

Modelling complex traits with ancestral recombination graphs

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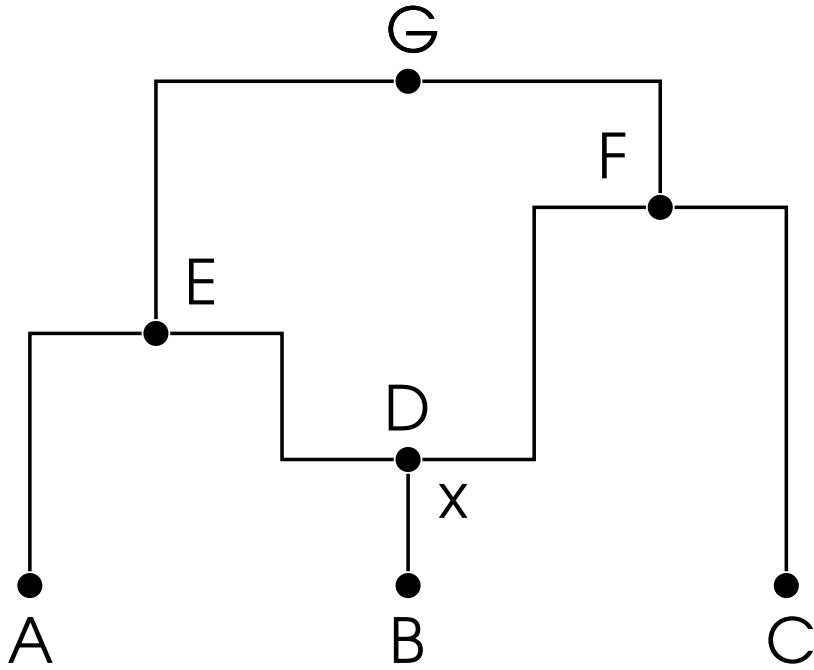
University of Michigan, Ann Arbor

Mar 7, 2025

Overview

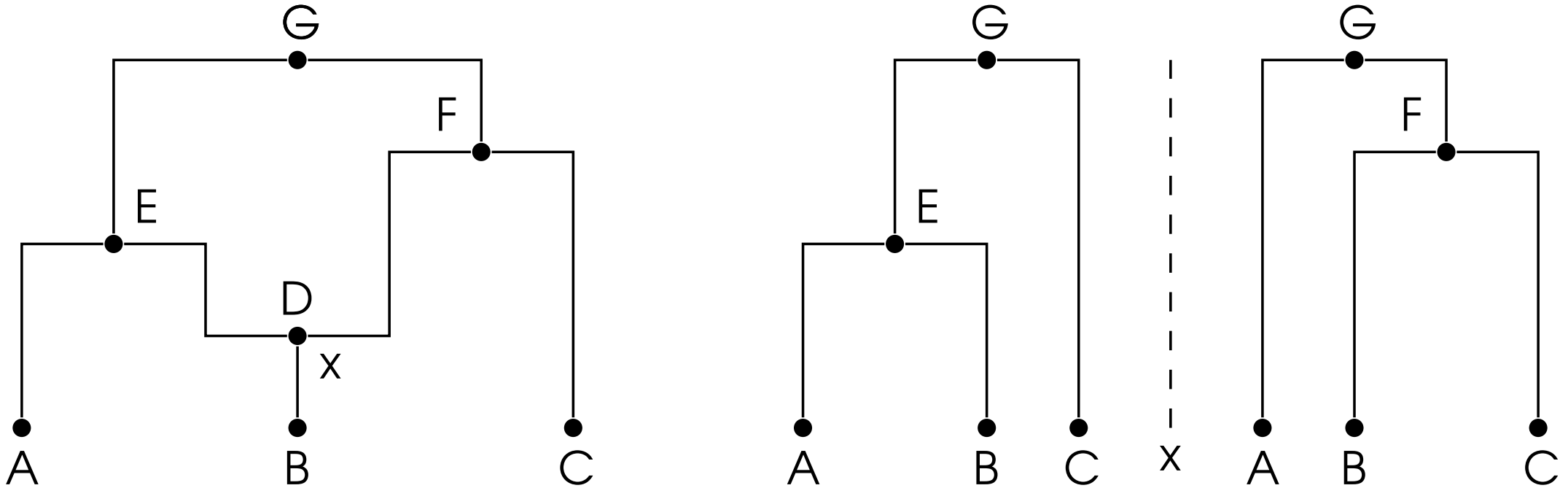
Overview

The ancestral recombination graph (ARG) describes the evolutionary relationship between genetic materials in the presence of recombination and drift



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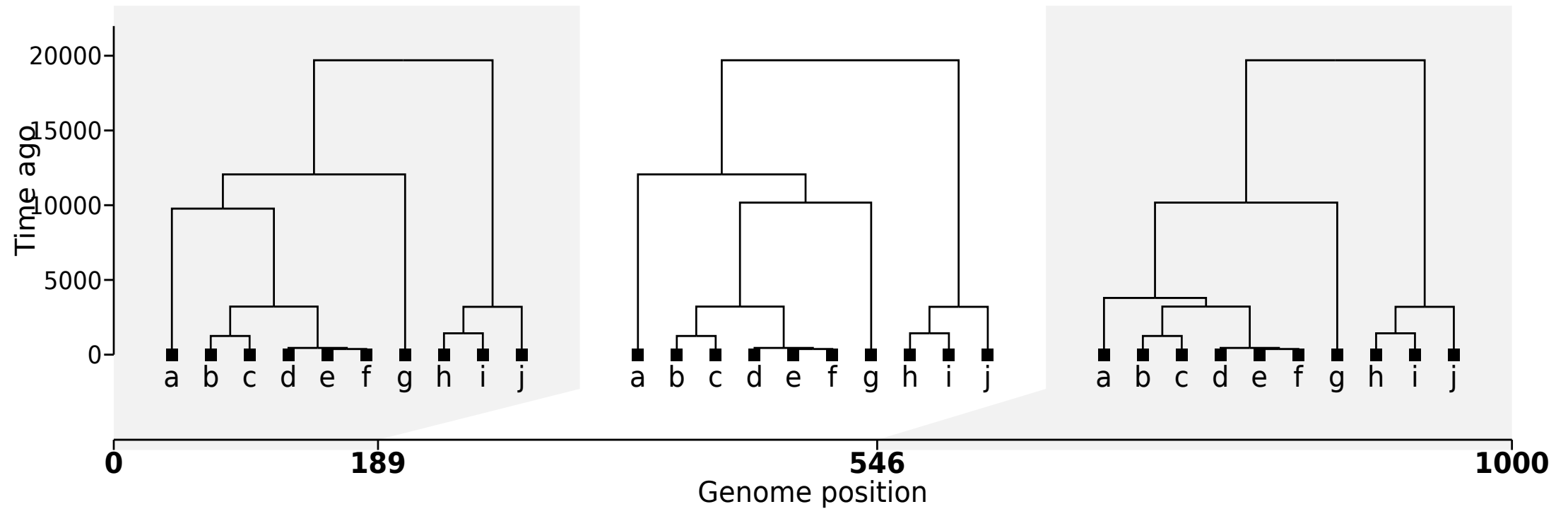
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Overview

The full probabilistic process is complicated

In this work, we condition on the realized ARG, resulting a sequence of local trees



From [tskit docs](#)

Overview

What is the conditional distribution of a trait given the trees?

Since the genealogy is fixed, the only randomness that remains is mutation

$$\text{Trait} \mid \text{Local trees} \sim ?$$

Linear mixed model

Linear mixed model

Linear mixed models are popular in quantitative genetics

$$\mathbf{y} = \underbrace{\mathbf{Z}\mathbf{u}}_{\text{random effects}} + \underbrace{\mathbf{X}\mathbf{b}}_{\text{fixed effects}} + \boldsymbol{\varepsilon}$$

where \mathbf{Z} includes genotyped variants and \mathbf{X} is the covariate matrix

In particular, the SNP effects $\mathbf{u} \sim p(\cdot)$ is *random*

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We answer these questions from a genealogical perspective

Setup and derivation

Setup and derivation

The trait \mathbf{y} is a linear function of the genotype \mathbf{G}

$$\mathbf{y} = \mathbf{G}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

$\mathbf{y} \in \mathbb{R}^N$, $\mathbf{G} \in \mathbb{R}^{N \times P}$, $\boldsymbol{\beta} \in \mathbb{R}^P$, and $\boldsymbol{\epsilon} \in \mathbb{R}^N$

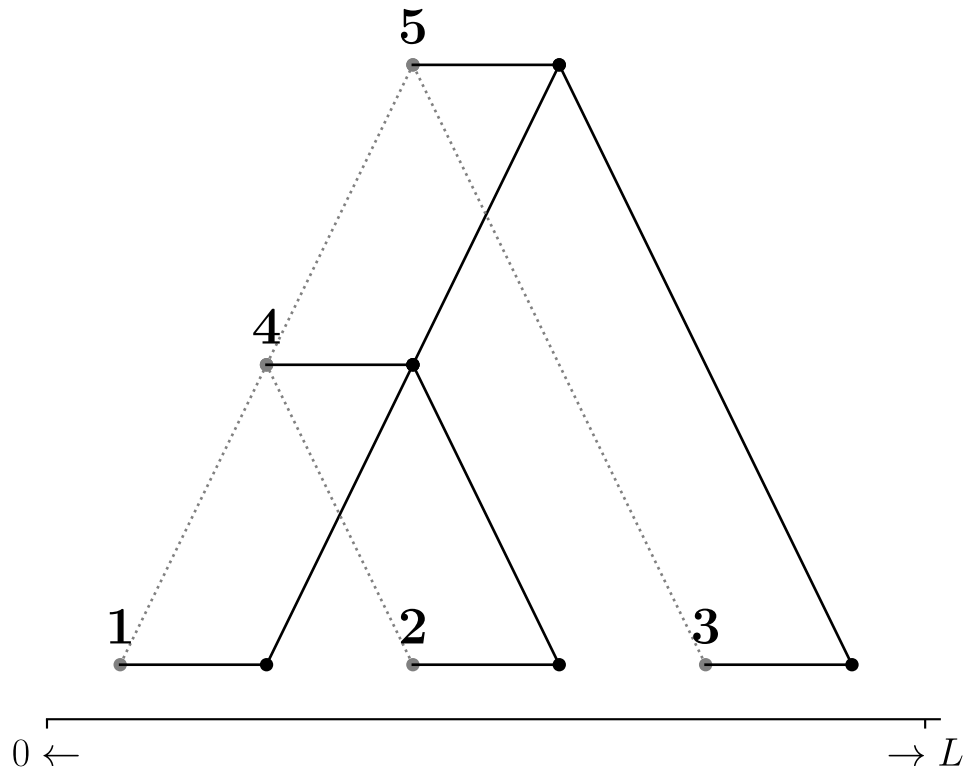
\mathbf{G} contains *all* positions the genome including genotyped ones

N : number of samples, P : length of the genome

How do we get traits?

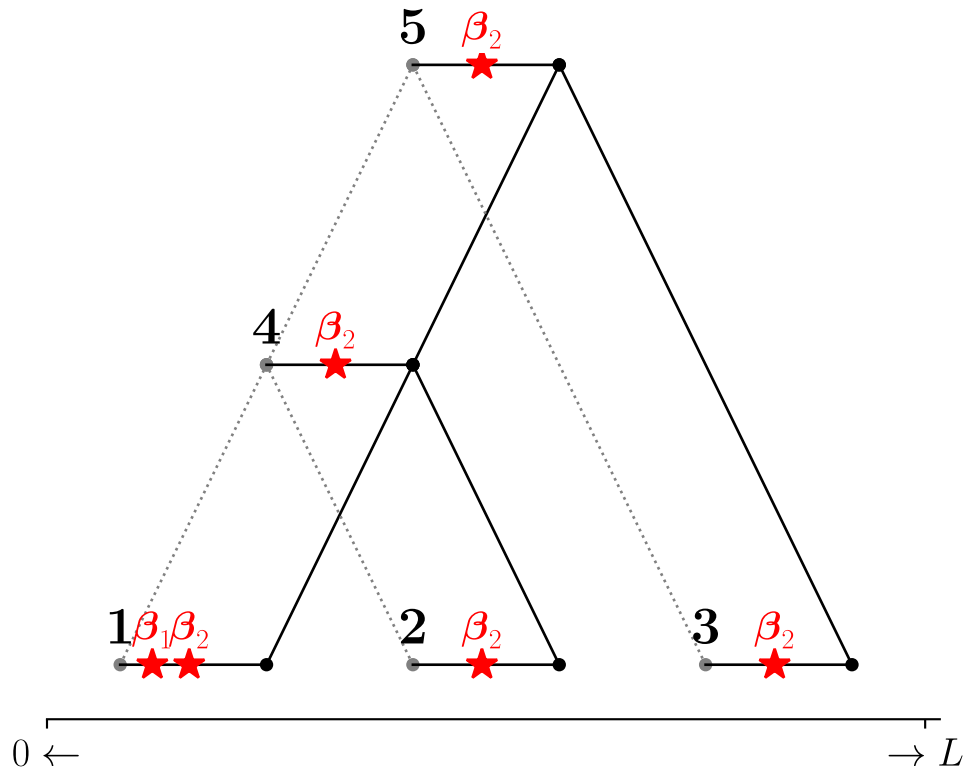
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Consider a local tree that spans over a region
We get trait values by adding up effect sizes (β)



How do we get traits?

$$\mathbf{y}_1 = \beta_1 + \beta_2, \mathbf{y}_2 = \beta_2, \mathbf{y}_3 = \beta_2$$

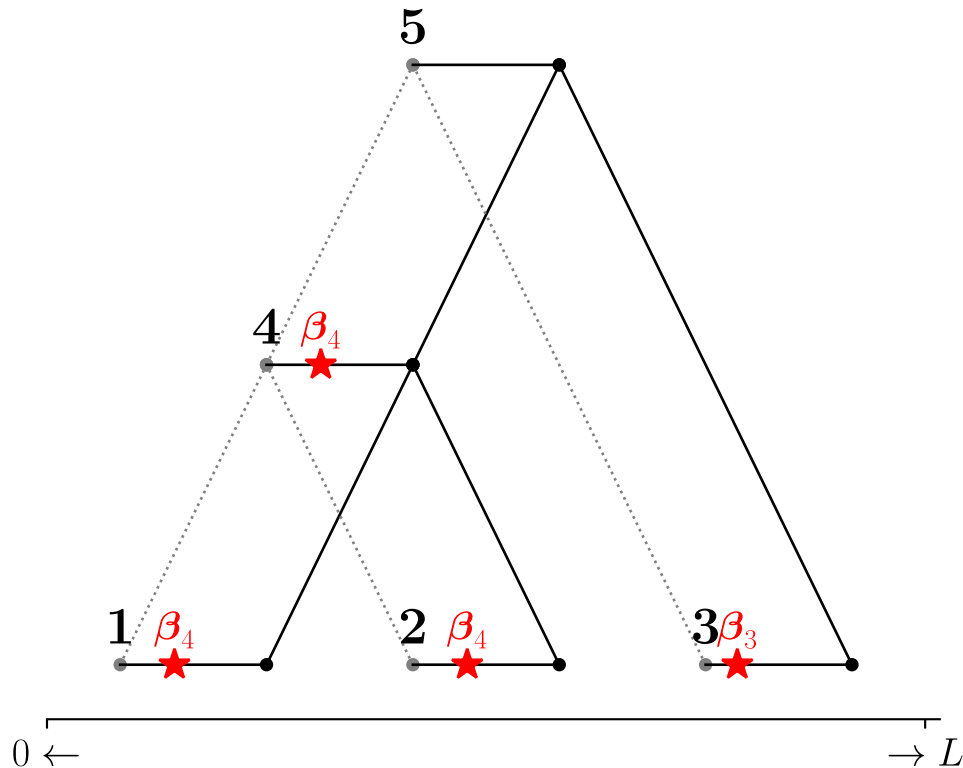


Consider a local tree that spans over a region
We get trait values by adding up effect sizes (β)

- $\mathbf{y}_n = \mathbf{G}_{n1}\beta_1 + \mathbf{G}_{n2}\beta_2$

How do we get traits?

$$y_1 = \beta_4, y_2 = \beta_4, y_3 = \beta_3$$



Consider a local tree that spans over a region
We get trait values by adding up effect sizes (β)

- $y_n = \mathbf{G}_{n1}\beta_1 + \mathbf{G}_{n2}\beta_2$
- $y_n = \mathbf{G}_{n3}\beta_3 + \mathbf{G}_{n4}\beta_4$

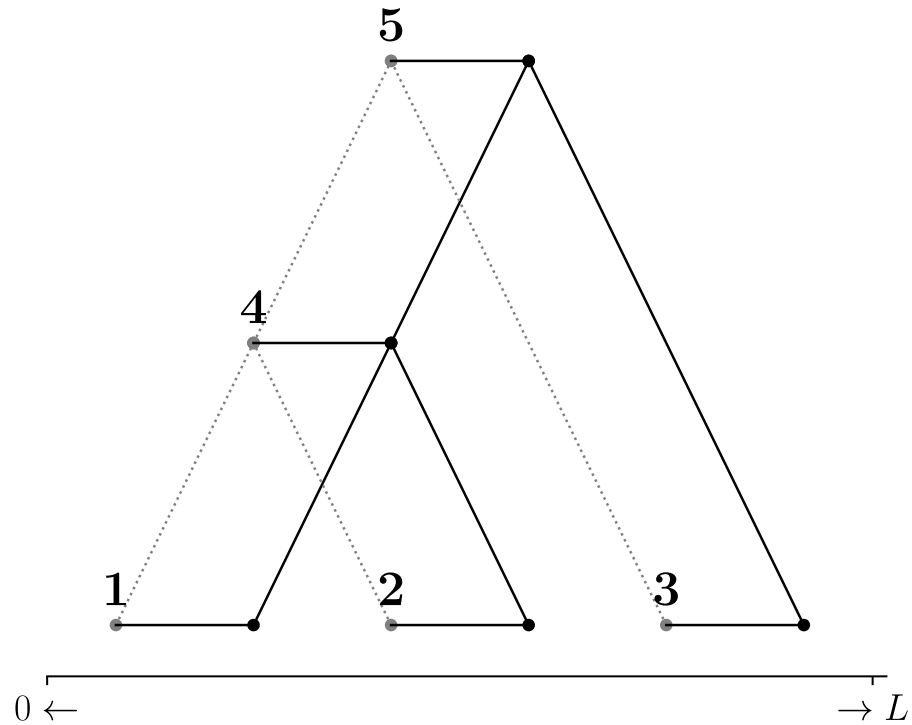
Branch-centric view of trait transmission

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Inherit a branch first, then a mutation

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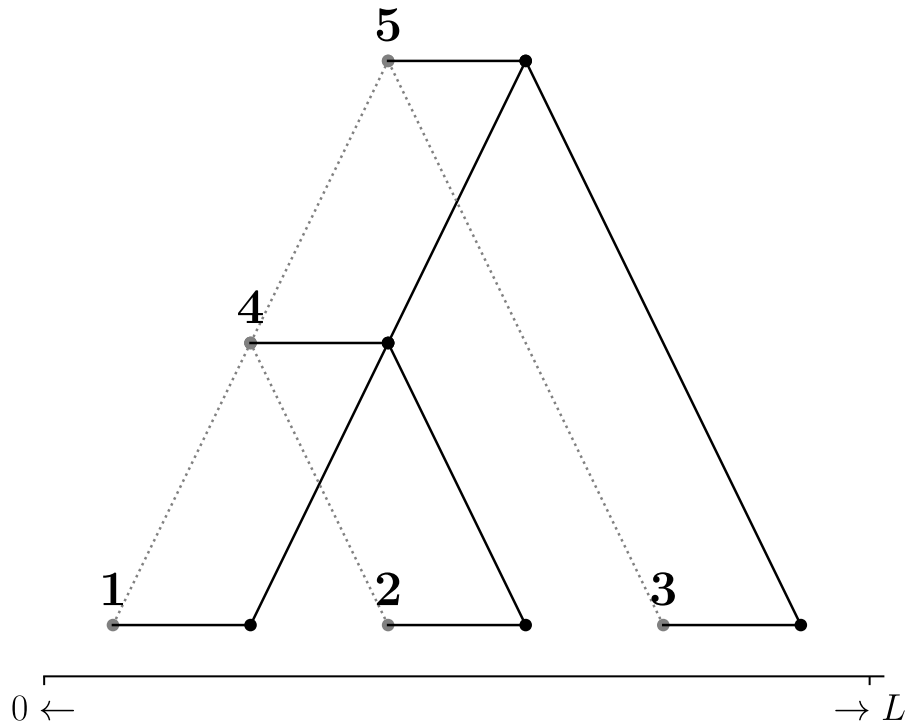
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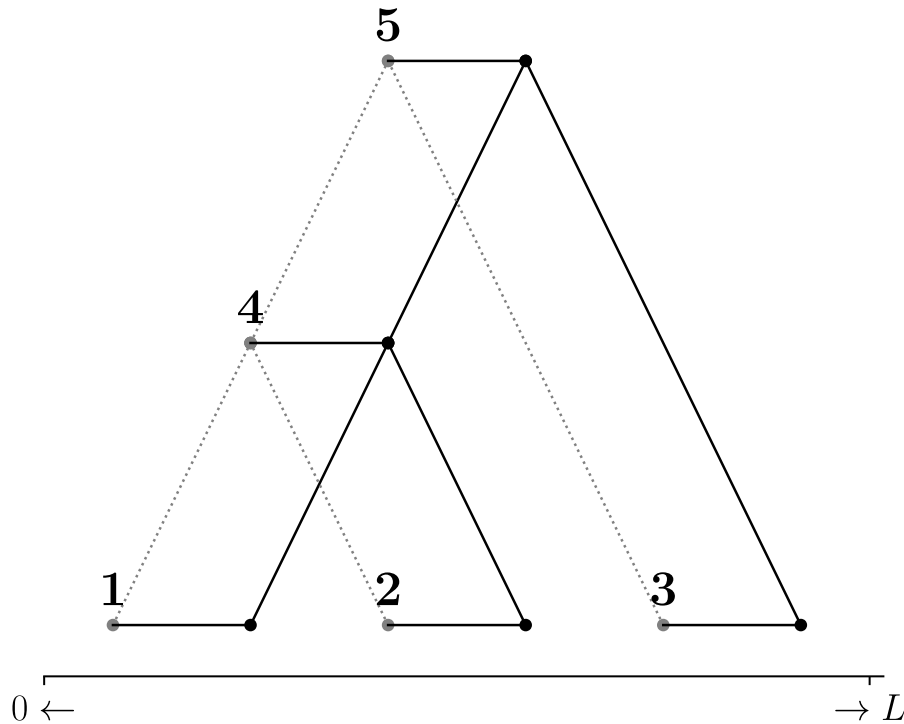
Branch-centric view of trait transmission

Inherit a branch first, then a mutation

- Sample 1 inherits edges $1 - 4$ and $4 - 5$



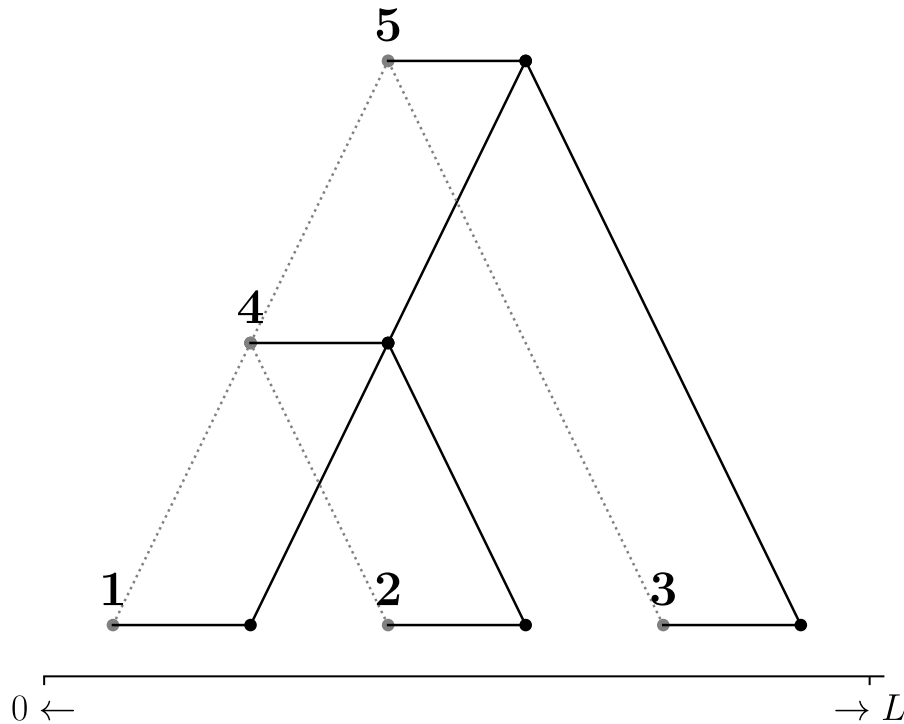
Branch-centric view of trait transmission



Inherit a branch first, then a mutation

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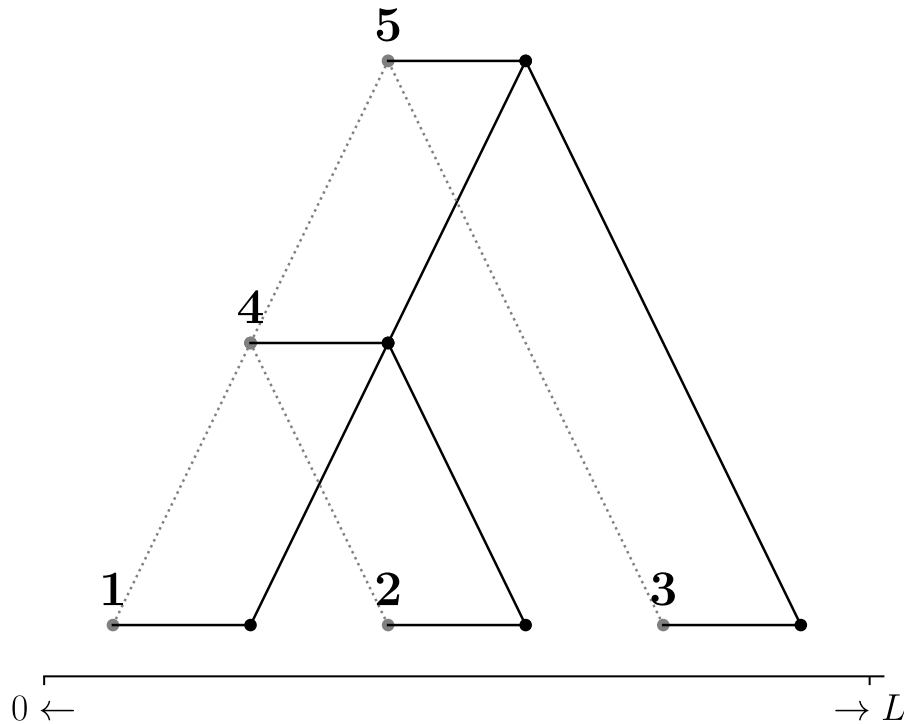
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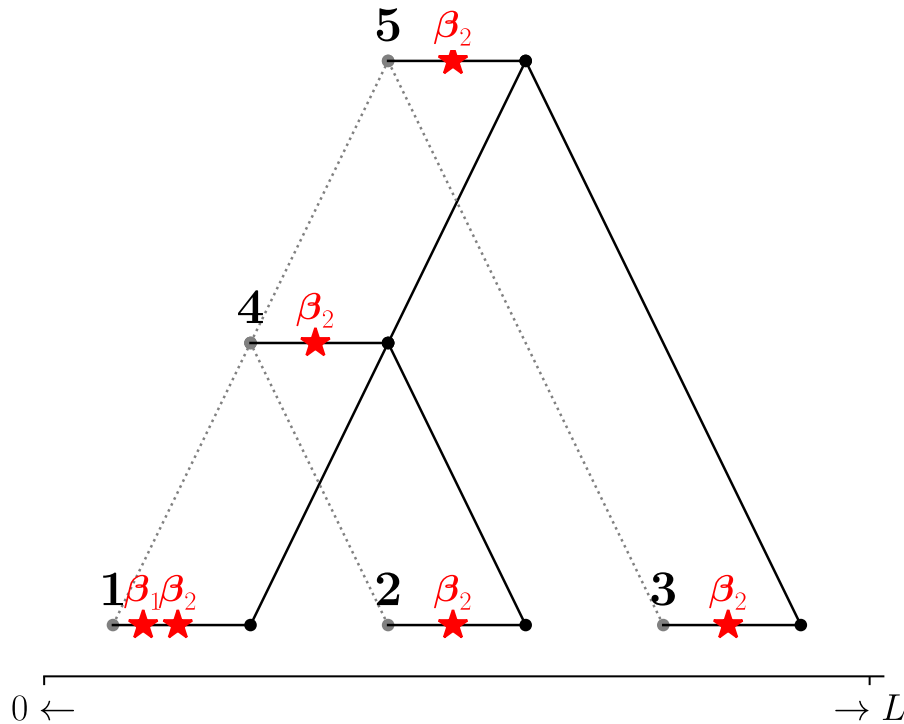


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Branch's effect = Sum of mutations' effect

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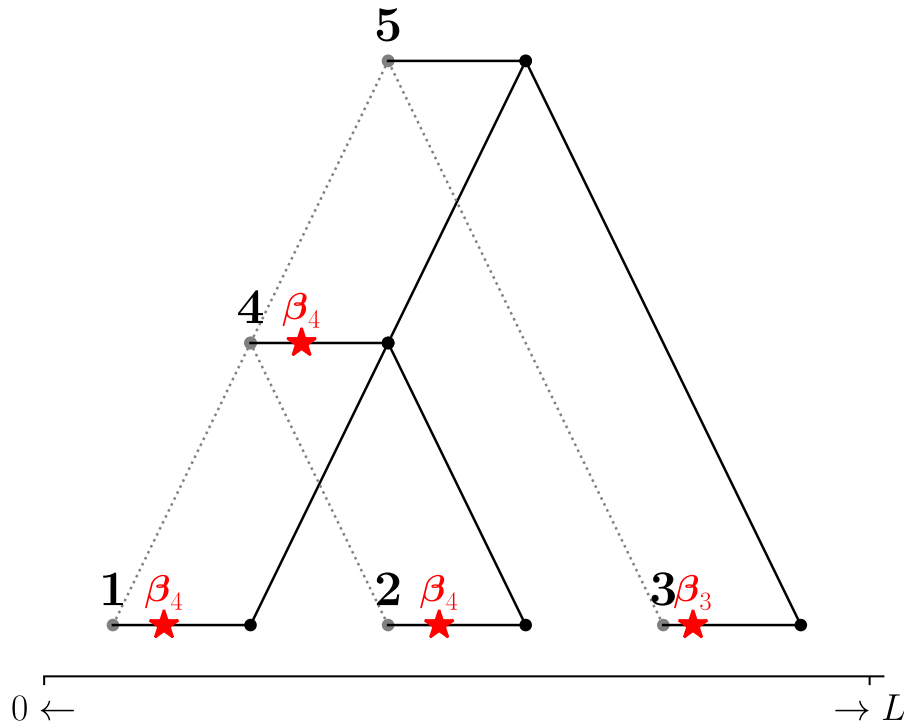
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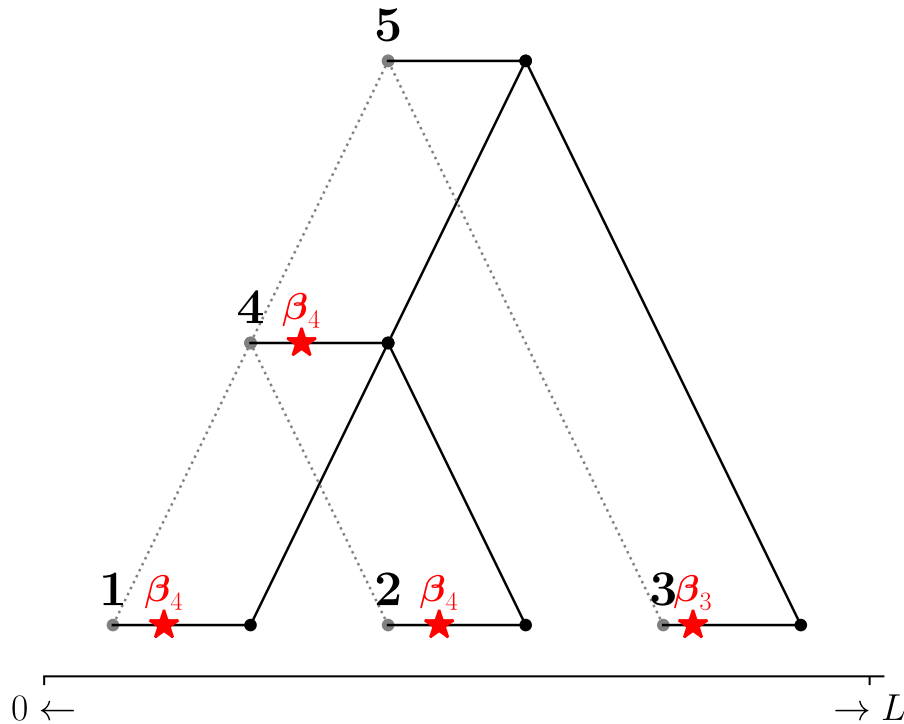
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Branch effect is a random variable!

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$$\mathbf{y} = \sum_p \mathbf{G}_p \boldsymbol{\beta}_p + \boldsymbol{\varepsilon} \quad \Rightarrow \quad \mathbf{y} = \sum_e \mathbf{Z}_e \mathbf{v}_e + \boldsymbol{\varepsilon}$$

where \mathbf{Z}_{ne} = the number of haplotypes of n that inherit e

Ancestral recombination graph linear mixed model (ARG-LMM)

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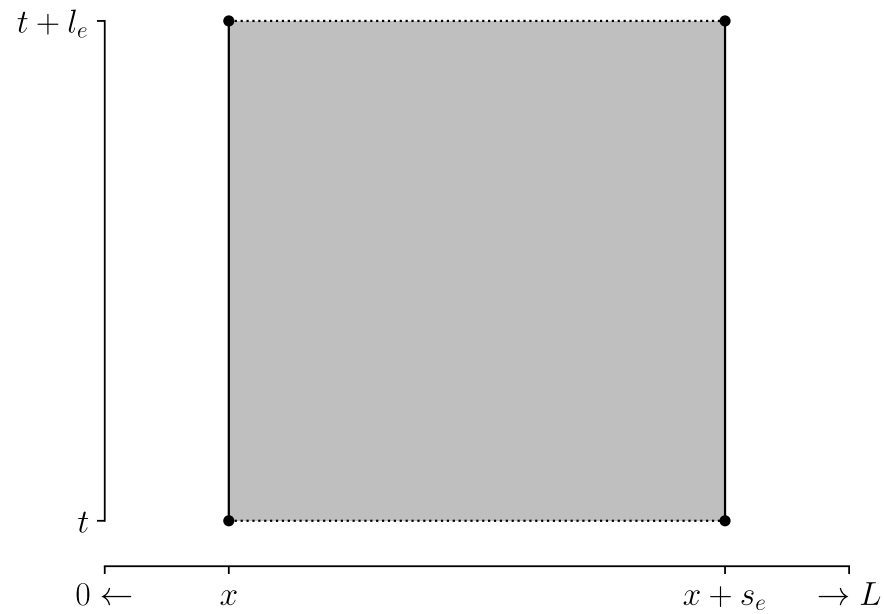
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- We start from more lower-level evolutionary statements to recover mixed model assumptions
- Independent random effects, random effect weights, normality, . . .

How do we weigh branches of the ARG?

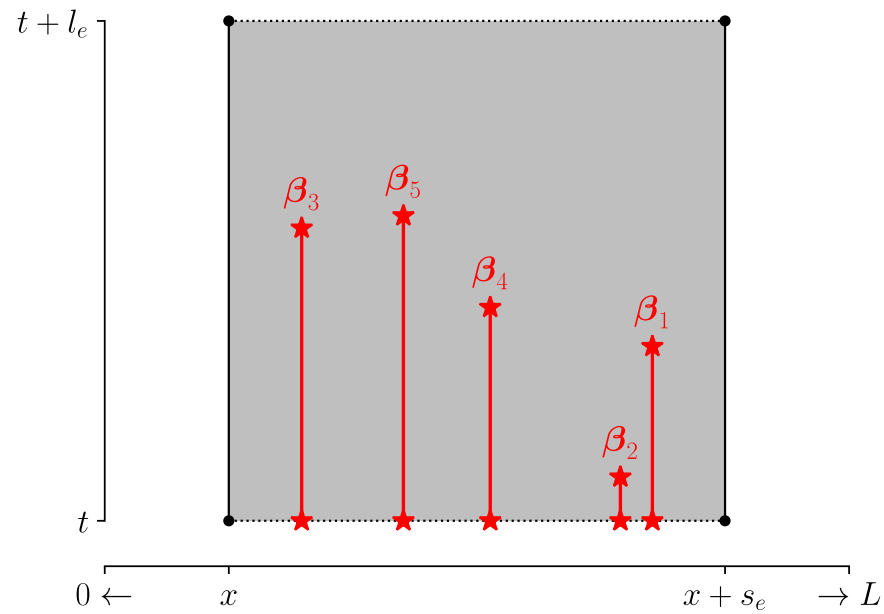
How do we weigh branches of the ARG?

l_e : length in time s_e : span in base pairs



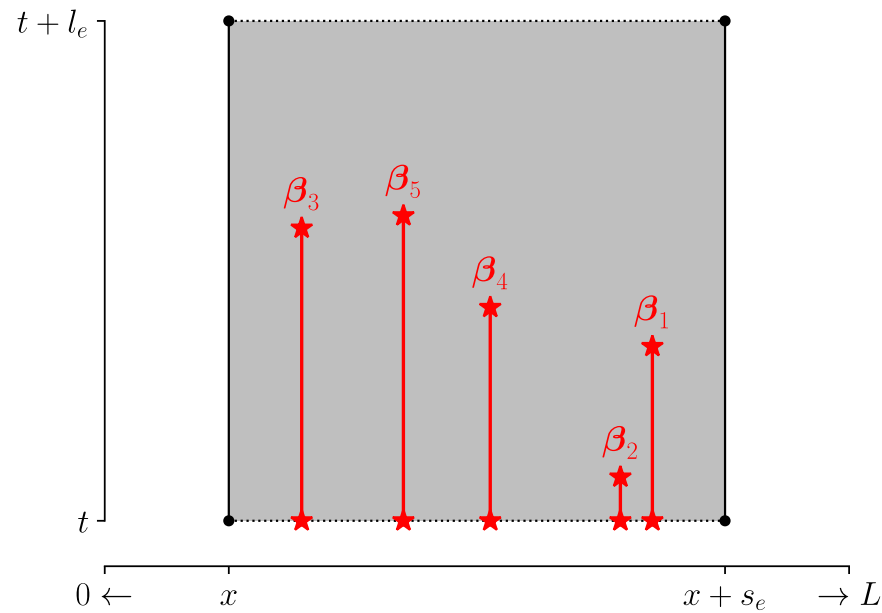
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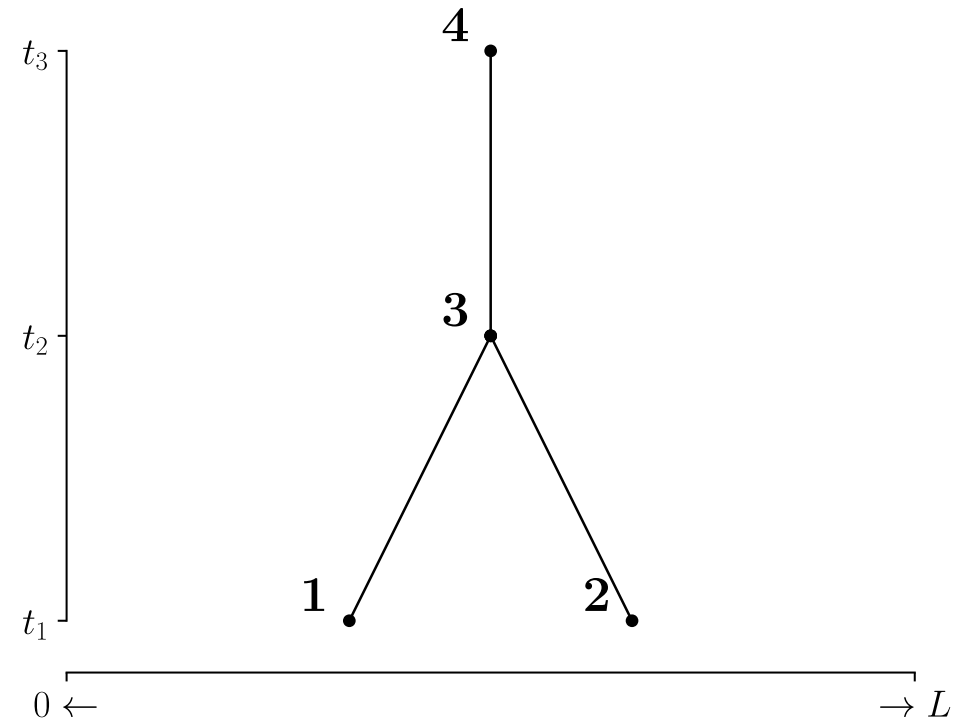
$$\text{Var}(\mathbf{u}_e) \propto \text{Number of mutations} \propto \text{Area} = l_e s_e$$

Complex traits through the lens of ARG-LMM

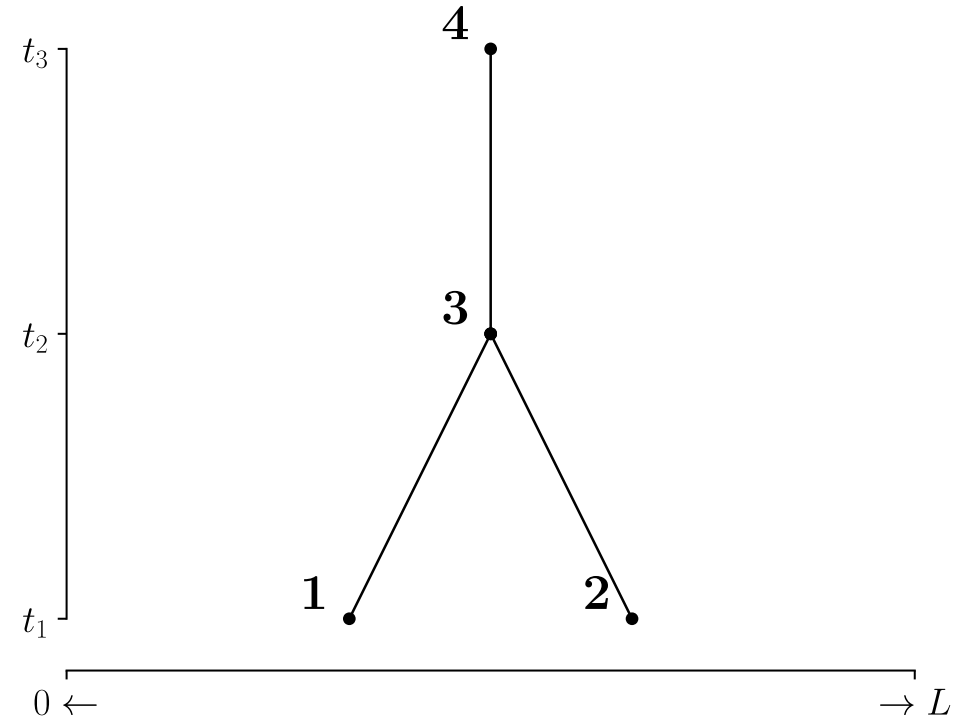
What does ARG-LMM tell us about complex trait analysis?

Genetic value covariance

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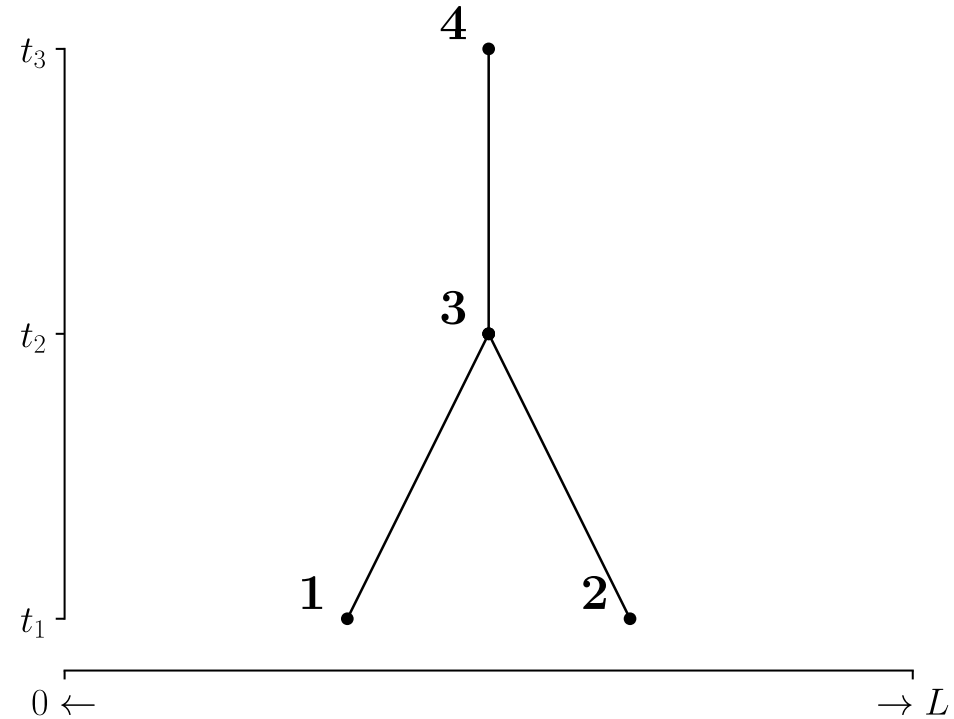


Genetic value covariance



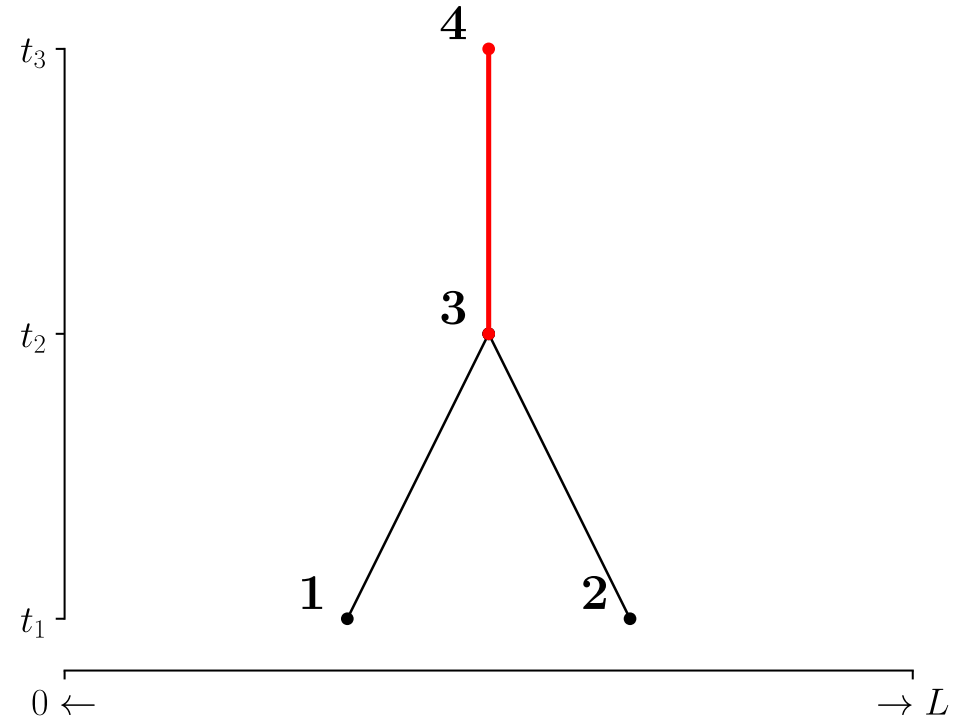
$$\mathbf{y}_1 = \mathbf{u}_{13} + \mathbf{u}_{34} \quad \text{and} \quad \mathbf{y}_2 = \mathbf{u}_{23} + \mathbf{u}_{34}$$

Genetic value covariance



$$\text{Cov}(\mathbf{y}_1, \mathbf{y}_2) = \text{Cov}(\mathbf{u}_{13} + \mathbf{u}_{34}, \mathbf{u}_{23} + \mathbf{u}_{34}) = \text{Cov}(\mathbf{u}_{34}, \mathbf{u}_{34}) = \text{Var}(\mathbf{u}_{34}) \propto t_3 - t_2$$

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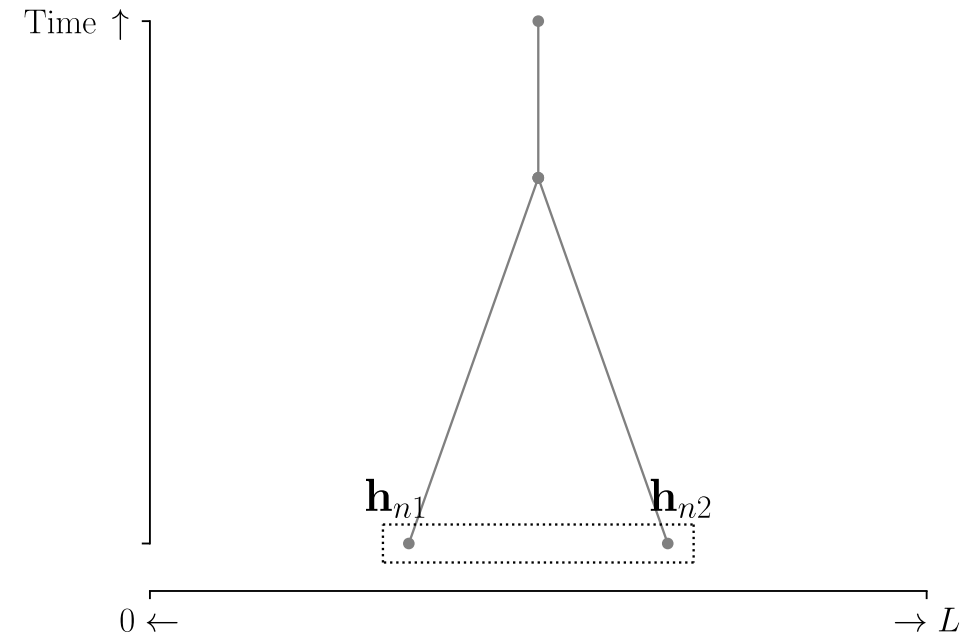
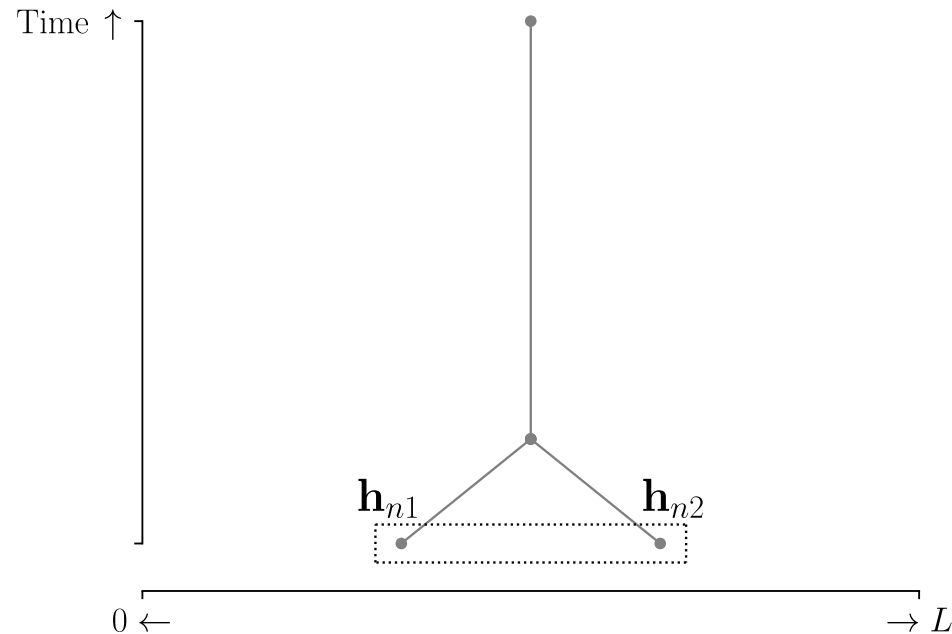
$$\text{Heritability: } h_g^2 = \frac{\text{Var}(\mathbf{g}_n)}{\text{Var}(\mathbf{y}_n)} = \frac{\text{Var}(\mathbf{g}_n)}{\text{Var}(\mathbf{g}_n) + \text{Var}(\boldsymbol{\epsilon}_n)}$$

This applies to all individuals $n \in \{1, \dots, N\}$

Heritability is *ill*-defined in ARG-LMM

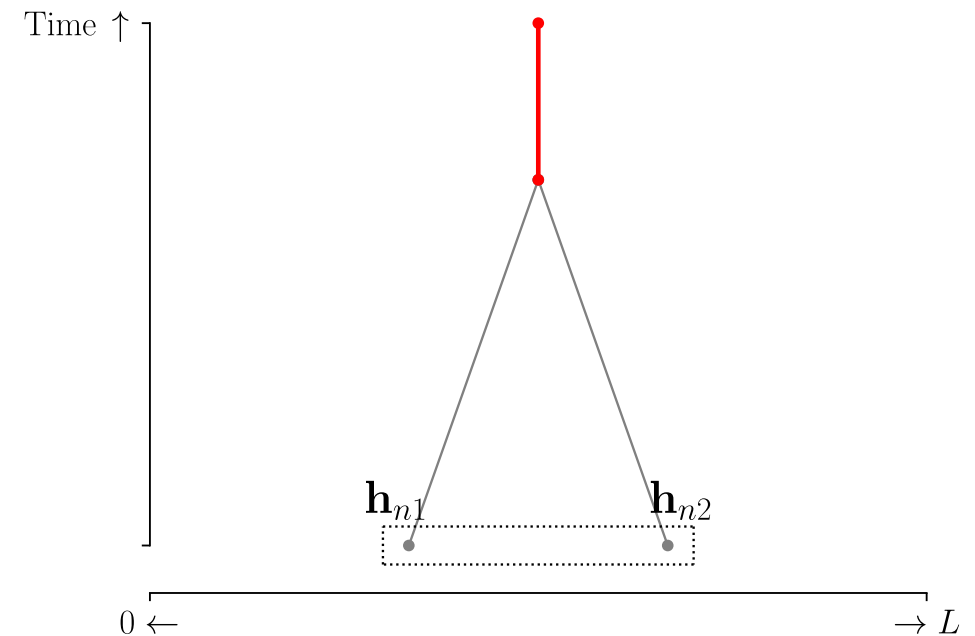
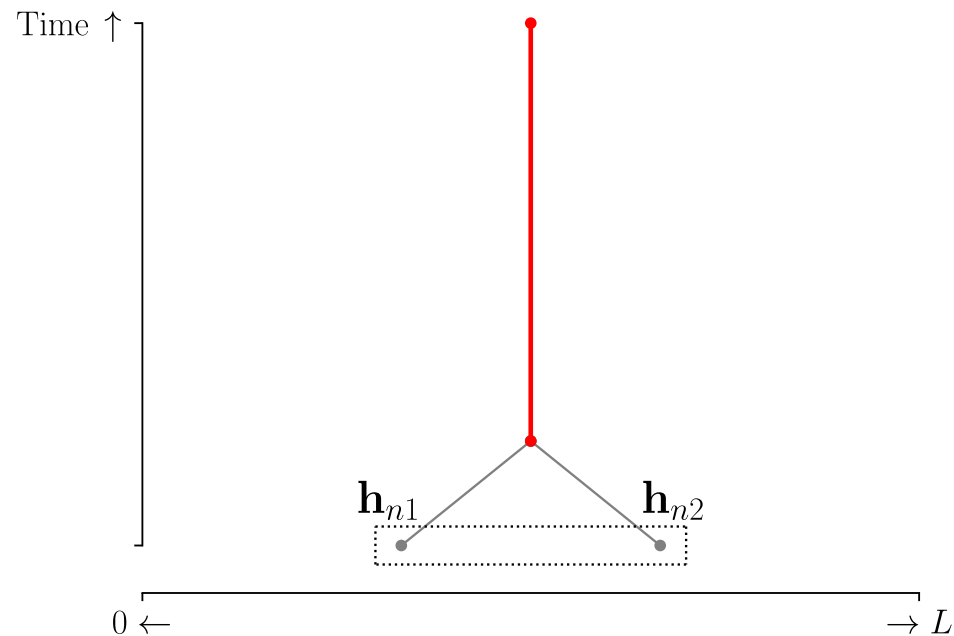
However, all individuals have a different amount of genetic variance (except haploids)

$$\text{Var}(\mathbf{g}_n) = \text{Var}(\mathbf{h}_{n1} + \mathbf{h}_{n2}) = \text{Var}(\mathbf{h}_{n1}) + \text{Var}(\mathbf{h}_{n2}) + 2\text{Cov}(\mathbf{h}_{n1}, \mathbf{h}_{n2})$$



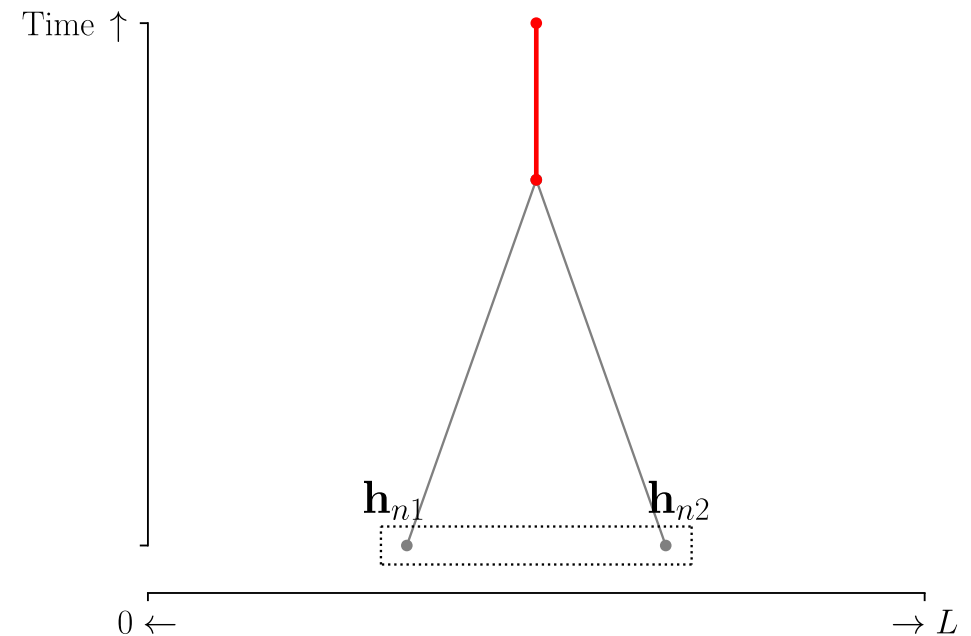
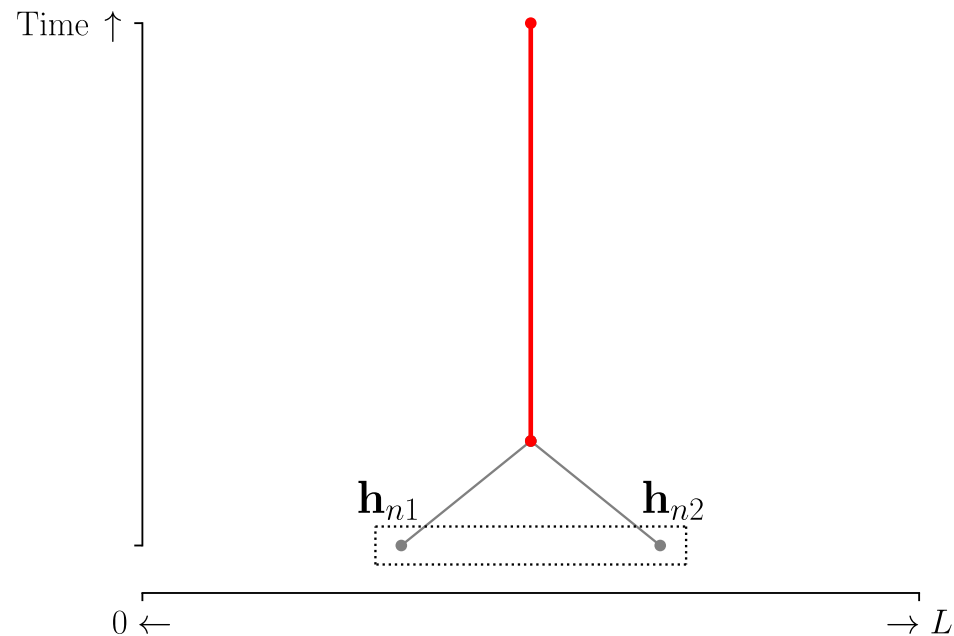
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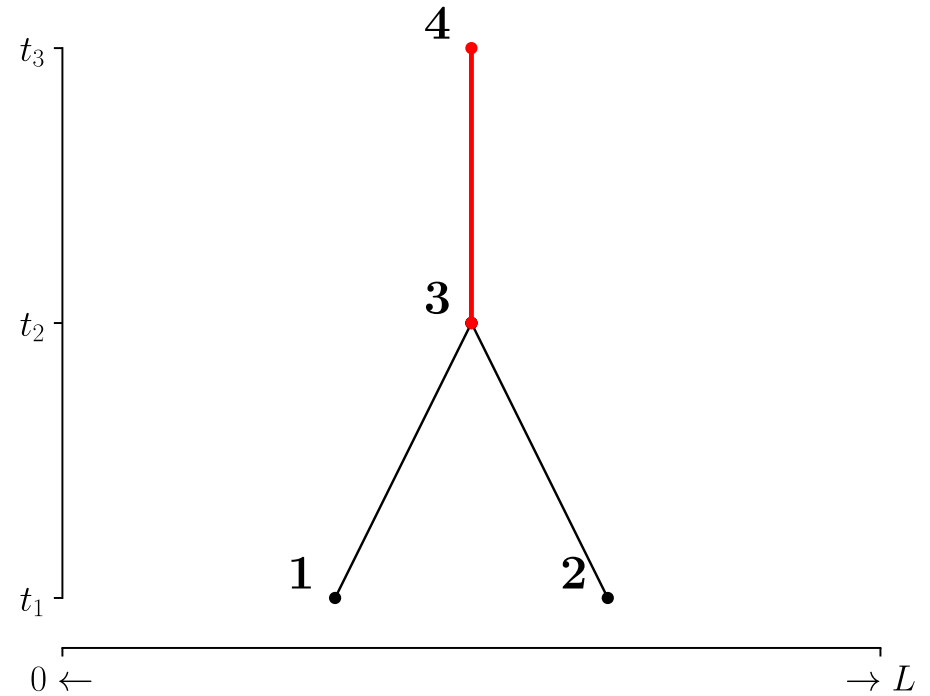


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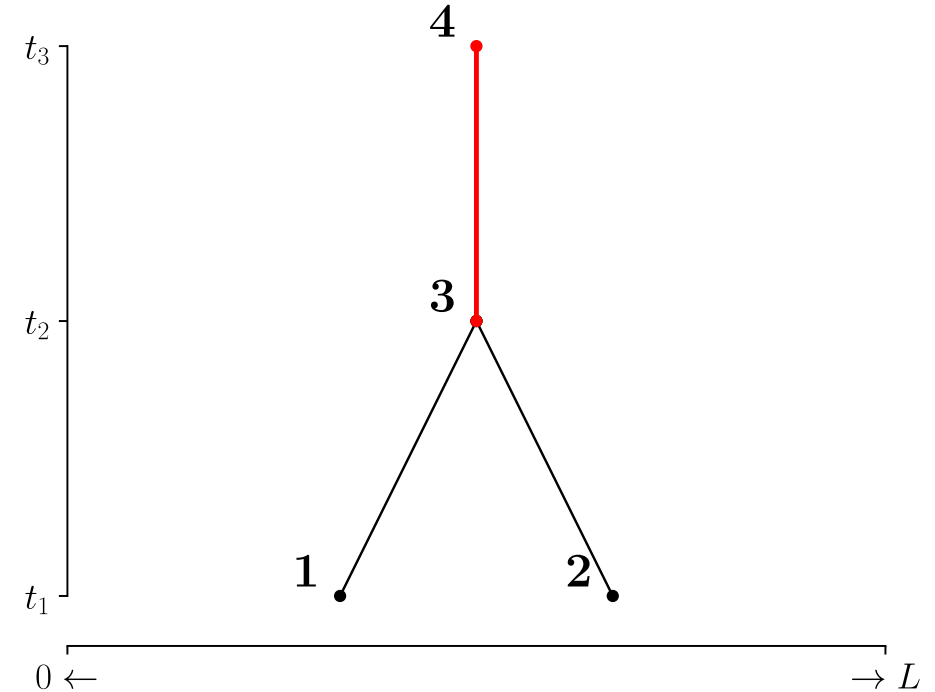
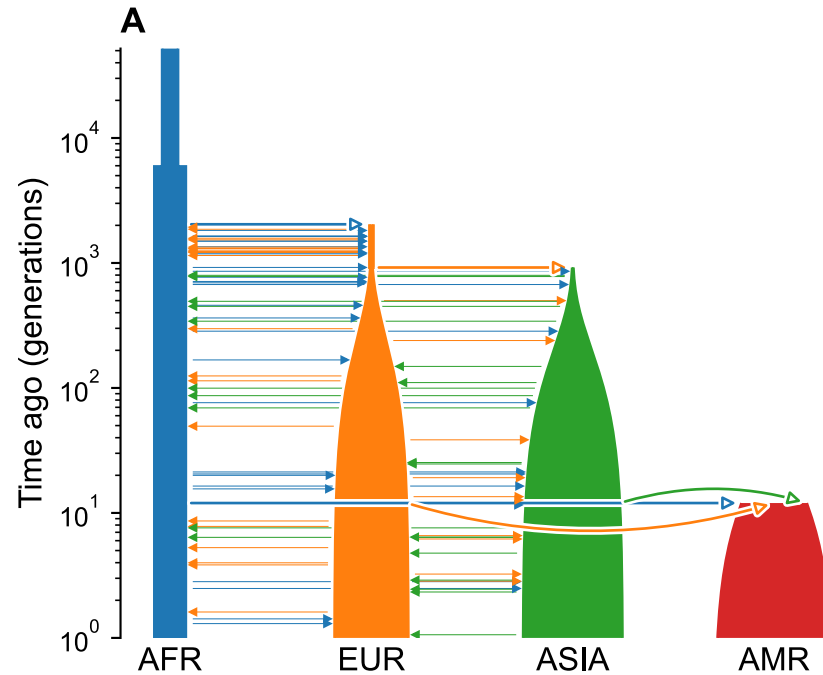
We can't define a single quantity $h_g^2 = \frac{\text{Var}(\mathbf{g}_n)}{\text{Var}(\mathbf{g}_n) + \text{Var}(\boldsymbol{\epsilon}_n)}$ for everyone



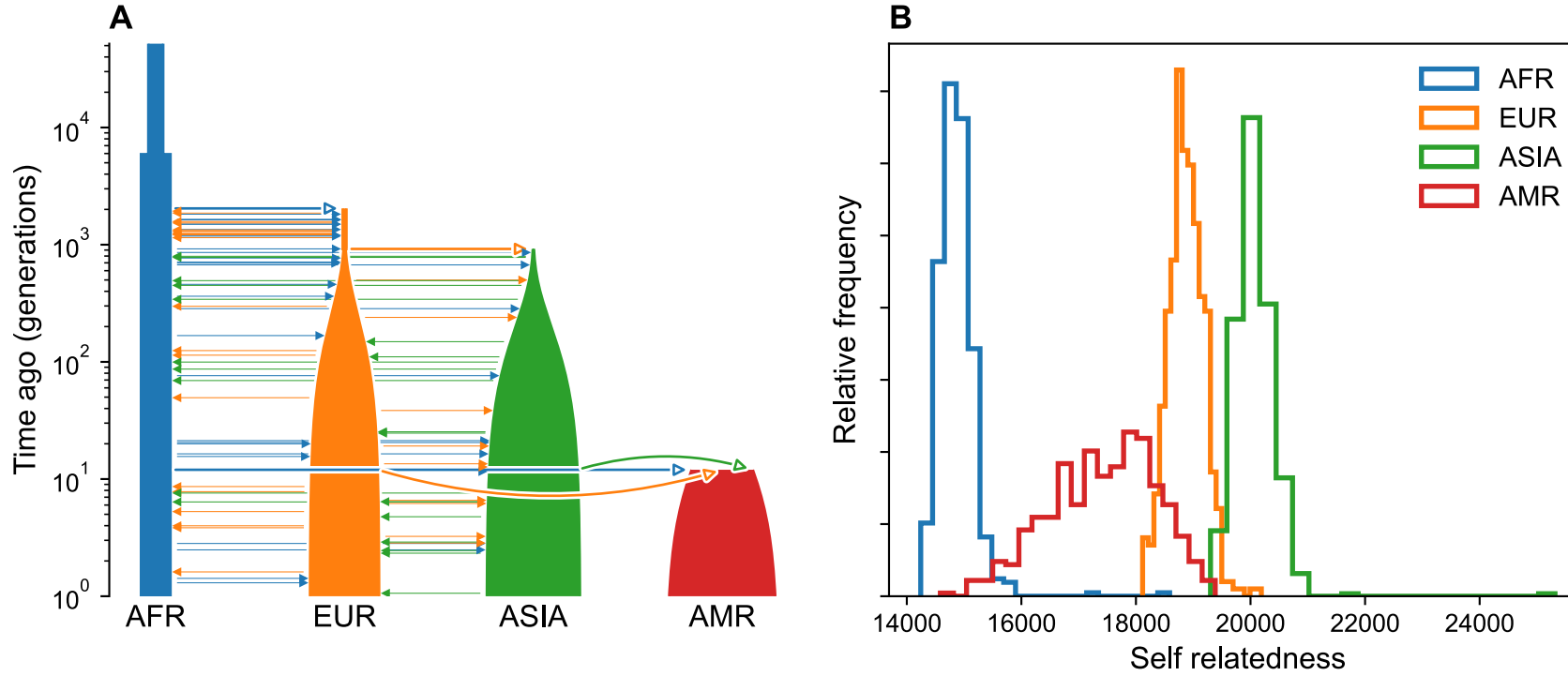
Polygenic prediction is constrained by demography



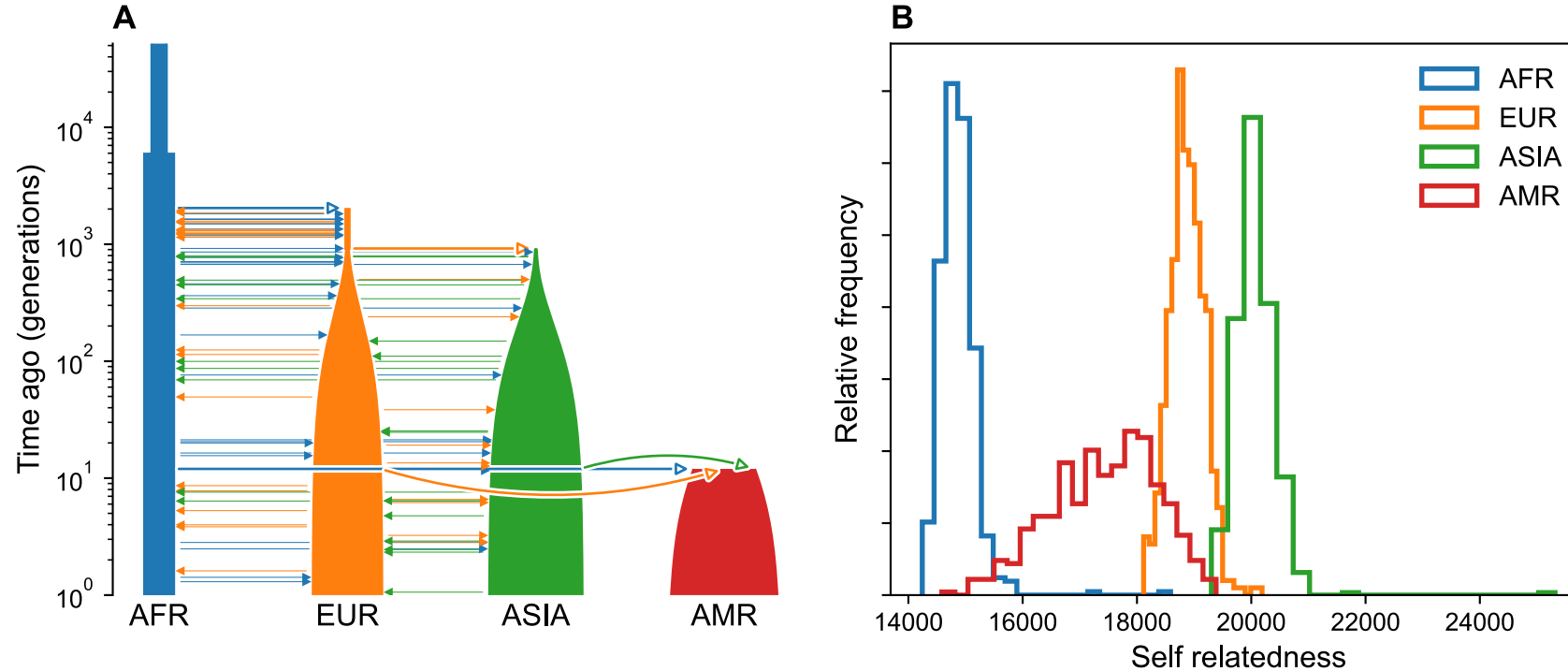
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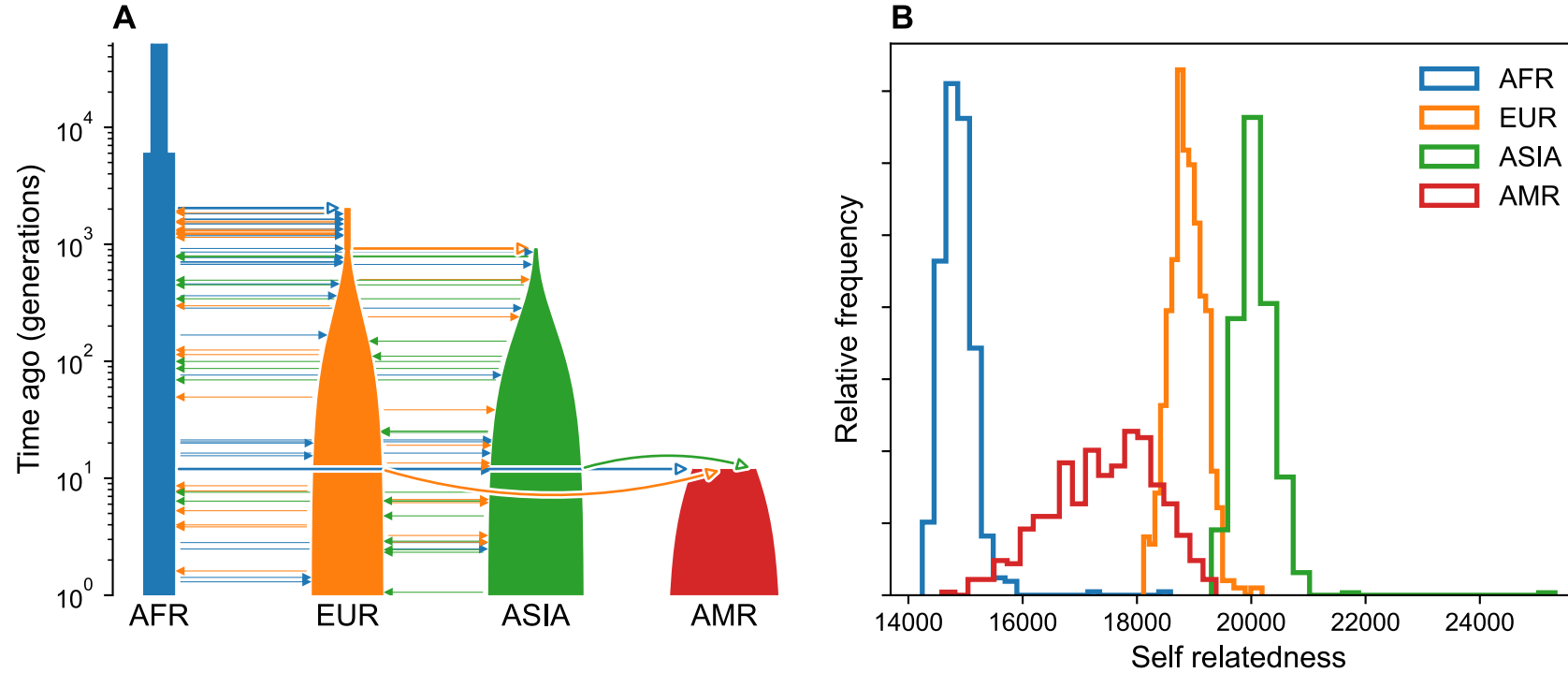


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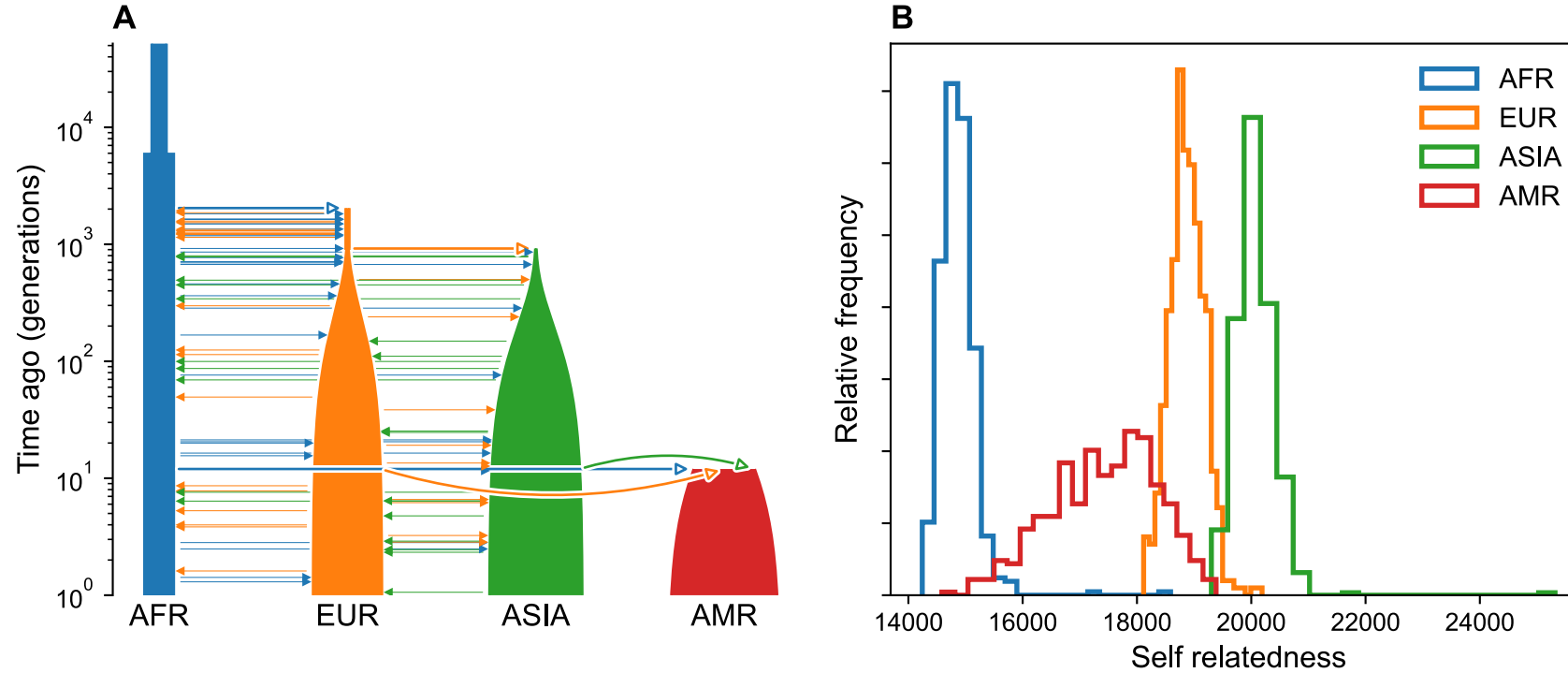
Some people are less genetically variable than others

Polygenic prediction is constrained by demography



Some people are harder to predict genetically than others

Polygenic prediction is constrained by demography



Some populations are **inherently harder** to predict!

tslmm, fitting ARG-LMM to tree sequences

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tslmm utilizes an efficient *genetic relatedness matrix - vector product* to fit the restricted maximum likelihood (REML) objective

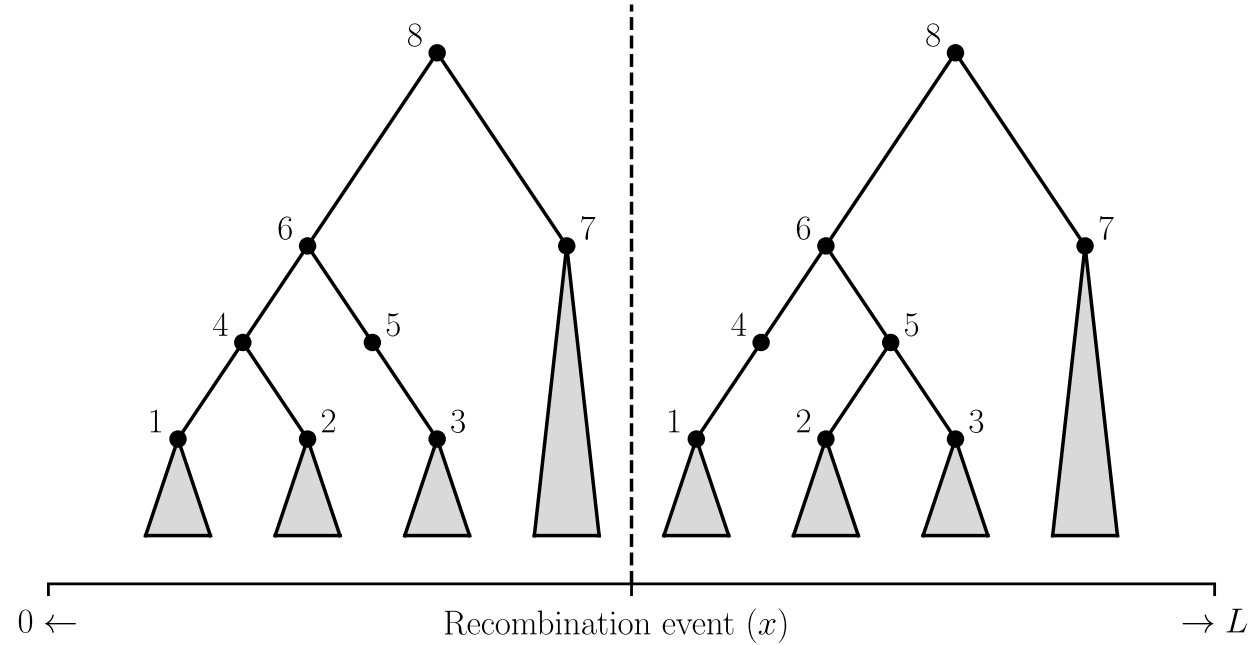
tslmm, fitting ARG-LMM to tree sequences

tslmm utilizes an efficient *genetic relatedness matrix - vector product* to fit the restricted maximum likelihood (REML) objective

It can estimate variance components and compute polygenic scores by best linear unbiased prediction (BLUP)

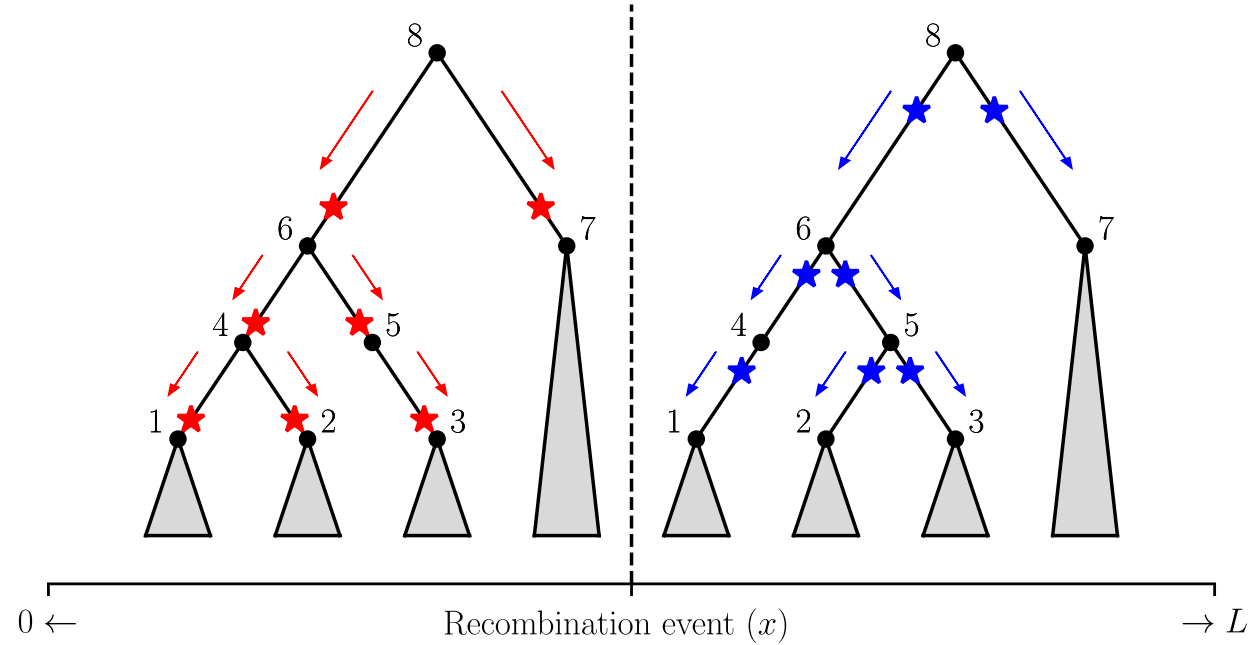
The matrix-vector product algorithm

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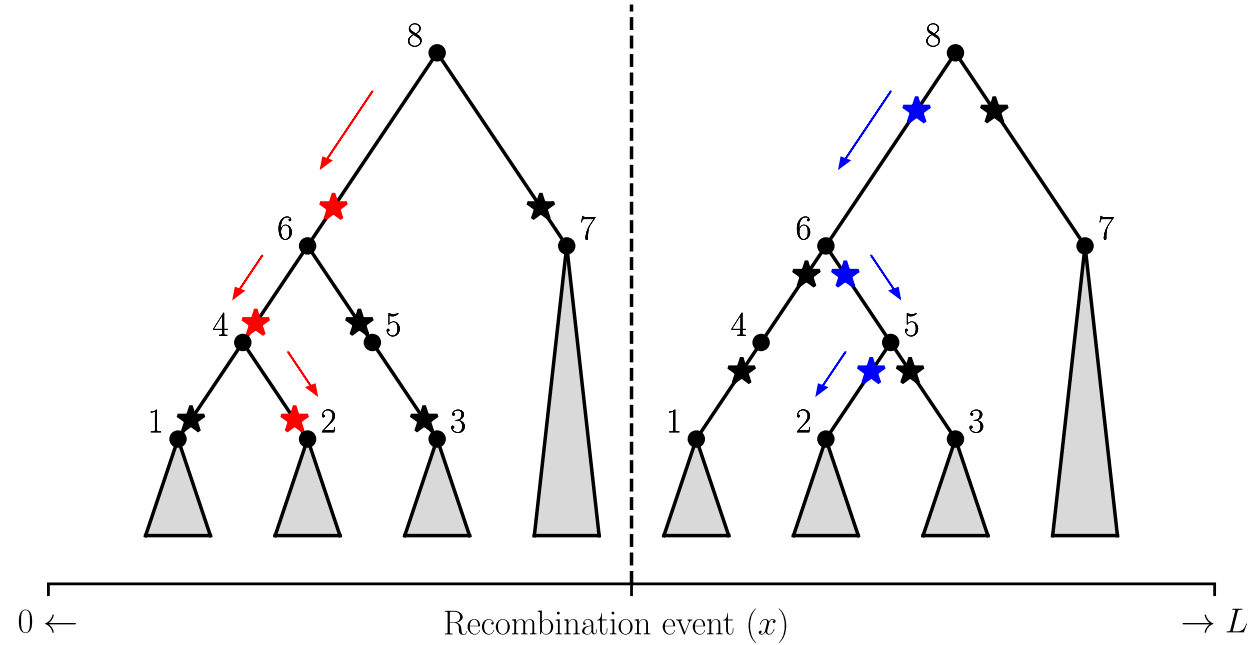
The algorithm needs to pass mutations to the correct samples

The matrix-vector product algorithm



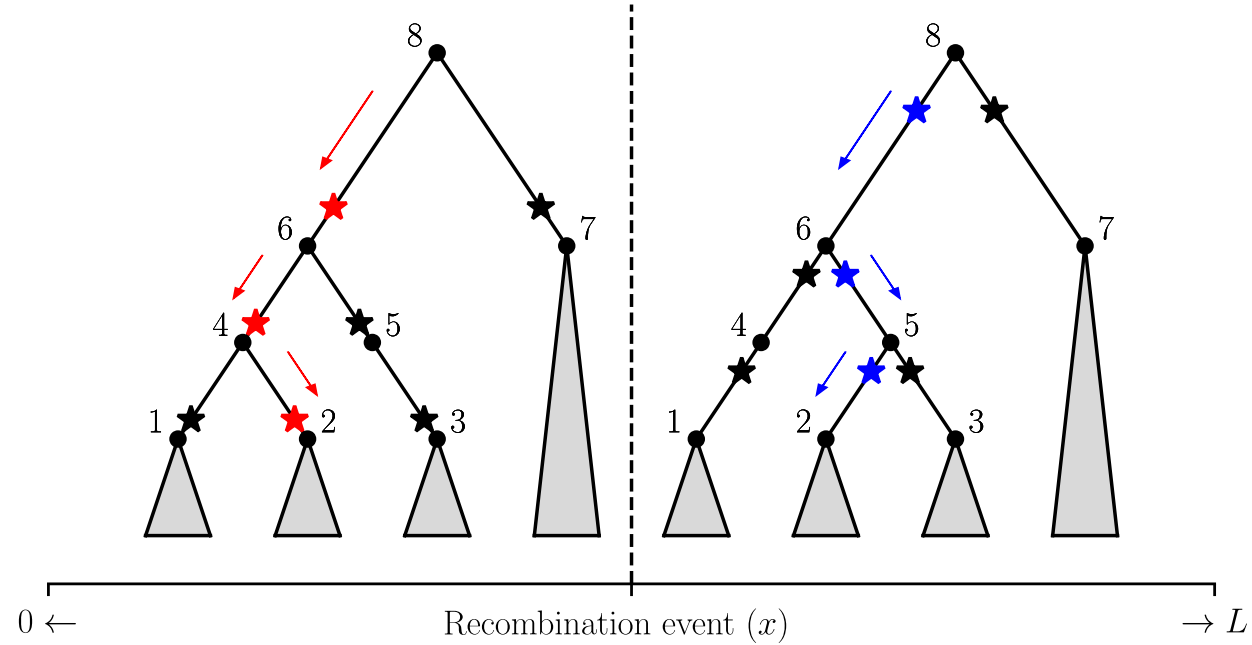
A naive approach is to push the mutations down to the leaves every time

The matrix-vector product algorithm



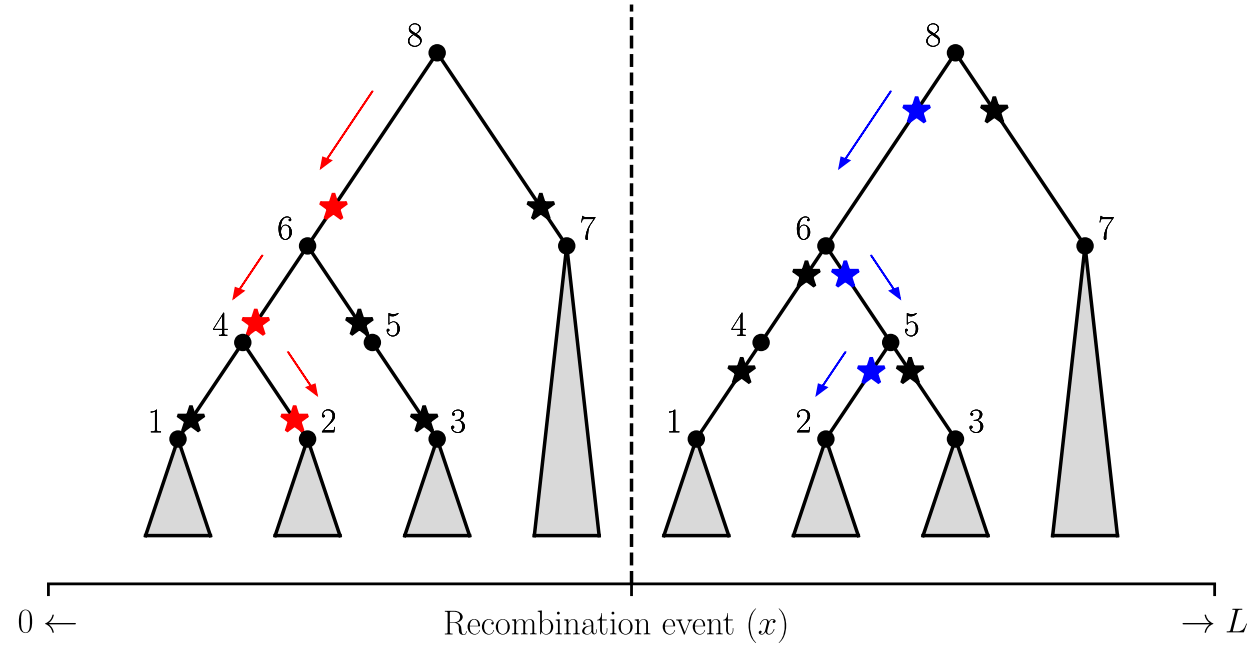
Wait until the subtree's topology changes due to edge insertion/deletion

The matrix-vector product algorithm



The wrong recipient will receive the mutations if we procrastinate further

The matrix-vector product algorithm

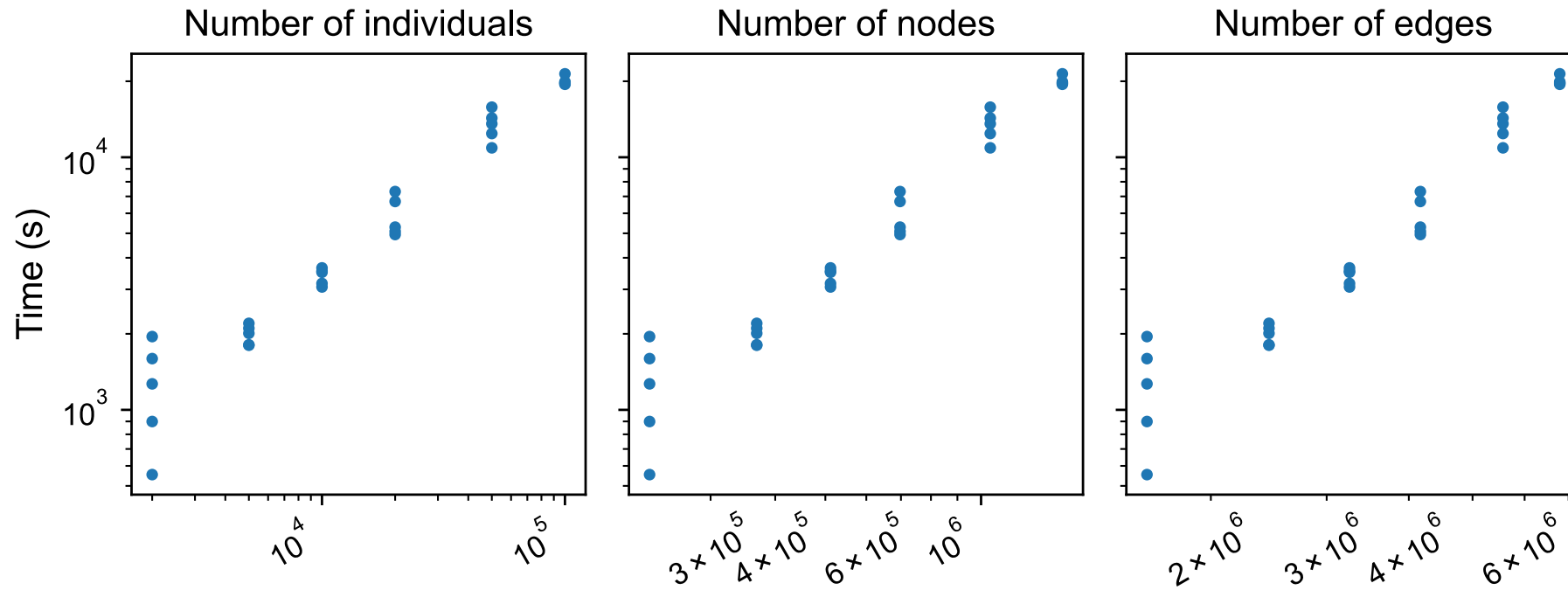


Fitting REML $\mathcal{O}(n_s^3) \Rightarrow \mathcal{O}(n_s + n_t \log n_s)$

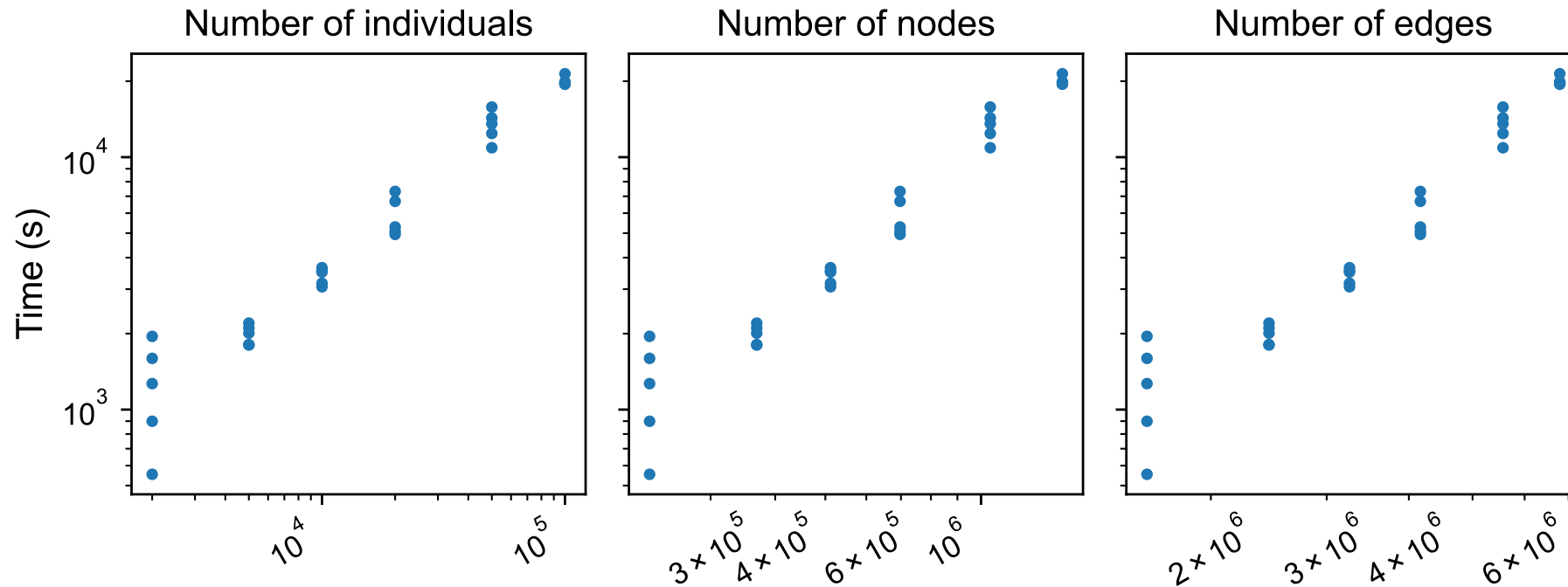
n_s : number of samples, n_t : number of trees

Runtime for variance component estimation

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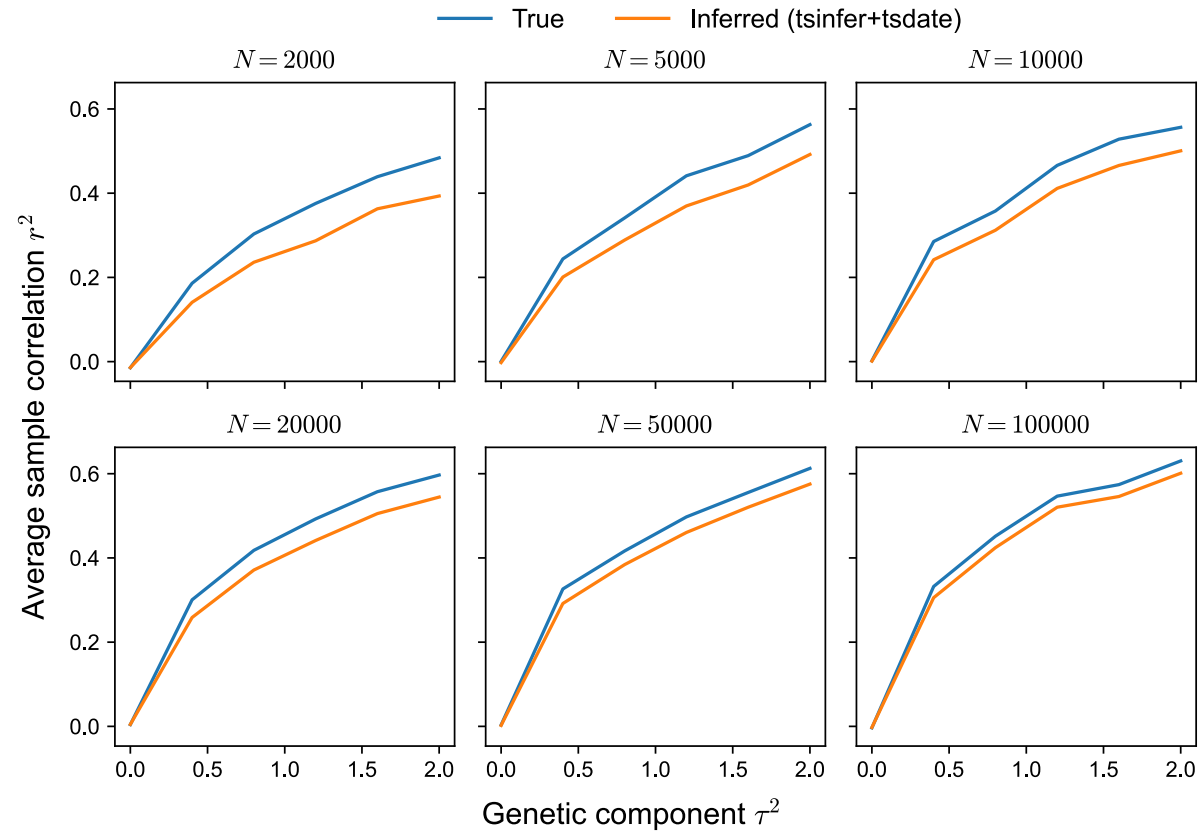
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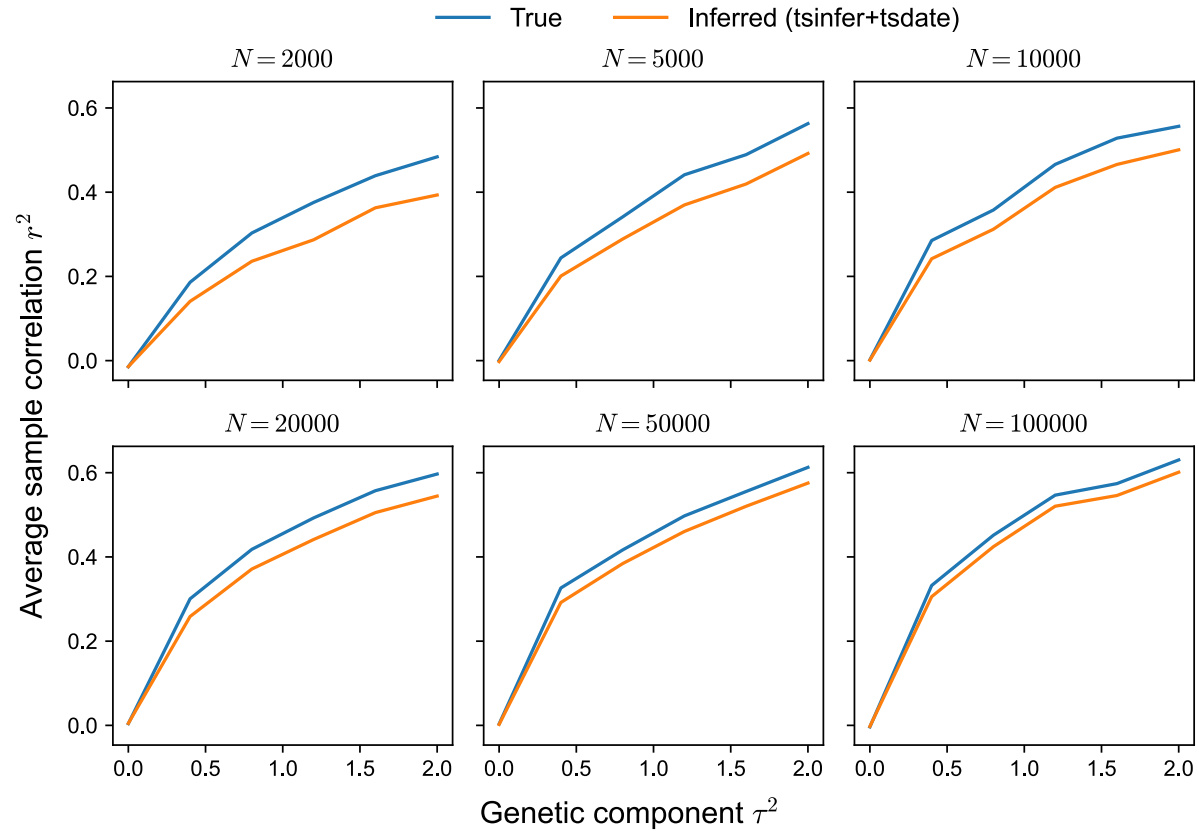
The runtime scales linearly with respect to the number of individuals (genome length = 10^8)

Best linear unbiased prediction (BLUP)

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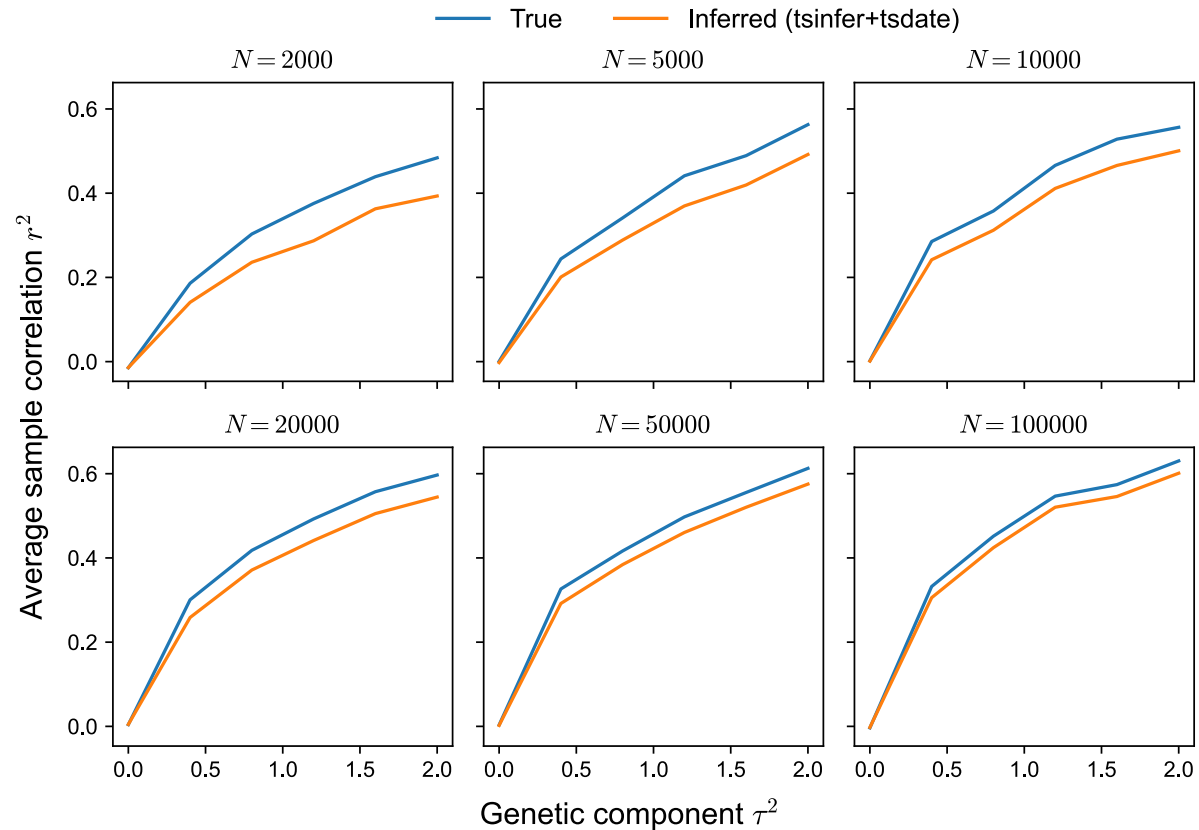


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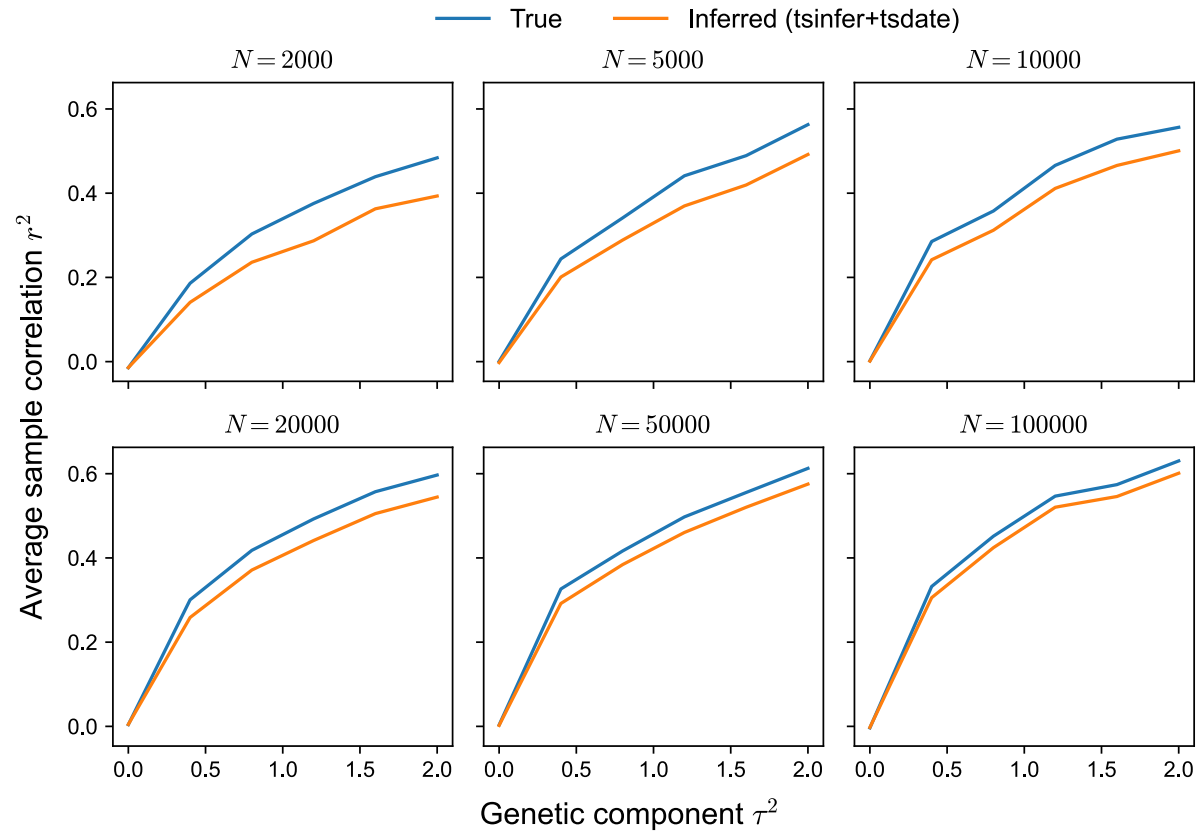
We measured the accuracy of polygenic scores computed from **tslmm**

Best linear unbiased prediction (BLUP)



Training and testing on two non-overlapping groups embedded in the same tree sequence

Best linear unbiased prediction (BLUP)



True trees are better, but inferred trees are not too behind!

Summary & Future directions

ARG-LMM lays an explicit connection between population and quantitative genetics

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Time conditioned analysis (random vs fixed effects) ([Fan, Mancuso, and Chiang 2022](#))

Thank you for listening



Link to ([Lehmann et al. 2025](#)), tslmm preprint coming soon

Collaborators: Nathaniel Pope (Oregon), Jerome Kelleher (Oxford), Gregor Gorjanc (Edinburgh), and Peter Ralph (Oregon)

References

- Browning, Sharon R., Brian L. Browning, Martha L. Daviglus, Ramon A. Durazo-Arvizu, Neil Schneiderman, Robert C. Kaplan, and Cathy C. Laurie. 2018. “Ancestry-Specific Recent Effective Population Size in the Americas.” Edited by Kirk E. Lohmueller. *PLOS Genetics* 14 (5): e1007385. <https://doi.org/10.1371/journal.pgen.1007385>.
- Cranmer, Kyle, Johann Brehmer, and Gilles Louppe. 2020. “The Frontier of Simulation-Based Inference.” *Proceedings of the National Academy of Sciences* 117 (48): 30055–62. <https://doi.org/10.1073/pnas.1912789117>.
- Edge, Michael D, and Graham Coop. 2018. “Reconstructing the History of Polygenic Scores Using Coalescent Trees.” *Genetics* 211 (1): 235–62. <https://doi.org/10.1534/genetics.118.301687>.
- Fan, Caoqi, Nicholas Mancuso, and Charleston W. K. Chiang. 2022. “A Genealogical Estimate of Genetic Relationships.” *The American Journal of Human Genetics* 109 (5): 812–24. <https://doi.org/10.1016/j.ajhg.2022.03.016>.
- Lehmann, Brieuc, Hanbin Lee, Luke Anderson-Trocme, Jerome Kelleher, Gregor Gorjanc, and Peter L. Ralph. 2025. “On ARGs, Pedigrees, and Genetic Relatedness Matrices,” March. <https://doi.org/10.1101/2025.03.03.641310>.
- Peng, Dandan, Obadiah J. Mulder, and Michael D. Edge. 2024. “Evaluating ARG-Estimation Methods in the Context of Estimating Population-Mean Polygenic Score Histories,” May. <https://doi.org/10.1101/2024.05.24.595829>.
- Rosenberg, Noah A., and Jenna M. VanLiere. 2009. “Replication of Genetic Associations as Pseudoreplication Due to Shared Genealogy.” *Genetic Epidemiology* 33 (6): 479–87. <https://doi.org/10.1002/gepi.20400>.
- Salehi Nowbandegani, Pouria, Anthony Wilder Wohns, Jenna L. Ballard, Eric S. Lander, Alex Bloemendal, Benjamin M. Neale, and Luke J. O’Connor. 2023. “Extremely Sparse Models of Linkage Disequilibrium in Ancestrally Diverse Association Studies.” *Nature Genetics* 55 (9): 1494–1502. <https://doi.org/10.1038/s41588-023-01487-8>.
- Wakeley, John. 2008. *Coalescent Theory*. Greenwood Village, CO: Roberts & Company.
- Wong, Yan, Anastasia Ignatieva, Jere Koskela, Gregor Gorjanc, Anthony W Wohns, and Jerome Kelleher. 2024. “A General and Efficient Representation of Ancestral Recombination Graphs.” Edited by G Coop. *GENETICS* 228 (1). <https://doi.org/10.1093/genetics/iyae100>.

Technical Notes

Edge splitting

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- Edges, in particular, may not have a unique set of samples along their span
- Salehi Nowbandegani and colleagues *bricked* the edges to divide them ([Salehi Nowbandegani et al. 2023](#))
- Henceforth, we assume that edges are splitted to have a unique subtopology

\mathbf{Z}_{ne} = The number of haplotypes of individual n that inherit e

The overall matrix \mathbf{Z} is an individual-edge design matrix.

Collapsing variants to edges

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$$\mathbf{G}_{np} = \sum_{e:p \in e} \mathbf{Z}_{ne} \mathbf{1}_{ep} \quad \Leftrightarrow \quad \mathbf{G}_p = \sum_{e:p \in e} \mathbf{Z}_e \mathbf{1}_{ep}$$

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- Assumes that there are no parent-child mutation pairs, but allows *some* recurrent mutations

Exchange the summations

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Recall $\mathbf{G}_p = \sum_{e:p \in e} \mathbf{Z}_e \mathbf{1}_{ep}$ and $\mathbf{y} = \sum_{p=1}^P \mathbf{G}_p \boldsymbol{\beta}_p + \boldsymbol{\varepsilon}$

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Substitute \mathbf{G}_p

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Exchange the inner and the outer summation

$$\sum_{e=1}^E \sum_{p:p \in e} \mathbf{Z}_e \boldsymbol{\beta}_p \mathbf{1}_{ep} + \boldsymbol{\varepsilon}$$

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Pull out \mathbf{Z}_e and group the positions nested in $p : p \in e$

$$\begin{aligned} & \sum_{e=1}^E \mathbf{Z}_e \left(\sum_{p:p \in e} \boldsymbol{\beta}_p \mathbf{1}_{ep} \right) + \boldsymbol{\varepsilon} \\ &= \sum_{e=1}^E \mathbf{Z}_e \boldsymbol{v}_e + \boldsymbol{\varepsilon} \\ &= \mathbf{Z} \boldsymbol{v} + \boldsymbol{\varepsilon} \end{aligned}$$

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\boldsymbol{v} is a random variable made up of mutation-driven random variables $\mathbf{1}_{ep}$!

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$$\text{Cov}(\mathbf{u}_e, \mathbf{u}_{e'}) = \sum_{p \in e, e'} \beta_p^2 \text{Cov}(\mathbf{1}_{ep}, \mathbf{1}_{e'p})$$

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- The covariance between the indicators are higher-order terms of mutation rates, so we ignore it ([Wakeley 2008](#))

$$\begin{aligned} \text{Cov}(\mathbf{1}_{ep}, \mathbf{1}_{e'p}) &= \mathbb{E}[\mathbf{1}_{ep} \mathbf{1}_{e'p}] - \mathbb{E}[\mathbf{1}_{e'p}] \mathbb{E}[\mathbf{1}_{ep}] \\ &= 0 - l_e u_{ep} l_{e'} u_{e'p} \approx 0 \end{aligned}$$

where l_e is the (time-)length of edge e .

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$$\mathbf{u}_e / \sqrt{l_e s_e} \cdot \sqrt{\frac{1}{s_e} \sum_{p:p \in e} \beta_p^2 u_{ep}} \rightarrow N(0, 1^2) \text{ as } s_e \rightarrow \infty$$

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- Fortunately, the variance is computable and is

$$\text{Var}(\mathbf{u}_e) = l_e s_e \cdot \frac{1}{s_e} \sum_{p:p \in e} \beta_p^2 u_{ep}$$

More on $\text{Var}(\mathbf{u}_e)$

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- As a measure of functional significance, variance components are confounded by the area

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$$[\mathbf{Zf}]_n = \sum_{e=1}^E \mathbf{z}_{ne} \mathbf{E} \left[\sum_{p:p \in e} \beta_p \mathbf{1}_{ep} \right]$$

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Conjecture: selection \Rightarrow fixed effects?

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- Non-overlapping samples are not independent

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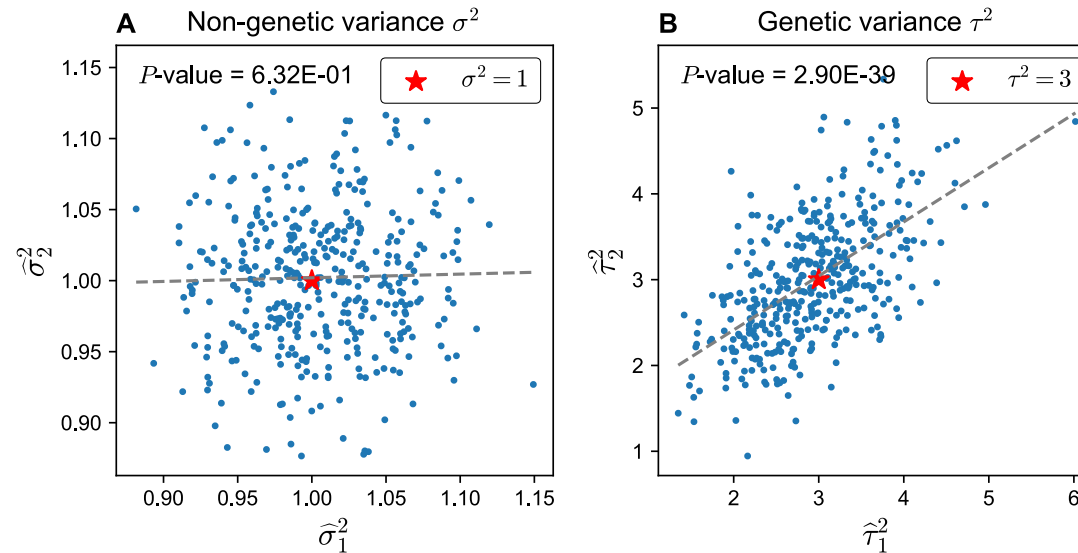
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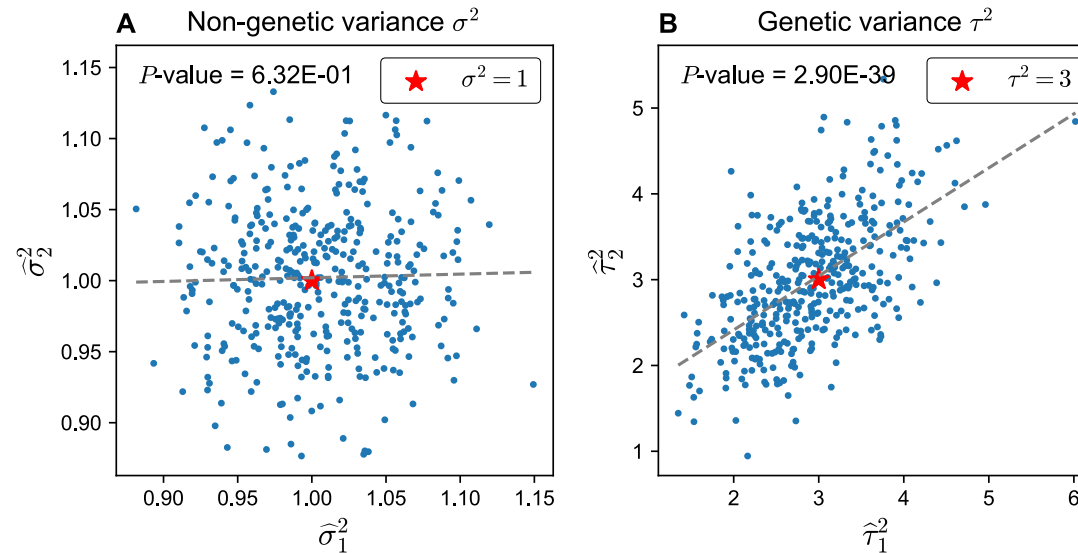
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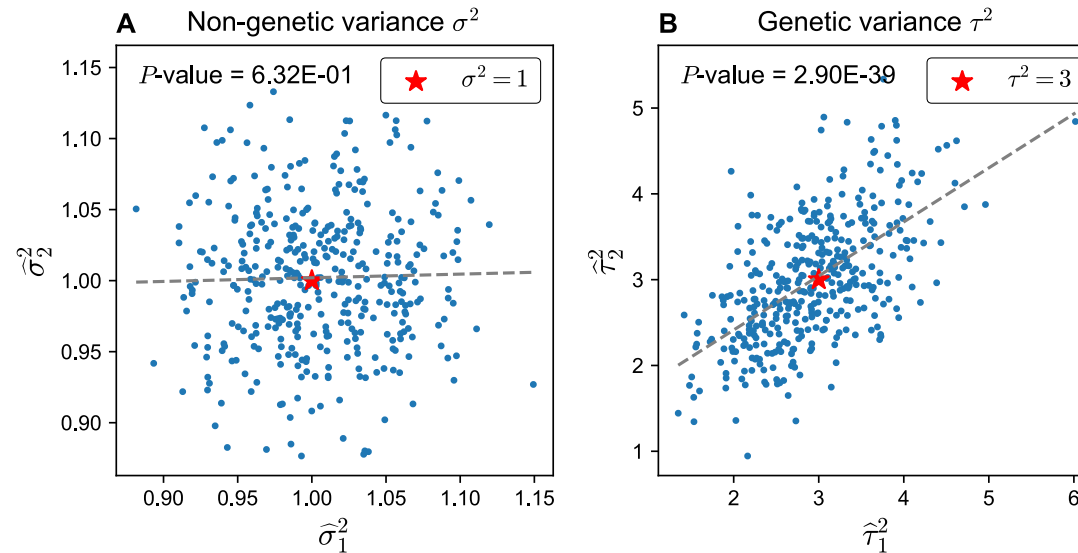
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This is also the very reason why BLUP works

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We are all correlated!

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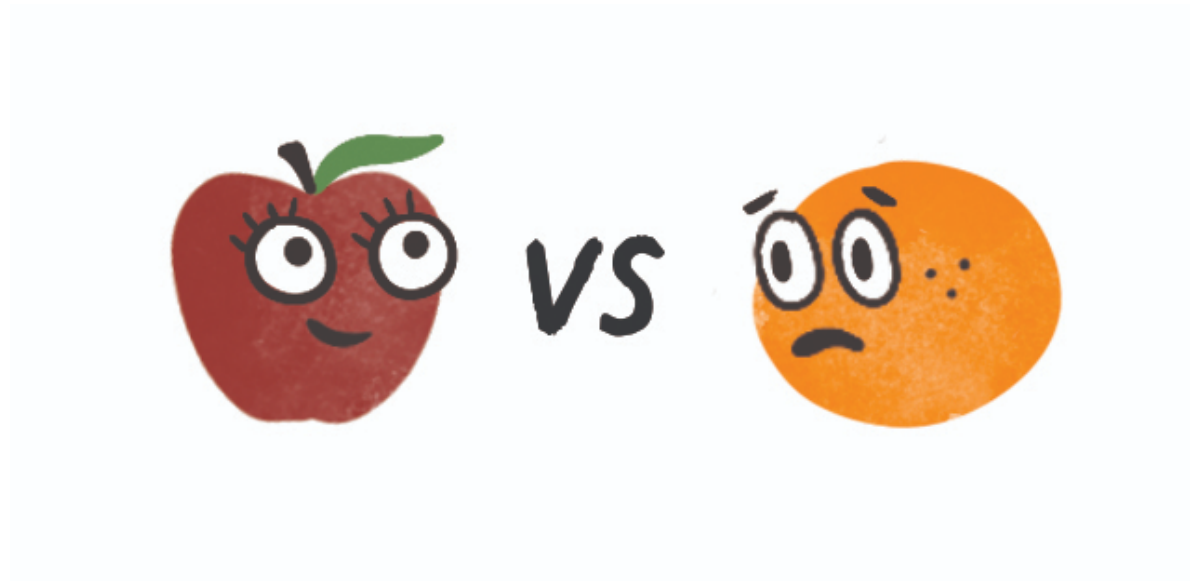
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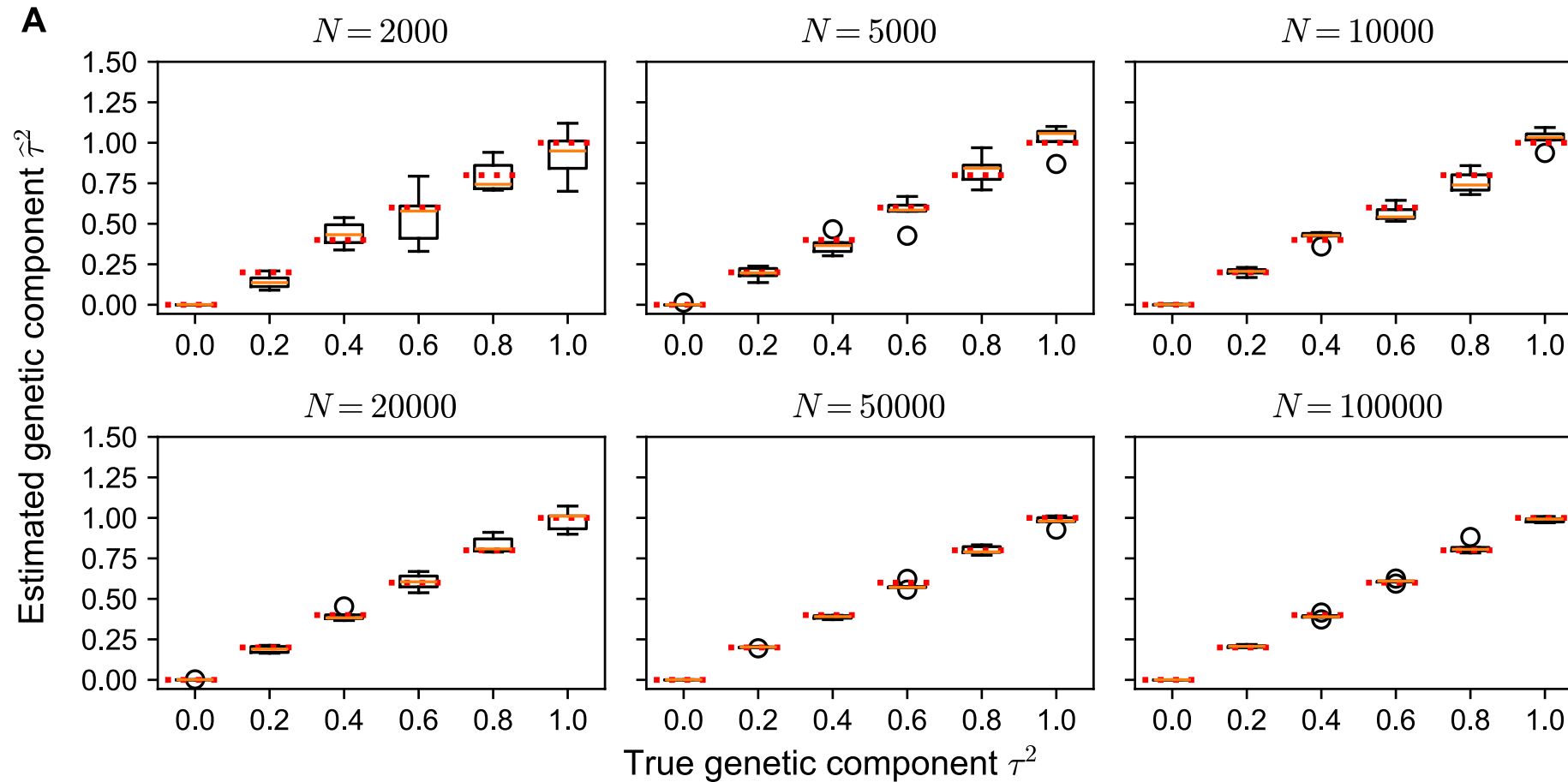
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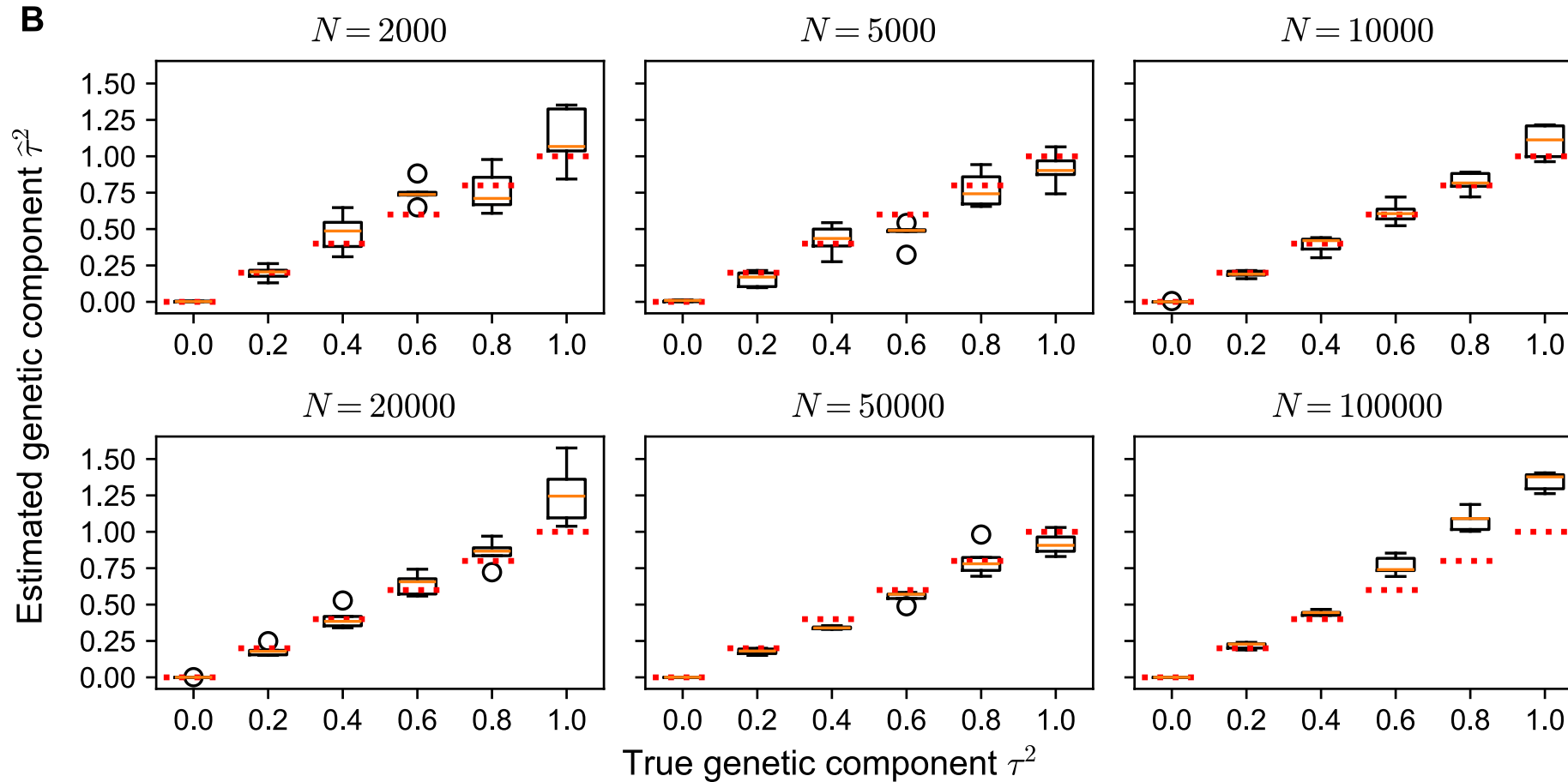
Estimation quality of variance components

Simulated trees



Estimation quality of variance components

Inferred (tsinfer+tsdate) trees



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- Time conditioning (reference population)