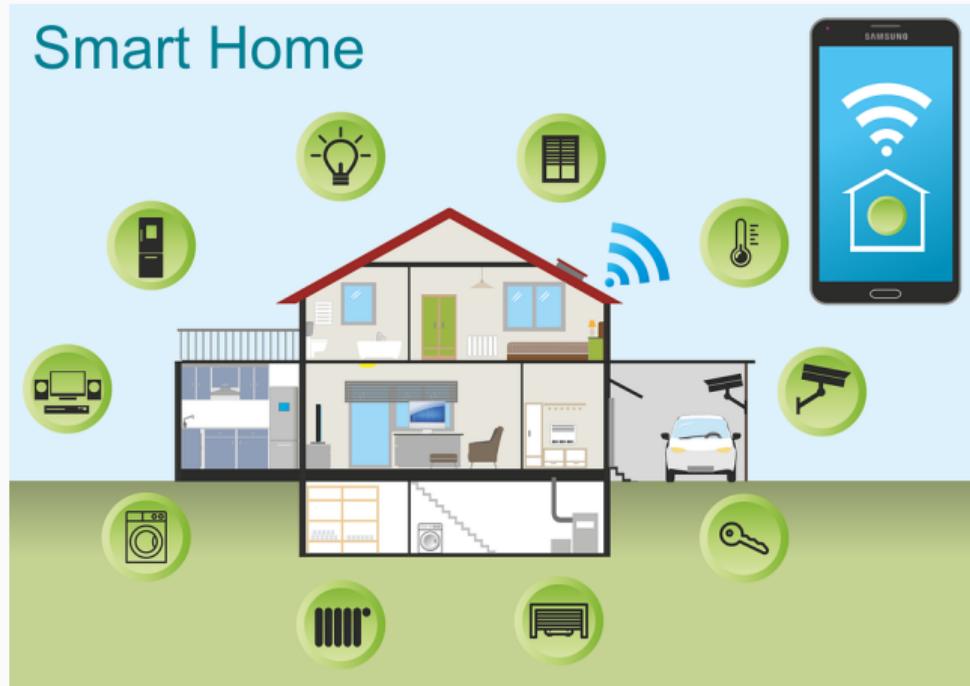
FUNDAMENTAL PERFORMANCE LIMITS OF STATISTICAL PROBLEMS: FROM DETECTION THEORY TO SEMI-SUPERVISED LEARNING

Ph.D. Thesis Defense

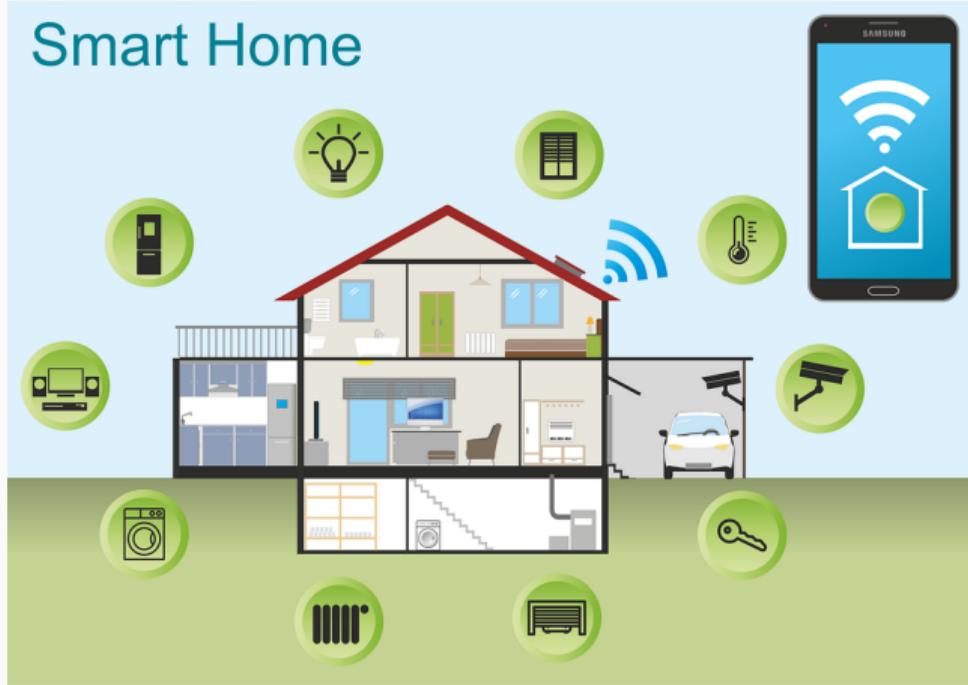
Candidate: Haiyun He

Department of Electrical and Computer Engineering

Smart Home

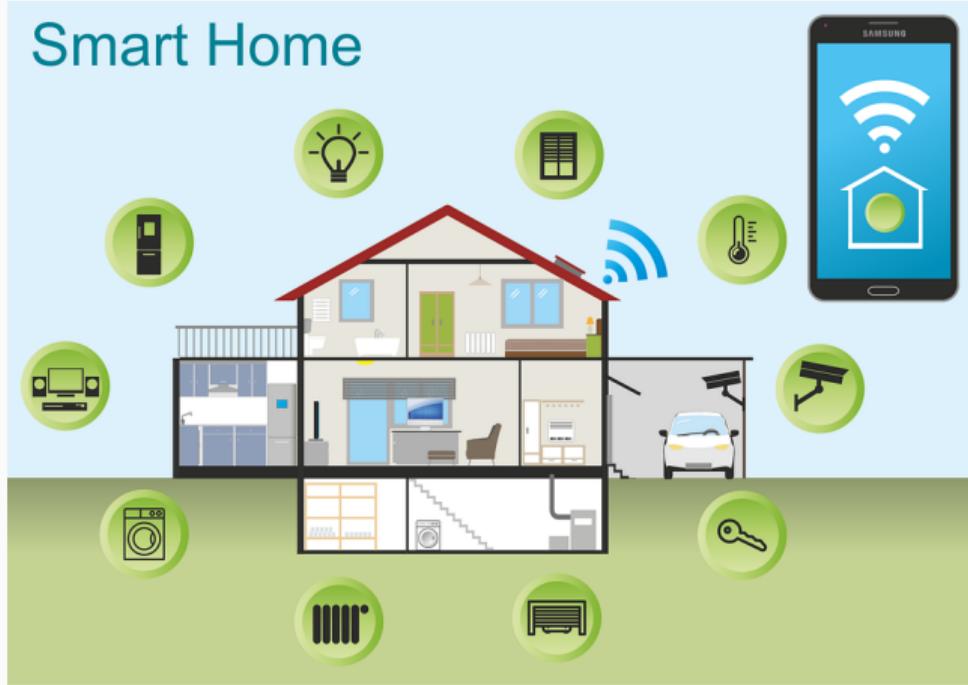


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One core problem: to design good mechanisms to infer or learn useful information from the raw data.

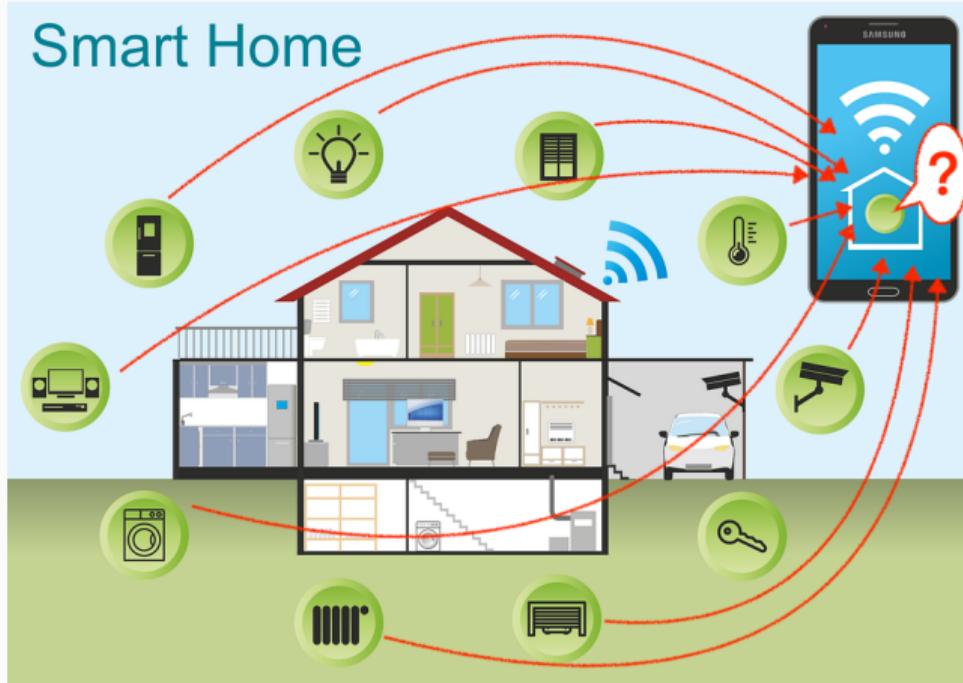
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Statistical viewpoint:

Motivation

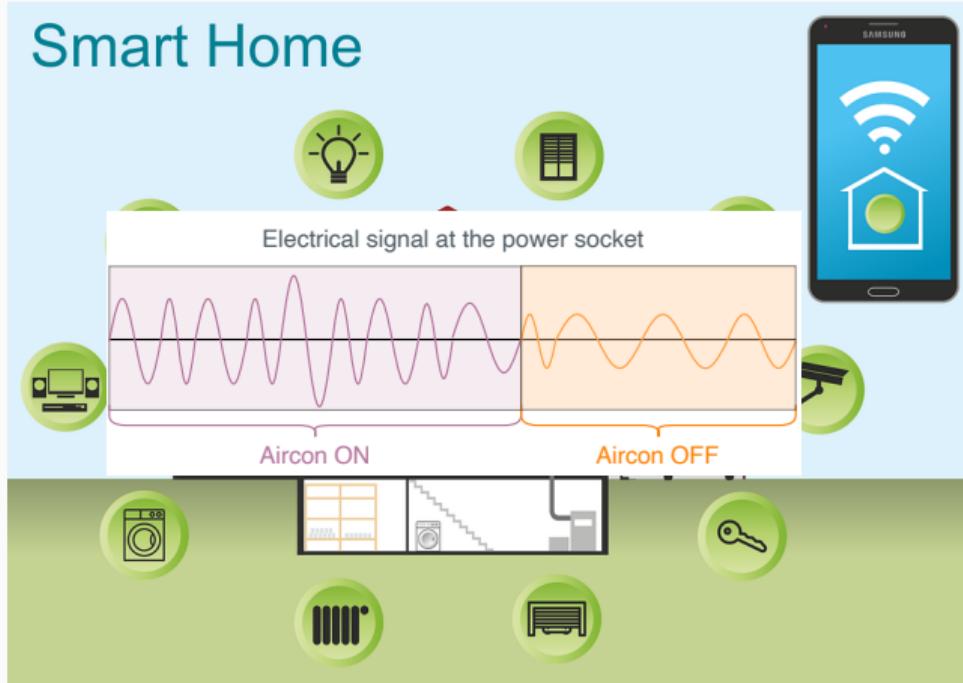


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Statistical viewpoint:

- Distributed detection

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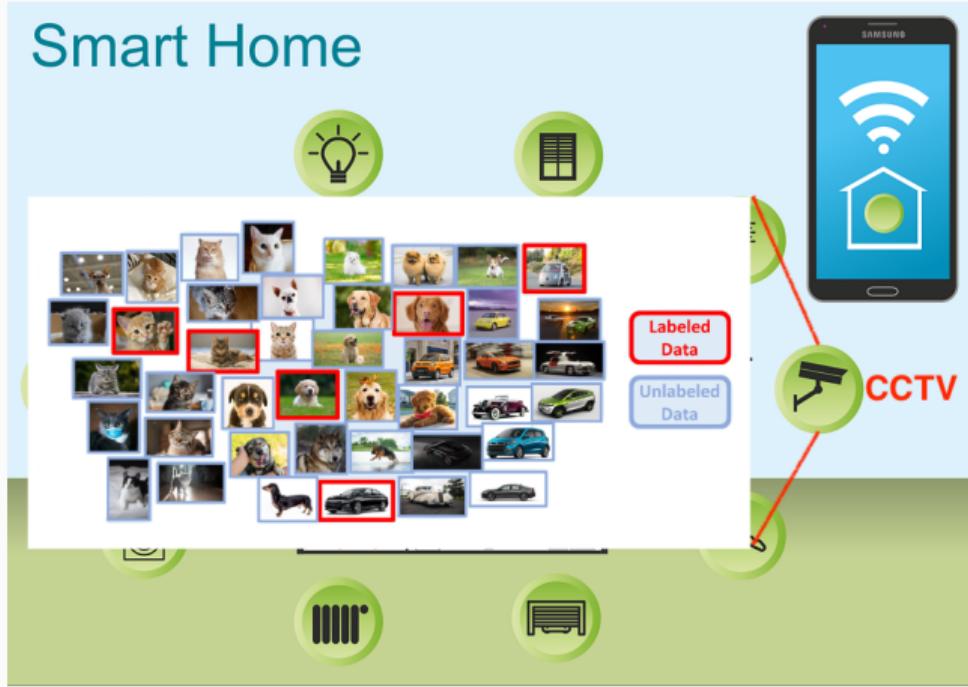


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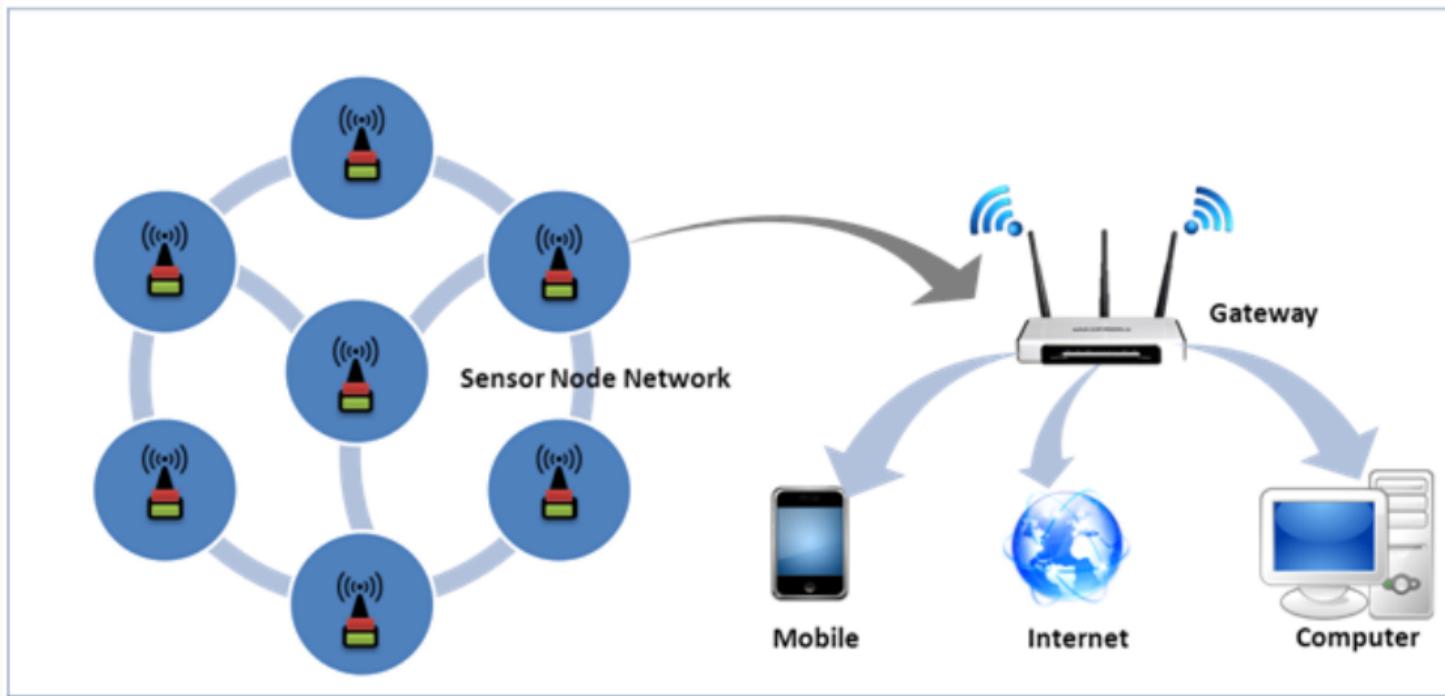
Statistical viewpoint:

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- Semi-supervised learning

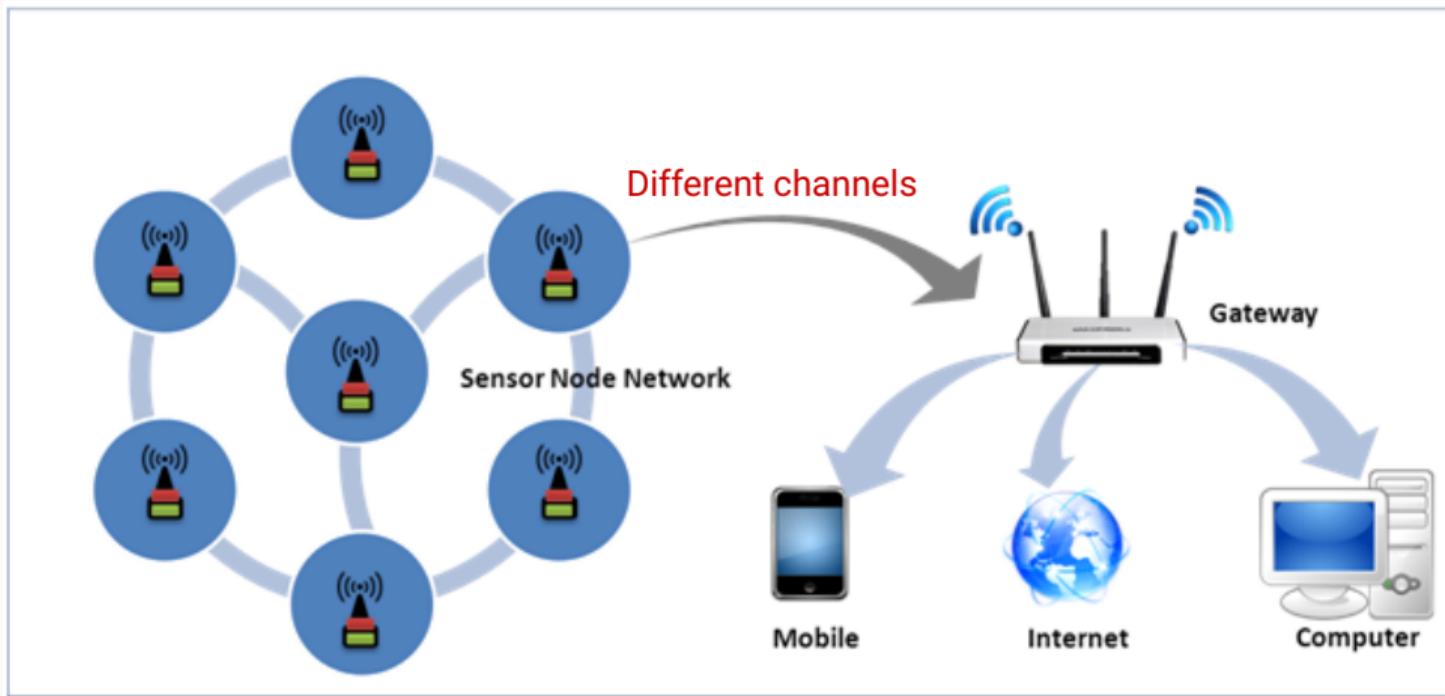
- 1 Distributed Detection with Empirically Observed Statistics**
- 2 Change-Point Detection with Training Sequences**
- 3 Information-Theoretic Generalization Error for Iterative Semi-Supervised Learning**

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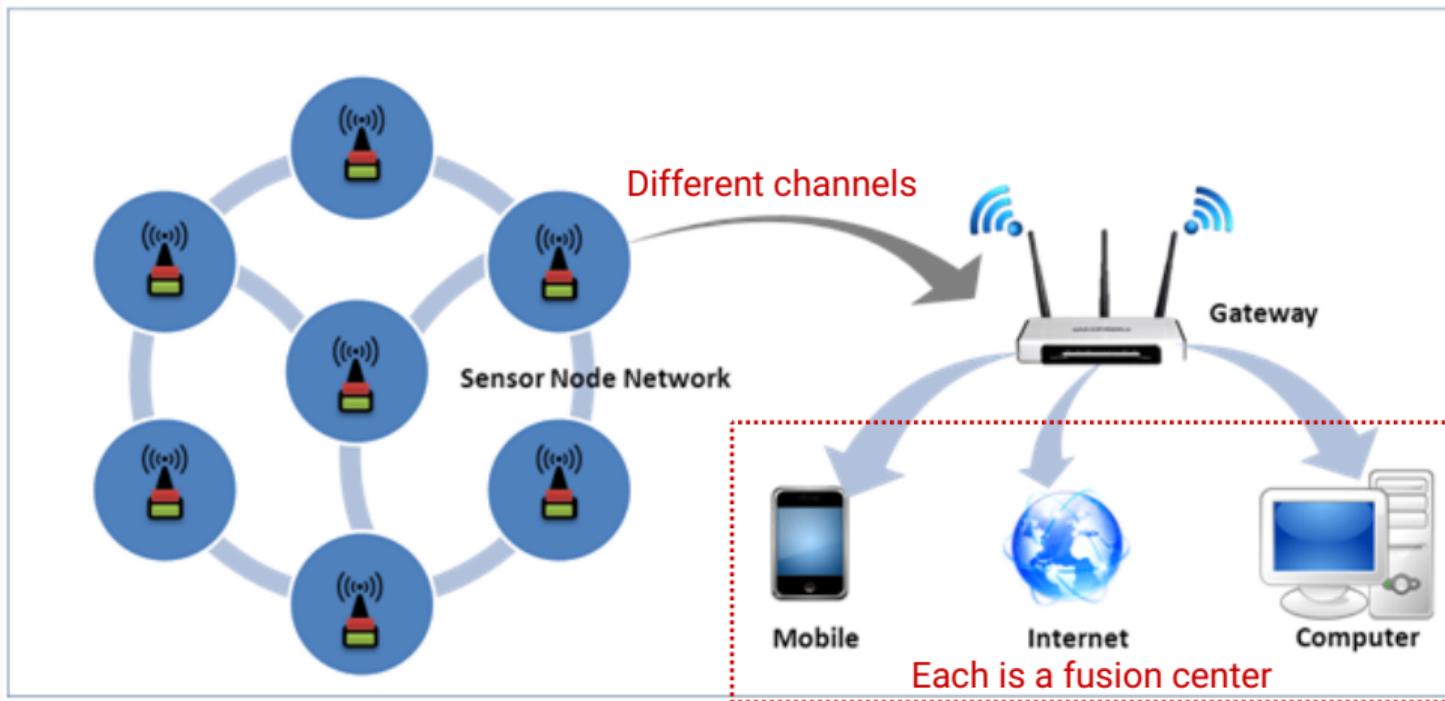
Background: Distributed Detection



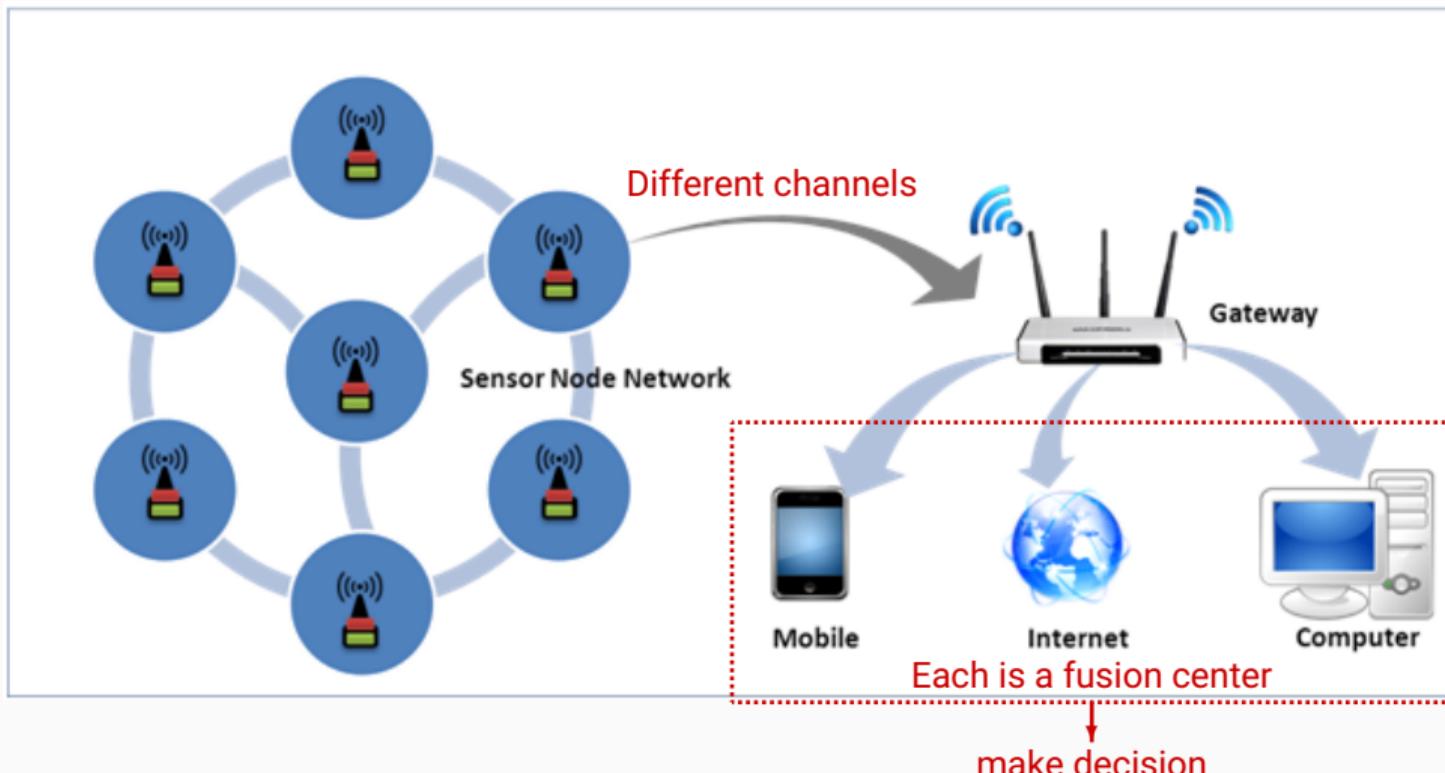
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Distributed Detection: Related Works

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Math. Control Signals Systems (1988) 1: 167–182

**Mathematics of Control,
Signals, and Systems**

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Decentralized Detection by a Large Number of Sensors*

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Abstract. We consider the decentralized detection problem, in which N independent, identical sensors transmit a finite-valued function of their observations to a fusion center which then decides which one of M hypotheses is true. For the case where the number of sensors tends to infinity, we show that it is asymptotically optimal to divide the sensors into $M(M - 1)/2$ groups, with all sensors in each group using the same decision rule in deciding what to transmit. We also show how the optimal number of sensors in each group may be determined by solving a mathematical programming problem. For the special case of two hypotheses and

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Theorem 1. *Subject to Assumption 1 below, $\lim_{N \rightarrow \infty} (Q_N - R_N) = 0$.*

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IEEE TRANSACTIONS ON INFORMATION THEORY, VOL. 35, NO. 2, MARCH 1989
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An approach which is closely related to information theory is to investigate asymptotically optimum tests, i.e., those tests with error exponents that asymptotically achieve optimal performance.

The optimum test for the case where the sources are known and where one of the misclassification probabilities is prescribed (case 1) was derived by Neyman and Pearson [2]. Its asymptotic behavior is described in [3] for independent identical distribution (i.i.d.) functions.

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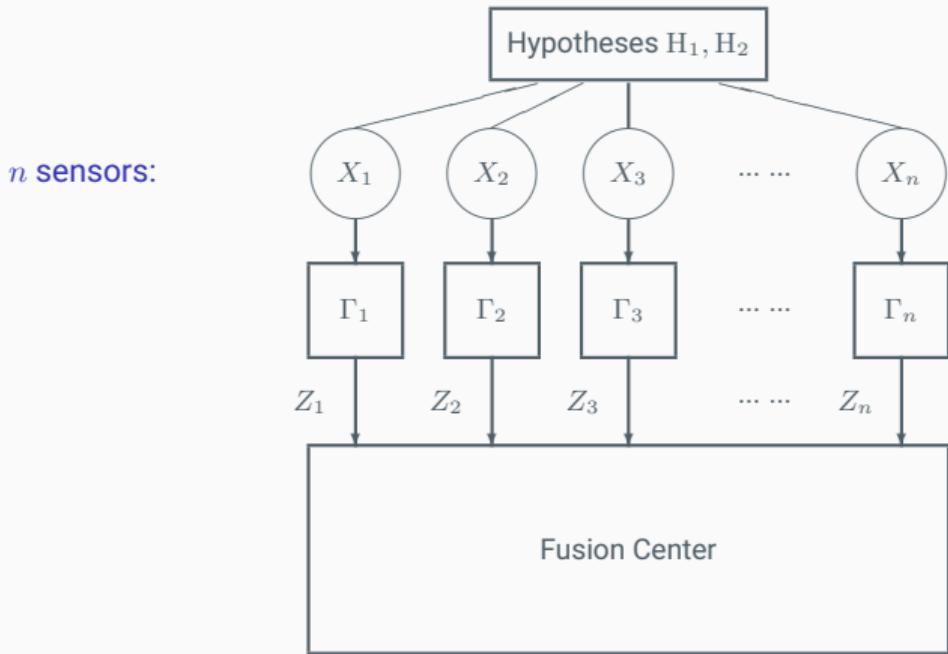
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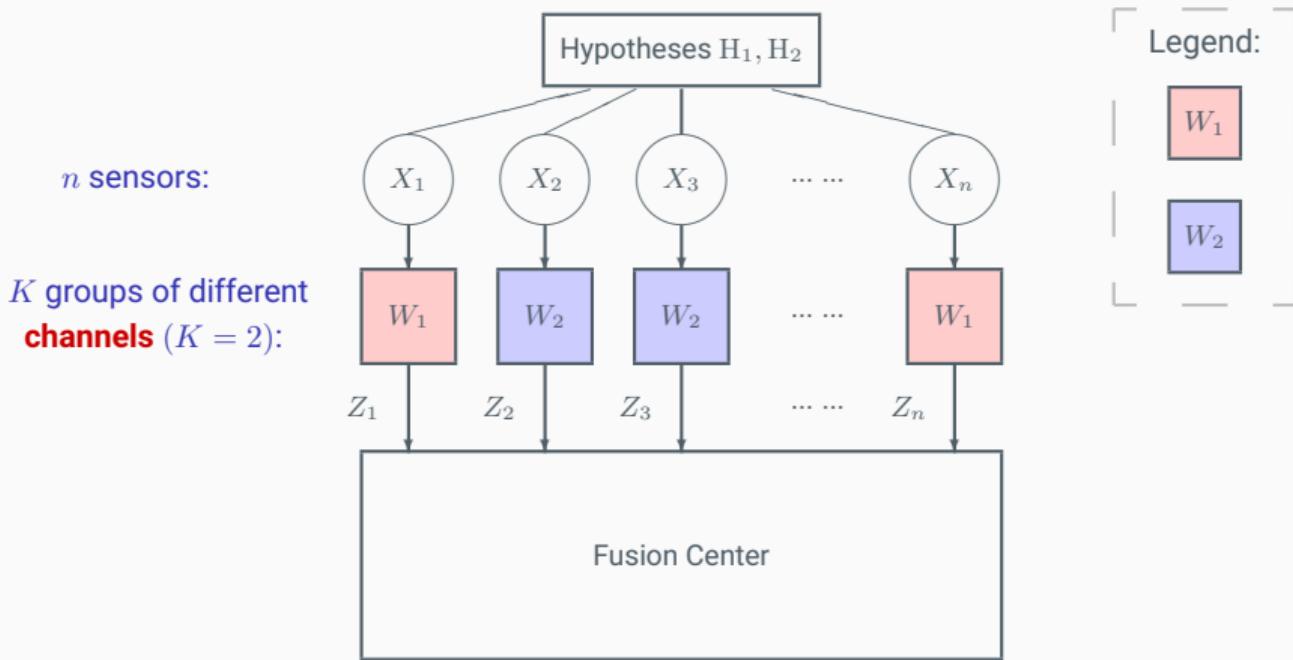
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★ Question: What is the optimal design of the channels and the decision rule at the fusion center?

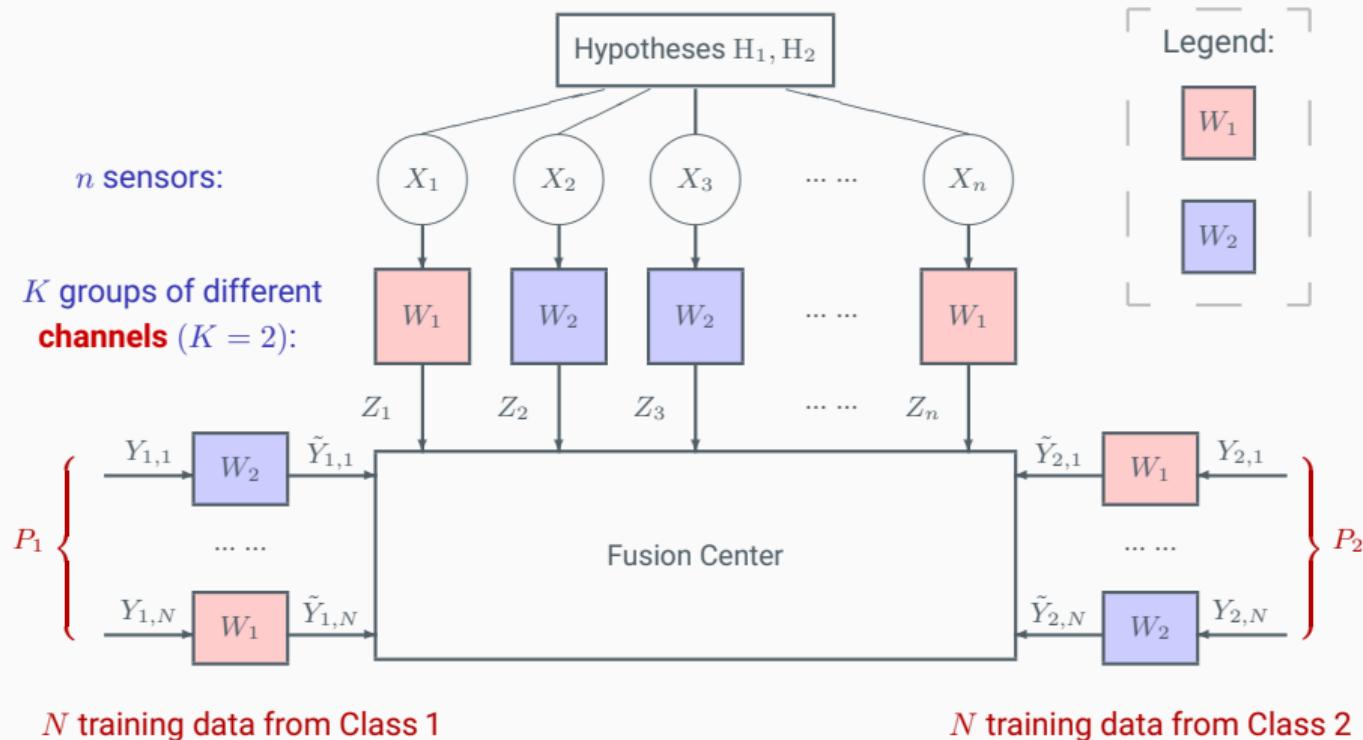
Distributed Detection: System Model



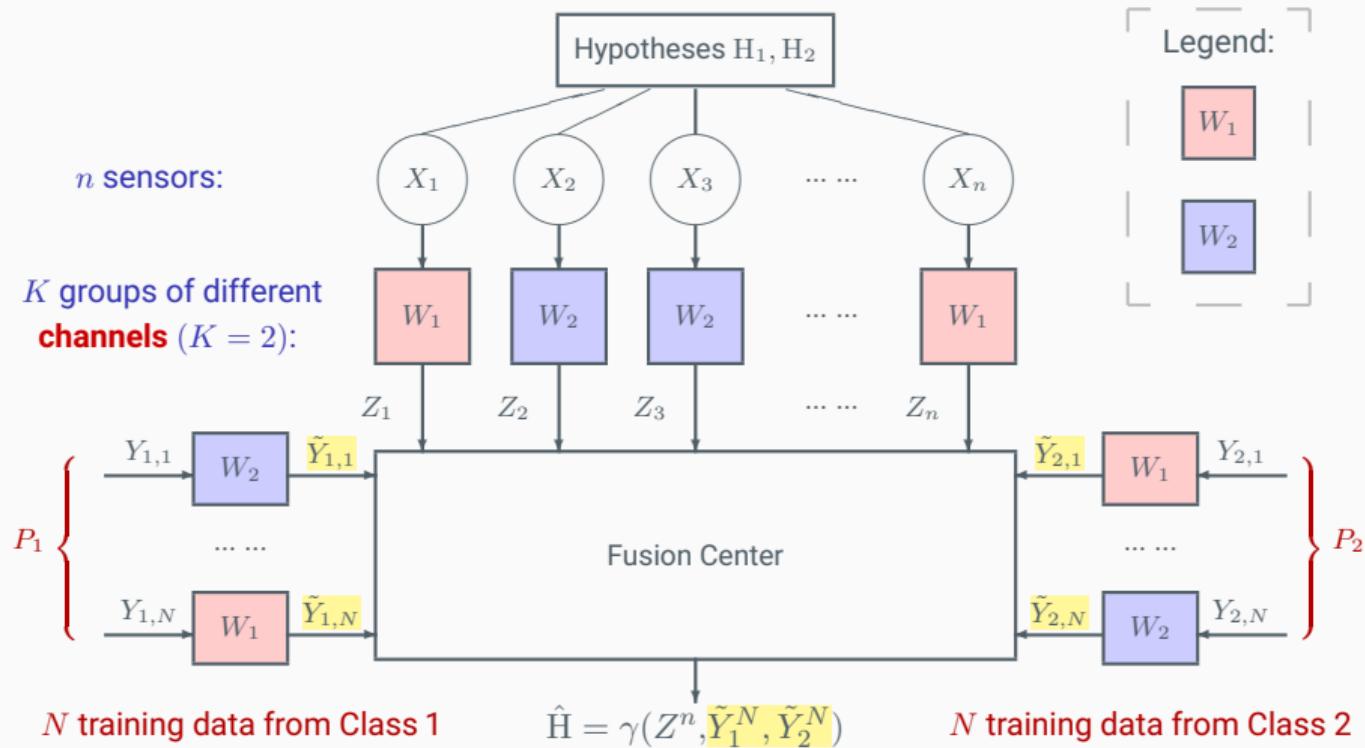
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Distributed Detection: System Model



- Ratio between lengths: $\alpha = \frac{N}{n}$

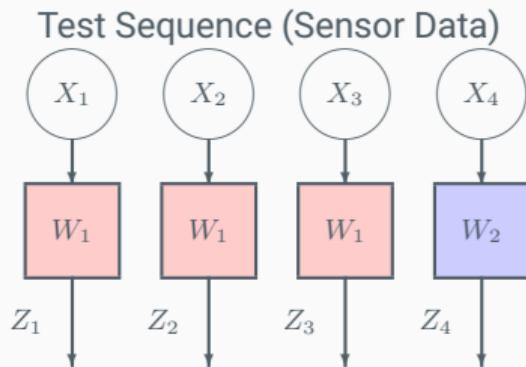
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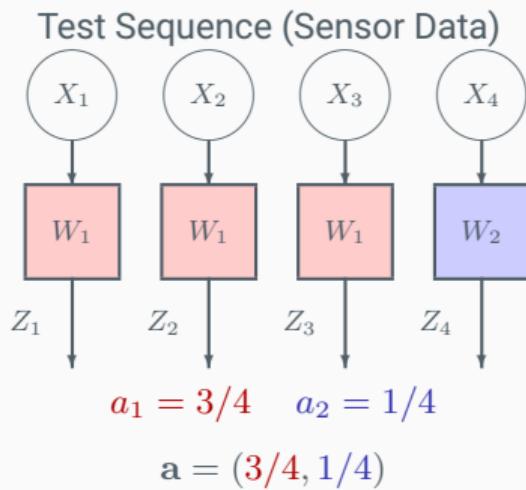
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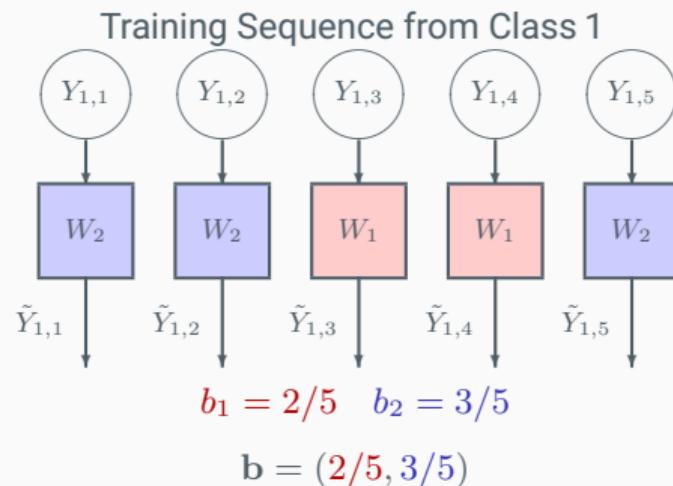
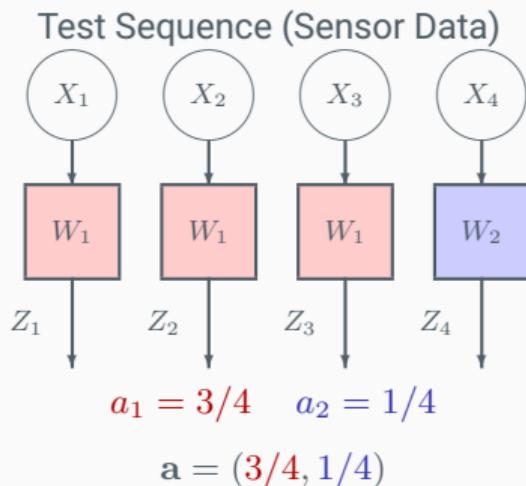
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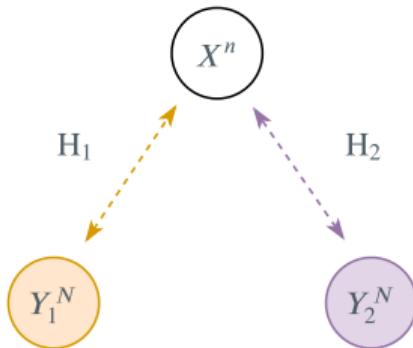
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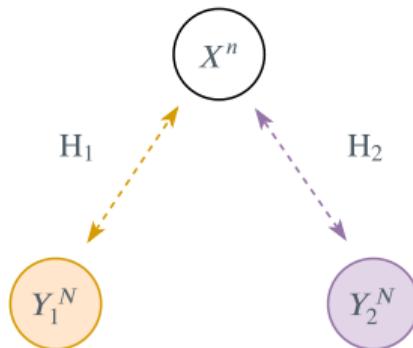


Distributed Detection: System Model

Fusion center decision rule γ : decide between the two hypotheses



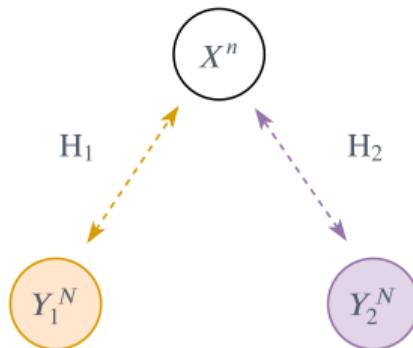
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- ★ Q1: Optimal fusion center decision rule γ given X^n, Y_1^N, Y_2^N and the channels $\{W_i\}_{i=1}^K$?

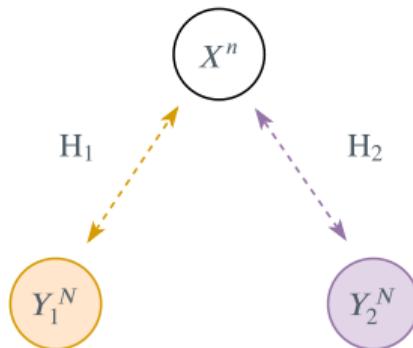
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- ★ Q2: Optimal error exponent?
- ★ Q3: Optimal proportions of different channels, i.e., $\mathbf{a} = (a_1, \dots, a_K), \mathbf{b} = (b_1, \dots, b_K)$?

Distributed Detection: Setup

- Type-I and type-II error probabilities:

$$\beta_j(\gamma, P_1, P_2) := \Pr\{\gamma(Z^n, \tilde{Y}_1^N, \tilde{Y}_2^N) \neq H_j \mid H_j\}, \quad j \in [2]$$

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- **Objective:** Consider the family $\Gamma_n(\lambda)$ of all tests γ s.t.

$$\max_{(\tilde{P}_1, \tilde{P}_2)} \beta_1(\gamma, \tilde{P}_1, \tilde{P}_2) \leq \exp(-n\lambda).$$

Given P_1, P_2 , we want to derive the **optimal type-II error exponent**

$$E^* := \liminf_{n \rightarrow \infty} \sup_{\gamma \in \Gamma_n(\lambda)} -\frac{1}{n} \log \beta_2(\gamma; P_1, P_2).$$

E^* depends on train/test ratio $\alpha = \frac{N}{n}$, **type-I error exponent** λ , ratios of channels $\mathbf{a} = (a_1, \dots, a_K)$, $\mathbf{b} = (b_1, \dots, b_K)$, and distributions P_1, P_2 (which will be suppressed).

Distributed Detection: Some Definitions

- Linear combinations of KL-divergences

$$\text{LD}(\mathbf{Q}, \tilde{\mathbf{Q}}, P, \tilde{P} | \alpha, \mathbf{a}, \mathbf{b}, \mathcal{W}) := \sum_{k \in [K]} (a_k D(Q_k \| PW_k) + \alpha b_k D(\tilde{Q}_k \| \tilde{P}W_k)),$$

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- Set of distributions:

$$\mathcal{Q}_\lambda(\alpha, \mathbf{a}, \mathbf{b}, \mathcal{W}) := \left\{ (\mathbf{Q}, \tilde{\mathbf{Q}}) : \min_{\tilde{P} \in \mathcal{P}(\mathcal{X})} \text{LD}(\mathbf{Q}, \tilde{\mathbf{Q}}, \tilde{P}, \tilde{P}) \leq \lambda \right\}.$$

When $K = 1$ and $W_1 = I_{|\mathcal{X}| \times |\mathcal{X}|}$ \implies recovers to Gutman's classification problem setup

Distributed Detection: Main Results ($n \rightarrow \infty$)

Theorem 1 (Asymptotically optimal type-II error exponent)

Given any pair of target distributions (P_1, P_2) , we have

$$E^*(\lambda, \alpha, \mathbf{a}, \mathbf{b}) = \min_{(\mathbf{Q}, \tilde{\mathbf{Q}}) \in \mathcal{Q}_\lambda(\alpha, \mathbf{a}, \mathbf{b}, V, \mathcal{W})} \text{LD}(\mathbf{Q}, \tilde{\mathbf{Q}}, P_2, P_1).$$

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In the achievability proof, we use the **asymptotically optimal** fusion center type-based test:

$$\gamma(Z^n, \tilde{Y}_1^N, \tilde{Y}_2^N) = \begin{cases} H_1 & \text{if } \min_{\tilde{P}} \text{LD}\left(\{T_{Z^{n a_k}}\}_{k \in [K]}, \{T_{\tilde{Y}_1^{N b_k}}\}_{k \in [K]}, \tilde{P}, \tilde{P}\right) \leq \lambda, \\ H_2 & \text{otherwise.} \end{cases}$$

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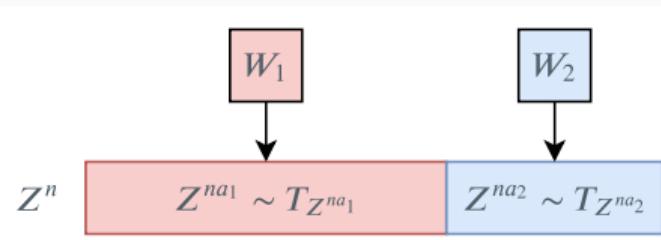
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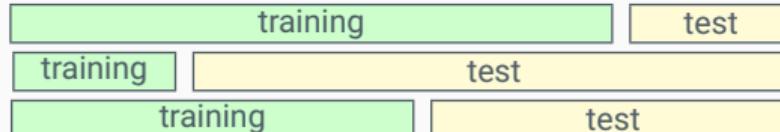
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- Maximized over (a, b)
- Three cases:

- $\alpha \rightarrow \infty$:
- $\alpha \rightarrow 0$:
- α moderate:



(Details in full thesis)

Further discussions on (\mathbf{a}, \mathbf{b}) : $\alpha \rightarrow \infty$

training test

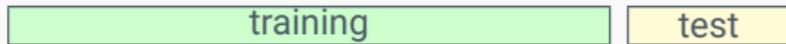
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Given any $\lambda \in \mathbb{R}_+$, as $\alpha \rightarrow \infty$, we have

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standard basis vector: \mathbf{e}_k

1	$k - 1$	k	$k + 1$	K
0	0	1	0	0

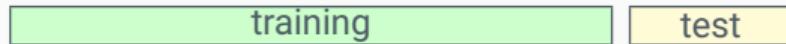
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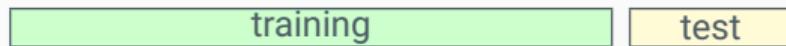
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⇒ analogous to Tsitsiklis' result

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training

test

Lemma 1

Given any $(\mathbf{a}, \mathbf{b}) \in \mathcal{P}([K])^2$ and any $\lambda \in \mathbb{R}_+$, $\exists \alpha_0(\mathbf{a}, \mathbf{b}, \lambda) > 0$, if $\alpha \leq \alpha_0(\mathbf{a}, \mathbf{b}, \lambda)$, then

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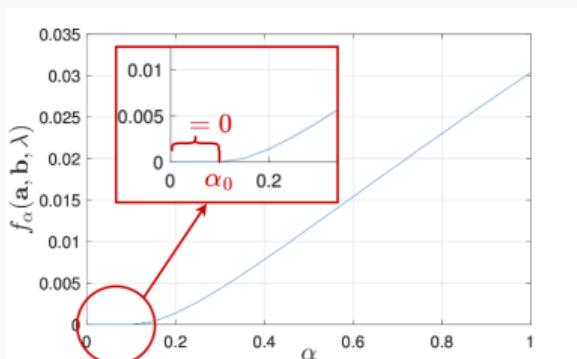
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H. He, L. Zhou, and V. Y. F. Tan, "Distributed detection with empirically observed statistics", *IEEE Transactions on Information Theory*, vol. 66, pp. 4349–4367, 2020.

1

Distributed Detection with Empirically Observed Statistics

2

Change-Point Detection with Training Sequences

3

Information-Theoretic Generalization Error for Iterative Semi-Supervised Learning

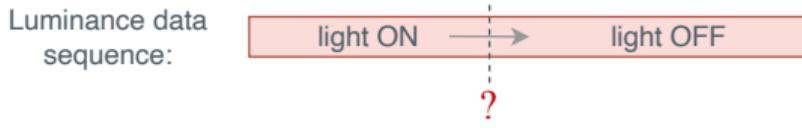
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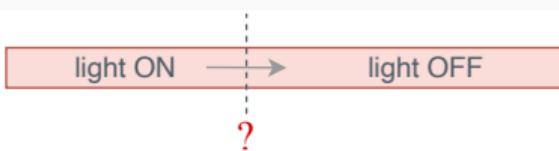


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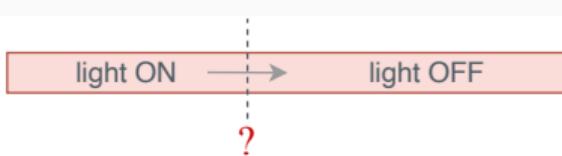
Luminance data
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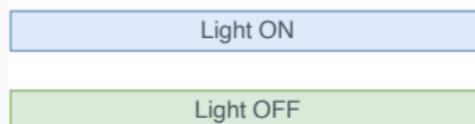
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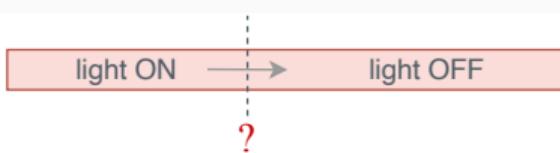
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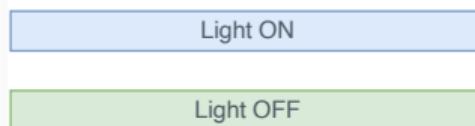
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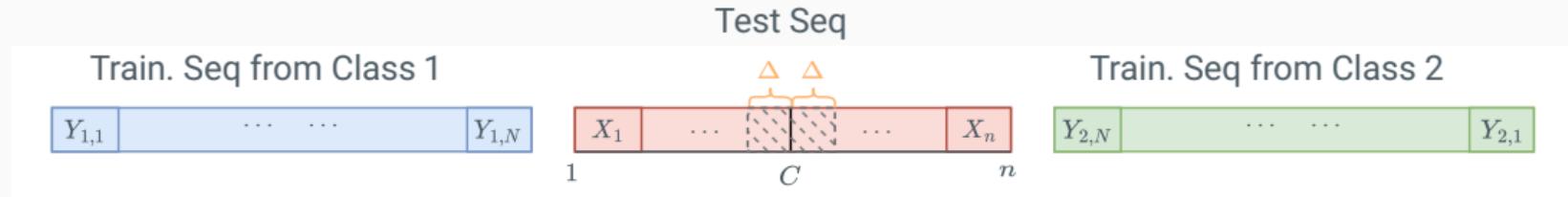
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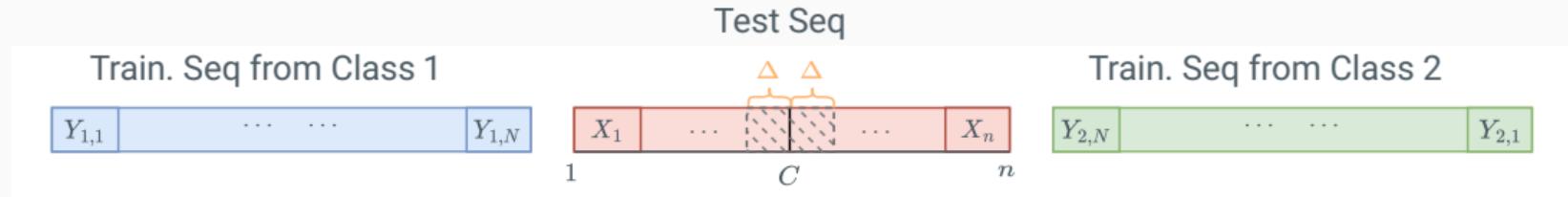
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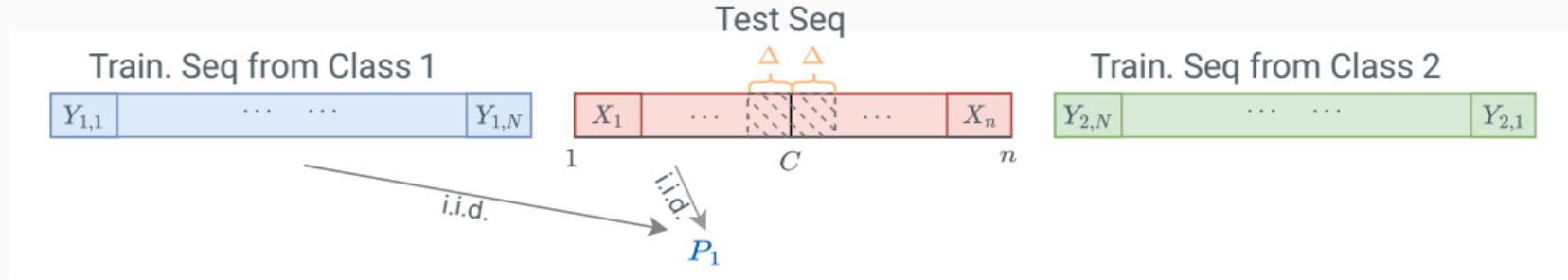
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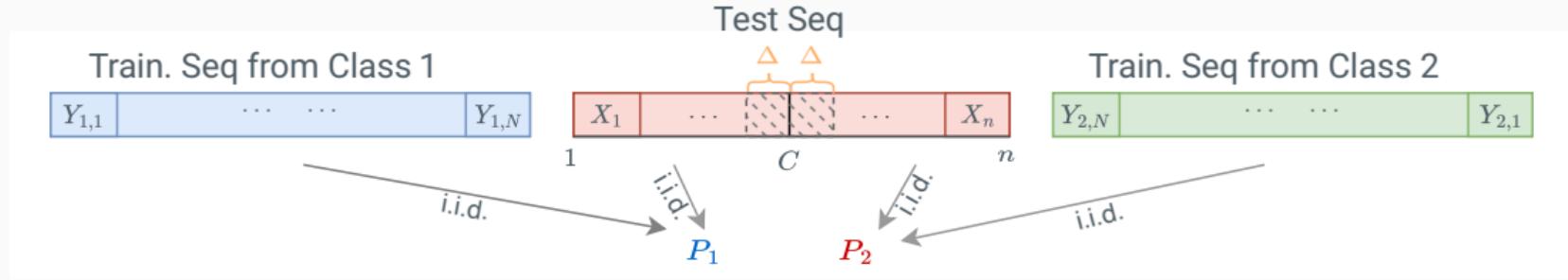
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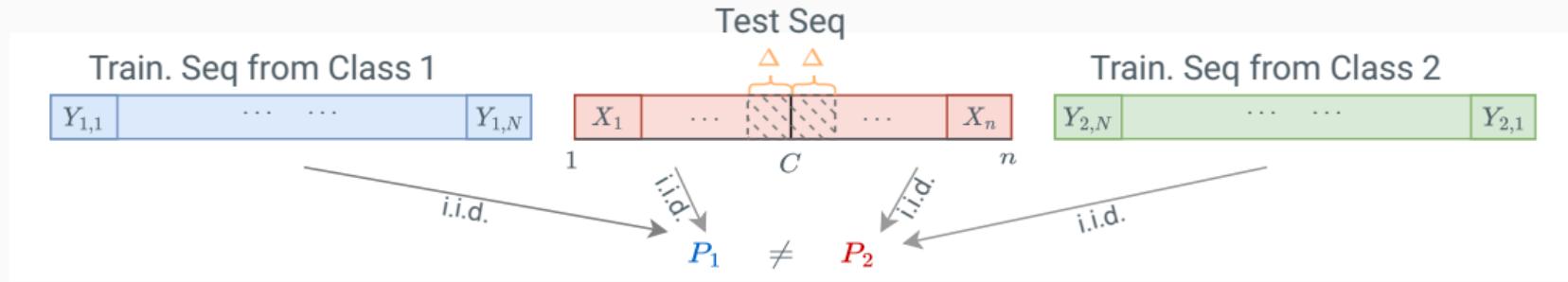
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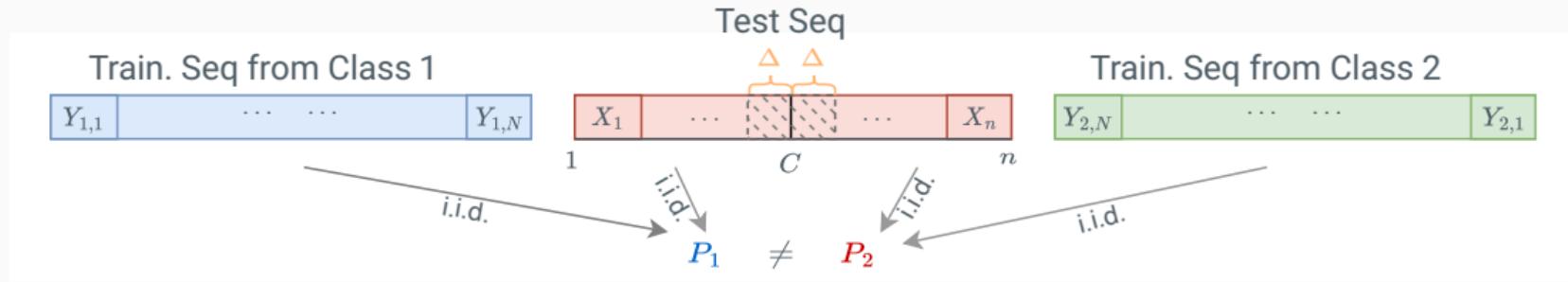
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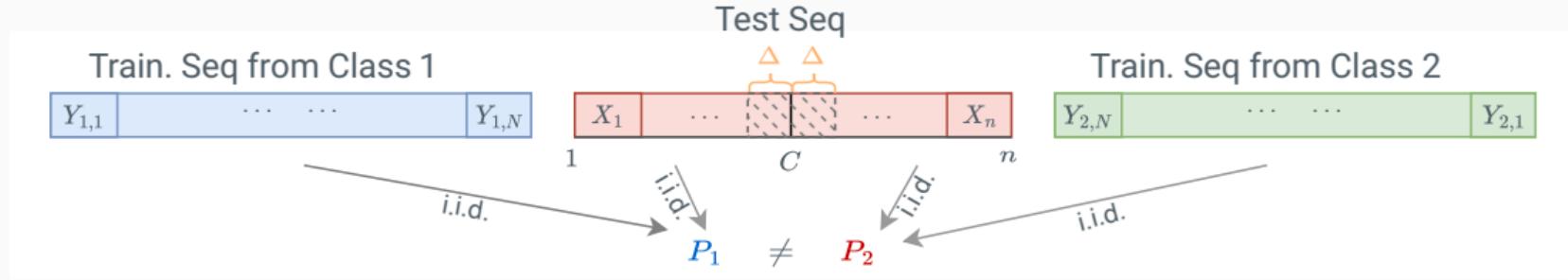
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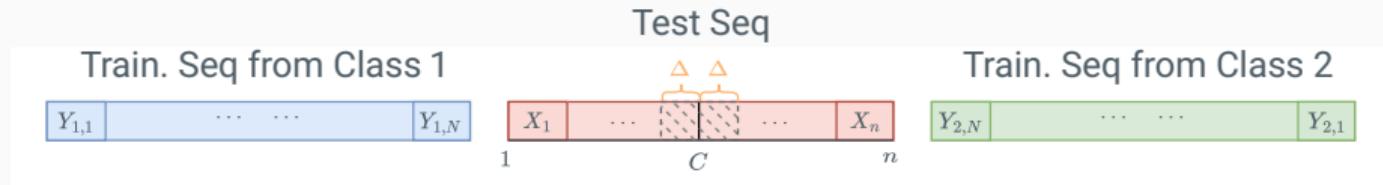


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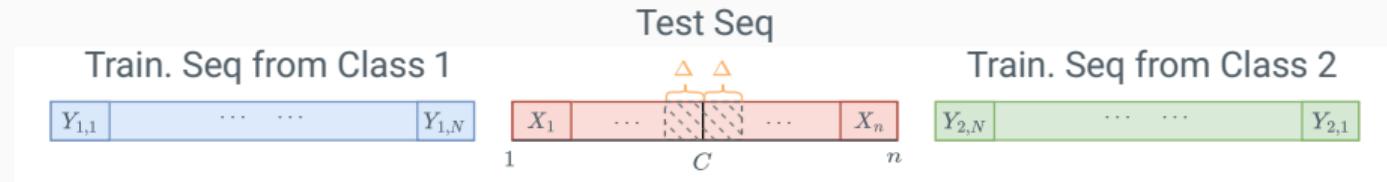


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- An estimator $\gamma : \mathcal{X}^{n+2N} \mapsto [n] \cup \{\text{e}\}$:

$$\begin{cases} \text{either declare one of } n \text{ points in the test sequence} \\ \text{or declare that an "erasure" has occurred} \end{cases}$$



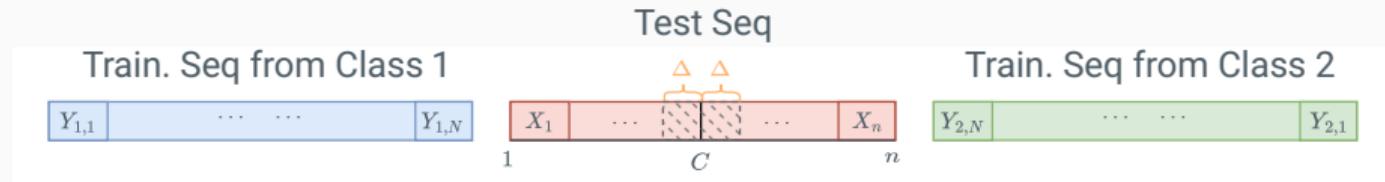
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Definition 1 (Good Estimator)

For any $\Delta \in [0, n/2]$, any $r \in \mathbb{R}_+$, any $(\lambda, \epsilon) \in \mathbb{R}_+ \times [0, 1]$, and any $t \in [0, 1/2]$, given any particular pair $(P_1, P_2) \in \mathcal{P}(\mathcal{X})^2$, an estimator $\gamma : \mathcal{X}^{n+2N} \mapsto [n] \cup \{\text{e}\}$ is said to be **$(n, \Delta, r, \lambda, \epsilon, t)$ -good** if

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- $t \in (0, 1/2)$: decay **subexponentially fast**, moderate deviations regime
- **Goal: what is the smallest Δ a good estimator can achieve?**

Theorem 2 (Optimal confidence width)

For any $r \in \mathbb{R}_+$, $\epsilon \in [0, 1]$, any pair of distributions $(P_1, P_2) \in \mathcal{P}(\mathcal{X})^2$, the optimal NCW is

$$\bar{\Delta}^*(r, \lambda, P_1, P_2) = \begin{cases} G_{\min}^{-1}(\lambda), & \lambda \in \left(0, G_{\min}\left(\frac{1}{2}\right)\right), \\ \frac{1}{2}, & \text{otherwise;} \end{cases} \quad (\text{G_{\min} is based on Jensen-Shannon divergence and P_1, P_2})$$

(\$\lambda\$ is the undetected error exponent)

In the moderate deviations regime, the *t-optimal NCW* for any $t \in (0, 1/2)$ and $\lambda > 0$ is

$$\bar{\Delta}_t^*(r, \lambda, P_1, P_2) = \max_{\alpha \in [0, 1]} \frac{\sqrt{\lambda} \left(\sqrt{\alpha(\alpha + r)} \chi_2(P_1 \| P_2) + \sqrt{(1 - \alpha)(1 - \alpha + r)} \chi_2(P_2 \| P_1) \right)}{\sqrt{2r \chi_2(P_1 \| P_2) \chi_2(P_2 \| P_1)}}.$$

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For any $t \in [0, 1/2]$, $\bar{\Delta}_t^*(r, \lambda, P_1, P_2)$ is independent of $\epsilon \implies$ strong converses hold.

* Refer to the full thesis for the asymptotically optimal estimator

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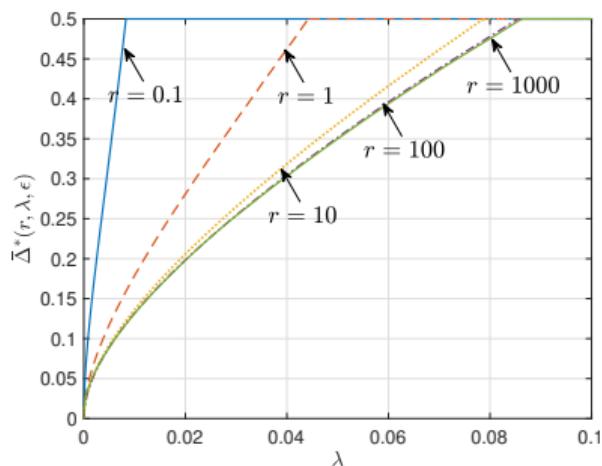


Fig: Large deviations regime.

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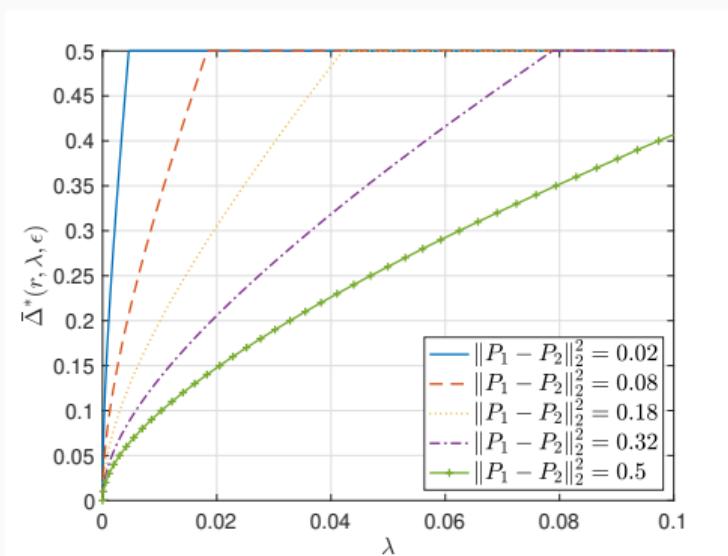
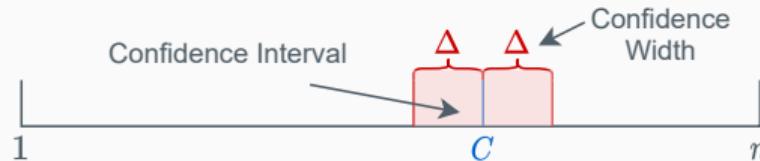


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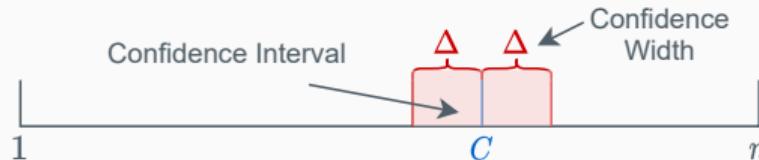
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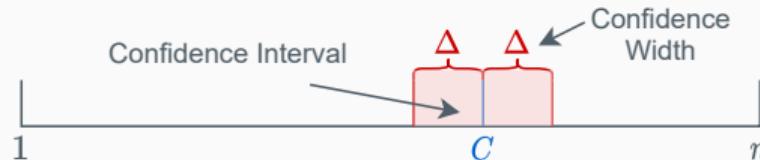
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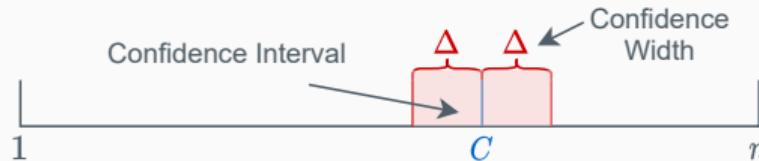
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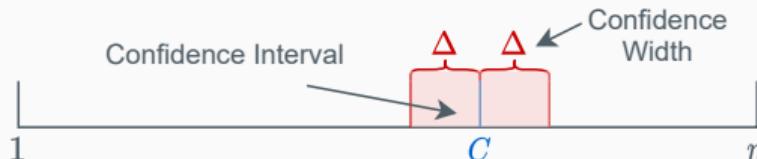
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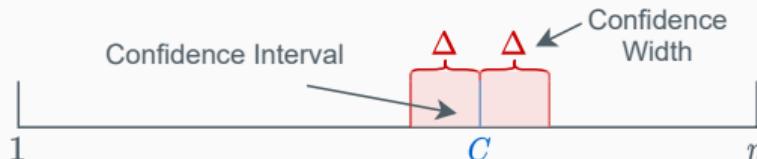


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H. He, Q. Zhang, and V. Y. F. Tan, "Optimal change-point detection with training sequences in the large and moderate deviations regimes", *IEEE Transactions on Information Theory*, vol. 67, no. 10, pp. 6758–6784, 2021.

- 1** Distributed Detection with Empirically Observed Statistics
- 2** Change-Point Detection with Training Sequences
- 3** Information-Theoretic Generalization Error for Iterative Semi-Supervised Learning

Semi-supervised learning (SSL) algorithms

a small amount of labelled data + a large amount of unlabelled data



Figure: An example of SSL.^{1,2}

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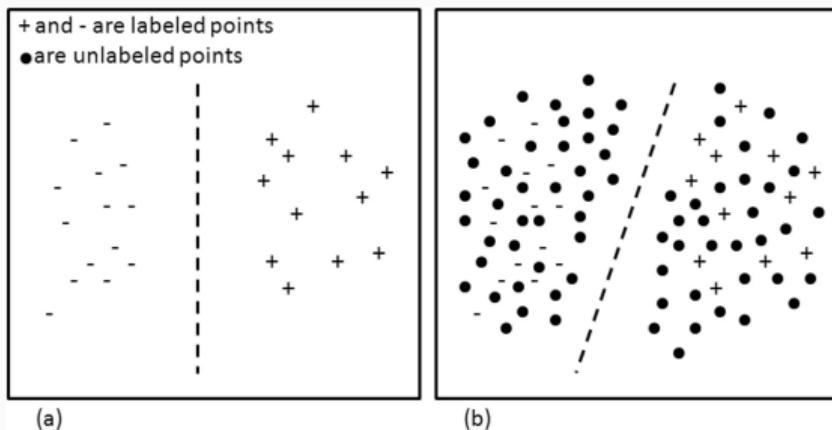
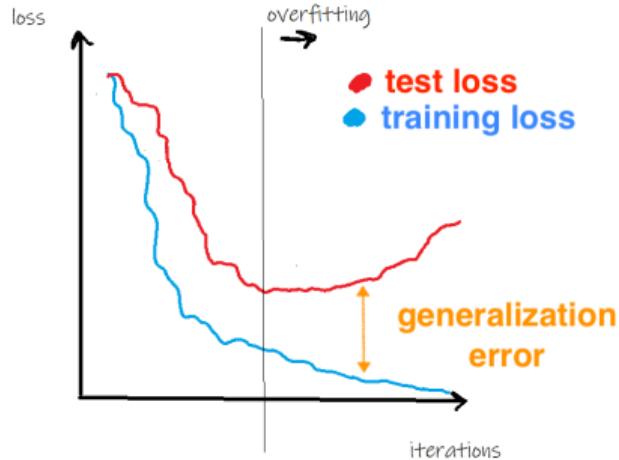


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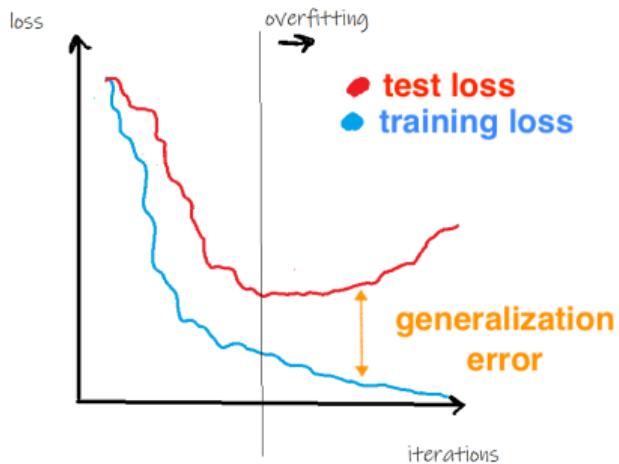
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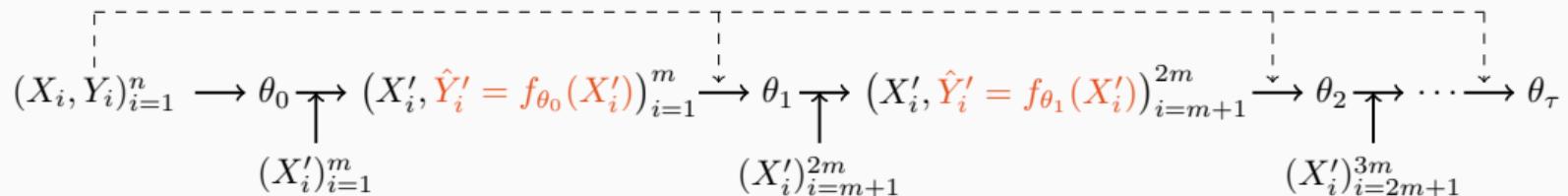
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Theorem 3 (Bu et al. 2020)

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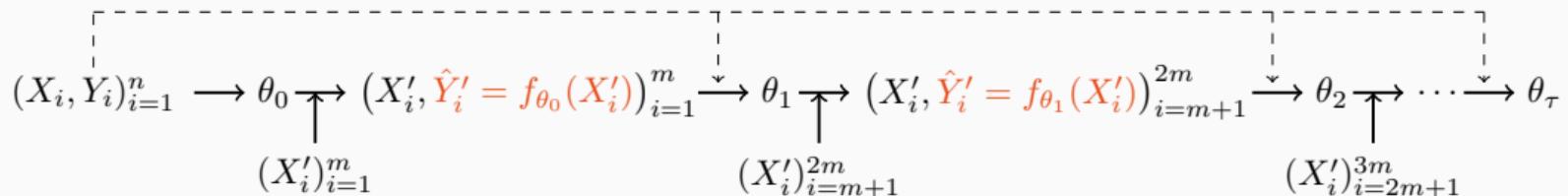
$$|\text{gen}| \leq \frac{1}{n} \sum_{i=1}^n \sqrt{2R^2 I(W; Z_i)}.$$

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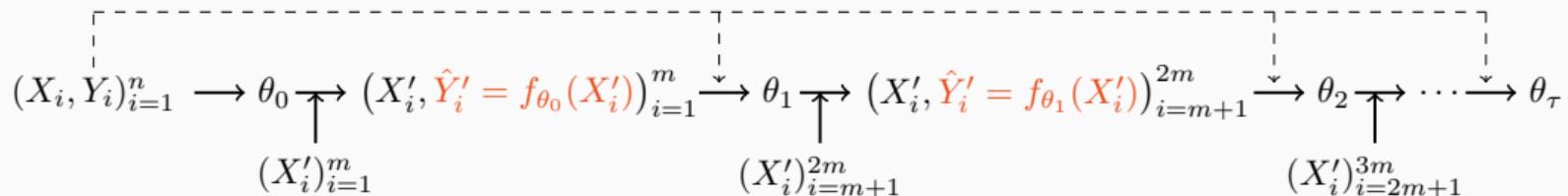
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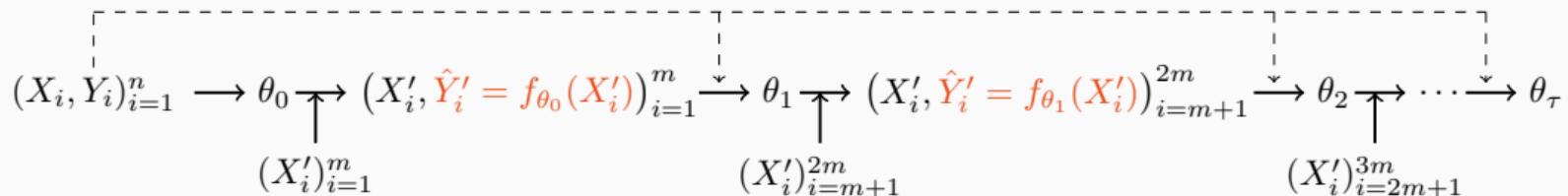
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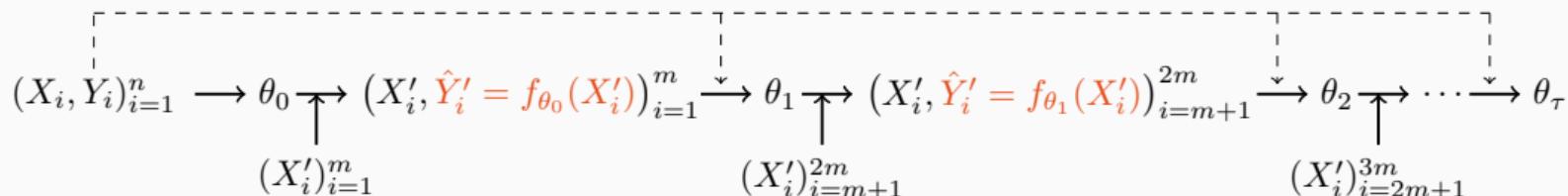
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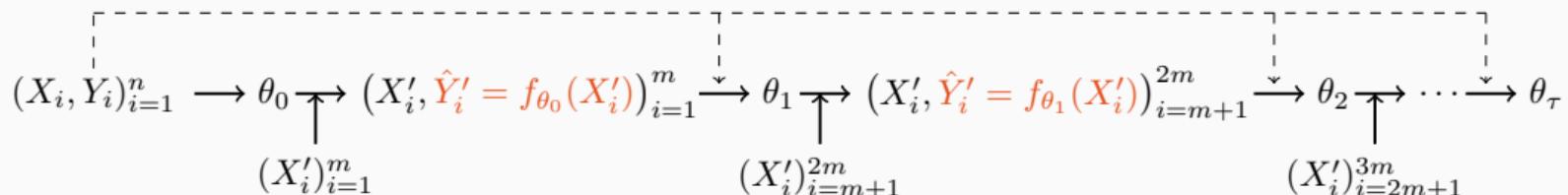
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Information-Theoretic Gen-Error for Iterative SSL Problem Setup

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Total empirical risk: $w = \frac{n}{n+m}$

$$\begin{aligned} L_{S_1, \hat{S}_{u,t}}(\theta_t) &:= w L_{S_1}(\theta_t) + (1 - w) L_{\hat{S}_{u,t}}(\theta_t) \\ &= \frac{1}{n+m} \left(\sum_{i=1}^n l(\theta_t, Z_i) + \sum_{i \in \mathcal{I}_t} l(\theta_t, (X'_i, \hat{Y}'_i)) \right). \end{aligned}$$

Generalization error at the t -th iteration: the expected gap between the **population risk** of θ_t and the **empirical risk** on the training data

$$\begin{aligned} \text{gen}_t(P_Z, P_X, \{P_{\theta_k|S_1, S_u}\}_{k=0}^t, \{f_{\theta_k}\}_{k=0}^{t-1}) &:= \mathbb{E}[L_{P_Z}(\theta_t) - L_{S_1, \hat{S}_u, t}(\theta_t)] \\ &= w \left(\mathbb{E}_{\theta_t} [\mathbb{E}_Z [l(\theta_t, Z) \mid \theta_t]] - \frac{1}{n} \sum_{i=1}^n \mathbb{E}_{\theta_t, Z_i} [l(\theta_t, Z_i)] \right) \\ &\quad + (1-w) \left(\mathbb{E}_{\theta_t} [\mathbb{E}_Z [l(\theta_t, Z) \mid \theta_t]] - \frac{1}{m} \sum_{i \in \mathcal{I}_t} \mathbb{E}_{\theta_t, X'_i, \hat{Y}'_i} [l(\theta_t, (X'_i, \hat{Y}'_i))] \right). \end{aligned}$$

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Questions

- ★ How does gen_t evolve as the iteration count t increases?
- ★ Do the unlabelled data examples in S_u help to improve the generalization error?

General Results for Generalization Error

Theorem 1.A (Gen-error upper bound for iterative SSL)

Suppose $l(\theta, Z) \sim \text{subG}(R)$ under $Z \sim P_Z$ for all $\theta \in \Theta$, then for any $t \in [0 : \tau]$,

$$\begin{aligned} |\text{gen}_t| &\leq \frac{w}{n} \sum_{i=1}^n \mathbb{E}_{\theta^{(t-1)}} \left[\sqrt{2R^2 I_{\theta^{(t-1)}}(\theta_t; Z_i)} \right] \\ &+ \frac{1-w}{m} \sum_{i=(t-1)m+1}^{tm} \mathbb{E}_{\theta^{(t-1)}} \left[\sqrt{2R^2 (I_{\theta^{(t-1)}}(\theta_t; X'_i, \hat{Y}'_i) + D_{\theta^{(t-1)}}(P_{X'_i, \hat{Y}'_i} \| P_Z))} \right]. \end{aligned}$$

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♣ Follows from Bu et al. (2020, Theorem 1) and Wu et al. (2020, Theorem 1)

General Results for Generalization Error

Theorem 1.B (**EXACT** gen-error for iterative SSL)

Consider the NLL loss function $l(\theta, Z) = -\log p_\theta(Z)$, where $p_\theta(Z)$ is the likelihood of Z under parameter θ . For any $t \in [0 : \tau]$,

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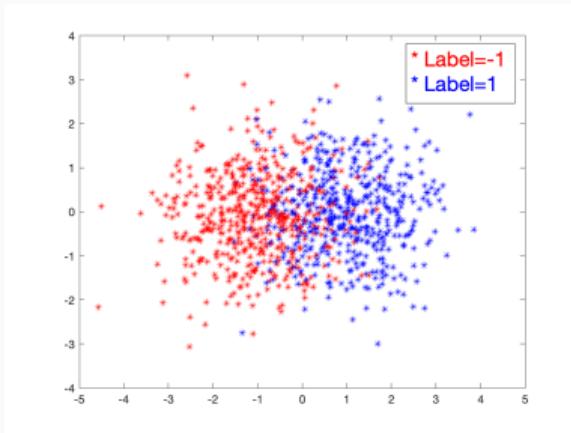
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- The term depends on the labelled training data.
- The term depends on the pseudo-labelled training data. The divergence is caused by pseudo-labelling.

Main Results on Binary Gaussian Mixture Model

♠ **Iterative SSL under bGMM:** Under the bGMM with mean μ and standard deviation σ ($b\text{GMM}(\mu, \sigma)$), assume $\mathcal{Y} = \{-1, +1\}$, $Y \sim P_Y = \text{unif}\{-1, +1\}$, and $X|Y \sim \mathcal{N}(Y\mu, \sigma^2 \mathbf{I}_d)$

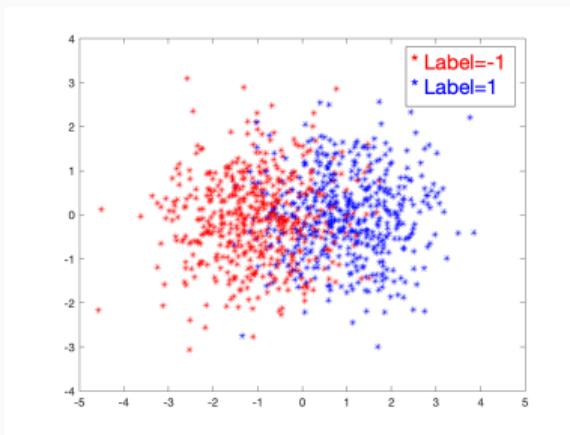


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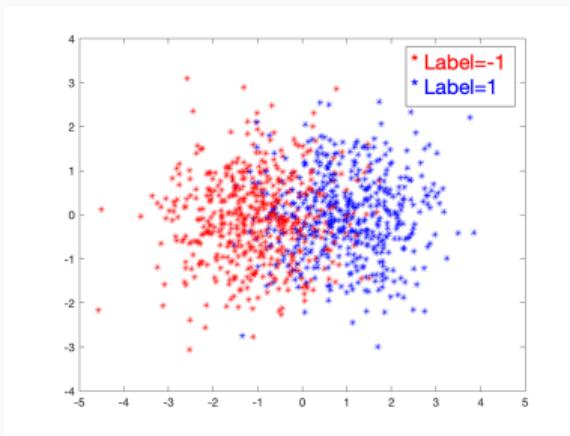
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Pseudo-labelling function: for any $t \in [0 : \tau]$,

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- **Step 1: Initial round $t = 0$ with S_l :** Estimate θ using labelled dataset $S_l = \{(X_i, Y_i)\}_{i=1}^n$, i.e.,

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- **Step 3: Refine the model:** Estimate new parameter using augmented dataset $S_l \cup \{(X'_i, \hat{Y}'_i)\}_{i \in \mathcal{I}_t}$, i.e.,

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If $t < \tau$, go back to Step 2.

Theorem 2 (Exact gen-error for iterative SSL under bGMM)

We derived exact characterization of gen-error gen_t for iterative SSL under bGMM **as a function of standard deviation σ** when the number of unlabelled data is large enough.

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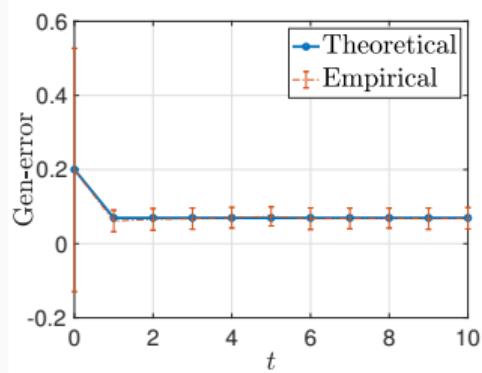


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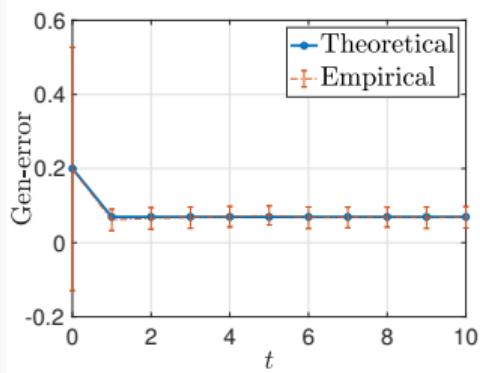


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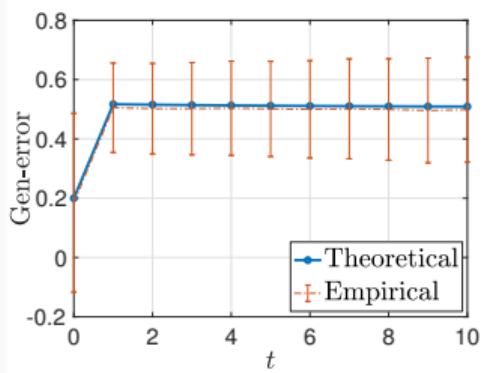


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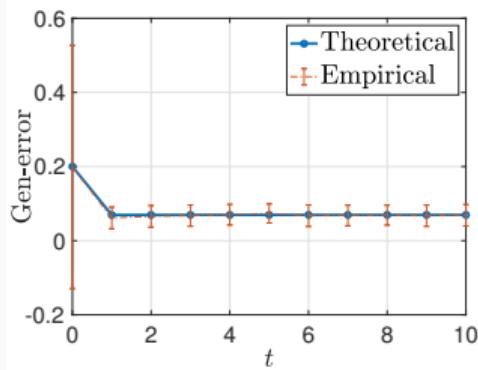


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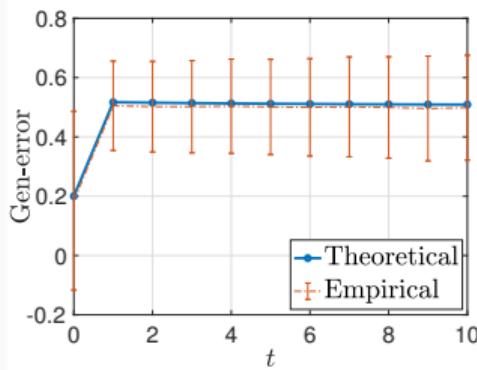


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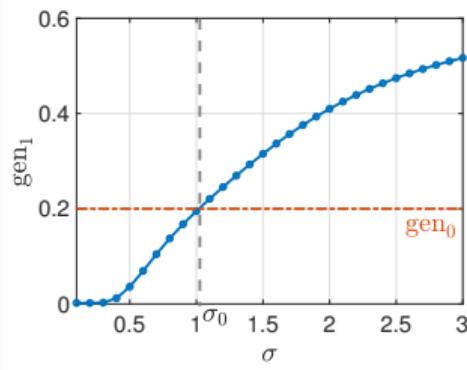


Fig.1.3 Gen-error at $t = 1$ vs σ

Main Results on Binary Gaussian Mixture Model (Continued)

To mitigate the undesirable increase of gen-error across the pseudo-labelling iterations, we prove that adding l_2 -regularization (add $\frac{\lambda}{2} \|\theta\|_2^2$ to loss function) to the loss function can help.

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Fix any $d \in \mathbb{N}$, and $\sigma, \lambda \in \mathbb{R}_+$. The gen-error at any $t \in [1 : \tau]$ is

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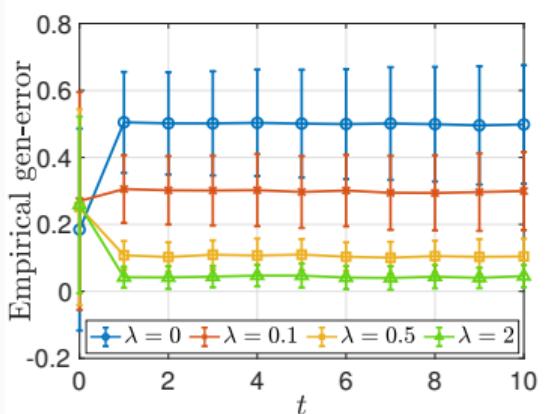


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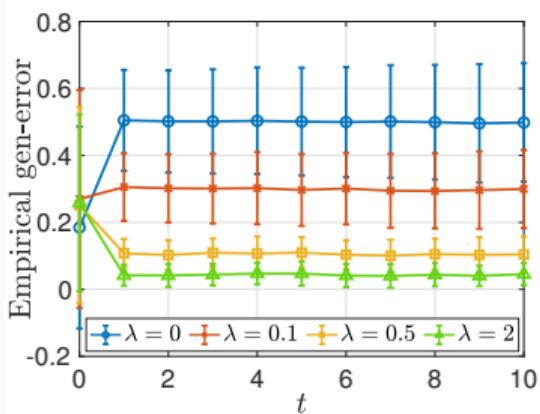


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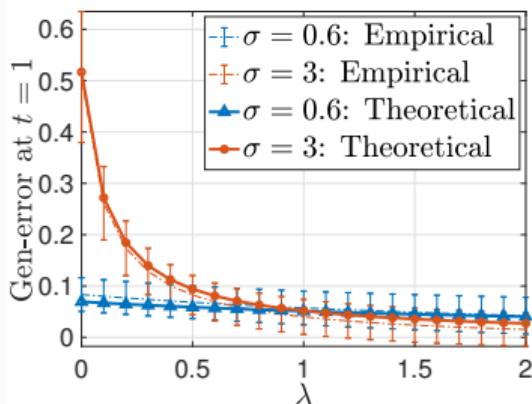


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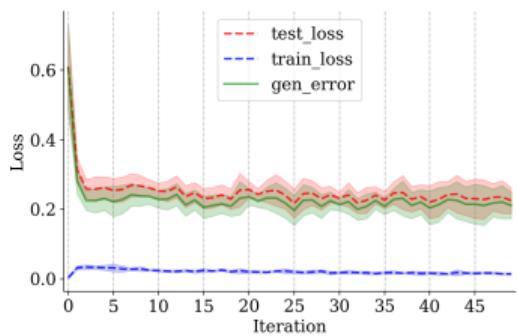


Fig.3.1 "horse-ship": gen-error

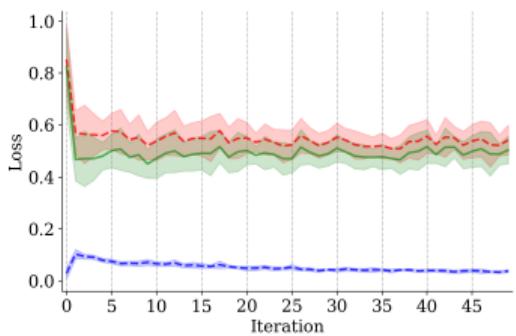


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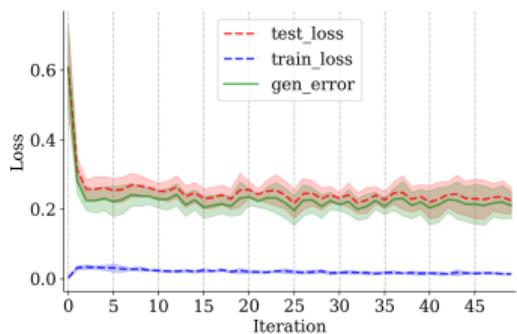


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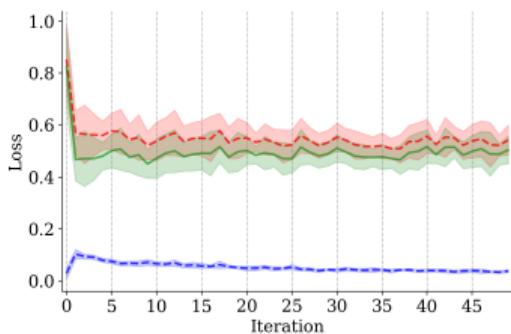


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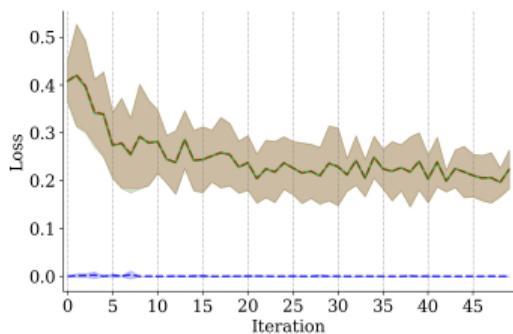


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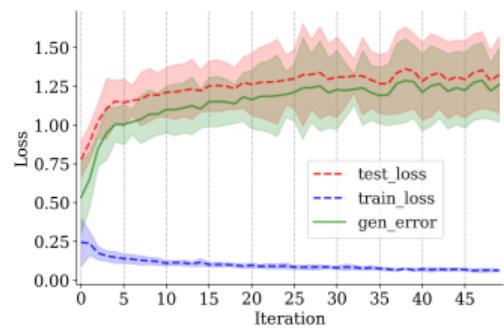


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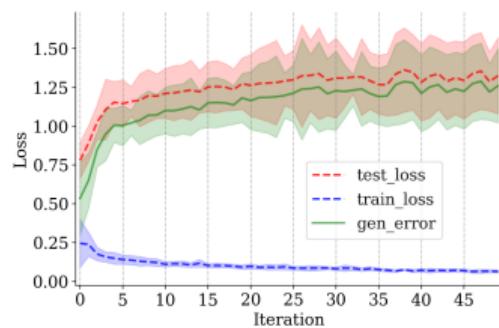


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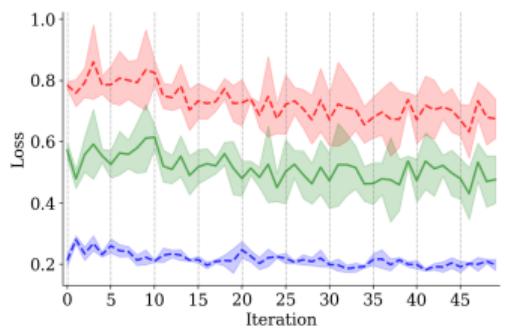


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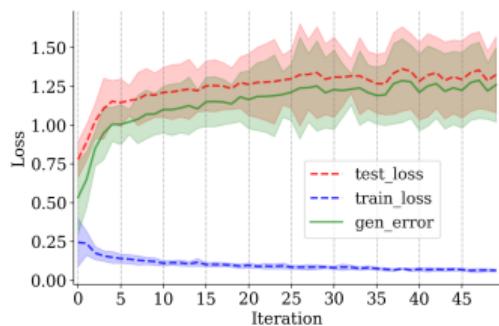


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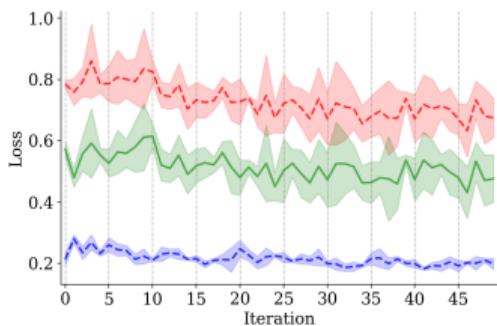


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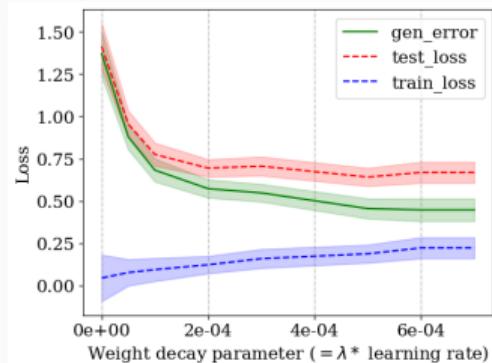


Fig.4.3 “cat-dog”: gen-error after convergence versus weight decay

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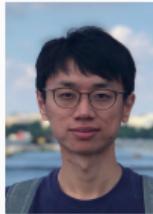
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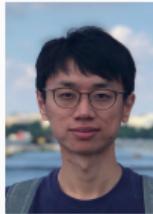
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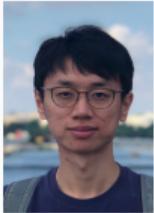
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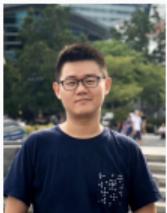
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Thesis advisor:
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SG BUAA Alumni Assoc.

Labmates in E4-06-12

Dear friends
&
My parents



People@
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THANKS FOR LISTENING

