

Goodness of Fit Tests and Categorical Data Analysis

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Plan

- Goodness of fit tests when all parameters are specified
- Goodness of fit tests when all parameters are unspecified
- Tests of independence in contingency tables

Introduction

- Interest is to determine whether or not a particular probability model is appropriate for a given random variable
- Tests whether a given sample comes from some specified or partially specified probability models are called *goodness of fit* tests

An experiment with rolling a dice

- Outcome of rolling a dice 120 times

3 3 2 6 4 4 1 3 2 6 2 5 3 4 1 2 6 3 3 2

4 1 4 1 4 6 1 5 2 3 4 5 3 5 2 5 6 2 5 6

3 1 1 1 2 4 1 5 6 3 3 6 3 3 2 3 6 3 2 3

4 2 5 2 4 3 4 4 3 2 4 3 2 5 2 2 1 1 1 6

4 2 5 5 6 2 1 3 3 1 5 3 3 4 5 4 4 6 6 1

3 4 6 3 2 2 5 2 6 4 1 1 6 3 5 6 5 1 6 1

- Is the dice biased?

- Frequency distribution

j	x
1	19
2	22
3	26
4	19
5	16
6	18

Goodness of fit tests when all parameters are specified

- Let Y_1, \dots, Y_n be a random sample, where

$$Y_j \in \{1, \dots, k\} \text{ and } P(Y_j = i) = p_i$$

- Let X_i be number of Y 's that are equal to i ($i = 1, \dots, k$) and

$$X_i \sim B(n, p_i)$$

Goodness of fit tests when all parameters are specified

- To test $H_0 : P(Y_j = i) = p_i \forall i$, where

$$\sum_{i=1}^k p_i = 1 \text{ and } \sum_{i=1}^k X_i = n$$

- The test statistic

$$T = \sum_{i=1}^k \frac{[X_i - E(X_i)]^2}{E(X_i)} = \sum_{i=1}^k \frac{(X_i - np_i)^2}{np_i} = \sum_{i=1}^k \frac{X_i^2}{np_i} - n \sim \chi_{k-1}^2$$

- At α level of significance, reject H_0 if $T \geq \chi_{k-1,\alpha}^2$ and $p = P(T \geq t)$, where $T \sim \chi_{k-1}^2$
- The test statistic T depends on X_1, \dots, X_k condition on $\sum_{i=1}^k X_i = n$

Test whether a dice is biased or not

- Let a dice is rolled n times and Y_j be the outcome of the j th roll
 - $Y_j \in \{1, \dots, 6\}$ and $p_i = P(Y_j = i), i = 1, \dots, 6$
- Let X_i be the number of face value i in n rolls of a dice and $\sum_{i=1}^6 X_i = n$
 - Assume $X_i \sim B(n, p_i)$, and $E(X_i) = np_i$
 - Under $H_0 : p_i = 1/6 = p_0 \forall i$, the test statistic

$$T = \sum_{i=1}^6 \frac{(X_i - np_0)^2}{np_0} \sim \chi_5^2$$

- We reject H_0 at α level of significance if $T \geq \chi_{5,\alpha}^2$

Chi-square distribution

- To show whether the test statistic T follows a chi-square distribution, we consider $k = 2$ for which $X_1 + X_2 = n$ and $p_1 + p_2 = 1$

$$\begin{aligned} T &= \frac{(X_1 - np_1)^2}{np_1} + \frac{(X_2 - np_2)^2}{np_2} \\ &= \frac{(X_1 - np_1)^2}{np_1} + \frac{[n - X_1 - n(1 - p_1)]^2}{n(1 - p_1)} \\ &= \frac{(X_1 - np_1)^2}{np_1} + \frac{(X_1 - np_1)^2}{n(1 - p_1)} = \frac{(X_1 - np_1)^2}{np_1(1 - p_1)} \sim \chi_1^2 \end{aligned}$$

Hypothesis of interest

- A famous person is less likely to die in the months before his or her birthday and more likely to die in months afterwards

Problem 1

- According to the Mendelian theory of genetics, a certain garden pea plant should produce either white, pink, or red flowers, with respective probabilities $1/4$, $1/2$, $1/4$.
- To test this theory, a sample of 564 peas was studied with the result that 141 produced white, 291 produced pink, and 132 produced red flowers.
- Using the chi-square approximation, what conclusion would be drawn at the 5 percent level of significance?

Goodness of fit tests when all parameters are unspecified

- In some problems, category-specific probabilities may not be specified by the null hypothesis and these probabilities need to be estimated from the data

- **Example 11.3a:** Weekly number of accidents over a 30-week period is available.

- Test the hypothesis that $X \sim Po(\lambda)$

$$T = \sum_{i=1}^k \frac{(X_i - n\hat{p}_i)^2}{n\hat{p}_i} \sim \chi_{k-1-m}^2$$

- $m \rightarrow$ number of parameters needed to estimate

no. of accidents	no. of weeks
0	6
1	5
2	4
3	4
4	4
5	2
7	1
8	2
9	1
12	1

no. of accidents, Y	Category, j	no. of weeks
0	1	6
1	2	5
2	3	4
3	3	4
4	4	4
5	4	2
7	5	1
8	5	2
9	5	1
12	5	1

$$\hat{\lambda} = \frac{95}{30} = 3.17$$

- $P(j = 1) = P(Y = 0)$
- $P(j = 2) = P(Y = 1)$
- $P(j = 3) = P(Y = 2) + P(Y = 3)$
- $P(j = 4) = P(Y = 4) + P(Y = 5)$
- $P(j = 5) = P(Y > 5)$

Y	Cat	X	prob
0	1	6	0.042
1	2	5	0.133
2	3	4	0.211
3	3	4	0.223
4	4	4	0.177
5	4	2	0.112
7	5	1	0.027
8	5	2	0.011
9	5	1	0.004
12	5	1	0.061

Cat	X	p	$e=np$	$(x-e)^2$	$(x-e)^2/e$
1	6	0.042	1.26	22.468	17.832
2	5	0.133	3.99	1.020	0.256
3	8	0.434	13.02	25.200	1.935
4	6	0.289	8.67	7.129	0.822
5	5	0.103	3.09	3.648	1.181

$$T = 22.026 \text{ and } \chi^2_{3,.05} = 7.8$$

Reject the null hypothesis, i.e., data don't follow a Poisson distribution

Tests of independence in contingency tables

- Suppose each member of a population can be classified according to two distinct characteristics:
 - X -characteristic and Y -characteristic
- The probability that a randomly selected member of the population will have X -characteristic i and Y -characteristic j is defined as

$$P_{ij} = P(X = i, Y = j), i = 1, \dots, r, \quad j = 1, \dots, s$$

- We can also define

$$p_i = P(X = i) = \sum_{j=1}^s P_{ij} \quad \text{and} \quad q_j = P(Y = j) = \sum_{i=1}^r P_{ij}$$

Tests of independence in contingency tables

- We are interested in the null hypothesis that X -characteristic and Y -characteristic are independent, i.e.

$$H_0 : P_{ij} = p_i q_j \text{ against } H_1 : P_{ij} \neq p_i q_j$$

- Suppose a sample n observations have N_{ij} subjects corresponds to the X -characteristic i and Y -characteristic j , and

$$N_i = \sum_{j=1}^s N_{ij} \text{ and } M_j = \sum_{i=1}^r N_{ij}$$

- We can estimate the marginal probabilities

$$\hat{p}_i = \frac{N_i}{n} \text{ and } \hat{q}_j = \frac{M_j}{n}$$

Tests of independence in contingency tables

- Expected frequency $e_{ij} = E(N_{ij}) = nP_{ij}$
- Expected frequency under $H_0 : P_{ij} = p_i q_j$

$$e_{ij} = E(N_{ij}) = nP_{ij} = np_i q_j \Rightarrow \hat{e}_{ij} = n\hat{p}_i \hat{q}_j = \frac{N_i M_j}{n}$$

- The test statistic

$$T = \sum_{i=1}^r \sum_{j=1}^s \frac{(N_{ij} - \hat{e}_{ij})^2}{\hat{e}_{ij}} \sim \chi^2_{(r-1)(s-1)}$$

- Reject H_0 if $T > \chi^2_{\alpha, (r-1)(s-1)}$

Problem 19

- An experiment is designed to study the relationship between hypertension and cigarette smoking from the following data.

	Nonsmoker	Moderate smoker	Heavy smoker
Hypertension	20	38	28
No hypertension	50	27	18

- Test the hypothesis that whether or not an individual has hypertension is independent of how much that person smokes

Problems

- 1-23