

The Exponential Distribution:

A continuous random variable X has an exponential distribution, and it is referred to as an exponential random variable, if and only if its p.density is given by

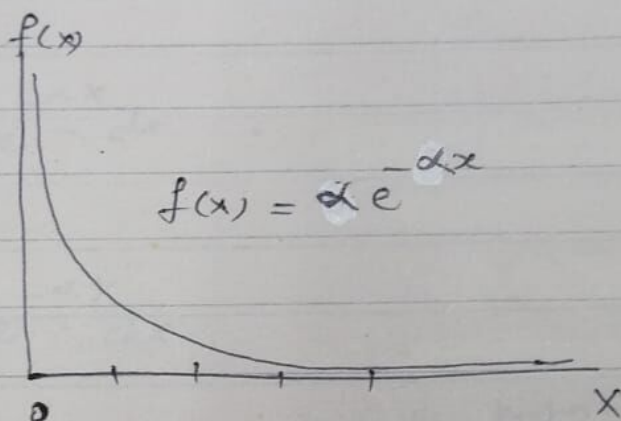
$$f(x) = \begin{cases} \frac{1}{\theta} e^{-\frac{x}{\theta}}, & \text{for } x > 0 \\ 0, & \text{elsewhere} \end{cases}$$

where $\theta > 0$.

✓ $f(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{for } x > 0 \\ 0, & \text{elsewhere} \end{cases}$

where $\lambda > 0$.

The p.d.f is



The Distribution function is

$$F(x) = \int_0^x f(x) dx$$

$$= \lambda \int_0^x e^{-\lambda x} dx$$

$$= \lambda \left[\frac{e^{-\lambda x}}{-\lambda} \right]_0^x$$

$$= \left[-e^{-\lambda x} \right]_0^x$$

$$F(x) = \left[1 - e^{-\lambda x} \right]$$

p.d.f $\int_0^{\infty} f(x) dx = 1$

$$= \lambda \int_0^{\infty} e^{-\lambda x} dx$$

$$= \lambda \left(\frac{e^{-\lambda x}}{-\lambda} \right)_0^{\infty}$$

$$= \left(-e^{-\lambda x} \right)_0^{\infty}$$

$$= 1 - 0$$

$$= 1 - 0$$

$$= 1 //$$

iii) $E(X)$ $= \int_0^{\infty} x \lambda e^{-\lambda x} dx$

$$= \lambda \int_0^{\infty} x^{2-1} e^{-\lambda x} dx$$

$$= \frac{\lambda \sqrt{2}}{\lambda^2} = \frac{1}{\lambda}$$

$$\boxed{E(X) = \frac{1}{\lambda}}$$

$$\underline{E(X^r)} = \int_0^{\infty} x^r \lambda e^{-\lambda x} dx$$

$$= \lambda \int_0^{\infty} x^{(r+1)-1} e^{-\lambda x} dx$$

$$= \frac{\lambda \sqrt{r+1}}{\lambda^{r+1}} = \boxed{\frac{\sqrt{r+1}}{\lambda^r} = \mu_r' = \frac{r!}{\lambda^r}}$$

using the properties of definite integral we have gamma function:

$$\Gamma n = \int_0^{\infty} x^{n-1} e^{-x} dx \quad ; n > 0$$

$$\Gamma \frac{1}{2} = \sqrt{\pi}$$

$$\frac{\lambda \sqrt{n}}{\lambda^n} = \lambda \int_0^{\infty} x^{n-1} e^{-\lambda x} dx ;$$

Put $\lambda = 2$

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$$\mu'_2 = E(x^2) = \frac{\sqrt{2+1}}{\lambda^2} = \frac{\sqrt{3}}{\lambda^2} = \frac{2}{\lambda^2}$$

$$\boxed{\mu'_2 = \frac{2}{\lambda^2}}$$

$$\sigma^2 = \text{Var}(x) = \mu'_2 - \mu_1'^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}$$

$$\boxed{\sigma^2 = \frac{1}{\lambda^2}}$$

Median

$$F(M) = \frac{1}{2} \Rightarrow 1 - e^{-\lambda M} = \frac{1}{2}$$

$$= \int_0^M e^{-\lambda x} dx = \frac{1}{2}$$

$$= \lambda \left[\frac{e^{-\lambda x}}{-\lambda} \right]_0^M = \frac{1}{2}$$

$$= 1 - e^{-\lambda M} = \frac{1}{2}$$

$$= e^{-\lambda M} = \frac{1}{2}$$

$$= e^{-\lambda M} = 2^{-1}$$

$$= -\lambda M = -\log_e 2$$

$$= \lambda M = \log_e 2$$

$$\boxed{M = \frac{1}{\lambda} \log_e 2}$$

$$F(Q_1) = \frac{1}{4}$$

$$1 - e^{-\alpha Q_1} = \frac{1}{2}$$

$$\int_0^{Q_1} \alpha e^{-\alpha x} dx = \frac{1}{4}$$

$$= \alpha \left[\frac{e^{-\alpha x}}{-\alpha} \right]_0^{Q_1} = \frac{1}{4}$$

$$1 - e^{-\alpha Q_1} = \frac{1}{4}$$

$$e^{-\alpha Q_1} = \frac{3}{4}$$

$$-\alpha Q_1 = -\log_e \frac{3}{4}$$

$$\alpha Q_1 = \log_e \frac{3}{4}$$

$$Q_1 = \frac{1}{\alpha} \log_e \frac{3}{4}$$

Similarly .

$$Q_3 = \frac{1}{\alpha} \log_e \left(\frac{1}{4} \right)$$

Mode:

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$$f(x) = \lambda e^{-\lambda x}$$

$$f'(x) = -\lambda^2 e^{-\lambda x} = 0$$

$$= e^{-\lambda x} = 0$$

$$= e^{-\infty} = 0 \Rightarrow x \rightarrow \boxed{\infty = \text{mode}}$$

but on the basis of sketch of pdf, it can be shown that mode exists at '0'

M.G.F

$$M_x(t) = E[e^{tx}] = \int_0^{\infty} e^{tx} f(x; \lambda) dx$$

$$= \lambda \int_0^{\infty} e^{tx} e^{-\lambda x} dx$$

$$= \lambda \int_0^{\infty} e^{-(\lambda-t)x} dx$$

$$= \lambda \left[\frac{e^{-(\lambda-t)x}}{-(\lambda-t)} \right]_0^{\infty}$$

$$\begin{aligned} \because e^{-0} &= 1 \\ e^{-\infty} &= 0 \end{aligned}$$

$$= \frac{\lambda}{(\lambda-t)} \left[-e^{-(\lambda-t)x} \right]_0^{\infty}$$

$$= \frac{\lambda}{(\lambda-t)} |1-0|$$

$$\boxed{M_x(t) = \frac{\lambda}{(\lambda-t)}}$$

for $t < \lambda$

$$\boxed{M_x(t) = \frac{1}{(1-\frac{t}{\lambda})}}; \lambda > 0$$

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$$M_x(t) = \frac{1}{(1-t/\alpha)} ; \alpha > 0.$$

$$\mu_1' = \left[\frac{d}{dt} M_x(t) \right]_{t=0}$$

$$= \left\{ \frac{d}{dt} \left[\left(1 - \frac{t}{\alpha} \right)^{-1} \right] \right\}_{t=0}$$

$$= \left| (-1) \left(1 - \frac{t}{\alpha} \right)^{-2} \cdot \left(-\frac{1}{\alpha} \right) \right|_{t=0}$$

$$= \left| \frac{1}{\alpha \left(1 - \frac{t}{\alpha} \right)^2} \right|_{t=0}$$

$$\boxed{\mu_1' = \frac{1}{\alpha}}$$

$$\mu_2' = \left\{ \frac{d^2}{dt^2} \left[M_x(t) \right] \right\}_{t=0}$$

$$= \left| (-2) \cdot \frac{1}{\alpha \left(1 - \frac{t}{\alpha} \right)^3} \cdot \left(-\frac{1}{\alpha} \right) \right|_{t=0}$$

$$= \left| \frac{2}{\alpha^2} \cdot \frac{1}{\left(1 - \frac{t}{\alpha} \right)^3} \right|_{t=0}$$

$$\boxed{\mu_2' = \frac{2}{\alpha^2}}$$

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Problem: Customers arrive in a certain shop according to an approximate Poisson process at a mean rate of 20 per hour. What is the probability that the shopkeeper will have to wait more than 5 minutes for the arrival of the customers?

Sol: Let X denote the waiting time in minutes until the first customer arrives and note that $\lambda = \frac{20}{60} = \frac{1}{3}$ is the expected number of arrivals per minutes. Thus.

$$\lambda = \frac{1}{3}$$

$$f(x) = \frac{1}{3} e^{-\frac{x}{3}} \quad x \geq 0$$

$$P(X > 5) = \int_5^{\infty} f(x) dx = \frac{1}{3} \int_5^{\infty} e^{-\frac{1}{3}x} dx$$

$$= \frac{1}{3} \left[\frac{e^{-\frac{1}{3}x}}{-\frac{1}{3}} \right]_5^{\infty}$$

$$= [e^{-5/3} - 0]$$

$$= e^{-5/3}$$

$$\boxed{P(X > 5) = 0.189}$$

Problem: Let the p.d.f of X be

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$$f(x) = \frac{1}{2} e^{-\frac{1}{2}x} ; 0 \leq x < \infty$$

- a) What are the mean, variance and m.g.f of X ?
b) Calculate $P(X > 3)$

$$f(x) = \frac{1}{2} e^{-\frac{1}{2}x}$$

$$E(X) = \frac{1}{2} \int_0^{\infty} x e^{-\frac{1}{2}x} dx.$$

$$y = \frac{1}{2}x.$$

$$= \frac{1}{2} \int_0^{\infty} (2y) e^{-y} \cdot 2dy.$$

$$x = 2y.$$

$$dx = 2dy.$$

$$= 2 \int_0^{\infty} y^{2-1} e^{-y} dy.$$

$$= 2 \cdot \Gamma_2$$

$$= 2 \cdot 1 = 2$$

$$\boxed{E(X) = 2}$$

$$E(X^2) = \frac{1}{2} \int_0^{\infty} x^2 e^{-\frac{1}{2}x} dx$$

$$= \frac{1}{2} \int_0^{\infty} (2y)^2 e^{-y} \cdot 2dy$$

$$= 4 \int_0^{\infty} y^{3-1} e^{-y} dy.$$

$$= 4 \Gamma_3$$

$$\boxed{E(X^2) = 8}$$

$$\text{Var}(X) = E[X^2] - \{E[X]\}^2$$

$$= 8 - 4$$

$$\boxed{\sigma^2 = 4}$$

$$\boxed{M_X(t) = \frac{1}{1-2t}}$$

Problem: The life time in years of a television tube of a certain make is a random variable T and its probability density function $f(t)$ is given by

$$f(t) = A e^{-kt} \quad \text{for } 0 \leq t < \infty \quad (k > 0)$$

$$= 0 \quad \text{e.w.}$$

obtain A in terms of k .

- If the manufacturer, after some research, finds that out of 1000 such tubes 371 failed within first two years of use, estimate the value of k .
- Using this value of k correct to 3 significant figures, calculate the mean and variance of T , giving answers correct to 2 significant figures.
- If two such tubes are bought, what is the probability that one fails within its first year and the other lasts longer than six years?

Sol. T is a random variable

$$\int_0^{\infty} f(t) dt = 1$$

$$A \int_0^{\infty} e^{-kt} dt = 1$$

$$A \left[\frac{e^{-kt}}{-k} \right]_0^{\infty} = 1$$

$$\frac{A}{k} [e^{-0} - e^{-\infty}] = 1$$

$$e^{-x} \rightarrow 0 \text{ as } x \rightarrow \infty$$

$$\frac{A}{k} = 1$$

$$\boxed{A = k}$$

$$\begin{aligned}
 P(T < 1) &= \int_0^1 k e^{-kt} dt \\
 &= [-e^{-kt}]_0^1 \\
 &= 1 - e^{-k} \\
 &= 1 - 0.793
 \end{aligned}$$

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and

$$P(T < 1) = 0.207$$

$$P(T > 6) = 1 - P(T \leq 6)$$

$$\begin{aligned}
 &= 1 - k \int_0^6 e^{-kt} dt \\
 &= 1 - [-e^{-kt}]_0^6 \\
 &= 1 + e^{-6k} - 1 \\
 &= e^{-6k}
 \end{aligned}$$

$$P(T > 6) = 0.249$$

\therefore If Two tubes are bought

$$P[(T_1 < 1) \cap (T_2 > 6)] + P[(T_2 < 1) \cap (T_1 > 6)]$$

$$= 2(0.207)(0.249)$$

$$= 0.103 //$$

Therefore the probability that one fails within its first year and the other lasts longer than 6 years is $0.103 //$

$$P(T < 1) = \int_0^1 k e^{-kt} dt$$

$$= [-e^{-kt}]_0^1$$

$$= 1 - e^{-k}$$

$$= 1 - 0.793$$

$$\boxed{P(T < 1) = 0.207}$$

and

$$P(T > 6) = 1 - P(T \leq 6)$$

$$= 1 - k \int_0^6 e^{-kt} dt$$

$$= 1 - [-e^{-kt}]_0^6$$

$$= 1 + e^{-6k} - 1$$

$$= e^{-6k}$$

$$\boxed{P(T > 6) = 0.249}$$

\therefore If Two tubes are bought

$$P[(T_1 < 1) \cap (T_2 > 6)] + P[(T_2 < 1) \cap (T_1 > 6)]$$

$$= 2(0.207)(0.249)$$

$$= 0.103 //$$

Therefore the probability that one fails within its first year and the other lasts longer than 6 years is 0.103.

MGF of Exponential Distribution

Date:

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$$M_X(t) = E[e^{tx}] = \int_0^{\infty} f(x) e^{tx} dx$$

$$= \lambda \int_0^{\infty} e^{-\lambda x} e^{tx} dx = \lambda \int_0^{\infty} e^{-x(\lambda-t)} dx$$

Let $u = x(\lambda-t)$

$du = (\lambda-t) dx$

$\Rightarrow dx = \frac{du}{(\lambda-t)}$

if $x=0, u=0$
 $x=\infty, u=\infty$

$$M_X(t) = \lambda \int_0^{\infty} e^{-u} \frac{du}{\lambda-t} = \frac{\lambda}{\lambda-t} \int_0^{\infty} e^{-u} du = \frac{\lambda}{\lambda-t} \int_0^{\infty} u^{0-1} e^{-u} du$$

$$= \frac{\lambda}{\lambda-t} \Gamma(0!) = \frac{\lambda}{\lambda-t} (1)$$

$$\boxed{M_X(t) = \frac{\lambda}{\lambda-t}} = \left(\frac{\lambda-t}{\lambda} \right)^{-1} = \left(1 - \frac{t}{\lambda} \right)^{-1}$$

Expanding we have

$$M_X(t) = \sum_{r=0}^{\infty} \frac{1}{r!} \left(\frac{t}{\lambda} \right)^r = \sum_{r=0}^{\infty} \left(\frac{r!}{\lambda^r} \right) \cdot \frac{t^r}{r!}$$

Comparing the coefficients of $\frac{t^r}{r!}$ we have

$$\boxed{\mu'_r = \frac{r!}{\lambda^r}}$$

Cumulant Generating function

$$K_X(t) = \log M_X(t) = \log \left(1 - \frac{t}{\lambda} \right)^{-1} = -\log \left(1 - \frac{t}{\lambda} \right)$$

$$= \sum_{r=1}^{\infty} \frac{(t/\lambda)^r}{r} = \sum_{r=1}^{\infty} \frac{(r-1)!}{\lambda^r} \cdot \frac{t^r}{r!}$$

$\ln(1-x) = -x - \frac{x^2}{2} - \dots$

Comparing the coefficients of $\frac{t^r}{r!}$,

$$k_1 = \frac{1}{\lambda}, k_2 = \frac{1}{\lambda^2}, k_3 = \frac{2}{\lambda^3}, k_4 = \frac{6}{\lambda^4}, \dots$$

Comparing with moments

$$\mu'_1 = k_1 = \frac{1}{\lambda}, \mu'_2 = k_2 = \frac{1}{\lambda^2}, \mu'_3 = k_3 = \frac{2}{\lambda^3}$$

$$\mu_4 = k_4 + 3k_2^2 = 9/\lambda^4$$

$$\beta_1 = \frac{\mu_3^2}{\mu_2^3} = 4$$

$$\gamma_1 = \sqrt{\beta_1} = 2$$

$$\beta_2 = \frac{\mu_4}{\mu_2} = 9$$

$$\gamma_2 = \beta_2 - 3 = 9 - 3 = 6$$

Problem:

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Telephone calls enter a college switchboard according to a Poisson process on the average of two every 3 minutes.

Let X denote the waiting time until the first call that arrives after 10 AM.

a) What is the P.d.f of X ?

b) Find $P(X > 2)$

a)

$$f(x) = \frac{1}{\theta} e^{-\frac{x}{\theta}}, \quad \text{for } x > 0$$

$$= 0$$

$$, \quad e^{-\infty}$$

~~ex~~

$$\theta = \frac{3}{2}$$

$$f(x) = \frac{2}{3} e^{-\frac{2}{3}x}, \quad \text{for } x > 0$$

$$b) \quad P(X > 2) = \int_2^{\infty} e^{-\frac{2}{3}x} dx.$$

$$= \frac{2}{3} \left[\frac{e^{-\frac{2}{3}x}}{-\frac{2}{3}} \right]_2^{\infty}$$

$$= \left[e^{-\frac{2}{3}x} - e^{-\infty} \right]$$

$$= e^{-3}$$