Detecting dependence

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Model selection tools for u-data

Now, that the u-scaled data are produced. we are looking for methods to find classes of suitable models, which are

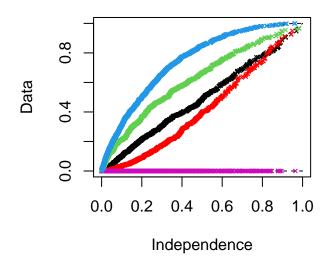
- Dependence structure
 - Kendall plot
 - ► Dependence measures
- ► Existence of tail dependence
- Exchangeability of model
- State of radial symmetry

Dependence structure via Kendall plot

In order to see dependence structure of a suitable model for the data, we use the well-known kendall plot:

```
#library(lcopula)
#m1 <- matria(runif(2000), ncol=2)#Independence
#K.plot(m1, main="Kendall plot")
#m2 <- rCopula(n=1000, claytonCopula(param=-0.5, dim=2))
#K.plot(m2, add=TRUE, col="red") #Negative dependence
#m3 <- rCopula(n=1000, claytonCopula(param=-1, dim=2))
#K.plot(m3, add=TRUE, col="magenta") #Perfect negative dependence
#m4 <- rCopula(n=1000, claytonCopula(param=iTau(claytonCopula(0.3),0.5), dim=2))
#K.plot(m4, add=TRUE, col="green3") #Positive dependence
#m5 <- rCopula(n=1000, claytonCopula(param=iTau(claytonCopula(0.3),1), dim=2))
#K.plot(m5, add=TRUE, col="blue") #Perfect positive dependence
#K.plot(m5, add=TRUE, col="blue") #Perfect positive dependence
```

Kendall plot



Dependence measures

The dependence structure of the suitable model for the data can be achieved via measures of central dependence.

Measures of centeral dependence

- Pearson correlation coefficient
- Spearman rank correlation coefficient
- Kendall's Tau coefficient

Dependence structure via Pearson correlation coefficient

Theoretical model based

For any random vector (X, Y), we have:

$$\rho_{X,Y} = \frac{Cov(X,Y)}{\sigma_X \sigma_Y}$$

Empirical estimator

For any set of joint observation (x_i, y_i) we have:

$$r_{xy} = \frac{n \sum x_i y_i - \sum x_i \sum y_i}{\sqrt{n \sum x_i^2 - (\sum x_i)^2} \sqrt{n \sum y_i^2 - (\sum y_i)^2}}$$

Dependence structure via Pearson correlation coefficient

```
data(danube, package = "lcopula")
cor(danube)
```

```
## donau inn
## donau 1.0000000 0.7374098
## inn 0.7374098 1.0000000
```

Dependence structure via Spearman rank correlation coefficient

The Spearman correlation between two variables is equal to the Pearson correlation between the rank values of those two variables; while Pearson's correlation assesses linear relationships, Spearman's correlation assesses monotonic relationships (whether linear or not).

Theoretical model based

$$\rho_{X,Y} = \frac{Cov(rank(X), rank(Y))}{\sigma_{rank(X)}\sigma_{rank(Y)}}.$$

Based on concordance measures:

$$\rho_{X,Y} = 3Q(F,\Pi) = 3Q(C,\Pi) = 12 \int_0^1 \int_0^1 C(u,v) du dv - 3.$$

Dependence structure via Spearman rank correlation coefficient

The empirical estimator is

$$r_{xy} = \frac{n \sum rank(x_i)rank(y_i) - \sum rank(x_i) \sum rank(y_i)}{\sqrt{n \sum rank(x_i)^2 - (\sum rank(x_i))^2} \sqrt{n \sum rank(y_i)^2 - (\sum rank(y_i))^2}}$$

```
#rho for data
data(danube, package = "lcopula")
cor(danube,method="spearman")[1,2]

## [1] 0.7374098

#rho for copula models in "Copula" package
gumbel.cop <- gumbelCopula(3)
rho(gumbel.cop)</pre>
```

[1] 0.848167

Dependence structure via Kendall's Tau coefficient

Theoretical model based

Based on concordance measures:

$$au_{X,Y} = Q(F,F) = Q(C,C) = 4 \int_0^1 \int_0^1 C(u,v) \ dC(u,v) - 1.$$

Empirical estimator

The Kendall's τ coefficient is defined as:

$$\tau = \frac{\left(\textit{number of concordant pairs}\right) - \left(\textit{number of discordant pairs}\right)}{\binom{n}{2}}$$

Any pair of observations (x_i, y_i) and (x_j, y_j) , where i < j, are said to be concordant if the ranks for both elements (more precisely, the sort order by x and by y) agree.

Dependence structure via Kendall's Tau coefficient

```
#tau for data
data(danube, package = "lcopula")
cor(danube, method="kendall")[1,2]
## [1] 0.5484731
#tau for copula models in "Copula" package
gumbel.cop <- claytonCopula(3)</pre>
tau(gumbel.cop)
## [1] 0.6
#tau for copula models in "VineCopula" package
BiCopPar2Tau(family = 3, par = 3)
## [1] 0.6
#family =3 is the clayton family
```

Existence of tail dependence

The lower tail dependence is defined as

$$\lambda_{\ell} = \lim_{q \to 0} \mathsf{P}\left(X_2 \le F_2^{\leftarrow}(q) \mid X_1 \le F_1^{\leftarrow}(q)\right),$$

where $F^{\leftarrow}(q) = \inf\{x \in \mathbb{R} : F(x) \geq q\}$, that is, the inverse of the cumulative probability distribution function for q.

The upper tail dependence is defined analogously as

$$\lambda_u = \lim_{q \to 1} \mathsf{P}\left(X_2 > F_2^{\leftarrow}(q) \mid X_1 > F_1^{\leftarrow}(q)\right).$$

Existence of tail dependence (example 1)

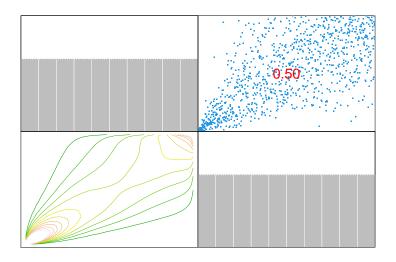
To see the existence of lower/upper tail dependence in the data we have to use the scatter plot & contour plot:

Example: Clayton copula has lower tail dependence.

```
#options(warn=-1) #warnings turned off
#library(copula)
#m1<- rCopula(n=1000, claytonCopula(
# param=iTau(claytonCopula(0.3), 0.5), dim=2))
#library(VineCopula)
#library(kdecopula)
#udata<-pobs(m1)
#pairs.copuladata(udata, method="kendall",
#
                  margins = "unif", col=4, cex=4)
#options(warn=0) #warnings turned on
```

Existence of tail dependence (example 1)

Example: Clayton copula has lower tail dependence.



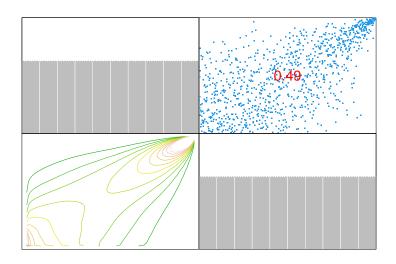
Existence of tail dependence (example 2)

Example: Joe copula has upper tail dependence.

```
#options(warn=-1) #warnings turned off
#library(copula)
\#m2 < - rCopula(n=1000, joeCopula(
# param=iTau(joeCopula(2), 0.5), dim=2))
#library(VineCopula)
#library(kdecopula)
#udata<-pobs(m2)
#pairs.copuladata(udata, method="kendall",
                  margins = "unif", col=4, cex=4)
#
#options(warn=0) #warnings turned on
```

Existence of tail dependence (example 2)

Example: Joe copula has upper tail dependence.



Test for tail independence

This is the tail dependence test described in Reiss and Thomas (2007) section 13.3. It is, unusually, a test whose null hypothesis is that the two random variables, X and Y, are dependent. So, for example, if a significance level alpha = 0.01 test is desired, then the null huypothesis (dependence) is rejected for values of the statistic with p-values less than 0.01.

Reference

Reiss, R.-D. and Thomas, M. (2007) Statistical Analysis of Extreme Values: with applications to insurance, finance, hydrology and other fields. Birkh"auser, 530pp., 3rd edition.

Test for tail independence (example 1)

Example: Joe copula has upper tail dependence.

```
library(copula)
m2<- rCopula(n=1000, joeCopula(
  param=iTau(joeCopula(2),0.5),dim=2))
library(extRemes)
taildep.test(m2)
##
##
   Reiss-Thomas (13.35)
##
## data: m2
## statistic = Inf, threshold = -0.5, m = 222.0, n = 1000.0
## rate (\%) = 22.2, p-value = 1
## alternative hypothesis: less
```

Test for tail independence (example 2)

Example: Normal copula has no dependence in the tails.

```
library(copula)
m2<- rCopula(n=1000,normalCopula(0.5,dim=2))</pre>
library(extRemes)
taildep.test(m2)
##
   Reiss-Thomas (13.35)
##
##
## data: m2
## statistic = -4.275, threshold = -0.5, m = 186.0, n = 100
## rate (%) = 18.6, p-value = 9.557e-06
## alternative hypothesis: less
```

Exchangeability of model

A copula C is exchangeable if

$$C(u, v) = C(u, v)$$
 for all $u, v \in (0, 1)$.

The hypothesis

$$H_0:(X,Y)$$
 Exchangeable vs $H_1:(X,Y)$ Non-exchangeable

Exchangeability of model (example 1)

Example: Joe copula is an exchangable model.

```
options(warn=-1) #warnings turned off
library(copula)
m2<- rCopula(n=1000, joeCopula(
   param=iTau(joeCopula(2),0.5),dim=2))
exchTest(m2)

##
## Test of exchangeability for bivariate copulas with argument 'm' set to
## 0
##
## data: m2
## statistic = 0.024403, p-value = 0.2263
options(warn=0) #warnings turned on</pre>
```

Exchangeability of model (example 2)

Example: khoudraji copula can construct non-exchangable model.

```
options(warn=-1) #warnings turned off
library(copula)
s \leftarrow c(0.6, 0.95)
copula1 <- gumbelCopula
copula2 <- claytonCopula
param <- function(tau) c(iTau(copula1(), tau),iTau(copula2(), tau))
kho <- khoudrajiCopula(copula1 = copula1(param(0.65)[1])
                       .copula2 = copula2(param(0.65)[2]).shapes=s)
U <- rCopula(1000,copula=kho)
exchTest(U)
##
   Test of exchangeability for bivariate copulas with argument 'm' set to
## 0
##
## data: U
## statistic = 0.14356, p-value = 0.0004995
options(warn=0) #warnings turned on
```

State of radial symmetry

A copula C is radially symmetric if and only if $C(u,v) = \bar{C}(u,v)$ for all $u,v \in (0,1)$, where $\bar{C}(u,v)$ is the so-called survival copula defined as

$$\bar{C}(u, v) = u + v - 1 + C(1 - u, 1 - v).$$

$$H_0: C(u,v) = \overline{C}(u,v)$$
 vs $H_1: C(u,v) \neq \overline{C}(u,v)$

There is relation betweendail dependence measures and radial symmetry.

$$\lambda_I = lim_{q o 0} rac{C(q,q)}{q}$$
 and $\lambda_u = lim_{q o 0} rac{ar{C}(q,q)}{q}$.

The equality of C and \hat{C} states that $\lambda_I = \lambda_u$ for the set of observations!

State of radial symmetry (example 1)

options(warn=-1) #warnings turned off

Example 1: Frank copula observations are radially symmetric

```
frank.cop <- rCopula(200, frankCopula(3))
radSymTest(frank.cop)

##

## Test of radial symmetry based on the empirical copula
##

## data: frank.cop
## statistic = 0.033675, p-value = 0.6019

options(warn=0) #warnings turned on</pre>
```

State of radial symmetry (example 2)

Example 2: Gumbel copula observations are not radially symmetric

```
options(warn=-1) #warnings turned off
gumbel.cop <- rCopula(500, gumbelCopula(2, dim=3))
radSymTest(gumbel.cop)

##
## Test of radial symmetry based on the empirical copula
##
## data: gumbel.cop
## statistic = 0.21547, p-value = 0.0004995
options(warn=0) #warnings turned on</pre>
```

See my homepage

To use the copula course materials, go to the web-page https://hamb8066.github.io/homepage and click on teaching section. Choose the "Copula Theory and Applications (Msc)".