

# Model of Super Spreading Events

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## Model

Let  $x_t = y_t + z_t$

where

$y_t$  = Number of infecteds not by a super-spreading event (NSSE)

$z_t = x_t - y_t$  = Number of infecteds infected by a super-spreading event (SSE)

## Probability

$$P(x_t) = \sum_{y_t=0}^{x_t} P(y_t) \cdot P(z_t)$$

### i. $y_t$ - Non super-spreading events

$y_t \sim \text{Poisson}(\alpha \cdot \lambda_t)$ ,

$$p(y_t) = \exp(-\alpha \cdot \lambda_t) \cdot \frac{1}{y_t!} \cdot (\alpha \cdot \lambda_t)^{y_t},$$

with

$$\lambda_t = \sum_{i=1}^{t-1} x_i \left( \text{Gamma}((t-i); k, \theta) - \text{Gamma}((t-i-1); k, \theta) \right)$$

### ii. $z_t$ - super-spreading events

For each SSE event we get  $\text{Poisson}(\gamma)$  infections

$z_t | n_t \sim \text{Poisson}(\gamma \cdot n_t)$  = Number of infecteds by all Super-Spreading Events at time t

$n_t \sim \text{Poisson}(\beta \cdot \lambda_t)$  = Number of super-spreading events

$$p(z_t) = \sum_{n_t=0}^{\infty} p(n_t = n_t) \cdot p(z_t | n_t = n_t)$$

$\therefore p(z_t)$  is a Poisson-Poisson Compound

## Determining the Mean and Variance of $z_t$

$$\mathbb{E}(z_t) = \mathbb{E}(\mathbb{E}(z_t | n_t))$$

$$= \mathbb{E}(\gamma \cdot n_t)$$

$$= \gamma \mathbb{E}(n_t)$$

$$= \gamma \cdot \beta \cdot \lambda_t$$

$$\text{Var}(z_t) = \text{Var}(\mathbb{E}(z_t | n_t)) + \mathbb{E}(\text{Var}(z_t | n_t))$$

As  $z_t|n_t$  and  $n_t$  are poisson rvs the mean and variance are equivalent giving;

$$\begin{aligned} &= \text{Var}(\gamma \cdot n_t) + \mathbb{E}(\gamma \cdot n_t) \\ &= \gamma^2 \cdot \beta \cdot \lambda_t + \gamma \cdot \beta \cdot \lambda_t \end{aligned}$$

### Poisson-Poisson Compound $z_t$ in the form of a Negative Binomial RV

The mean and variance of the Poisson-Poisson Compound  $z_t$  is equated to that of a Negative Binomial distribution of size  $n$  with probability of success  $p$  and density given by;

$$\frac{\Gamma(x+n)}{\Gamma(n) \cdot x!} \cdot p^n (1-p)^x$$

for  $x = 0, 1, 2, \dots, n > 0$  and  $0 < p \leq 1$ .

This represents the number of failures which occur in a sequence of Bernoulli trials before a target number of successes is reached. The mean and variance are respectfully;

$$\mu = \frac{n(1-p)}{p}, \quad \text{Var} = \frac{n(1-p)}{p^2}$$

Equating the poisson-poisson compound to that of the negative binomial distribution gives;

$$\text{I. } \mu = \frac{n(1-p)}{p} = \gamma \cdot \beta \cdot \lambda_t$$

$$\text{II } \text{Var} = \frac{n(1-p)}{p^2} = \gamma^2 \cdot \beta \cdot \lambda_t + \gamma \cdot \beta \cdot \lambda_t$$

From **I** we can say that the size

$$n = \frac{p \cdot \gamma \cdot \beta \cdot \lambda_t}{1-p} \quad \text{III.}$$

Substituting **III** into **II** gives;

$$\frac{\left(\frac{p \cdot \gamma \cdot \beta \cdot \lambda_t}{1-p}\right) \cdot (1-p)}{p^2} = \gamma^2 \cdot \beta \cdot \lambda_t + \gamma \cdot \beta \cdot \lambda_t$$

Cancelling like terms gives;

$$\frac{\cancel{\left(\frac{p \cdot \gamma \cdot \beta \cdot \lambda_t}{1-p}\right)} \cdot \cancel{(1-p)}}{p^{\cancel{2}}} = \gamma^{\cancel{2}} \cdot \beta \cdot \cancel{\lambda_t} + \gamma \cdot \beta \cdot \cancel{\lambda_t}$$

$$\frac{1}{p} = \gamma + 1$$

$$1 = p \cdot \gamma + p$$

$$1 = p \cdot \gamma + p$$

$$\therefore p = \frac{1}{\gamma + 1} \quad \text{IV.}$$

Substituting **IV** into **III** gives;

$$n = \frac{\frac{1}{\gamma + 1} \cdot \gamma \cdot \beta \cdot \lambda_t}{1 - \frac{1}{\gamma + 1}}$$

$$n = \frac{\frac{\gamma \cdot \beta \cdot \lambda_t}{\gamma + 1}}{\frac{\gamma + 1 - 1}{\gamma + 1}}$$

$$n = \frac{\frac{\cancel{\gamma} \cdot \beta \cdot \lambda_t}{\cancel{\gamma} + 1}}{\frac{\cancel{\gamma}}{\cancel{\gamma} + 1}}$$

$$\therefore n = \beta \cdot \lambda_t$$

$z_t$  can thus be written as a Negative Binomial of size  $n$  with probability of success  $p$ , i.e

$$z_t \sim \text{NB}(\beta \cdot \lambda_t, \frac{1}{\gamma + 1})$$

$$p(z_t) = \frac{\Gamma(x + \beta \cdot \lambda_t)}{\Gamma(\beta \cdot \lambda_t) \cdot z_t!} \cdot \frac{1}{\gamma + 1}^{\beta \cdot \lambda_t} \cdot (1 - \frac{1}{\gamma + 1})^{z_t} \quad \text{for } z_t = 0, 1, 2, \dots, n > 0 \text{ and } 0 < p \leq 1$$

## Likelihood

$$P(x_t) = \sum_{y_t=0}^{x_t} P(y_t) \cdot P(z_t)$$

$$\begin{aligned} L(\alpha, \beta, \gamma, \lambda | \mathbf{x}_t) &= \prod_{t=1}^{Ndays} \sum_{y_t=0}^{x_t} P(y_t) \cdot P(z_t) \\ &= \prod_{t=1}^{Ndays} \sum_{y_t=0}^{x_t} \exp(-\alpha \cdot \lambda_t) \cdot \frac{1}{y_t!} \cdot (\alpha \cdot \lambda_t)^{y_t} \times \frac{\Gamma(z_t + \beta \cdot \lambda_t)}{\Gamma(\beta \cdot \lambda_t) \cdot z_t!} \cdot \frac{1}{\gamma + 1}^{\beta \cdot \lambda_t} \cdot (1 - \frac{1}{\gamma + 1})^{z_t} \\ &= \prod_{t=1}^{Ndays} \sum_{y_t=0}^{x_t} \exp(-\alpha \cdot \lambda_t) \cdot \frac{1}{y_t!} \cdot (\alpha \cdot \lambda_t)^{y_t} \times \frac{\Gamma(z_t + \beta \cdot \lambda_t)}{\Gamma(\beta \cdot \lambda_t) \cdot z_t!} \cdot \frac{1}{\gamma + 1}^{\beta \cdot \lambda_t} \cdot (\frac{\gamma}{\gamma + 1})^{z_t} \end{aligned}$$

## Log Likelihood

$$\begin{aligned} l(\alpha, \beta, \gamma, \lambda | \mathbf{x}_t) &= \ln \left( \prod_{t=1}^{Ndays} \sum_{y_t=0}^{x_t} P(y_t) \cdot P(z_t) \right) \\ &= \sum_{t=1}^{Ndays} \ln \left( \sum_{y_t=0}^{x_t} (\exp(-\alpha \cdot \lambda_t) \cdot \frac{1}{y_t!} \cdot (\alpha \cdot \lambda_t)^{y_t} \times \frac{\Gamma(z_t + \beta \cdot \lambda_t)}{\Gamma(\beta \cdot \lambda_t) \cdot z_t!} \cdot \frac{1}{\gamma + 1}^{\beta \cdot \lambda_t} \cdot (\frac{\gamma}{\gamma + 1})^{z_t}) \right) \end{aligned}$$

## Log exp sum implementation

**Simplest case**  $x_1, x_2$

Let  $L_i = \log(x_i)$

$$\begin{aligned}
 \log(x_1 + x_2) &= \log\left(\exp(L_1) + \exp(L_2)\right) \\
 &= \log\left((\exp(L_1) + \exp(L_2)) * \frac{\exp(M)}{\exp(M)}\right) \\
 &= \log(\exp(M)) + \log\left(\frac{\exp(L_1)}{\exp(M)} + \frac{\exp(L_2)}{\exp(M)}\right) \\
 &= M + \log\left(\exp(L_1 - M) + \exp(L_2 - M)\right)
 \end{aligned}$$

In general;

$$LSE(x_1, \dots, x_n) = x^* + \log\left(\exp(x_1 - x^*) + \dots + \exp(x_n - x^*)\right)$$

where

$$x^* = \max(x_1, \dots, x_n)$$

For this super-spreading events model  $x_i$  is underlined in the likelihood;

$$l(\alpha, \beta, \gamma, \lambda | \mathbf{x}_t) = \sum_{t=1}^{Ndays} \ln \left( \sum_{y_t=0}^{x_t} \frac{\exp(-\alpha \cdot \lambda_t) \cdot \frac{1}{y_t!} \cdot (\alpha \cdot \lambda_t)^{y_t} \times \frac{\Gamma(z_t + \beta \cdot \lambda_t)}{\Gamma(\beta \cdot \lambda_t) \cdot z_t!} \cdot \frac{1}{\gamma + 1}^{\beta \cdot \lambda_t} \cdot \left(\frac{\gamma}{\gamma + 1}\right)^{z_t}} \right)$$

$$x_i = \exp(-\alpha \cdot \lambda_t) \cdot \frac{1}{y_t!} \cdot (\alpha \cdot \lambda_t)^{y_t} \times \frac{\Gamma(z_t + \beta \cdot \lambda_t)}{\Gamma(\beta \cdot \lambda_t) \cdot z_t!} \cdot \frac{1}{\gamma + 1}^{\beta \cdot \lambda_t} \cdot \left(\frac{\gamma}{\gamma + 1}\right)^{z_t} \text{ for a given } y_t$$

and

$$\begin{aligned}
 L_i &= \log(x_i) \\
 &= -(\alpha \cdot \lambda_t) - \text{lfactorial}(y_t) + y_t \cdot \log(\alpha \cdot \lambda_t) + \text{lgamma}(z_t + \beta \cdot \lambda_t) - \text{lgamma}(\beta \cdot \lambda_t) \\
 &\quad - \text{lfactorial}(z_t) - \beta \cdot \lambda_t \cdot \log(\gamma + 1) + z_t \cdot \log(\gamma) - z_t \cdot \log(\gamma + 1)
 \end{aligned}$$

Replacing  $z_t$  with  $x_t - y_t$  gives;

$$\begin{aligned}
 L_i &= -(\alpha \cdot \lambda_t) - \text{lfactorial}(y_t) + y_t \cdot \log(\alpha \cdot \lambda_t) + \text{lgamma}((x_t - y_t) + \beta \cdot \lambda_t) - \text{lgamma}(\beta \cdot \lambda_t) \\
 &\quad - \text{lfactorial}(x_t - y_t) - \beta \cdot \lambda_t \cdot \log(\gamma + 1) + (x_t - y_t) \cdot \log(\gamma) - (x_t - y_t) \cdot \log(\gamma + 1)
 \end{aligned}$$