

**Problem Set 5**

**Issued:** Monday 30<sup>th</sup> November, 2020

**Due:** Monday 14<sup>th</sup> December, 2020

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**Notations:** We use  $\mathbf{x}, \mathbf{y}, \mathbf{w}$  and  $\underline{\mathbf{x}}, \underline{\mathbf{y}}, \underline{\mathbf{w}}$  to denote random variables and random vectors.

5.1. *Cramer-Rao inequality with a bias term.* Let  $\mathbf{y} \sim f(\mathbf{y}; x)$  and let  $\hat{x}(\mathbf{y})$  be an estimator for  $x$ . Let  $b(x) = \mathbb{E}[\hat{x}(\mathbf{y})] - x$  be the bias of the estimator. Show that

$$\mathbb{E}[(\hat{x}(\mathbf{y}) - x)^2] \geq \frac{[1 + b'(x)]^2}{J_{\mathbf{y}}(x)} + b^2(x)$$

5.2. (a) Let

$$p_{\mathbf{y}}(y; x) = \begin{cases} x & \text{if } 0 \leq y \leq 1/x \\ 0 & \text{otherwise} \end{cases}$$

for  $x > 0$ . Show that there exist no unbiased estimators  $\hat{x}(\mathbf{y})$  for  $x$ . (Note that because only  $x > 0$  are possible values, an unbiased estimator need only be unbiased for  $x > 0$  rather than all  $x$ .)

(b) Suppose instead that

$$p_{\mathbf{y}}(y; x) = \begin{cases} 1/x & \text{if } 0 \leq y \leq x \\ 0 & \text{otherwise} \end{cases}$$

for  $x > 0$ . Does a minimum-variance unbiased estimator for  $x$  based on  $\mathbf{y}$  exist? If your answer is yes, determine  $\hat{x}_{\text{MVU}}(\mathbf{y})$ . If your answer is no, explain.

5.3. Suppose, for  $i = 1, 2$

$$\mathbf{y}_i = x + \mathbf{w}_i$$

where  $x$  is an unknown but non-zero constant,  $\mathbf{w}_1$  and  $\mathbf{w}_2$  are statistically independent, zero-mean Gaussian random variables with

$$\begin{aligned} \text{var}(\mathbf{w}_1) &= 1 \\ \text{var}(\mathbf{w}_2) &= \begin{cases} 1 & x > 0 \\ 2 & x < 0 \end{cases}. \end{aligned}$$

(a) Calculate the Cramér-Rao bound for unbiased estimators of  $x$  based on observation of

$$\underline{\mathbf{y}} = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix}.$$

(b) Show that a minimum variance unbiased estimator  $\hat{x}_{\text{MVU}}(\underline{\mathbf{y}})$  does not exist.

*Hint:* Consider the estimators

$$\begin{aligned} \hat{x}_1(\underline{\mathbf{y}}) &= \frac{1}{2}y_1 + \frac{1}{2}y_2, \\ \hat{x}_2(\underline{\mathbf{y}}) &= \frac{2}{3}y_1 + \frac{1}{3}y_2. \end{aligned}$$

5.4. Let  $\underline{y} = [y_1 \ y_2]^T$  be a vector random variable whose components are i.i.d. Bernoulli random variables with parameter  $x$ ,  $0 < x < 1$ , i.e.,  $\mathbb{P}(y_i = 1) = x, i = 1, 2$ .

- (a) Show that  $t(\underline{y}) = y_1 + y_2$  is a sufficient statistic.
- (b) Let  $\hat{x}(\underline{y}) = y_1$  be an estimator of the parameter  $x$  from the observation  $\underline{y}$ . Find  $\text{MSE}_{\hat{x}}(x)$ , the mean-square error of this estimator.
- (c) Let  $\hat{x}'(t) = \mathbb{E}[\hat{x}(\underline{y}) | \underline{t} = t]$  be an estimator of the parameter  $x$  that uses the sufficient statistic  $t$  instead of the observations  $\underline{y}$ .
  - i. Show that  $\hat{x}'(t)$  is a valid estimator, i.e., it does not depend on  $x$ .
  - ii. Show that  $\text{MSE}_{\hat{x}'}(x) = \gamma \text{MSE}_{\hat{x}}(x)$  and find the constant  $\gamma$ .
- (d) We now consider a generalization of this problem. Let  $\underline{y}$  be a random variable generated by a distribution  $p_{\underline{y}}(\cdot; x)$  and  $\underline{t}(\underline{y})$  be a sufficient statistic. Let  $\hat{x}(\underline{y})$  be an estimator of the parameter  $x$  based on the observation  $\underline{y}$ . We define an alternate estimator  $\hat{x}'(\underline{t}) = \mathbb{E}[\hat{x}(\underline{y}) | \underline{t} = \underline{t}]$ .
  - i. Show that  $\hat{x}'(\underline{t})$  is a valid estimator, i.e., it does not depend on  $x$ .
  - ii. Show that for any cost function  $C(x, \hat{x})$  that is convex in  $\hat{x}$ , the following inequality holds:

$$\mathbb{E}[C(x, \hat{x}'(\underline{t}))] \leq \mathbb{E}[C(x, \hat{x}(\underline{y}))].$$

5.5. For a non-bayesian case  $p_{\underline{y}}(y; x)$ , we do a binary hypothesis testing where  $x \in \{H_0, H_1\}$ . Please prove that  $t(y) = \frac{p_{\underline{y}}(y; H_1)}{p_{\underline{y}}(y; H_0)}$  is a complete sufficient statistics.

5.6. In class we developed the EM algorithm for maximum likelihood estimation (EM-ML). That is, we gave an iterative procedure to compute

$$\hat{x}_{ML}(y) = \arg \max_a p_{\underline{y}}(y; a).$$

and showed that the likelihood was non-decreasing with each iteration.

Please develop the EM-MAP algorithm for MAP estimation:

$$\hat{x}_{MAP}(y) = \arg \max_a p_{\underline{x}|\underline{y}}(a|y)$$

where the complete data  $\underline{z}$  is an arbitrary random vector. (Please follow the procedures in the lecture note)