

# Model of Super Spreading Events

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## Model

Let  $x_t = y_t + z_t$

where

$y_t$  = Number of infecteds not by a super-spreading event (NSSE)

$z_t = x_t - y_t$  = Number of infecteds infected by a super-spreading event (SSE)

## Probability

$$P(x_t) = \sum_{y_t=0}^{x_t} P_{NSSE}(y_t) \otimes P_{SSE}(z_t)$$

where

$$i. P_{NSSE}(y_t) = \exp(-\lambda_t) \frac{1}{y_t!} \lambda_t^{y_t},$$

with

$$\lambda_t = r0_1 \sum_{i=1}^{t-1} y_i ( \text{Gamma}((t-i); k, \theta) - \text{Gamma}((t-i-1); k, \theta)$$

$$ii. P_{SSE}(z_t) = e^{-\lambda'_t} \frac{1}{z_t!} \lambda'_t^{z_t},$$

with  $\lambda_t \sim \text{Poisson}()$

## Super-spreading events

For each super-spreading event there are;

-  $\text{Poisson}(\mu)$  infecteds (For a given rate  $\mu$ , e.g = 5)

-  $n_t$  = Number of super-spreading events

$z_t$

$z_t \sim \text{Poisson}(n_t \mu)$  = Number of infecteds by all Super-Spreading Events at time t

$n_t \sim \text{Poisson}(\lambda_t \cdot k)$  = Number of infecteds by a super-spreading person at time t

where

$k$  = some multiplicative factor  $> 1$

Probability( $z_t$ )

$$P(z_t) = \sum_{n_t=0}^{\infty} p(n_t = n_t) \times p(z_t | n_t)$$

$$P(z_t) = \sum_{n_t=0}^{\infty} \frac{1}{n_t!} e^{-\lambda_t k} (\lambda_t k)^{n_t} \times \frac{1}{z_t!} e^{-n_t \mu} (n_t \mu)^{z_t}$$

$$P(z_t) = \sum_{n_t=0}^{\infty} \frac{1}{n_t!} (\lambda_t k)^{n_t} e^{-n_t \mu} (n_t)^{z_t} \times \frac{1}{z_t!} e^{-\lambda_t k} \mu^{z_t}$$

## Other - Likelihood

$$L(\lambda_1, \lambda_2, p | \mathbf{x}) = \prod_{t=1}^{Ndays} p \cdot \left[ \exp(-\lambda_{1t}) \frac{1}{x_t!} \lambda_{1t}^{x_t} \right] + (1-p) \cdot \left[ \exp(-\lambda_{2t}) \frac{1}{x_t!} \lambda_{2t}^{x_t} \right]$$

where

$$\lambda_{1t} = r0_1 \sum_{i=1}^{t-1} x_i \left( \text{Gamma}((t-i); k_1, \theta_1) - \text{Gamma}((t-i-1); k_1, \theta_1) \right)$$

$$\lambda_{2t} = r0_2 \sum_{i=1}^{t-1} x_i \left( \text{Gamma}((t-i); k_2, \theta_2) - \text{Gamma}((t-i-1); k_2, \theta_2) \right)$$

## Log Likelihood

$$l(\lambda_1, \lambda_2, p | \mathbf{x}) = \prod_{t=1}^{Ndays} \log \left( p \cdot \left[ \exp(-\lambda_{1t}) \frac{1}{x_t!} \lambda_{1t}^{x_t} \right] \right) + \log \left( (1-p) \cdot \left[ \exp(-\lambda_{2t}) \frac{1}{x_t!} \lambda_{2t}^{x_t} \right] \right)$$

$$l(\lambda_1, \lambda_2, p | \mathbf{x}) = \sum_{t=1}^{Ndays} \log(p) + \log(\exp(-\lambda_{1t})) + \log\left(\frac{1}{x_t!}\right) + \log(\lambda_{1t}^{x_t}) + \log(1-p) + \log(\exp(-\lambda_{2t})) + \log\left(\frac{1}{x_t!}\right) + \log(\lambda_{2t}^{x_t})$$

$$l(\lambda_1, \lambda_2, p | \mathbf{x}) = \sum_{t=1}^{Ndays} \log(p) - \lambda_{1t} + \log\left(\frac{1}{x_t!}\right) + \log(\lambda_{1t}^{x_t}) + \log(1-p) - \lambda_{2t} + \log\left(\frac{1}{x_t!}\right) + \log(\lambda_{2t}^{x_t})$$

Terms that do not contain one of the model parameters are removed giving;

$$l(\lambda_1, \lambda_2, p | \mathbf{x}) = \sum_{t=1}^{Ndays} \log(p) - \lambda_{1t} + \log(\lambda_{1t}^{x_t}) + \log(1-p) - \lambda_{2t} + \log(\lambda_{2t}^{x_t})$$

## Bayesian Model

### Likelihood

$$L(\lambda_1, \lambda_2, p | \mathbf{x}) = \prod_{t=1}^{Ndays} p \cdot \left[ \exp(-\lambda_{1t}) \frac{1}{x_t!} \lambda_{1t}^{x_t} \right] + (1-p) \cdot \left[ \exp(-\lambda_{2t}) \frac{1}{x_t!} \lambda_{2t}^{x_t} \right]$$

where

$$\lambda_{1t} = r0_1 \sum_{i=1}^{t-1} x_i \left( \text{Gamma}((t-i); k_1, \theta_1) - \text{Gamma}((t-i-1); k_1, \theta_1) \right)$$

$$\lambda_{2t} = r0_2 \sum_{i=1}^{t-1} x_i \left( \text{Gamma}((t-i); k_2, \theta_2) - \text{Gamma}((t-i-1); k_2, \theta_2) \right)$$

### Priors

$$\text{prior}(r0_1) = \text{Gamma}(\alpha_1, \beta_1) = \frac{\beta_1^{\alpha_1}}{\Gamma(\alpha_1)} r0^{\alpha_1-1} e^{-\beta_1 r0}$$

$$\text{prior}(r0_2) = \text{Gamma}(\alpha_2, \beta_2) = \frac{\beta_2^{\alpha_2}}{\Gamma(\alpha_2)} r0^{\alpha_2-1} e^{-\beta_2 r0}$$

$$\text{prior}(p) = \text{Beta}(\alpha_3, \beta_3) = \frac{1}{\alpha_3 \beta_3} \cdot p^{\alpha_3-1} (1-p)^{\beta_3-1}$$

## Draft - Likelihood

$$L(r0_1, k_1, \theta_1, r0_2, k_2, \theta_2, p | \mathbf{x}) = \prod_{t=1}^{Ndays} p \cdot [\exp(-\lambda_{1t}) \frac{1}{x_t!} \lambda_{1t}^{x_t}] + (1-p) \cdot [\exp(-\lambda_{2t}) \frac{1}{x_t!} \lambda_{2t}^{x_t}]$$

# Draft

## Bayesian Epidemic Modelling

### Mathematics

Likelihood

$$L(r_0, k, \theta | \mathbf{x}) = \prod_{t=1}^{Ndays} \exp(-\lambda_t) \frac{1}{x_t!} \lambda_t^{x_t}$$

where

$$\lambda_t = r_0 \sum_{i=1}^{t-1} x_i ( \text{Gamma}((t-i); k, \theta) - \text{Gamma}((t-i-1); k, \theta)$$

Log likelihood

$$l = \sum_{t=1}^{Ndays} x_t \ln(\lambda_t) - \lambda_t$$

$$l = \sum_{t=1}^{Ndays} x_t \ln(r_0 \lambda_t) - r_0 \lambda_t$$

$$l = \sum_{t=1}^{Ndays} (x_t \ln(r_0) + x_t \ln(\lambda_t) - r_0 \lambda_t)$$

where

$$\lambda_t = \sum_{i=1}^{t-1} x_i ( \text{Gamma}((t-i); k, \theta) - \text{Gamma}((t-i-1); k, \theta)$$

$$l = \sum_{t=1}^{Ndays} (x_t \ln(r_0 \lambda_t) - r_0 \lambda_t)$$

### MLE

$$\frac{dl}{dr_0} = \sum_{t=1}^{Ndays} \frac{x_t}{r_0} - \sum_{t=1}^{Ndays} \lambda_t$$

$$\hat{r_0} = \frac{\sum_{t=1}^{Ndays} x_t}{\sum_{t=1}^{Ndays} \lambda_t}$$

### Bayesian Inference

Likelihood

$$L(r_0, k, \theta | \mathbf{x}) = \prod_{t=1}^{Ndays} \exp(-\lambda_t) \frac{1}{x_t!} \lambda_t^{x_t}$$

where

$$\lambda_t = r0 \sum_{i=1}^{t-1} x_i ( \text{Gamma}((t-i); k, \theta) - \text{Gamma}((t-i-1); k, \theta)$$

Prior

$$p(r0) = \text{Gamma}(\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} r0^{\alpha-1} e^{-\beta r0}$$

### Bayesian Inference

- Shape, scale\*\* To write out  
Posterior  $p(r0|x)$

Given  $\text{Gamma}(1, 1)$  prior on  $r0$ ;

$$\begin{aligned} p(r0|x) &\propto \prod_{t=1}^{Ndays} \exp(-r0\lambda_t) \frac{1}{x_t!} (r0\lambda_t)^{x_t} \times r0^{\alpha-1} e^{-\beta r0} \\ &\propto \exp(-r0(\sum_{t=1}^{Ndays} \lambda_t + 1)) \times (r0\lambda_t)^{\sum_{t=1}^{Ndays} x_t} \\ \therefore p(r0|x) &\propto \text{Gamma}(\sum_{t=1}^{Ndays} x_t + 1, \sum_{t=1}^{Ndays} \lambda_t + 1) \\ &\propto \text{Gamma}(\sum_{t=1}^{Ndays} x_t + \alpha, \sum_{t=1}^{Ndays} \lambda_t + \beta) \end{aligned}$$

**where** (\*Above is shape, rate. Below is shape, scale)

$$\lambda_t = \sum_{i=1}^{t-1} x_i ( \text{Gamma}((t-i); k, \theta) - \text{Gamma}((t-i-1); k, \theta)$$

### Bayesian Inference v0

Posterior  $p(r0|x)$

Given  $\text{Gamma}(1, 1)$  prior on  $r0$ ;

$$\begin{aligned} p(r0|x) &\propto \prod_{t=1}^{Ndays} \exp(-r0\lambda_t) \frac{1}{x_t!} (r0\lambda_t)^{x_t} \times r0^{\alpha-1} e^{-\beta r0} \\ &\propto \exp(-r0(\sum_{t=1}^{Ndays} \lambda_t + 1)) \times (r0\lambda_t)^{\sum_{t=1}^{Ndays} x_t} \\ \therefore p(r0|x) &\propto \text{Gamma}(\sum_{t=1}^{Ndays} x_t + 1, \sum_{t=1}^{Ndays} \lambda_t + 1) \\ &\propto \text{Gamma}(\sum_{t=1}^{Ndays} x_t + \alpha, \sum_{t=1}^{Ndays} \lambda_t + \beta) \end{aligned}$$

**where** (\*Above is shape, rate. Below is shape, scale)

$$\lambda_t = \sum_{i=1}^{t-1} x_i ( \text{Gamma}((t-i); k, \theta) - \text{Gamma}((t-i-1); k, \theta)$$

Prior

$$p(r_0) = \text{Gamma}(\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} r_0^{\alpha-1} e^{-\beta r_0}$$

## Explanation of Mathematics

- Simulation of Epidemic - Infectiousness (Discrete gamma) - I.e 'Infectiousness Pressure' = Sum of all people - Explanation: Gamma is a continuous function so integrate over the density at that point in time (today - previous day) - Assumption: Number of daily cases follows a Poisson distribution. Reason why there is spikes from day to day

- Metropolis Hastings step - The logarithm of the acceptance probability  $\lambda$  includes; the sum of the log likelihood, prior and the proposal for the current time Y and the previous time step  $r_0[t-1]$

- Symmetrical proposal distribution so the proposal for the current time Y and the previous time step  $r_0[t-1]$  cancel out

## Assumptions

- Number of daily cases follows a Poisson distribution - Poisson-Gamma conjugacy leads to nice posterior  
- although not realistic

\* Assume the following Poisson model of two regimes for n random variables