

# Exercise 7 & 8

## Probability Theory 2020 Autumn

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### 1 Problem 1

Denote the entries in  $U$  as  $u_{ij}$  and entries in  $Y$  as  $Y_{ij}$ , so

$$Y_{ij} = \sum_{r,s} u_{ri} u_{sj} X_{rs} \quad (1)$$

Also we have,

$$\text{Cov}(X_{ij}, X_{mn}) = \begin{cases} 2, & i = j = m = n \\ 1, & (i, j) = (m, n) \text{ or } (i, j) = (n, m), i \neq j \\ 0, & \text{otherwise} \end{cases} \quad (2)$$

Thus

$$\begin{aligned} \text{Cov}(Y_{ij}, Y_{mn}) &= \text{Cov} \left( \sum_{r,s} u_{ri} u_{sj} X_{rs}, \sum_{p,q} u_{pm} u_{qn} X_{pq} \right) \\ &= 2 \sum_{r=1}^n u_{ri} u_{rj} u_{rm} u_{rn} + \sum_{r \neq s} u_{ri} u_{sj} u_{rm} u_{sn} + \sum_{r \neq s} u_{ri} u_{sj} u_{rn} u_{sm} \\ &= \sum_{r,s} [u_{ri} u_{sj} u_{rm} u_{sn} + u_{ri} u_{sj} u_{rn} u_{sm}] \\ &= \left( \sum_r u_{ri} u_{rm} \right) \left( \sum_s u_{sj} u_{sn} \right) + \left( \sum_r u_{ri} u_{rn} \right) \left( \sum_s u_{sj} u_{sm} \right) \end{aligned} \quad (3)$$

Denote the column vectors in  $U$  as  $\mathbf{u}_i, i = 1, 2, 3, \dots, N$ , so  $\mathbf{u}_i \cdot \mathbf{u}_j = \sum_r u_{ri} u_{rj} = \delta_{ij}$ .

$$\text{Cov}(Y_{ij}, Y_{mn}) = \delta_{im} \delta_{jn} + \delta_{jm} \delta_{in} = \begin{cases} 2, & i = j = m = n \\ 1, & (i, j) = (m, n) \text{ or } (i, j) = (n, m), i \neq j \\ 0, & \text{otherwise} \end{cases} \quad (4)$$

Because  $X_{ij}, j \geq i$  are independent Gaussian variables, so the joint distribution of  $Y_{ij}$  is joint Gaussian distribution, which means,

$$\text{Cov}(Y_{ij}, Y_{mn}) = 0 \iff Y_{ij}, Y_{mn} \text{ are independent} \quad (5)$$

So all the entries on and above the diagonal of  $Y$  are independent, and  $Y_{ii} \sim N(0, 2), i = 1, 2, 3, \dots, N$  and  $Y_{ij} \sim N(0, 1), 1 \leq i < j \leq n$ . (It is easy to see that  $\mathbb{E}[Y_{ij}] = 0$ )

## 2 Problem 2

Notice that

1. If  $X_n = 1$ , then  $X_{n+1}, X_{n+2}, \dots$  are independent of  $X_1, X_2, \dots, X_n$
2. There is at least one 1 in any five-in-a-row  $X_i$ s as  $\{X_n, X_{n+1}, \dots, X_{n+4}\}$

So we can split  $X_1, X_2, \dots, X_n$  into a series of epsisodes, each episode  $L_j = [0, \dots, 0, 1]$  is consisted of  $n$  zeros ( $n$  can be  $0, 1, 2, 3, 4$ ) and 1 one. And  $L_j, j = 1, 2, \dots, m$  are independent. (For the last episode, if it is ended with 0, we can append 1 to its end and let  $n = n + 1$ .) Denote the length of each episode as  $l_j$ , so  $\sum_{j=1}^m l_j = n$ .

Consider the distribution of  $l_j$ , it can only take values in  $1, 2, 3, 4, 5$ ,

- $P(l_j = 1) = P(X_1 = 1) = 0.2$
- $P(l_j = 2) = P(X_1 = 0, X_2 = 1) = 0.16$
- $P(l_j = 3) = P(X_1 = 0, X_2 = 0, X_3 = 1) = 0.128$
- $P(l_j = 4) = P(X_1 = 0, X_2 = 0, X_3 = 0, X_4 = 1) = 0.1024$
- $P(l_j = 5) = P(X_1 = 0, X_2 = 0, X_3 = 0, X_4 = 0) = 0.4096$

So

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = \lim_{m \rightarrow \infty} \frac{m}{l_1 + l_2 + \dots + l_m} = \lim_{m \rightarrow \infty} \frac{1}{\frac{1}{m} \sum_{j=1}^m l_j} \quad (6)$$

According to Strong Law of Large Numbers,

$$\frac{1}{m} \sum_{j=1}^m l_j \xrightarrow{a.s.} E[l_j] = 3.3616 \quad (7)$$

So

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} \xrightarrow{a.s.} \frac{1}{3.3616} \quad (8)$$

## 3 Problem 3

### 3.1 (i)

Suppose the corresponding  $k$  of  $X_n$  is  $k_n$ , i.e.  $\sum_{i=1}^{k_n} Y_i = X_n + n$ . If  $X_n \geq 1$ ,  $\sum_{i=1}^{k_n} Y_i \geq n + 1$ , so  $k_{n+1} = k_n, X_{n+1} = X_n - 1$ . If  $X_n = 0$ ,  $\sum_{i=1}^{k_n} Y_i = n, \sum_{i=1}^{k_n+1} Y_i = n + Y_{n+1} \geq n + 1$ , so  $k_{n+1} = k_n, X_{n+1} = Y_{n+1} - 1$ .

So given  $X_n$ ,  $X_{n+1}$  is independent of  $X_{n-1}, \dots, X_1$ .  $\{X_n\}_{n=1}^\infty$  forms a Markov Chain. And the transition probability is,

$$P(X_{n+1} = i | X_n = 0) = p_{i+1}, i = 0, 1, \dots \quad (9)$$

$$P(X_{n+1} = i | X_n = j, j \geq 1) = \begin{cases} 1, & i = j - 1 \\ 0, & \text{otherwise} \end{cases} \quad (10)$$

### 3.2 (ii)

Notice that  $f(n) = P(X_n = 0)$ , so  $\lim_{n \rightarrow \infty} f(n) = \lim_{n \rightarrow \infty} P(X_n = 0)$ . If we want  $\lim_{n \rightarrow \infty} f(n)$  exists, the Markov chain must be irreducible, aperiodic and positive recurrent.

It is irreducible obviously. Consider the support set  $\mathcal{Y} = \{i : p_i > 0\}$  of  $Y$ , if  $\inf \mathcal{Y} = N < \infty$ , the state space  $\mathcal{S}$  of the Markov Chain is finite  $\{0, 1, \dots, N\}$ . Obviously  $N$  can be reached from 0. And because  $N-1, N-2, \dots, 0$  can be reached from  $N$ , so it is irreducible. If  $\inf \mathcal{Y} = \infty$ , for any state  $n$ , there exists a state  $m > n$ , and  $m$  can be reached from 0, so  $n$  can be reached from 0. In that case, the Markov chain is also irreducible.

For it to be aperiodic, if it comes from 0 to  $i$ , it will return to 0 in  $i$  steps. So if  $\mathcal{Y} = \{i : p_i > 0\}$  is like  $\{2, 4, \dots, 2k, \dots\}$  or  $\{3, 6, 9, \dots, 3k, \dots\}$ , for certain steps it will not arrive at 0. So the Markov chain is aperiodic if and only if  $\gcd(\mathcal{Y}) = 1$

And it is positive recurrent if and only if  $\mathbb{E}[T_0] < \infty$ . It is easy to see that  $P(T_0 = i+1) = p_i, i \geq 1$ , so

$$\mathbb{E}[T_0] = \sum_{i=1}^{\infty} i p_i = \mathbb{E}[Y_1] \quad (11)$$

So the necessary and sufficient condition for  $\lim_{n \rightarrow \infty} f(n)$  to exist is  $\gcd(\{i+1 : p_i > 0\}) = 1$  and  $\sum_{i=1}^{\infty} i p_i < \infty$

### 3.3 (iii)

The limit equals to the steady-state probability,

$$\pi_0 = \lim_{n \rightarrow \infty} f(n) = \frac{1}{\mathbb{E}[T_0]} = \frac{1}{\mu} \quad (12)$$

## 4 Problem 4