

EC508: Econometrics

Sampling Distribution of the OLS Estimator

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The Sampling Distribution of the OLS Estimator (SW 4.5)

- The OLS estimator is computed from a sample of data. A different sample yields a different value of $\hat{\beta}_1$. This is the source of the “sampling uncertainty” of $\hat{\beta}_1$. We want to:
 - quantify the sampling uncertainty associated with use $\hat{\beta}_1$ to test hypotheses such as $\beta_1 = 0$
 - construct a confidence interval for β_1
 - All these require figuring out the sampling distribution of the OLS estimator. Two steps to get there:
 - Probability framework for linear regression
 - Distribution of the OLS estimator

Probability Framework for Linear Regression

The probability framework for linear regression is summarized by the three least squares assumptions.

- **Population:**

The group of interest (ex: all possible school districts)

- **Random variables:** Y, X

Ex: (Test Score, STR)

- **Joint distribution of (Y, X) .** We assume:

- The population regression function is linear
- $\mathbb{E}(u|X) = 0$ (1st Least Squares Assumption)
- X, Y have nonzero finite fourth moments (3rd L.S.A.)

- **Data Collection by simple random sampling** implies:

$\{(X_i, Y_i)\}, i = 1, \dots, n$, are i.i.d. (2nd L.S.A.)

- Mean Independence: $\mathbb{E}(u_i|X_i) = 0$
- Implies $\text{cov}(u_i, X_i) = 0$: no correlation between X and u
- Implied by X and u independent + u mean zero

The Sampling Distribution of $\hat{\beta}_1$

- Like \bar{Y} , $\hat{\beta}_1$ has a sampling distribution.
- What is $\mathbb{E}(\hat{\beta}_1)$?
 - If $\mathbb{E}(\hat{\beta}_1) = \beta_1$, then OLS is unbiased – a good thing!
- What is $\text{var}(\hat{\beta}_1)$? (measure of sampling uncertainty)
 - We need to derive a formula so we can compute the standard error of $\hat{\beta}_1$.
- What is the distribution of $\hat{\beta}_1$ in small samples?
 - It is very complicated in general
- What is the distribution of $\hat{\beta}_1$ in large samples?
 - In large samples, $\hat{\beta}_1$ is normally distributed.

The mean and variance of the sampling distribution of $\hat{\beta}_1$

Some preliminary algebra:

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$

$$\bar{Y} = \beta_0 + \beta_1 \bar{X} + \bar{u}$$

$$\Rightarrow Y_i - \bar{Y} = \beta_1 [X_i - \bar{X}] + u_i - \bar{u}$$

This implies that:

$$\begin{aligned}\hat{\beta}_1 &= \frac{\sum_{i=1}^n (X_i - \bar{X}_n)(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X}_n)^2} \\ &= \frac{\sum_{i=1}^n \beta_1 (X_i - \bar{X}_n)^2}{\sum_{i=1}^n (X_i - \bar{X}_n)^2} + \frac{\sum_{i=1}^n (X_i - \bar{X}_n)(u_i - \bar{u})}{\sum_{i=1}^n (X_i - \bar{X}_n)^2}\end{aligned}$$

The mean and variance of the sampling distribution of $\hat{\beta}_1$

$$\begin{aligned}\hat{\beta}_1 &= \frac{\sum_{i=1}^n \beta_1 (X_i - \bar{X}_n)^2}{\sum_{i=1}^n (X_i - \bar{X}_n)^2} + \frac{\sum_{i=1}^n (X_i - \bar{X}_n)(u_i - \bar{u})}{\sum_{i=1}^n (X_i - \bar{X}_n)^2} \\&= \beta_1 + \frac{\sum_{i=1}^n (X_i - \bar{X}_n)(u_i - \bar{u})}{\sum_{i=1}^n (X_i - \bar{X}_n)^2} \\&= \beta_1 + \frac{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n) u_i}{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2}\end{aligned}$$

Now we can calculate $\mathbb{E}(\hat{\beta}_1)$ and $\text{var}(\hat{\beta}_1)$:

$$\begin{aligned}\mathbb{E}(\hat{\beta}_1) &= \beta_1 + \mathbb{E} \left(\frac{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n) u_i}{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2} \right) \\ &= \beta_1 + \mathbb{E} \left(\mathbb{E} \left(\frac{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n) u_i}{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2} \middle| X_1, \dots, X_n \right) \right) \\ &= \beta_1 + \mathbb{E} \left(\frac{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n) \mathbb{E}(u_i | X_1, \dots, X_n)}{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2} \right)\end{aligned}$$

LSA 2: (Y_i, X_i) iid implies $\mathbb{E}(u_i | X_1, \dots, X_n) = \mathbb{E}(u_i | X_i)$

LSA 1: $\mathbb{E}(u_i | X_i) = 0$

Together: $\mathbb{E}(\hat{\beta}_1) = \beta_1$, $\hat{\beta}_1$ is an **unbiased** estimator of β_1

Weak Law of Large Numbers

- Let Z_1, \dots, Z_n be iid with $\mathbb{E}(|Z_i|^2) < \infty$
- Then:

$$\bar{Z}_n = \frac{1}{n} \sum_{i=1}^n Z_i \xrightarrow{P} \mathbb{E}(Z_i)$$

- \xrightarrow{P} is the convergence in probability:

$$\mathbb{P}(|\bar{Z}_n - \mathbb{E}(Z_i)| > \varepsilon) \rightarrow 0, \text{ as } n \rightarrow \infty$$

- The WLLN can be proved using Chebyshev's inequality:

$$\mathbb{P}(|\bar{Z}_n - \mathbb{E}(Z_i)| > \varepsilon) \leq \frac{\mathbb{E}(|\bar{Z}_n - \mathbb{E}(Z_i)|^2)}{\varepsilon^2}$$

Consistency of OLS

$$\hat{\beta}_1 = \beta_1 + \frac{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n) u_i}{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2}$$

- Let $Z_i = X_i u_i$. Z has mean zero, finite variance: $\bar{Z}_n \xrightarrow{P} 0$
- $\bar{X}_n \bar{u}_n \xrightarrow{P} \mathbb{E}(X_i) \mathbb{E}(u_i) = 0$
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$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2 &= \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}_n^2 \\ &\xrightarrow{P} \mathbb{E}(X_i^2) - \mathbb{E}(X_i)^2 \\ &= \text{var}(X_i) > 0 \end{aligned}$$

- Together these imply:

$$\hat{\beta}_1 = \beta_1 + \frac{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n) u_i}{\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2} \xrightarrow{P} \beta_1 + 0$$