

# Survey Data Analysis Final Report

*Gerbrich Ferdinands*

*Sanne Meijering*

*Hanne Oberman*

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## Abstract

In this technical report, a single wave of the multi-year panel study ‘Understanding Society’ is analyzed. More specifically, this research project focussed on the use of design and adjustment weights for estimating population quantities from a subset of the ‘Understanding Society’ data: the ‘Innovation Panel’. The data was analyzed using the statistical software *R*, and the ‘survey’ package.

*Key terms:* survey, design weights, adjustment weights, nonresponse.

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## Introduction

To estimate population quantities from sample data, broadly two strategies can be employed: model-based inference and design-based inference (Lohr, 2010). In this technical report, the effect of the latter on population estimates is investigated, by means of analyzing a single wave of Great Britain’s multi-year panel study ‘Understanding Society’. This annual household panel survey is used to assess topics like housing, education, health, and employment among the general population of the United Kingdom. More specifically, we focus on the a subset of first wave, conducted between 2008 and 2009: the ‘Innovation Panel’ (IP).

The purpose of the IP was to enable methodological research, before initiating the main ‘Understanding Society’ survey. As a forerunner, the target sample size of the IP was 1500 households across Great Britain, whereas the intended sample size for the main survey was 40,000 households within the United Kingdom (Boreham, & Constantine, 2008). Therefore, the sampling frame of the IP, The Postcode Address File, encompassing England, Wales and Scotland. Other regions of the United Kingdom like Northern Ireland were not part of the target population.

To obtain the sample from the sampling frame, multistage stratified cluster sampling was used. Postal sectors acted as primary sampling units (PSUs) in the cluster design, and were sorted by Government Office Region. Within these regions, strata were constructed based on a.o. non-manual occupations, ethnic minority density, and population density. After stratification, a sample was drawn from the sampling frame (the Postcode Address File), with probabilities proportional to the number of postal delivery points within each postcode sector (the stratified PSUs). Per postcode sector, 23 addresses were selected, yielding a total of

2760 addresses. However, interviews were only achieved in 1489 households (Boreham, & Constantine, 2008). This discrepancy – the sampling, non-sampling, and nonresponse errors – is one of the main topics that is investigated in the current research project.

The aim of this research project is to investigate several properties of the IP data, with respect to the sampling procedure, the use of weights in survey data analysis, and to estimate population quantities from the sample. The report is based on eight research questions (RQs) that were provided in the research proposal (Appendix I: Research Proposal). The RQs are structured according to three topics; Part I: Sampling, Part II: Weighting, and Part III: Estimation. In the sampling part, we investigate how within-household selection of respondents was performed (RQ1), and how many people within the sample personally conducted an interview (RQ6). The part concerning survey weights encompasses a theoretical explanation for the variation between design weights across observations (RQ2), computation of enumerated weights (RQ3), and an effort to suggest variables that could be used to construct nonresponse weights (RQ8). Finally, in the third part population quantities are estimated: the relationship between age and household composition (RQ5), and the proportion of employed people of working age in the population with and without accounting for nonresponse (RQ4 and RQ7, respectively). The exact formulation of the research questions can be found in Appendix I: Research Proposal.

## Methods

The data in this research project was obtained by ‘The National Centre for Social Research’ from the United Kingdom (Boreham, & Constantine, 2008, p. 3), and retrieved via Dr. Peter Lugtig (Utrecht University). Data analysis is performed using the statistics freeware R (R Core Team, 2003). Preprocessing of the data consisted of removing variables irrelevant to the current study from the dataframe and subtracting 6 years from the variable ‘a\_dvage’ so it correctly reflects the age of each person at the time of the first wave rather than the average age over all waves (see Appendix II: R Script). Additional preprocessing was done for RQ8 specifically. The details of this are described under this question’s section.

To obtain accurate estimates from survey data with a multistage sampling design, regular SRS estimators do not suffice (Lohr, 2010). Therefore, the R package ‘survey’ (Lumley, 2018) is used. By specifying the survey design and appropriate weights, estimates of population quantities and their variances are adjusted for the complex nature of the survey data. The survey package functions ‘svydesign’, ‘svymean’, and ‘svychisq’ are employed to estimate population quantities. To investigate the relationship between age and household\_type, the survey package function ‘svytable’ was used.

Moreover, stepwise logistic regression is performed to investigate which variables might be used to construct nonresponse weights. This exploratory analysis is conducted using the R package ‘MASS’ (Venables, & Ripley, 2002).

## Results

### Part I: Sampling

#### Within-household selection of respondents (RQ1)

For the Understanding Society study, up to three households per dwelling and up to three people per household were selected for interviews. If more than three households were found at a single dwelling, or if a single household contained more than three people, the Kish Grid method was used to select households or people at random.

This method considered to lead to a random sample of household members, and avoids selection bias in the survey. When using the Kish Grid method, the researcher has to list all eligible household members and assign each of them a number  $i$ . The Kish Grid then helps you pick a random person  $i$  within the  $n^{\text{th}}$  household you’re investigating, see Figure 1.

It is a popular method because when carried out correctly, it leads to (almost) equal probability sampling,

|           | Eligible People |   |   |   |   |   |   |    |
|-----------|-----------------|---|---|---|---|---|---|----|
| Household | 1               | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
| 1st       | 1               | 1 | 1 | 1 | 1 | 1 | 1 | 1  |
| 2nd       | 1               | 2 | 2 | 2 | 2 | 2 | 2 | 2  |
| 3rd       | 1               | 1 | 3 | 3 | 3 | 3 | 3 | 3  |
| 4th       | 1               | 2 | 1 | 4 | 4 | 4 | 4 | 4  |
| 5th       | 1               | 1 | 2 | 1 | 5 | 5 | 5 | 5  |
| 6th       | 1               | 2 | 3 | 2 | 1 | 6 | 6 | 6  |
| 7th       | 1               | 1 | 1 | 3 | 2 | 1 | 7 | 7  |
| 8th       | 1               | 2 | 2 | 4 | 3 | 2 | 1 | 8  |
| 9th       | 1               | 1 | 3 | 1 | 4 | 3 | 2 | 1  |
| 10th      | 1               | 2 | 1 | 2 | 5 | 4 | 3 | 2  |

Figure 1: Kish Grid. *Note.* Reprinted from <https://www.statisticshowto.datasciencecentral.com/wp-content/uploads/2017/09/kish-grid.png>

something other selection methods do not obtain (Kumar, 2014). This is because of the way the grid is set up, each household member has an equal chance of being selected. Moreover, the only information you need in order to select respondents is a list containing the names of the persons in the household you want to sample, because selection is based on the number you’ve assigned to them. Other selection methods usually need more information, like the date of birth of the household members.

### Proportion of personal interviews within the sample (RQ6)

Within the data, there is a proportion of sample members that did not personally take part in an interview, potentially leading to nonresponse bias. However, not all of these observations count as nonresponse. Of the 3600 individuals in the sample, there are 459 children under the age of ten. These sample members are enumerated individuals, but ineligible for an interview. Accordingly, the number of eligible individuals in the sample is 3141. Of these 3141 eligible individuals, only 2656 respondents personally completed an interview. This number consists of 2399 adults who performed a full interview and 257 interviews with children. Taken together, 485 cases of (partial) unit nonresponse exist within the subset of eligible individuals.

Moreover, there are 169 sample members for whom a proxy interview was conducted with a close relative or cohabitant (generally a spouse, partner or adult child; Understanding Society, ‘Questions about variables’, N.D.). Questions answered on the nonresponders’ behalf were used to estimate factual information about the nonrespondent – information about their beliefs is not collected by proxy, and are therefore indicated as missing (item nonresponse). In the current dataset, no variables about beliefs or opinions of respondents are included, leading to a lower nonresponse rate, but potentially higher nonresponse bias when proxy interview data is used to compensate partial nonresponse.

Since (partial) nonresponse can lead to nonresponse bias, it is important to investigate the effect of nonresponse on the estimates generated above. Ideally, we would both compare respondents to partial nonresponse and unit nonresponse. Because of time constraints within this research project, we only investigate how weights may be constructed to adjust for unit nonresponse (see Part III: Estimation).

## Part II: Weighting

### Variation in design weights (RQ2)

Following from the sampling design, all sample members were assigned a design weight. As Table 1 shows, the variance in the design weight variable is 0.02. With a mean of 1.01, the coefficient of variation is only 0.14.

Table 1. *Descriptive statistics of design weight variable ‘a\_psnenip\_xd’.*

| Mean | SD   | Min. | Max. |
|------|------|------|------|
| 1.01 | 0.14 | 1.00 | 4.00 |

The low variance within the design weight variable can be due to rescaling of the weights, as explained in the IP User Guide (p. 56): “Each set of weights has been scaled by a constant factor to produce a mean of one amongst cases eligible to receive the weight”. That is, the design weights are standardized on one. This procedure does, however, not affect the coefficient of variation.

Moreover, the maximum value of this variable suggests trimming of weights greater than 4. Trimming, truncating, or smoothing of weights are techniques to inhibit the effect of a single observation on population estimates and its variance. Trimming may yield biased estimates, but is shown to decrease the mean square error (MSE; Lohr, 2010). It is therefore reasonable to assume that weight trimming was performed within the current dataset.

A histogram of the design weight variable (see Figure 2) shows a different reason for the low variance: all but 26 respondents have a design weight of exactly 1. This can be attributed to the sampling design, in which, among other things, population density within PSUs was used to construct strata. With that, inclusion probabilities of households were roughly equal within strata. The resulting design weights thus represent the discrepancy between the number of observation units in the population, and the proportion of households per stratum selected from the sampling frame.

### Computation of enumerated weights (RQ3)

Apart from design weights, each respondent is also assigned an enumerated weight (the variable ‘a\_psnenip\_xw’). This weight was the design weights post-stratified for sex, age and government office region. In the innovation panel study, population values were retrieved from the UK government and the weights were adjusted based on the differences between the sample proportions and these population values. These population values could not be obtained for the current study due to the data not being available to the general public. Another hurdle that had to be overcome was that age and government office region were grouped into respectively seven and four groups (IP User Guide, p.57), but no information was given on how these groups were constructed. An attempt was made to reconstruct these groups using linear regression.

The first step taken was to construct a linear model that contained the design weights, sex, age (numeric, ‘a\_dvage’) and the government office regions (GOR):

$$Weight_{adjusted} \sim Weight_{design} + Sex + Age + GOR$$

To find the grouping of government office regions, the coefficients that resulted from this regression, including the baseline, were ordered from lowest to highest, which resulted in the following graph:

While there are clear groupings in the graph, they do not divide the data into four groups. To be able to investigate the data from a theoretical point of view, the numbers were drawn into a map of the regions, as shown below.

The three government office regions with high weights are Scotland, London and the West Midlands. Scotland and London are prime candidates to be categories of their own, as Scotland has a strong regional identity and London is the capital of the country. From a data-driven perspective, combining the West Midlands with

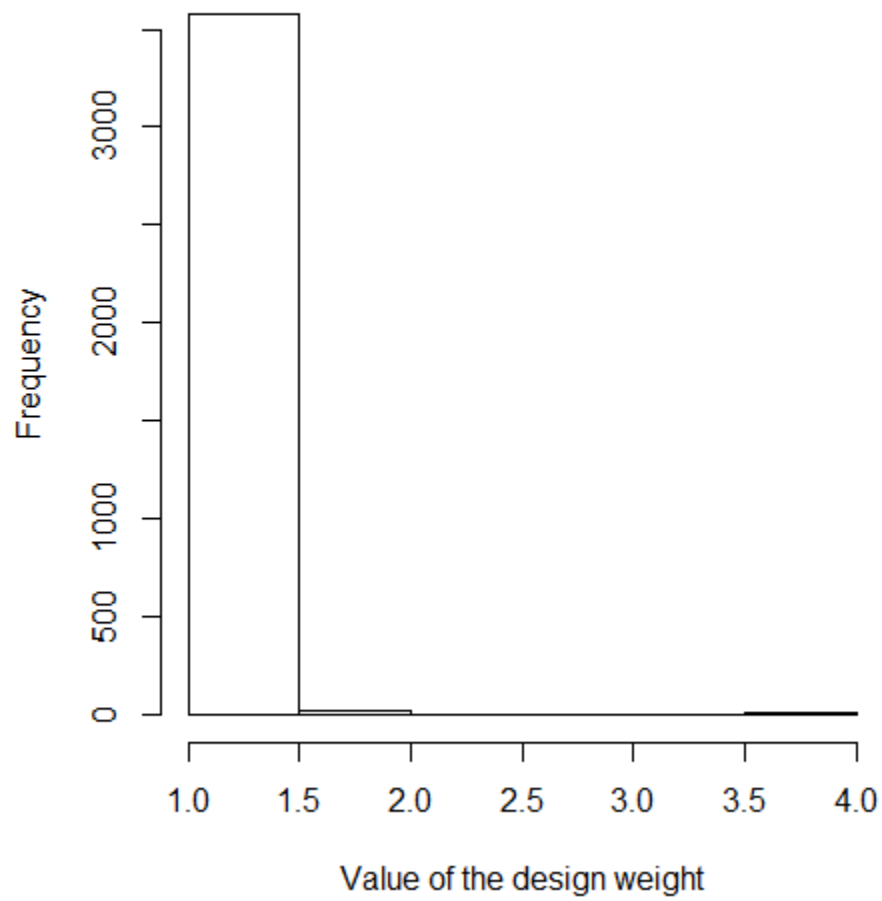


Figure 2: Histogram of design weight variable 'a\_psenip\_xd'.

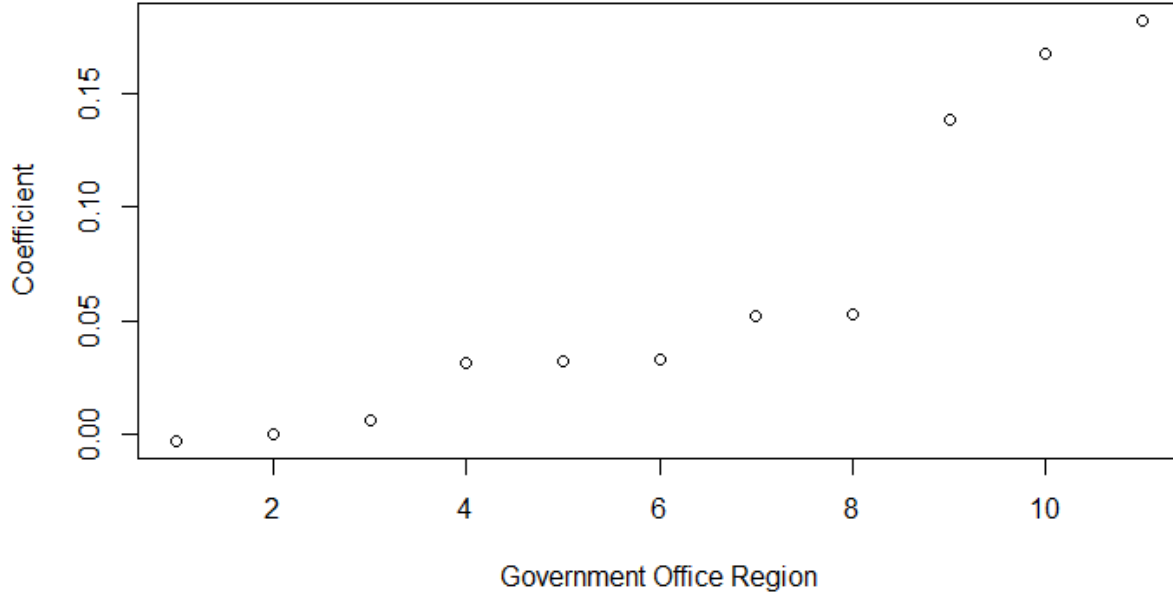


Figure 3: Coefficients per Government Office Region.

any region other than London or Scotland does not make sense, as the weight difference between those region and all other regions is relatively big. From a theory-driven perspective, we cannot think of any good reason to combine the West Midlands with London or Scotland: it is not filled with cities nor is it close enough to Scotland to assume a similar population.

When assuming that the regions with near identical weights are in the same group and Scotland and London each form one group on their own, the division between groups that makes most sense is that London, Scotland and the West Midlands are each categories of their own and Wales and the rest of England form the last group. Any other combination is unlikely due to inconsistencies with the found coefficients or with the theoretical background: any other combination would either require the grouping of the West Midlands with regions with a far lower weights, or would require London to be combined with Scotland or the West Midlands, which have very different demographics.

To find the age groups, a new variable for grouped government office region was created, divided into the groups described above. A factor-version of the 'a\_dvage' variable was also created in order to obtain the coefficients per age. A second linear regression was run similar to the previous one, but with the grouped GOR and factored age replacing the previously used age and GOR variables. The resulting graph, shown below, was used to split the age category into seven groups. In the graph, these groups are depicted with red lines.

These categories make sense both in theory and weight-wise (see Figure 3). Children are grouped together (0-14), as are teens and students (15-24). Adults are divided in three categories (25-34, 35-44, 45-64) and the elderly in two (65-74, 75+).

As a final step, a third linear regression analysis was done to find out which groups had higher weights and were thus underrepresented in the sample.

A regression analysis revealed which groups were under- and overrepresented in the sample (compared to the population of Great Britain), with higher weights being assigned to underrepresented groups. While the



Figure 4: Government Regions in Great Britain.

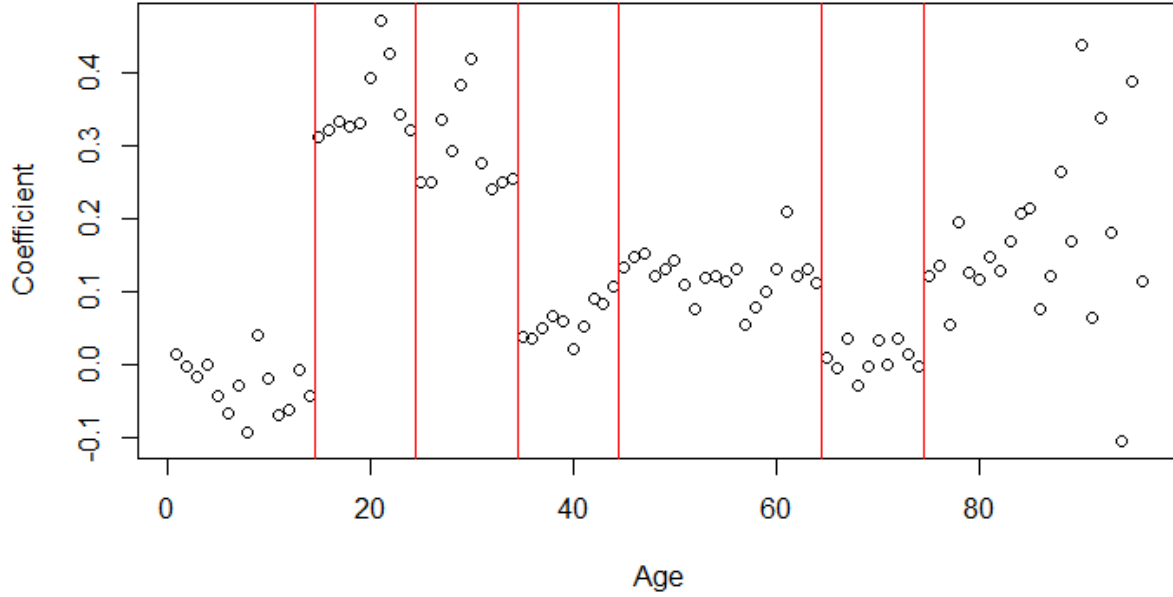


Figure 5: Coefficients per Age.

weights are mostly decided by the design weights, the post-stratification does alter them.

Females have slightly lower weights than males, indicating that they were overrepresented. For government office region groups, the people from England (excluding London and the Western Midlands) and Wales are the baseline and have the lowest of all weight, and thus were also overrepresented, with Scotland being the most underrepresented, followed by London and then the Western Midlands. With regards to age, the 0-20 group was surprisingly the most overrepresented, forming the baseline for this variable. The 21-30 and 31-40 groups were the most underrepresented and thus had the highest weights, followed by the 80+ group, the 51-70 group, the 41-50 group and finally the 71-80 group, which was by far the most underrepresented.

### Construction of nonresponse weights (RQ8)

Nonresponse can seriously threaten the quality of estimates if the nonresponders differ from the responders in the study. We can adjust for this by the use of auxiliary variables that hold information about the entire sample. We can compare the distribution of the auxiliary variables of the population with its response distribution, and assess whether the response distribution is representative for the whole population.

If the distributions differ, we must conclude that nonresponse has lead to a nonrepresentative sample. We can adjust for this by constructing nonresponse weights from the auxiliary variables. Then the weighted sample is representative with respect to the auxiliary variables used. Different choices of auxiliary variables lead to different weights. In order to the weights to be useful, it is important to use powerful auxiliary variables (Brick, 2013).

The IP User Guide states that “all models used to predict response propensities as described in the Technical Details are fitted using stepwise backward logistic regression with  $p=0.05$ ” (p. 56). Compared to other selection strategies, this is a fairly simple one. Considering these other strategies for selecting auxiliary variables for weighting adjustments for nonresponse, we advise to look into Schouten’s (2007) proposed strategy of a forward inclusion-backward elimination algorithm and/or Saarnald and Lundstrom’s (2010)



paper. We opt for a stepwise backwards logistic regression of because this method is similar to, yet easier to apply than the selection strategies described in Schouten (2007) Saarnadal and Lundstrom's (2010). Most important, this selection strategy was also used in constructing the original weights in the dataset. With this method, we hope to identify variables that predict nonresponse.

There are two stages of nonresponse we would have wanted to consider: \* person nonresponse: whether a personal interview was conducted yes or no; \* household nonresponse: defined as completion of at least the household grid.

However, as the levels of variable 'a\_hhresp\_dv' show, all households within the sample completed at least the household grid. We do not have any information about the 1271 addresses that were selected from the sampling frame but did not take part in an interview. Therefore, we only account for individual unit nonresponse, which occurred in 26.2% of the observations in the sample.

The forwards stepwise logistic regression model shows us that the variables 'a\_racel\_dv', 'a\_agegr5\_dv', 'a\_agegr13\_dv', 'hh\_size', 'a\_sex', and 'a\_urban\_dv' are useful in predicting nonresponse. That is, a linear combination of variables indicating race, sex, and age of respondents in combination with living situation (urban or rural area, and household size) yield a significant prediction of the logit of the nonresponse indicator. A backward stepwise analysis also indicated the variables household type (hh\_type) and government office region (a\_gor\_dv) could explain some additional variance in nonresponse. Since this type of analysis is exploratory, no inferences will be drawn from these results.

Evaluate fit/prediction of model: The predictive accuracy of the stepwise model is ..., where for the full model this was ..

half full model:  $NR \sim a\_gor\_dv + a\_urban\_dv + a\_sex + a\_racel\_dv + a\_employ + a\_agegr5\_dv + a\_agegr10\_dv + a\_agegr13\_dv + hh\_size + n\_child + hh\_type$  half forward model:  $NR \sim a\_racel\_dv + a\_agegr5\_dv + a\_agegr13\_dv + hh\_size + a\_sex + a\_urban\_dv$  half backwards model: (gor & hh\_type are additional compared to forward)  $NR \sim a\_gor\_dv + a\_urban\_dv + a\_sex + a\_racel\_dv + a\_agegr5\_dv + a\_agegr13\_dv + hh\_size + hh\_type$

We investigate the most important predictors individually: >>> plot relationships (I did a few analyses in the finalcode file but since the model isn't definitive I didn't write up any concluding remarks about them yet)

## Part III: Estimation

### Relationship between age and household composition (RQ5)

To investigate the relationship between age and household composition, three steps were performed: choosing which age variable would be used, constructing a new variable to represent household composition, and including the appropriate weight in the analysis.

Firstly, we investigated which variable to choose representing the age of respondents. There are four variables in the data expressing the age of respondents: a\_agegr5\_dv, a\_agegr10\_dv, a\_agegr13\_dv, a\_dvage. The first three are categorical variables, divided into groups as seen in the UK Labour Force Survey. The fourth variable, a\_dvage, is continuous.

The first categorical variable, a\_agegr5\_dv, contains 15 categories indicating age in five-year partitions up to '70 years or older'. The variable a\_agegr10\_dv is highly similar, indicating age in ten-year intervals, and consisting of 8 different categories between ages zero and seventy-plus. Finally, a\_agegr13\_dv consists 13 levels, with unequal age intervals under the age of 20 (namely 0-15, 16-17, and 18-19 years old), five-year intervals between the ages 20 and 64, and a single interval containing all ages from 65 years old upward.

The continuous variable a\_dvage denotes the age of respondents in completed years at the moment the interview took place. This was computed by comparing the date of birth and the interview date, or using an estimate (see estimated age variable a\_ageest (n = 22), and imputed age variable a\_ageif (n = 55)).

To investigate the relationship between age and household composition, it is necessary to know which households consist of children living at home, and which don't. Presumably, the largest variation between

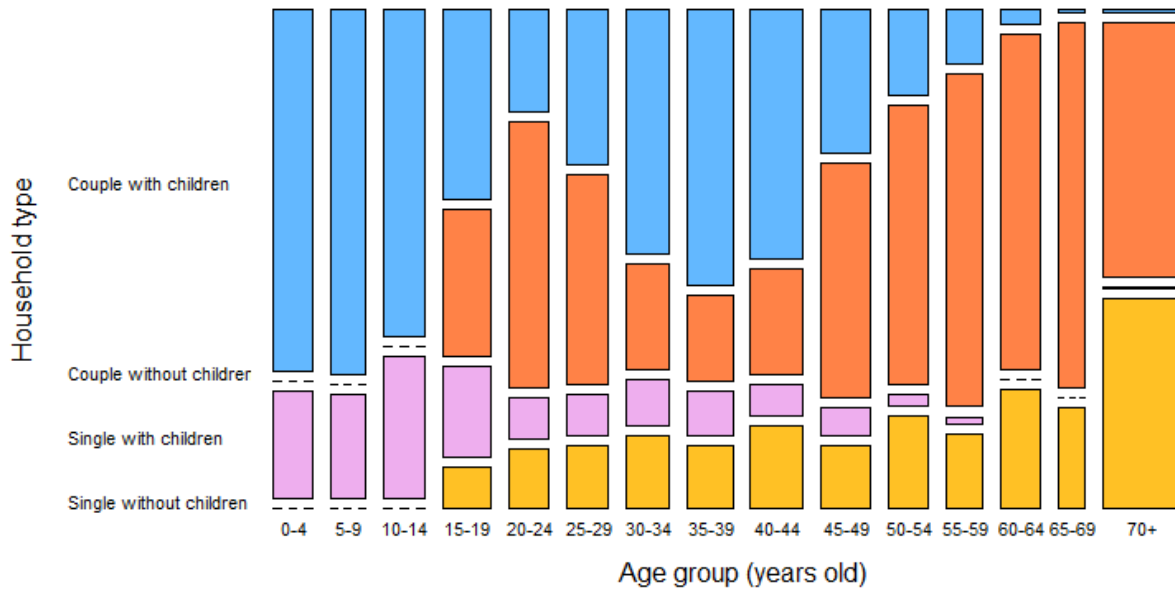


Figure 6: Household type per age category.

children living at home and those living independently occurs within the age group 15-25 year-olds. If we would use the ten-year intervals to indicate age, we might get less precise estimates, because of the large variance within the age category '10-19 years old'. Therefore, we decided to use an age variable with smaller intervals. Because of the unequal intervals in the age variable with 13 categories, it should be analysed as ordinal variable. To interpret the relationship with household composition, this is not ideal. Ultimately, we decided to use `agegr5_dv`, which is categorical and allows us for contingency tables and nice graphical representations.

Secondly, to construct a new variable denoting household composition, we used Eurostat as reference. Eurostat distinguishes three types of variables used to describe household composition:

- Household size;
- Number of children;
- Household type, with levels: couple (with/without children), single adult (with/without children), other type of household (with/without children).

All three household composition indicators were constructed from the data and included per household into a household composition matrix (see Appendix II: R Script).

Thirdly, we determined which weighting variable is appropriate to investigate the relationship between age and household composition. We want to use a weight per person (PSN), as it is impossible to calculate age per household. We also want to use the enumeration grid weights (EN), as both age and household composition have no missing values. We want the weights of the first wave (IP), as that is the one we are investigating. Finally, we want to use the stratified weights (XW) rather than the design weights, as we are investigating population values and not sample values. Thus, we want to use the weight `a_psnip_xw`, as this is the only weight that covers the correct data in the correct way.

Our objective was to investigate the relationship between age and household composition. We therefore plotted the composition of variable `hh_type` per age category (from `a_agegr5_dv`). The plot shows the relationships between `hh_type` and age you'd expect, from which we will highlight a few:

- Persons in the youngest three age categories (0-4, 5-9 and 10-14 years old) where all members of either

households of type “couple with children”, or “single parent with children”, obviously.

- The reverse pattern is found in the eldest age categories, where almost all persons are in a household of type “couple without children” or “single without children”.
- There are also a few interesting developments of categories over age. \*\* As age increases, overall the occurrence of households with children (pink and blue colored bars) declines. \*\* Looking at household type “couple with children”, the transition from being child in such a household and becoming a parent shows from the decline and subsequent increase of occurrences of this householdtype over age. \*\* Similarly, also the transition from living in a household with no children to living in a household with children, to an “empty nest” is visible by the development of the orange bars (couple without children) and the blue bars (couple with children). This is also true for the yellow and pink bars (single without children and single parent with children)

These are obviously just general observations, further statistical investigation is required if there are more specific research questions. Therefore we created a new age variable in which all household types were represented within an age group (see Appendix II: R Code). With this new variable, we performed a Chi square test of independence to evaluate the statistical relationship between age and household type. It appears that age and household type are not independent within the current sample;  $\chi^2(30) = 1763.1$ ,  $p < .001$ . This result strongly suggests that age and household type are related in the population.

#### **Proportion of employed people of working age (RQ4)**

43,3% of the population is employed, with a 95% confidence interval of 41.4%-45.2%.

#### **Proportion of employed people of working age, accounting for nonresponse (RQ7)**

After post-stratification (as we also included post-stratification in the previous analysis), 46% of the people who did a personal interview is employed, with a confidence interval of 44.9%-49.3%. For the previous question, where all but two people were included, the estimated proportion of the population that is employed was 43,3%, with a 95% confidence interval of 41.4%-45.2%. As neither mean is contained in the 95% confidence interval of the other, it can be concluded that there is a difference between respondents and non-respondents in employment, with respondents being more likely to be employed than non-respondents.

## **Discussion**

**In conclusion, . . .**

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## Appendix I: Research Proposal

Understanding Society is a large, multi-year panel study of the population in the United Kingdom and Northern Ireland. The dataset that you will analyse is wave 1 of the Understanding Society Innovation Panel, which is used for testing instruments. The Innovation panel (IP) uses the same sampling procedure as the main Understanding Society Survey, apart from the fact that Northern Ireland is not included. Sampling is done using a multistage stratified cluster sample. For more information, see [https://www.understandingsociety.ac.uk/sites/default/files/downloads/documentation/innovation-panel/user-guides/ip\\_user\\_guide.pdf](https://www.understandingsociety.ac.uk/sites/default/files/downloads/documentation/innovation-panel/user-guides/ip_user_guide.pdf)

Within this file, further references are made to the sampling and weighting guidelines. Do read these, because this will be important for the rest of the analyses. The data file you need for this assignment is ‘understanding society innovation panel wave A.RDS’.

Questions for report: (for questions 1-5), please ignore nonresponse errors. 1. The Innovation panel would interview multiple people from every household, but in the case households are very large, a Kish Grid is used. Can you explain why the Kish Grid is a popular method to do within-household selection of respondents?

2. Despite the complex data structure, the design weights as included in the variable ‘a\_psnenip\_xd’ have a very low variance. Can you explain why this is the case?
3. Apart from the design weight, an enumerated design weight ‘a\_psnenip\_xw’ is also included. Please show using syntax in R how you can calculate this weight.
4. What is the proportion of employed people of working age (15-64) in the population? (use variable ‘a\_employ’)
5. Please investigate the relationship between age and household composition for the population of Great Britain. Take the following steps:
  - Investigate which age variable you should use (and whether choosing a particular age variable at all matters)
  - Construct a new variable ‘household\_composition’ from the following four variables: ‘a\_livesp\_dv, a\_cohab\_dv, a\_single\_dv, a\_mastat\_dv’. You may also use information from across household members to do this.
  - Which weight should you use here? From here on, take nonresponse error into account.
6. Not all sample members you included in questions 1-5 participated in the survey. Use the variable ‘a\_ivfio’ to investigate how many people personally conducted an interview.
7. Repeat your analysis under question 4. Use only the people who personally did an interview. What is your conclusion? Is there a difference between respondents and nonrespondents when it comes to their employment?
8. Using only variables from this dataset, can you find variables which could potentially be useful for constructing nonresponse weights? Why these variables?