

JTMS-12: Probability and Random Processes

Fall 2020

M. Bode

Lecture 13:

Recap and Exam-type questions

Lecture 13

Recap:

Chapter 3: Functions of Random Variables

3.1 Functions of Random Variables

3.2 $Y = g(X)$

3.3 $Z = g(X, Y)$

3.4 $V = g(X, Y), W = h(X, Y)$

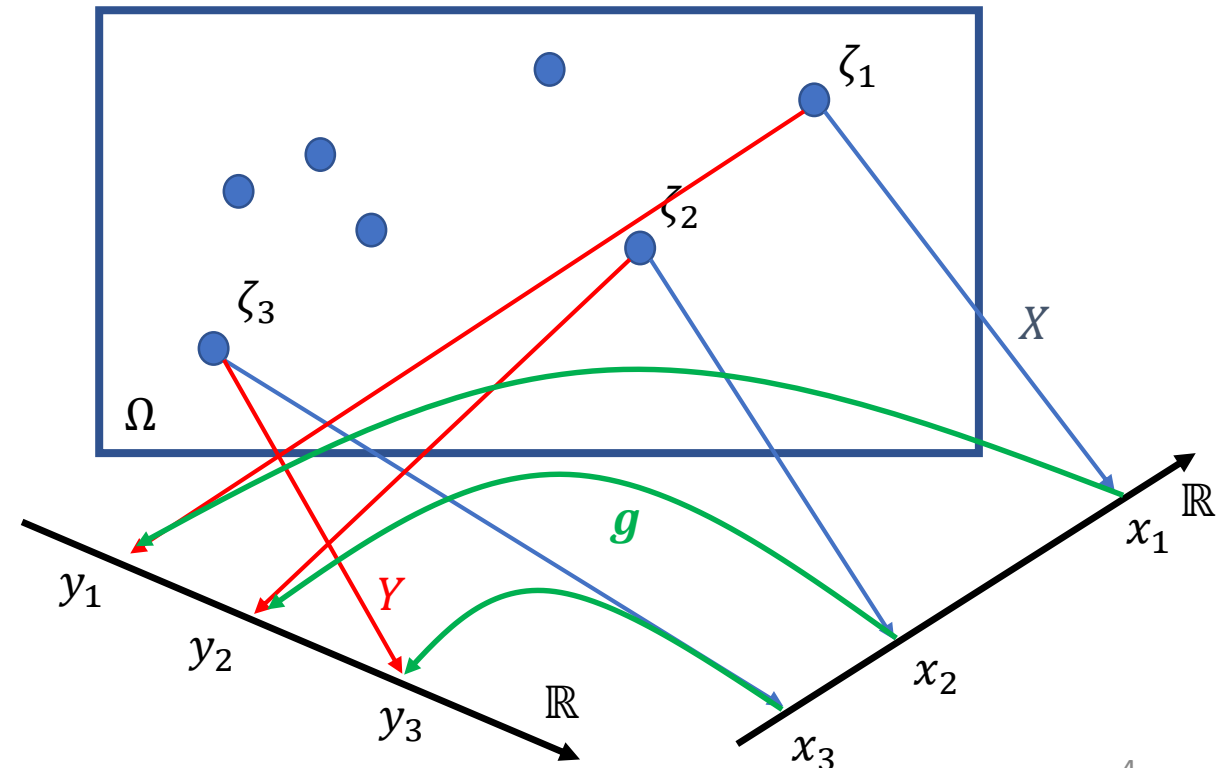
Functions of Random Variables

Idea: Map outcomes to (real) numbers.

The **random variable** $X: \Omega \rightarrow \mathbb{R}$ maps all outcomes from the sample description space to a real number.

Re-label: $y = g(x)$

Re-interpret: $Y: \Omega \rightarrow \mathbb{R}, Y(\zeta) = g(X(\zeta))$



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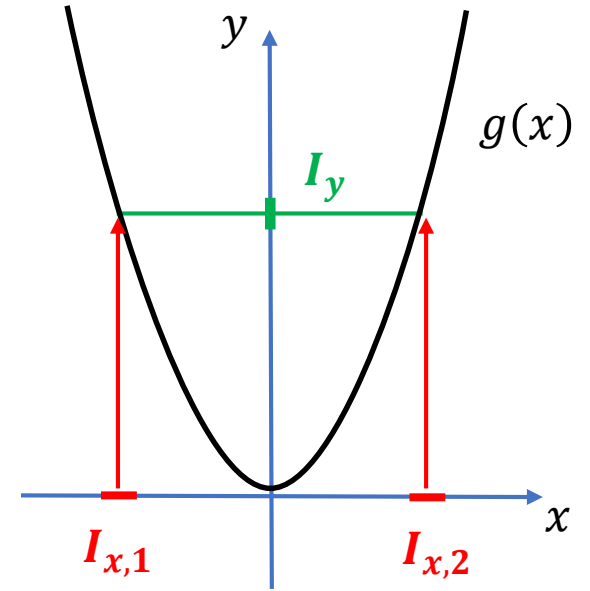
Functions of Random Variables

Quick way to find the density:

The direct formula (for cont. r.v.s)

$$P[Y \in I_y] = P[X \in I_{x,1}] + P[X \in I_{x,2}]$$

$$f_Y(y)|\Delta y| \approx f_X(x_1)|\Delta x_1| + f_X(x_2)|\Delta x_2|$$



Mind the relative orientations of I_y and its (partial) pre-images $I_{x,1}$ and $I_{x,2}$.

$$\Rightarrow f_Y(y) \approx f_X(x_1) \left| \frac{\Delta x_1}{\Delta y} \right| + f_X(x_2) \left| \frac{\Delta x_2}{\Delta y} \right|$$

Hence, in the limit $\Delta y \rightarrow 0$:

$$f_Y(y) = f_X(x_1) \left| \frac{1}{g'(x_1)} \right| + f_X(x_2) \left| \frac{1}{g'(x_2)} \right|$$

...sum over all pre-images

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Functions of Random Variables

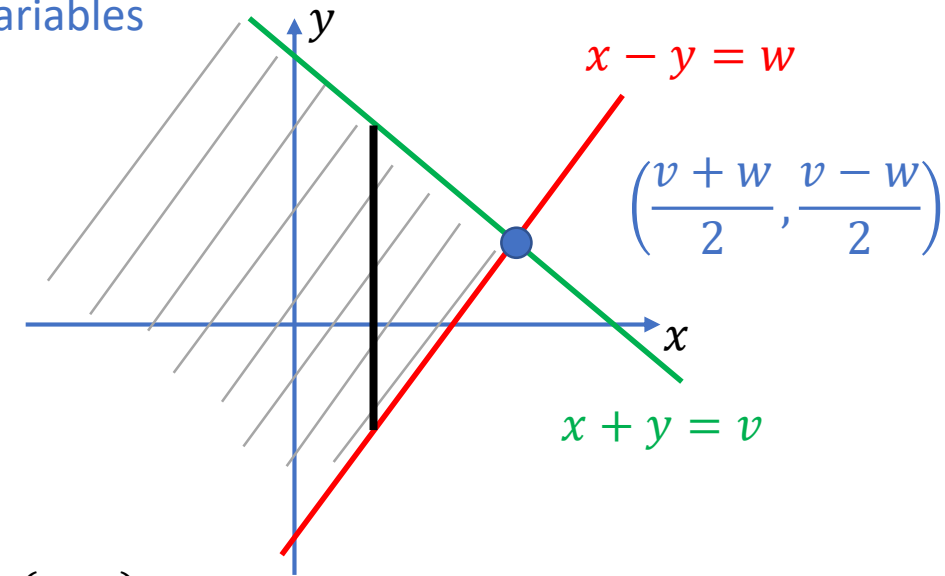
Example:

Consider continuous
r.v.s X, Y , and

$$V = g(X, Y) = X + Y$$

$$W = h(X, Y) = X - Y$$

Find $F_{VW}(v, w)$ and $f_{VW}(v, w)$.



Standard approach ...

$$F_{VW}(v, w) = P[V \leq v, W \leq w] = P[g(X, Y) \leq v, h(X, Y) \leq w]$$

$$= P[X + Y \leq v, X - Y \leq w] = \iint_{\text{shaded area}} f_{XY}(x, y) dx dy$$

... integrate & derive to find the density

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Jacobian matrix:

$$\frac{\Delta(v, w)}{\Delta(x, y)} = \left| \det \begin{pmatrix} 1 & 1 \\ 1 & -1 \end{pmatrix} \right| = 2$$

Functions of Random Variables

Determine densities like $f_{UV}(u, v)$ directly

Same Example - Alternative Perspective:

$$V = g(X, Y) = X + Y, \quad W = h(X, Y) = X - Y$$

$$\Leftrightarrow X = \frac{V + W}{2}, \quad Y = \frac{V - W}{2}$$

→

$$f_{VW}(v, w) = \frac{1}{2} f_{XY} \left(\frac{v + w}{2}, \frac{v - w}{2} \right)$$

Mind:

$$f_{VW}(v, w) = f_{XY}(x(v, w), y(v, w)) \frac{\Delta(x, y)}{\Delta(v, w)}$$

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Recap & Exam type questions

TASK 8:

Consider a continuous r.v. X , and

$$Y = g(X) = \exp(X)$$

Find $F_Y(y)$ and $f_Y(y)$.

Standard approach ...

$$\begin{aligned} F_Y(y) &= P[Y \leq y] = P[g(X) \leq y] = P[\exp(X) \leq y] \\ &= P[X \leq \ln(y)] = F_X(\ln(y)) \end{aligned}$$

For which y -range does that make sense?

$y > 0$, only, $y \leq 0$ is impossible.

$$\Rightarrow F_Y(y) = \begin{cases} F_X(\ln(y)); & y > 0 \\ 0; & \text{else} \end{cases}$$

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Recap & Exam type questions

Solve:

Consider a continuous r.v. X , and

$$Y = g(X) = \exp(X)$$

Find $F_Y(y)$ and $f_Y(y)$.

Chapter 3: Functions of Random Variables

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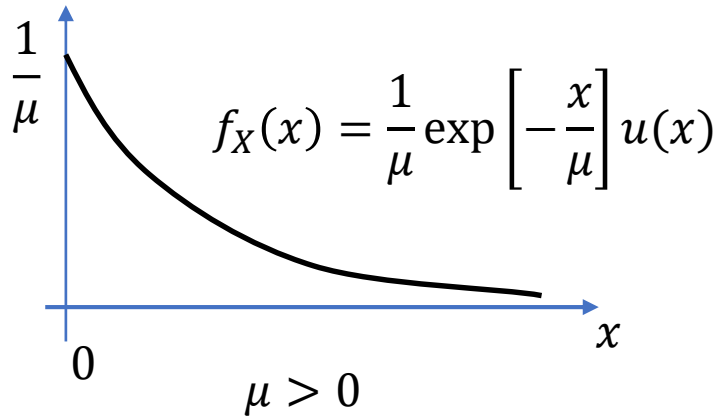
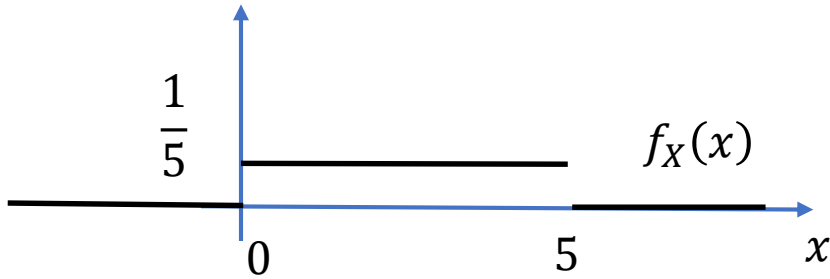
3.4 $V = g(X, Y), W = h(X, Y)$

$$F_Y(y) = \begin{cases} F_X(\ln(y)); & y > 0 \\ 0; & \text{else} \end{cases}$$

$$\Rightarrow f_Y(y) = \begin{cases} \frac{1}{y} f_X(\ln(y)); & y > 0 \\ 0; & \text{else} \end{cases}$$

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As in an exam: Find the mistakes!
Double check via integration of $f_Y(y)$



Recap & Exam type questions

Solve:

Consider a continuous r.v. X , and $Y = g(X) = \exp(X)$

Special case: uniform over $(0, 5)$

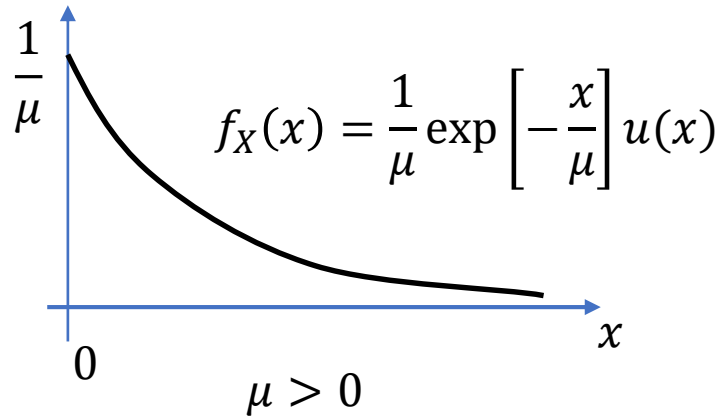
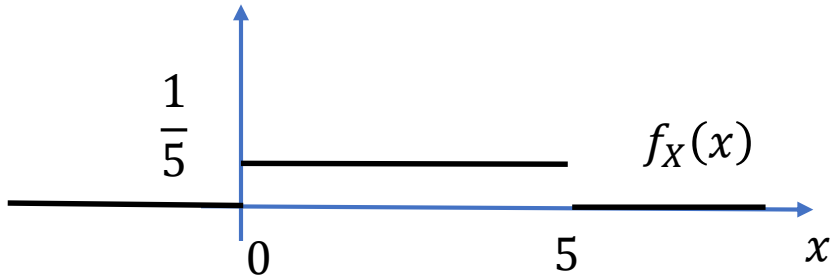
$$\Rightarrow f_Y(y) = \begin{cases} \frac{1}{y} f_X(\ln(y)); & y > 0 \\ 0; & \text{else} \end{cases} = \begin{cases} \frac{1}{5y}; & y > 0 \\ 0; & \text{else} \end{cases}$$

Special case: exponential

$$\Rightarrow f_Y(y) = \begin{cases} \frac{1}{\mu y} \exp\left[-\frac{\ln(y)}{\mu}\right]; & y > 0 \\ 0; & \text{else} \end{cases} = \begin{cases} \frac{1}{\mu y^{\mu+1}}; & y > 0 \\ 0; & \text{else} \end{cases}$$

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Found the mistakes?



Recap & Exam type questions

Solve:

Consider a continuous r.v. X , and $Y = g(X) = \exp(X)$

Special case: uniform over $(0, 5)$... maps to $(1, e^5)$

$$\Rightarrow f_Y(y) = \begin{cases} \frac{1}{y} f_X(\ln(y)); & y > 0 \\ 0; & \text{else} \end{cases} = \begin{cases} 0; & y > e^5 \\ \frac{1}{5y}; & 1 \leq y \leq e^5 \\ 0; & y \leq 1 \end{cases}$$

Special case: exponential over $(0, \infty)$... maps to $(1, \infty)$

$$\Rightarrow f_Y(y) = \begin{cases} \frac{1}{\mu y} \exp\left[-\frac{\ln(y)}{\mu}\right]; & y > 0 \\ 0; & \text{else} \end{cases} = \begin{cases} \frac{1}{\mu y^{\mu+1}}; & y > 1 \\ 0; & y \leq 1 \end{cases}$$

Mind the boundaries
Resulting from the support of $f_X(x)$

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Functions of Random Variables

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TASK 9:

Consider continuous two independent standard normal r.v.s X, Y

$$f_{XY}(x, y) = \frac{1}{2\pi} \exp\left[-\frac{x^2 + y^2}{2}\right]$$

$$V = g(X, Y) = X + 2Y, \quad W = h(X, Y) = X - 2Y$$

Find $f_{VW}(v, w)$. Are V and W still independent?

Standard approach ...

$$\begin{aligned} F_{VW}(v, w) &= P[V \leq v, W \leq w] = P[g(X, Y) \leq v, h(X, Y) \leq w] \\ &= P[X + 2Y \leq v, X - 2Y \leq w] \dots \end{aligned}$$

... We *might* integrate & derive to find the density

... but we don't

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Functions of Random Variables

Solve:

Consider two jointly standard normal r.v.s X, Y

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$$f_{XY}(x, y) = \frac{1}{2\pi} \exp \left[-\frac{x^2 + y^2}{2} \right]$$

$$V = g(X, Y) = X + 2Y, \quad W = h(X, Y) = X - 2Y$$

Direct approach ...

$$f_{VW}(v, w) = f_{XY}(x(v, w), y(v, w)) \frac{\Delta(x, y)}{\Delta(v, w)}$$

$$\frac{\Delta(v, w)}{\Delta(x, y)} = \left| \det \begin{pmatrix} 1 & 2 \\ 1 & -2 \end{pmatrix} \right| = 4$$

$$\Rightarrow f_{VW}(v, w) = \frac{1}{4} f_{XY} \left(\frac{v+w}{2}, \frac{v-w}{4} \right) = \frac{1}{8\pi} \exp \left[-\frac{\left(\frac{v+w}{2} \right)^2 + \left(\frac{v-w}{4} \right)^2}{2} \right] = \frac{1}{8\pi} \exp \left[-\frac{5v^2 + 6vw + 5w^2}{32} \right]$$

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Compare: Normal with $\mu = -\frac{6w}{10}$, $\sigma^2 = \frac{16}{5}$

$\Rightarrow f_W(w) =$

$$\frac{1}{8\pi} \int_{-\infty}^{\infty} \exp\left[-\frac{5\left(v + \frac{6w}{10}\right)^2}{32}\right] dv \cdot \exp\left[-\frac{w^2}{10}\right]$$

Use the normalization
factor of a normal

$$= \frac{\sqrt{2\pi \frac{16}{5}}}{8\pi} \cdot \exp\left[-\frac{w^2}{10}\right]$$

$$= \frac{1}{\sqrt{2\pi \cdot 5}} \cdot \exp\left[-\frac{w^2}{2 \cdot 5}\right]$$

Functions of Random Variables

Solve:

Are V and W still independent?

Find the marginals of

$$f_{VW}(v, w) = \frac{1}{8\pi} \exp\left[-\frac{5v^2 + 6vw + 5w^2}{32}\right]$$

$$f_W(w) = \int_{-\infty}^{\infty} f_{VW}(v, w) dv = \frac{1}{8\pi} \int_{-\infty}^{\infty} \exp\left[-\frac{5v^2 + 6vw + 5w^2}{32}\right] dv$$

$$= \frac{1}{8\pi} \int_{-\infty}^{\infty} \exp\left[-\frac{5v^2 + 10v \frac{6w}{10} + 5w^2}{32}\right] dv$$

Complete the square

$$= \frac{1}{8\pi} \int_{-\infty}^{\infty} \exp\left[-\frac{5\left(v + \frac{6w}{10}\right)^2 - 5\left(\frac{6w}{10}\right)^2 + 5w^2}{32}\right] dv$$

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Functions of Random Variables

Solve:

Are V and W still independent?

Discuss:

Can we go in the reverse direction?

That is, can we use a transformation to obtain independent r.v.s?

How can we find such a transformation?

Techniques coming, soon!

$$f_{VW}(v, w) = \frac{1}{8\pi} \exp \left[-\frac{5v^2 + 6vw + 5w^2}{32} \right]$$

$$\Rightarrow f_W(w) = \frac{1}{\sqrt{2\pi \cdot 5}} \cdot \exp \left[-\frac{w^2}{2 \cdot 5} \right]$$

Similarly:

$$\Rightarrow f_V(v) = \frac{1}{\sqrt{2\pi \cdot 5}} \cdot \exp \left[-\frac{v^2}{2 \cdot 5} \right]$$

Hence,

$$f_V(v)f_W(w) = \frac{1}{10\pi} \cdot \exp \left[-\frac{v^2 + w^2}{10} \right]$$

Consequence: R.v.s V and W are **not** independent.

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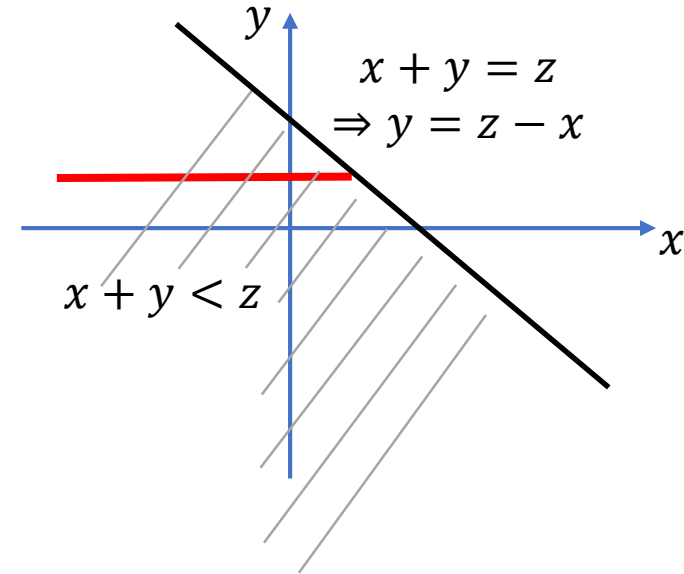
Functions of Random Variables

TASK 10 (re-visit an old example):

Consider continuous r.v.s X, Y ,

and $Z = g(X, Y) = X + Y$

Find $f_Z(z)$.



Alternative approach (dummy variable):

$$Z = X + Y, \quad U = X - Y \text{ (for instance)}$$

1) Find $f_{ZU}(z, u)$

$$f_{ZU}(z, u) = f_{XY}(x(z, u), y(z, u)) \frac{\Delta(x, y)}{\Delta(z, u)}$$

$$= f_{XY}\left(\frac{z+u}{2}, \frac{z-u}{2}\right) \cdot \frac{1}{2}$$

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Interesting Option:

Another change of variables: $(z, u) \mapsto (z, t)$

That is,

keep z , and use $t = \frac{z+u}{2}$ or $u = 2t - z$

$$\begin{aligned} f_{ZT}(z, t) &= f_{ZU}(z, u) \frac{\Delta(z, u)}{\Delta(z, t)} \\ &= f_{XY}(t, z - t) \end{aligned}$$

$$\frac{\Delta(z, u)}{\Delta(z, t)} = \left| \det \begin{pmatrix} 1 & 0 \\ -1 & 2 \end{pmatrix} \right| = 2$$

Can you simplify this derivation?

Functions of Random Variables

Solve:

Alternative approach (dummy variable):

$$Z = X + Y, \quad U = X - Y \text{ (for instance)}$$

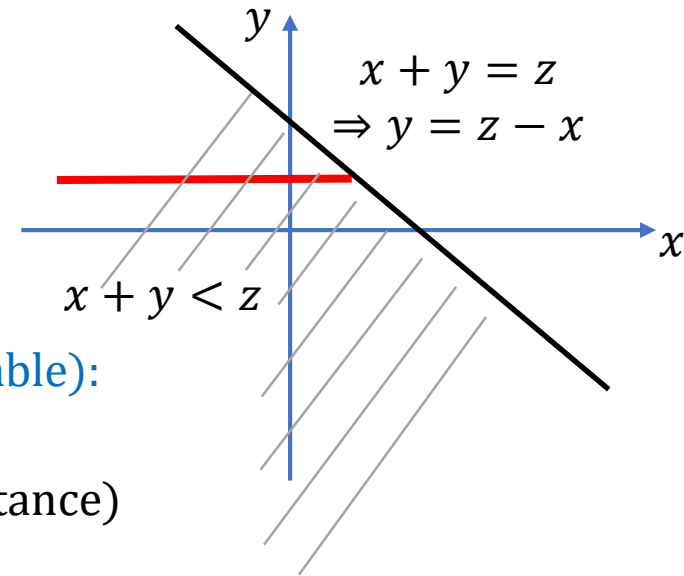
2) Marginalize to find $f_Z(z)$

$$f_{ZU}(z, u) = \frac{1}{2} f_{XY}\left(\frac{z+u}{2}, \frac{z-u}{2}\right)$$

$$\Rightarrow f_Z(z) = \int_{-\infty}^{\infty} f_{ZU}(z, u) du = \frac{1}{2} \int_{-\infty}^{\infty} f_{XY}\left(\frac{z+u}{2}, \frac{z-u}{2}\right) du$$

Hence, same thing:

$$f_Z(z) = \int_{-\infty}^{\infty} f_{XY}(t, z - t) dt$$



The End

Next time: Chp. 4