Dependence Modelling: Exploratory data analysis

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Load data

```
First, let's make fake data:
```

```
x <- rgamma(350,1,1)
y <- rexp(350,2)
mydata <- cbind(x,y)
head(mydata)</pre>
```

```
## x y
## [1,] 2.2125930 1.9442141
## [2,] 3.2956063 0.1796873
## [3,] 0.5109615 0.2774951
## [4,] 0.2076685 0.1970217
## [5,] 0.6231303 0.3548611
## [6,] 0.4290022 0.1560212
```

A summary

Having a summary of data is useful in understanding the variation of data:

```
options(digits = 3)
summary(x)
##
     Min. 1st Qu. Median Mean 3rd Qu.
                                       Max.
     0.00 0.24 0.63 1.01
                                       7.37
##
                                1.28
summary(y)
##
     Min. 1st Qu. Median
                         Mean 3rd Qu.
                                       Max.
##
    0.001 0.138 0.282
                         0.468
                               0.678
                                      2.806
options(digits = 5)
```

Sklar theorem

Sklar's theorem states that every multivariate cumulative distribution function

$$H(x_1,\ldots,x_d)=\Pr[X_1\leq x_1,\ldots,X_d\leq x_d]$$

of a random vector

$$(X_1,X_2,\ldots,X_d)$$

can be expressed in terms of its marginals $F_i(x_i) = \Pr[X_i \leq x_i]$ and a copula C. Indeed:

$$H(x_1,...,x_d) = C(F_1(x_1),...,F_d(x_d)).$$

Probability integral transform

Suppose that a random variable X has a continuous distribution for which the cumulative distribution function (CDF) is $F_X(x)$. Then

$$Y = F_X(X) \sim U(0,1).$$

Procedure

For example, to find $H(x_1, x_2) = \Pr[X_1 \le x_1, X_2 \le x_2]$, we have to find the

- Step 1. Find the marginal distributions $(F_1(x_1))$ and $F_2(x_2)$ or use a non-parametric method, and if needed transform the data using the probability integral transform $u_{1i} = F_1(x_{1i})$ and $u_{2i} = F_2(x_{2i})$,
- Step 2. Find the copula associated with (u_{1i}, u_{2i}) .

What's the plan ?!!!

Do inference on the joint distibution $H(x_1, x_2)$. For Example, $E[X_1|X_2=x_2]$ and $F_{X_1|X_2=x_2}^{-1}(p)$.

Step 1.

The first step can be performed based on the following methods:

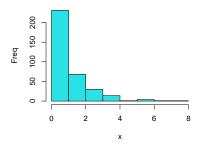
- Parametric method
- Non-parmetric method (Rank estimator)
- Non-parmetric method (Kernel density estimator)

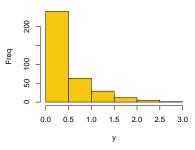
Parametric method

To have a guess of the density function of data, we can use the histogram of data:

```
#par(mfrow=c(1,2))
#hist(x, col = 5, main="", xlab="x", ylab="Freq")
#hist(y, col = 7, main="", xlab="y", ylab="Freq")
#par(mfrow=c(1,1))
```

Histogram





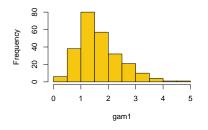
Gamma distribution: A good guess!?

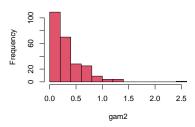
The gamma distribution might be a good candid for both \boldsymbol{x} and \boldsymbol{y} .

```
#gam1 <- rgamma(250, 5 , 3)
#gam2 <- rgamma(250, 1 , 3)
#par(mfrow=c(1,2))
#hist(gam1)
#hist(gam2)
#par(mfrow=c(1,1))</pre>
```

Gamma distribution: A good guess!?

The gamma distribution is a good candid for both x and y.





Model fitting

Now we want to estimate the parameters of gamma distribution (using MLE) and do a goodness of fit for the data:

```
library (MASS)
options(warn=-1) #warnings turned off
par.x <- fitdistr(x, "gamma") $estimate</pre>
par.x
##
     shape rate
## 0.92969 0.92503
par.y <- fitdistr(y, "gamma") $estimate</pre>
par.y
##
     shape rate
## 0.97645 2.08727
options(warn=0) #warnings turned on
```

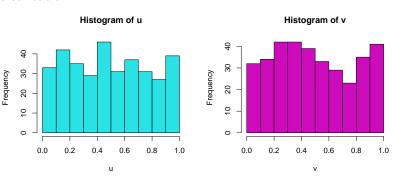
Transform data

Now, we are going to make the u-scaled data using the fitted distribution:

```
#par(mfrow=c(1,2))
#u <- pgamma(x, shape=par.x[1], rate=par.x[2])
#hist(u)
#v<- pgamma(y, shape=par.y[1], rate=par.y[2])
#hist(v)
#par(mfrow=c(1,1))</pre>
```

Transform data

Now, we are going to make the u-scaled data using the fitted distribution:

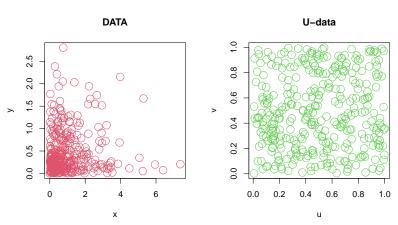


The U-data

In the following you can see the raw data vs the u-scaled data:

The U-data

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Non-parmetric methods

Why Non-parametric method?!!

Noh et al. (2013) point out that modeling the marginals as well as the copula parametrically might cause the resulting fully parametric estimator to be biased and inconsistent if one of the parametric models is misspecified.

Noh, H., Ghouch, A. E., and Bouezmarni, T. (2013), "Copula-based regression estimation and inference," Journal of the American Statistical Association, 108, 676–688.

Non-parmetric method (The Rank estimator)

Definition (Empirical distribution function)

Let x_1, \ldots, x_n be an i.i.d. sample from a distribution function F, then the empirical distribution function is defined as

$$\hat{F}(x) := \frac{1}{n+1} \sum_{i=1}^{n} I_{[x_i \le x]}, \quad \text{for all } x.$$

Division by n+1 instead of n is used to avoid boundary problems of the estimator $\hat{F}(x)$.

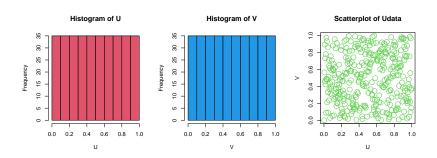
Ranks and empirical distributions:

Let R_i be the rank of observation x_i , i.e., $R_i = k$ if the observation x_i is the kth largest observation among the observations $x_1, ..., x_n$. In this case, it follows that $\hat{F}(x_i) = \frac{R_i}{n+1}$ for i = 1, ..., n.

Non-parmetric method (The Rank estimator)

```
#par(mfrow=c(1,3))
#library(copula)
#udata <- pobs(mydata)
#hist(udata[,1],col=2, main="Histogram of U", xlab="U")
#hist(udata[,2],col=4, main="Histogram of V", xlab="V")
#plot(udata,main="Scatterplot of Udata", xlab="U"
#, ylab="V",col=3,cex=2)
#par(mfrow=c(1,1))</pre>
```

Non-parmetric method (The Rank estimator)



Non-parmetric method (The Kernel estimator)

We want to use the continuous kernel smoothing estimator, which is, given a sample $(x^{(i)})_{i=1,\dots,n}$, defined as

$$\hat{F}(x) = \frac{1}{n} \sum_{i=1}^{n} K(\frac{x - x^{(i)}}{h}), \quad x \in R,$$

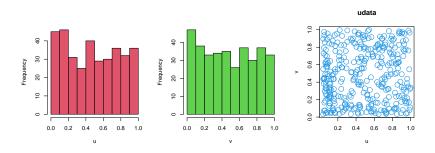
where $K(x) := \int_{-\infty}^{x} k(t)dt$ with k(.) being a symmetric probability density function and h>0 a bandwidth parameter. Usually, we choose $k=\phi$, i.e. a Gaussian kernel, and the plugin bandwidth developed in Duong (2016) which minimizes the asymptotic mean integrated squared error.

Duong (2016), "Non-parametric smoothed estimation of multivariate cumulative distribution and survival functions, and receiver operating characteristic curves," Journal of the Korean Statistical Society, 45, 33–50.

Non-parmetric method (The Kernel estimator)

```
\#par(mfrow=c(1,3))
#library(kde1d)
#fit.x <- kde1d(x)
\#u < -pkde1d(x, fit.x)
\#hist(u,main="",col=2)
#fit.y <- kde1d(y)
\#v < -pkde1d(y, fit.y)
\#hist(v.main="".col=3)
\#plot(cbind(u,v), main="udata", col=4, cex=2)
\#par(mfrow=c(1,1))
```

Non-parmetric method (The Kernel estimator)



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