Generalized Reciprocity: Theory and Experiment

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Abstract

Those who receive help may be more likely to help a third party. This pay-it-forward behavior, called generalized reciprocity, can be powerful in spreading kindness from person to person. We use laboratory games to explore the psychological motivations that contribute to generalized reciprocity. We rationalize our experimental results using a psychological gametheoretic framework that incorporates and extends Dufwenberg and Kirchsteiger (2004), Battigalli and Dufwenberg (2009), and Fehr and Schmidt (1999). We find that altruism, reciprocity

incentives, and inequity aversion are all necessary to explain subjects' behaviors in general-

ized reciprocal exchange.

Keywords: prosocial behavior, reciprocity, pay-it-forward, altruism, psychological game

theory

JEL: C79, C90, C91

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1 Introduction

Prior work has shown that people who have received help are more likely to help a third party, even at their own expense (Ben-Ner et al., 2004; Bartlett and DeSteno, 2006; Desteno et al., 2010; Herne et al., 2013; Gray et al., 2014; Tsvetkova and Macy, 2014; van Apeldoorn and Schram, 2016; Mujcic and Leibbrandt, 2018; Simpson et al., 2018; Melamed et al., 2020). This pay-it-forward behavior, which we call *generalized reciprocity*, is core to the propagation of kindness from one person to another. At the workplace, an employee who was mentored by a superior may elect to advise a new coworker in turn. In fast food restaurant drive-throughs, when one customer pays for a second customer's order, the second customer may then pay for the next customer after her. Generalized reciprocity can transform social norms and create cultures of cooperation within organizations, social groups, and communities (Binmore, 1994; McCullough et al., 2008), since it can propagate to include hundreds of people. For example, at a Dairy Queen restaurant drive-through in Minnesota, over 900 consecutive cars chose to pay for the order of the car behind them (Ebrahimji, 2020).

Generalized reciprocity cannot be explained by theories in which individuals only care about their own material payoffs. Both theoretical and empirical work have shown that psychological components are necessary to explain giving behavior in multiplayer interactions (Rabin, 1993; see Fehr and Gächter, 2000 for an overview). Dufwenberg and Kirchsteiger (2004) develop a workhorse model for dynamic games that involve *direct* reciprocity, in which players can pay back another player's actions. Battigalli and Dufwenberg (2009) present a more general framework of dynamic psychological games and apply it in subsequent work to incorporate emotions, reciprocity, image concerns, and self-esteem in economic analysis.² Due to their focus on the

¹Under Mujcic and Leibbrandt (2018) and van Apeldoorn and Schram (2016), generalized reciprocity is also called *upstream indirect reciprocity*, in which a player pays forward kindness because she has received kindness in the past. It is not our goal to address *downstream indirect reciprocity*, where a player decides whether to benefit another based on the other player's past actions (Takahashi, 2000; Yoeli et al., 2013; Engelmann and Fischbacher, 2009; Ong and Lin, 2011; Khadjavi, 2017). Upstream indirect reciprocity is much less studied compared to downstream indirect reciprocity, especially in the economics literature; see studies on downstream indirect reciprocity: Bolton et al. (2005); Seinen and Schram (2006); Zeckhauser et al. (2006); Berger (2011); Charness et al. (2011); Heller and Mohlin (2017); Gong and Yang (2019); Gaudeul et al. (2021).

²See Battigalli et al. (2019) and Battigalli and Dufwenberg (2021) for recent surveys on theoretical, experimental,

strategic nature of direct reciprocity interactions, these models cannot be readily applied to *generalized* reciprocity, in which people cannot pay back their benefactors but must choose whether to pay it forward by benefiting an unrelated third party.

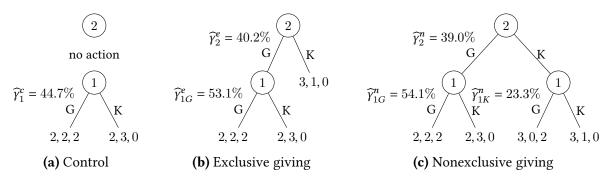
Investigating generalized reciprocity yields unique insights regarding altruistic behavior. In direct reciprocity games, a player may help another because she expects her beneficiary to pay her back. It is therefore difficult to distinguish whether a player helps another due to altruism or to maximize own expected payoffs. In games without direct reciprocity considerations, players know that their beneficiaries cannot pay them back, so the decision to give goes against maximizing their own material payoffs.

This enables us to more cleanly investigate two channels in how reciprocity motives impact altruism. First, we explore how receiving a gift influences giving to a third party. Second, we examine how giving decision is influenced by the possibility that their beneficiary will pay it forward, propagating their kindness to others. Beliefs that one's beneficiary will reciprocate the generosity by helping a third party may make one more willing to give, since it expands the impact of the generosity.

Specifically, we embed altruism, inequity aversion, generalized reciprocity, and higher-order beliefs into a psychological game-theoretic framework that extends Dufwenberg and Kirchsteiger (2004) and Fehr and Schmidt (1999). We test the model's theoretical predictions in a laboratory experiment involving three-player games, depicted in Figure 1. In all games, players choose whether to pass a chip worth \$1 to the next player. Following the multiplier methods used for investment and public goods games, each chip that is passed turns into two chips for the recipient. Each player's index denotes the number of players behind them in the chain. P0 is the last potential recipient, P1 is the last player to decide on giving, and P2 (if allowed to give) is the penultimate player to decide on giving.

Figure 1a displays the *control game*, in which P2 is endowed with 2 chips and cannot give a chip, P1 is endowed with 3 chips, and P0 with no chip. Only P1 makes a giving decision. In the applied, and methodological work, and also Falk and Fischbacher (2006) for an alternative but related theory of reciprocity.

Figure 1: Games and giving rates in the experiment



Note: Material payoffs are (π_2, π_1, π_0) . The $\widehat{\gamma}$ s indicate the giving rates in our experiment. The superscript denotes game type, where c stands for control, e for exclusive, and n for nonexclusive. The subscript G stands for P1's decision after P2 gives, and K for P1's decision after P2 keeps.

treatment games depicted in Figures 1b and 1c, P2 is endowed with three chips, P1 with one chip, and P0 with no chip. P2 and P1 make giving decisions. P2 can give a chip to P1 so that P1 has 3 chips in total. If P2 gives, all three games have the interim allocation $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$ before P1 makes a giving decision. If P1 keeps, the game concludes with payoffs $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$, and if P1 gives, the game concludes with payoffs $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$. We thus keep payoff distributions the same in the three games, so that differences in P1's giving behavior across games cannot arise from absolute or relative allocation concerns.

The two treatment games differ in whether P0's channel of receiving is exclusive. In the *exclusive* game depicted by Figure 1b, P1 cannot give P0 a chip unless P2 gives P1 a chip first. If P2 keeps, the game concludes with payoffs $(\pi_2, \pi_1, \pi_0) = (3, 1, 0)$. However in the *nonexclusive* game depicted by Figure 1c, P1 can give P0 a chip regardless of whether P2 gives a chip to P1 first. If P1 chooses to keep after P2 keeps, the game concludes with payoffs $(\pi_2, \pi_1, \pi_0) = (3, 1, 0)$. If P1 gives even after P2 keeps, the game concludes with payoffs $(\pi_2, \pi_1, \pi_0) = (3, 0, 2)$.

We first compare the giving decisions of the last mover, P1. By comparing P1's actions across games, we establish the generalized reciprocity effect: receiving a gift makes you more likely to give to a third party. Figure 1 summarizes empirical giving rates. P1 is (i) most likely to give when P2 has given in the exclusive game (54.1%); (ii) second most likely to give when P2 has given in the nonexclusive game (53.1%); (iii) third most likely to give when P2 cannot give in control game

(44.7%); and (iv) least likely to give when P2 can give but decides not to in the nonexclusive game (23.3%). Compared with control, P1 is 8.4-9.4 percentage points (18-21%) more likely to give in the exclusive and nonexclusive games after receiving a chip from P2 (p < 0.005). The game structure holds income effects, distributional preferences, and social image concerns constant across games, enabling us to rule out these alternative explanations.

We next compare the behavior of first movers: P2 in the treatment games and P1 in the control game. This allows us to investigate whether your decision to give is influenced by the possibility that your beneficiary will pay forward your generosity, magnifying your impact. If so, we would expect giving to be greater among P2 in the treatment game than P1 in the control game. However, we find the opposite. P1's giving rate in the control game is 44.7%, which is significantly greater than P2's giving rates of 40.2% and 39.0% in the exclusive and nonexclusive games, respectively (p < 0.05). It appears that expectations about P1 paying forward P2's generosity plays a negligible role in guiding P2's giving decision. Rather, as our model will show, inequity aversion is necessary to explain P2's behavior. Intuitively, P1 can directly equalize payoffs through giving in the control game, resulting in allocations of (2,2,2) rather than (2,3,0). However, P2 risks an unequal allocation of (2,3,0) by giving in the treatment games since she cannot control what P1 will do after she gives.

Lastly, we quantify the importance of altruism, reciprocity, and inequity aversion in explaining subject behavior. We find that the model with altruism, reciprocity, and inequity aversion performs best overall, explaining the behavior of 90% of subjects. The model excluding altruism can only explain the behavior of 30% of subjects, while the model excluding reciprocity can only explain the behavior of 70% of subjects. In contrast, the model which excludes inequity aversion performs almost as well as the full model overall, in that it explains the behavior of 90% of our subjects. Together, the results indicate that altruism and reciprocity are key to explaining payit-forward behavior, while inequity aversion plays a less important role. Moreover, our results indicate that altruism and reciprocity incentives rationalize P1's decision to pay forward P2's generosity. In contrast, inequity aversion explains why P2's giving does *not* grow in situations

where P1 could pay her generosity forward.

Our contributions are threefold. Most game-theoretic frameworks explain direct reciprocity between two individuals that directly interact (Rabin, 1993; Dufwenberg and Kirchsteiger, 2004; Falk and Fischbacher, 2006; Seinen and Schram, 2006; Cox et al., 2007; Berger, 2011; Gong and Yang, 2019; Gaudeul et al., 2021), but cooperative communities frequently involve three or more individuals who do not necessarily directly interact.³ To the best of our knowledge, our paper is the first to develop a behavioral game-theoretic framework for the systematic investigation of generalized reciprocity, promoting reciprocal exchange in environments where not all parties directly interact. The model allows us to evaluate the importance of altruism, inequity aversion, and reciprocity incentives in rationalizing pay-it-forward behavior. Since generalized reciprocity is the channel by which kind actions beget further kind actions, our study constitutes a first step toward understanding how helping behavior can spread from person to person. In doing so, our paper informs how to foster cultures of cooperation within workplaces, neighborhoods, and communities.

Second, we introduce a simple, novel experiment that establishes the role of generalized reciprocity motives while controlling for alternative explanations. Many prior papers on generalized reciprocity cannot rule out the income effect (Herne et al., 2013; van Apeldoorn and Schram, 2016; Simpson et al., 2018; Mujcic and Leibbrandt, 2018), where the act of receiving a gift itself can make subjects more likely to give through increasing their wealth. Furthermore, to the best of our knowledge, our paper is the first to experimentally account for relative wealth differences, which could lead to pay-it-forward behavior if subjects exhibit inequity aversion. Since our experiment holds social concerns constant across games, we leverage our within-subject design to difference out the impact of social image or social pressure considerations, which (Charness and Rabin, 2002; Sobel, 2005; Cox et al., 2008) theoretically argue and Malmendier et al. (2014) experimentally find to be important in reciprocal interactions.

³Wu (2018) and Jiang and Wu (2019) discuss models involving indirect interactions among more than two players. Reciprocal behavior has also been investigated in other disciplines, including evolutionary biology (Nowak and Sigmund, 1998a,b; Ohtsuki and Iwasa, 2006; Iwagami and Masuda, 2010) and psychology (Hu et al., 2019; Nava et al., 2019).

Third, our theory and experiment work together to distinguish models of fairness based on outcomes, types, and intentionality. Outcome-based models propose that fairness depends on players' relative payoffs, so inequity aversion and minimax preferences should drive how subjects allocate wealth between themselves and others (Fehr and Schmidt, 1999; Bolton and Ockenfels, 2000). Type-based models posit that giving behavior depends on one's innate altruism parameter (Levine, 1998; Cox et al., 2007; Malmendier et al., 2014). Intentions-based models argue that utility also depends on beliefs about others' kindness intentions (Rabin, 1993; Dufwenberg and Kirchsteiger, 2004; Battigalli and Dufwenberg, 2009; Gul and Pesendorfer, 2016). As our paper demonstrates, outcome- and type-based models which ignore intentionality cannot rationalize our experimental findings. By comparing within subjects across games, we show that P1 is more likely to give in the treatment games than the control game, even though all games have the same fairness parameters and payoff distribution. Such behavior can only arise if subjects' conceptions of fairness depend on their beliefs of others' intentions, namely if P1's utility depends on her beliefs of P2's kindness.

The rest of the paper is organized as follows. Section 2 introduces the psychological gametheoretic framework, defines the solution concept of *dynamic reciprocity equilibrium*, describes the games, and derives the predicted equilibrium individual giving rates. Section 3 describes the experiment. Section 4 compares our experimental results with theoretical predictions, shows the necessity of altruism, fairness, and reciprocity in rationalizing our results, and suggests future research avenues (e.g., incorporating credit attribution). Section 5 concludes, and the appendices collect omitted proofs and additional experimental results.

2 Theory

2.1 Dynamic reciprocity equilibrium

We consider finite-action multistage games with observable actions and without moves of nature. That is, play proceeds in stages in which each player, along any path reaching that stage, (i) knows all preceding choices, (ii) moves exactly once, and (iii) obtains no information about other players' choices in that stage. Since we extend the direct reciprocity framework of Dufwenberg and Kirchsteiger (2004), we adopt and modify their notation.

Formally, let $N=\{1,\ldots,n\}$ denote the set of players. Let h denote a history of preceding choices represented by a node in the extensive-form representation of games, and let H denote the set of histories of a game. The set of behavioral strategies of player $i \in N$ is denoted by Σ_i , where a strategy $\sigma_i \in \Sigma_i$ of player i assigns a probability distribution over the set of possible choices of player i for each history $h \in H$. Let $\Sigma = \prod_{i \in N} \Sigma_i$ denote the collection of behavioral strategy profiles σ of all players, and $\Sigma_{-i} = \prod_{j \in N \setminus \{i\}} \Sigma_j$ the collection of behavioral strategy profiles σ_{-i} of all players other than i. Let Σ'_{ij} be the set of beliefs of player i about the strategy of player j (i.e., i's first-order beliefs). Let Σ''_{ijk} be the set of beliefs of player i about the belief of player j about the strategy of player k (i.e., i's second-order beliefs). By definition, $\Sigma'_{ij} = \Sigma_j$ and $\Sigma''_{ijk} = \Sigma'_{ik} = \Sigma_k$.

With $\sigma_i \in \Sigma_i$ and $h \in H$, let $\sigma_i(h)$ denote the updated strategy that prescribes the same choices as σ_i , except for the choices that define history h. Note that $\sigma_i(h)$ is uniquely defined for any history h. For any beliefs $\sigma'_{ij} \in \sigma'_{ij}$ or $\sigma''_{ijk} \in \Sigma''_{ijk}$, define updated beliefs $\sigma'_{ij}(h)$ and $\sigma''_{ijk}(h)$ analogously.

Player i's utility function depends on strategies, first-order beliefs, and second-order beliefs, which we summarize by a vector $\vec{\sigma} \equiv (\sigma, \sigma', \sigma'')$. These strategies and beliefs in turn determine the expected payoffs of players, which are determined by her own material payoff and three psychological components: (i) the altruistic payoff, (ii) the reciprocity payoff, and (iii) the fairness payoff. The utility function then takes the following form:

$$u_{i}(\vec{\sigma}) = \pi_{i}(\sigma) + A_{i} \sum_{j \neq i} \pi_{j}(\sigma) + \sum_{j,k \neq i} Z_{i} \lambda_{iki}(\vec{\sigma}) \kappa_{ij}(\vec{\sigma})$$

$$\text{altruism} \qquad \text{generalized reciprocity}$$

$$\text{disadvantageous inequity aversion} \qquad \text{advantageous inequity aversion}$$

$$-\sum_{s \in S} \sigma(s) \left[\alpha_{i} \frac{1}{n-1} \sum_{j \neq i} \max\{\pi_{j}(s) - \pi_{i}(s), 0\} + \beta_{i} \frac{1}{n-1} \sum_{j \neq i} \max\{\pi_{i}(s) - \pi_{j}(s), 0\} \right]$$

$$\text{inequity aversion}$$

where $\pi_i(\sigma)$ is i's material payoff; $A_i \in [0, 1]$ is i's altruistic factor that dictates how much utility i derives from the material payoffs of other players regardless of the distribution of relative wealth; and Z_i is i's generalized reciprocity parameter. From Fehr and Schmidt (1999), we incorporate inequity aversion parameters α and β . Player i receives disutility α_i for each unit of lower payoff than others ("disadvantageous inequity aversion") and disutility β_i for each unit of higher payoff than others ("advantageous inequity aversion"), where $\alpha_i \geq \beta_i$ and $0 \leq \beta_i \leq 1$.

The function $\kappa_{ij}: \Sigma_i \times \prod_{j \neq i} \Sigma'_{ij} \to \mathbb{R}$ is *i*'s kindness to *j* from choosing strategy σ_i while other players choose σ_{-i} :

$$\kappa_{ij}\left(\sigma_i(h),(\sigma'_{ij}(h))_{j\neq i}\right)=\pi_j\left(\sigma_i(h),(\sigma'_{ij}(h))_{j\neq i}\right)-\pi_j^{Q_i}\left((\sigma'_{ij}(h))_{j\neq i}\right)$$

where $\pi_j^{Q_i}((\sigma'_{ij})_{j\neq i}) = \frac{1}{2}\left[\max_{\sigma_i\in\Sigma_i}\pi_j(\sigma_i,(\sigma'_{ij})_{j\neq i}) + \min_{\sigma_i\in\Sigma_i}\pi_j(\sigma_i,(\sigma'_{ij})_{j\neq i})\right]$ is player j's equitable payoff with respect to i. It is the average between j's lowest and highest possible material payoff based on i's strategy. Since kindness is defined relative to j's equitable payoff, i's kindness is positive (negative) if i chooses an action that gives a strictly higher (lower) expected payoff for j than j's equitable payoff.

The function $\lambda_{iji}: \Sigma'_{ij} \times \prod_{k \neq j} \Sigma''_{ijk} \to \mathbb{R}$ is *i*'s belief of *j*'s kindness to *i* given *i*'s belief of *j*'s

belief of other players' strategies:

$$\lambda_{iji}(\sigma'_{ij}(h),(\sigma''_{ijl}(h))_{l\neq j})=\pi_i\left(\sigma'_{ij}(h),(\sigma_{ijl}(h))_{l\neq j}\right)-\pi_i^{Q_j}\left((\sigma''_{ijl}(h))_{l\neq j}\right)$$

All objects are discussed in greater detail in Appendix A.

Compared with Dufwenberg and Kirchsteiger (2004), there are three additions. First, we incorporate an altruistic payoff component. Second, we change the direct reciprocity component to a generalized reciprocity component in which subjects can reciprocate kind acts by helping a third party. Third, we follow Fehr and Schmidt (1999) and incorporate inequity aversion parameters.

Below, we define alternative specifications of the utility function in which some psychological components are assumed away. We formulate key predictions based on these alternative utility specifications and test them against our experimental results in Section 4.

Definition 1. Altruistic, inequity averse, and reciprocal (AIR) utility assumes that $A_i \ge 0$, $\alpha_i \ge 0$, $\beta_i \ge 0$, and $Z_i \ge 0$ for all i. Standard/selfish (S) utility ignores psychological components and assumes that $A_i = 0$, $\alpha_i = 0$, $\beta_i = 0$, $Y_i = 0$, and $Z_i = 0$ for all i. Altruistic (A), Reciprocal (R), Inequity averse (I), AI, AR, and IR utilities respectively assume the relevant utility components to be nonnegative and other utility components to be zero. See the top rows of Table 1 for the complete parametric specification of each utility function.

Note that the reciprocity utility component depends on strategies, beliefs, and other players' material payoffs. Therefore, the equilibrium is defined with respect to both strategies and beliefs.

Definition 2. Strategies and beliefs $\vec{\sigma}$ constitute a *dynamic reciprocity equilibrium* if and only if (i) (consistency) players have correct beliefs about other players' actions, i.e., $\sigma = \sigma'_i = \sigma''_{ij}$ for any players i and j; and (ii) (utility maximization) strategy profile σ maximizes players' utilities at each information set given first-order and second-order beliefs σ'_i and σ''_{ij} .

Theorem 1. A dynamic reciprocity equilibrium always exists.

We design our three experimental games to have generically unique equilibrium outcomes.⁴

2.2 Predictions

The theoretical framework and experimental design complement each other in generating eight predictions—proven as eight propositions in the online appendix—regarding subjects' strategies. In the experiment, we elicit subjects' choices at all nodes of all games using the strategy method. This enables us to compare behavior at different nodes for each subject.

Table 1 summarizes the eight predictions under each of the utility functions listed in Definition 1, which turn on or off each of the three psychological components we consider. Our predictions involve comparing players' actions at the decision nodes under the LHS and RHS columns of Table 1. To compare giving rates, we say that one is more inclined to give than another in the following sense.

Definition 3. Player i is *more (less) inclined* to take action s at node H than player j to take action s' at node H', $\sigma_i(s|H) > (<) \sigma_j(s'|H')$, if for $\alpha_i = \alpha_j$ and $Z_i = Z_j$, $\sigma_i(s|H) \ge (\le) \sigma_j(s'|H')$ for all parameters, and the inequality holds strictly for some parameters.

In the most general AIR utility specification where $A_i > 0$, $\alpha_i > 0$, $\beta_i > 0$, and $Z_i > 0$ (Column 8), the giving rates by the last mover are ordered: $\gamma_{1G}^e \sim \gamma_{1G}^n > \gamma_1^c > \gamma_{1K}^n$ (Predictions 1-5 in Column 8). The initial mover's giving rate is higher in the exclusive game than the nonexclusive game (Prediction 6 in Column 8), but it is unclear whether the initial mover in the control game is more inclined to give than in the treatment games, since altruism pushes for greater giving and inequity aversion pushes for lower giving in the treatment games (Predictions 7 and 8 in Column 8).

A standard model without the aforementioned psychological components predicts no giving under any circumstance (Column 1). We also consider the predictions with alternative utility

⁴In general, a dynamic reciprocity equilibrium is not necessarily unique. However in our games, except for a measure zero set of parameters, equilibrium strategies are uniquely determined. For a measure zero set of parameters, a player may be indifferent between giving and keeping, and hence any probability of giving can constitute an equilibrium, resulting in multiple equilibria.

specifications in which only one psychological component is considered at a time (Columns 2-4). Under altruism alone (Column 2), there would be a positive giving rate for each of the six decisions, but no distinction in the likelihood of giving for last movers (P1) in different games (Predictions 1-5 in Column 2). Initial movers would be most likely to give in the exclusive game, and equally likely to give in the nonexclusive and control game (Predictions 6-8 in Column 2). Under inequity aversion alone (Column 3), P1's giving rates after P2 gives in the treatment games would equal that in the control game (Predictions 2, 3, and 5 in Column 3), and P1 would not give in the nonexclusive game after P2 keeps (Predictions 1 and 4 in Column 3). There are ambiguous predictions in the comparison of first movers in the control versus treatment games (Predictions 7-8 in Column 3). Under reciprocity motives alone (Column 4), P1 would only give after P2 gives and not after P2 keeps in the nonexclusive game (Prediction 4 in Column 3). This would make the nonexclusive and exclusive games effectively the same to P2, so P2 would be equally inclined to give in the two games (Prediction 6 in Column 3).

We also consider the predictions when one psychological component is omitted (Columns 5-7). Without reciprocity (Column 5), P1 would be equally inclined to give in the control game and after P2 gives in the treatment games (Predictions 6-8 in Column 5 vs 8). Without inequity aversion (Column 6), predictions for the giving rate by last mover are the same as in the general AIR model, but the initial mover would be less inclined to give in the control game than in the treatment games (Predictions 7 and 8 in Column 6). Finally, without altruism (Column 7), the predictions are the same as in the general model, but there would be no reason for P1 to give after P2 keeps in the nonexclusive game (Predictions 1 and 4 in Column 7).

3 Experimental procedure and data

Experimental sessions were implemented on Amazon Mechanical Turk (MTurk), an online platform commonly used by experimental social scientists to collect information about choices, attitudes, and opinions. The study was administered between February 23 and March 26 of 2021.⁵ The study was approved by the Institutional Review Board at Michigan State University.

We conducted the experiment using Qualtrics software. Since our study involved three-player games, we held sessions of 9 subjects each. Recruitment materials informed subjects that they would receive \$3 for completing the study and up to \$5 in bonus payments. Subjects could preview all experimental materials before choosing to participate. Experimental materials informed subjects that upon study completion, they would be randomly assigned to a game and a group with other players from their session. Their bonus earnings were calculated based on their giving decisions, as well as the giving decisions of their group mates. Subjects received their payments via MTurk within 24 hours of study completion.

We used the strategy method to elicit subjects' actions at all nodes of each game. For example, in the nonexclusive game we asked subjects whether they would give as P1 in the case that P2 gave and whether they would give in the case that P2 kept. We leverage this within-subject variation to examine how each subject's actions differ across player roles and across games. Importantly, subjects made their giving decisions after being told how they would be compensated but before they knew which game, group, or player role they would be compensated for. Subjects could not contact each other or know whom they would be playing with when they made their decisions.

All subjects proceeded through the experiment as follows. First, they were taken to the consent page, which described their rights as study participants. Next, they viewed a video that described the study and all games.⁶ Prior to the beginning of each game, subjects viewed the extensive-form diagram of the game they were about to play, which contained information about endowments and payoffs for each realization of the game. Throughout the session, they could click on a link that displayed the extensive-form diagram and video describing the relevant game. To incentivize subjects to pay attention during the study, we asked quality check questions be-

⁵The experiment was conducted online since it took place during the Covid pandemic, and nonessential in-person studies were prohibited by Michigan State University.

⁶The experiment involved three three-player games and two four-player games. This paper compares giving results between the three-player games, since our goal is to develop a framework to explain pay-it-forward behavior. We plan to explore the predictions of the theory on giving results in the four-player games in subsequent studies.

fore the exclusive and nonexclusive games. Subjects were informed they would earn an additional \$0.50 per game if they answered all questions correctly on their first attempt.⁷

All subjects played the control game first. After the control game, the order of the exclusive and nonexclusive games was randomized.⁸ At the end of the study, subjects completed a demographic questionnaire. Further details of the experiment, including screenshots, are available in Appendix B.

Table 2 displays summary statistics. The top panel summarizes subject performance. Bonus payments ranged from \$0 to \$5, with a median payment of \$3 and an average payment of \$2.55. In the full sample, the average number of quality check questions answered incorrectly on the first try out of four was 2.16, with a median of 1. However, the distribution is positively skewed: Of 403 total subjects, 140 had no incorrect questions, 87 had one incorrect question, and 97 had two incorrect questions on the first try. The remaining 78 had 3 or more incorrect questions on the first try. It is likely that those who answered more than two questions incorrectly were not paying attention or had difficulty understanding the games, so their choices may not reflect their true preferences. To exclude subjects who demonstrably struggled to understand the game, we define a separate subsample of subjects who answered two or more questions correctly on the first try, called the *accurate responders* sample. We present summary statistics for the full sample in the first column (N = 403) and the accurate responders sample in the second column (N = 324).

In the full sample, subjects took 28-29 minutes on average to complete the study. Although the study was designed to be completed within an hour, two subjects took 69.53 and 124.90 minutes. Twenty-eight subjects took between 45 and 60 minutes. As Section 4.4 will show, excluding these 30 subjects from the two samples does not appreciably change results.

Responses to the demographic questionnaire reveal no significant difference across the two samples (bottom panel). Less than a third of subjects (27.5%) are women, 80-83% have at least an

⁷In addition to the three player games, subjects concluded the study with two four player games. Subjects could earn \$0.50 each for four sets of quality check questions. They could earn \$2 maximum on the quality check questions and \$3 maximum from their giving decision, so their maximum bonus earnings were \$5.

⁸Robustness checks in Section 4.4 show that almost all our results hold independent of game order.

⁹Our full sample is not a multiple of 9 since two subjects in a session did not answer all questions and were omitted from the analysis.

associate's degree, and over 90% are employed full- or part-time. Around 67-68% are US citizens, 70-73% are US residents, and over 75% are native English speakers. In terms of race and ethnicity, about half of subjects are white, 30% are Asian, 10-13% are Black, 5% are Hispanic, and the remaining 3% are categorized as other race or ethnicity. In terms of age, half of the subjects are 26-35 years old, 22-23% are 36-45 years old, 15-16% are 16-25 years old, and about 10% of subjects are 46-65 years old. Only 1.2-1.5% of subjects are 65 or older.

4 Results

We first assess the results of our experiment against our eight propositions using the paired one-tailed t-test results summarized in Table 3. We then tabulate the number of subjects whose behaviors align with model predictions in Table 4. Although the two methods evaluate subject behavior in different ways, they arrive at the same conclusion. Table 3 compares the likelihood that subjects will give at two different nodes, while Table 4 focuses on the number of subjects that choose a given set of strategies. Both tables demonstrate that altruism, reciprocity, and inequity aversion are necessary to rationalize pay-it-forward behavior.

In Table 3, panel a uses the full sample of 403 subjects and panel b restricts the sample to accurate responders. We first compare the actions of last movers (P1) across games in Section 4.1, which establishes the generalized reciprocity effect and points to the need for both altruism and reciprocity incentives in explaining giving behavior. We then compare the actions of first movers (P1 in the control game and P2 in the treatment games) in Section 4.2, which establishes inequity aversion as a necessary psychological component to align theoretical predictions with experimental results. In particular, we find that the possibility P1 would magnify P2's impact by paying P2's generosity forward does *not* increase P2's giving likelihood. Rather, the inability to ensure equal payoffs for all players lowers P2's giving likelihood relative to the P1's actions in the control game.

4.1 Giving rates by last movers

Proposition 1 compares P1's giving in the nonexclusive game after P2 gives versus after P2 keeps. Across the two samples, P1's giving likelihoods are 54.1%-56.2% after P2 gives and 19.8-23.3% after P2 keeps, with the difference significant at the 0.01% level. This behavior is consistent with models that account for altruism and inequity aversion (AI), altruism and reciprocity (AR), and altruism, inequity aversion, and reciprocity (AIR). Altruism is necessary to explain why P1 has a positive probability of giving even when P2 keeps $(\widehat{\gamma}_{1K}^n > 0)$, but is alone insufficient to explain why P1 would give more after P2 gave than after P2 kept $(\widehat{\gamma}_{1G}^n > \widehat{\gamma}_{1K}^n)$. Altruism with reciprocity incentives provides one explanation for this behavior, since reciprocity incentives make P1 more likely to give after P2 gave and to keep after P2 kept. Another explanation is that inequity aversion leads P1 to experience disutility from unequal payoffs. After P2 gives, P1 can equalize payoffs by giving to P0 such that all players earn \$2. If P2 kept, however, nothing P1 chooses will equalize payoffs. P1 would therefore be more likely to give after P2 gave than after P2 kept. Proposition 1 shows the necessary role of altruism in explaining P1's behavior in the nonexclusive game, but cannot distinguish whether reciprocity incentives, inequity aversion, or both contribute to why P1's giving likelihood is greater after P2 gave than after P2 kept.

Propositions 2 and 3 establish the generalized reciprocity effect. They compare P1's behavior in the control game versus the treatment games after P2 gave. We find that P1's giving rates are 52.2-53.1% in the exclusive game and 54.1-56.2% in the nonexclusive game after P2 gave. Both values are significantly greater than P1's giving rate of 44.7-45.2% in the control game (p < 0.01 across both samples). The pattern of giving establishes that reciprocity incentives are necessary in explaining behavior in our games (Models R, AR, IR, and AIR), since it contradicts the predictions of all models that exclude reciprocity. In all cases, P1 chooses between payoffs of (2,2,2) if she were to give and (2,3,0) if she were to keep. The game design holds constant social concerns, the number of players behind P1, P1's own income, and the relative payoffs across all players. The only difference between the treatment and control conditions is that P1's endowment is partly attributable to P2's kindness, rather than experimental conditions. Receiving the gift increases

P1's giving likelihood by 15-21 percent, indicating that benefiting from another person's kindness make subjects more likely to pay it forward.

Proposition 4 compares P1's giving in the control game with P1's giving in the nonexclusive game if P2 were to keep. Empirical giving rates are 21-25 percentage points higher in the control game (44.7-45.4%) than in the nonexclusive game case after P2 kept (19.8-23.3%, p < 0.0001 across all three samples). This result aligns with all models that incorporate either reciprocity or inequity aversion (all models except S and A). Under the reciprocity explanation, like actions beget like actions, so P1 is more inclined to keep her chip if P2 kept hers. Under the inequity aversion explanation, P1 can equalize payoffs in the control game by giving her chip to P0 but not in the nonexclusive game if P2 kept her chip.

Proposition 5 examines how P1's behavior differs in the exclusive and nonexclusive games in the case in which P2 gave. Giving rates are 54.1-56.2% in the nonexclusive game and 52.2-53.1% in the exclusive game. The difference is insignificant at the 5% level, as predicted by all models we consider except for the one where subjects only care about own material payoffs.

Comparing last mover's decisions leads us to reject all models except AR and AIR, since reciprocity is necessary to explain why P1 gives more in the treatment games after P2 gives than in the control game while altruism is necessary to explain why P1 would give after P2 kept in the nonexclusive game. To evaluate the AR and AIR models, we next turn to examining first movers' decisions by comparing P2 in the treatment games and P1 in the control game.

4.2 Giving rates by first movers

Proposition 6 compares P2's giving in the exclusive and nonexclusive games. Results differ between the full sample and the accurate responders sample. The full sample results can only be explained by models that we have already rejected, while the results from the accurate responders sample can be explained by models that we have not yet rejected. We therefore prefer the accurate responders sample to the full sample, since they provide a more coherent system for evaluating the role of psychological components.

More specifically, in the full sample, giving rates are not statistically different, with P2's giving propensity equal to 40.2% in the exclusive game and 39.0% in the nonexclusive game. These results only align with the inequity aversion model (I) and the inequity aversion and reciprocity model (IR). We have already rejected both models, since they cannot explain why P1 would give after P2 kept (Proposition 1). In contrast, in the sample of accurate responders in panel b, P2 is more likely to give in the exclusive game compared with the nonexclusive game (44.1% versus 40.7%, p < 0.05), which is consistent with all models that incorporate altruism. Intuitively, P2 knows that keeping in the exclusive game shuts down giving by P1, and therefore definitely harms P0. However, in the nonexclusive game P1 can theoretically still give even if P2 were to keep, leading to less expected harm to P0. If P2 had altruistic concerns for P1 and P0, she should be more likely to give in the exclusive game than in the nonexclusive game.

Finally, Propositions 7 and 8 compare P1's giving in the control group with P2's giving in the treatment groups. We find significantly greater giving by P1 in the control game than P2 in the treatment games (p < 0.05). In the control game, P1's giving rate is 44.7-45.4%. In the exclusive and nonexclusive games, P2's giving rates are 40.2-44.1% and 39.0-40.7%, respectively. The experimental results go against the predictions of all models that do not incorporate inequity aversion. Propositions 7 and 8 thus show that inequity aversion is necessary to align theoretical predictions with the behavior of first movers. Intuitively, P1 in the control game knows that by giving, she can equalize everyone's payoffs. However, P2 in the treatment games cannot equalize everyone's payoffs, since she cannot control what P1 will do. This difference appears to encourage greater giving in the former case.

4.3 Summary of t-test results

Table 3 reports how aggregate giving rates align with theoretical predictions. Reciprocity motives explain why P1 is more likely to give to P0 after receiving a gift from P2, controlling for income effects, distributional payoffs across all players, social concerns, and the number of players behind P1 in the chain (Propositions 2 and 3). However, knowing that P1 may pay forward P2's generosity

does not increase P2's chances of giving. Rather, inequity aversion explains why P2's giving rates are lower in the treatment games than P1's giving in the control game (Propositions 7 and 8). Lastly, altruism alone can explain why P1 would give in the nonexclusive game even after P2 kept (Propositions 1 and 4). Thus, the model which incorporates altruism, inequity aversion, and reciprocity (AIR) best rationalizes behavior for both samples.

It performs especially well in predicting behavior in the accurate responders sample, which excludes subjects who failed to demonstrate comprehension of the games. We next show that these results hold for various other restrictions on our sample.

4.4 Robustness checks

We conduct a number of robustness checks by examining results across different samples. Table 3b shows results for subjects who answered at least half of the accuracy check questions correctly on the first try. We next present t-test results for subjects who answered one or fewer questions incorrectly on the first try. The results, presented in Appendix Table B1a, show that all findings hold even though the sample reduces to only 227 subjects.

As a related robustness check, in Appendix Table B1b we restrict the sample to subjects who submitted six or fewer incorrect *answers*. The number of incorrect answers differs from the number of questions answered incorrectly, since a subject can submit multiple incorrect answers to the same question. For example, someone who submits four incorrect answers to one question but answers every other question correctly on the first try would count as having one question answered incorrectly and four incorrect answers. Again, we find that this alternative restriction does not change our main results in Table 3b.

Next, we explore whether time spent on the study affects results. We designed the study to be comfortably completed within an hour, and almost all subjects finished within 45 minutes. Excluding the 30 subjects who took more than 45 minutes from the sample of accurate responders brings the sample to 298. Despite the small sample size, our results do not change, as shown in Table B1c.

Lastly, we investigate whether the order of games affects behavior. After subjects played the control game, game order was randomized. Roughly half of the subjects were shown the nonexclusive game and then the exclusive game, and the other half played the games in the opposite order. Separately assessing the results based on order of games cuts the sample in half, yet our findings barely change. Table B1d shows results for accurate responders who saw the nonexclusive game first, while Table B1e shows results for accurate responders who saw the exclusive game first. Among the 156 subjects who saw the nonexclusive game first, empirical results are the same as in Table 3b.

Among the 168 subjects who saw the exclusive game first, all propositions still align with the AIR model predictions, but the comparisons regarding Proposition 6 become directional. Despite this, the AIR model still most consistently describes overall giving strategies across all subsamples. To see this, recall that Proposition 6 compares P2's giving rate in the exclusive and nonexclusive games. Among the accurate responders who saw the exclusive game first, we find a directional but insignificant effect. Only the inequity aversion (I) model and inequity aversion with reciprocity (IR) model would predict no significant differences in this comparison, but these models do not incorporate altruism and cannot explain why P1 would give after P2 keeps in the nonexclusive game (Propositions 1 and 4). They perform worse than the AIR model in rationalizing the totality of subjects' strategies across all nodes of our games.

4.5 How well does each model explain subject behavior?

In this section, we tabulate the number of subjects whose choices are consistent with each model's predictions. This alternate way of assessing model performance has two advantages over the paired t-tests in Table 3. First, it better leverages within-subject variation by counting the number of subjects that choose a given strategy, rather than computing aggregate giving likelihoods at each node. It generates slightly different results from Table 3 for Propositions 7 and 8, which depend on subject-level altruism and inequity aversion parameters (A, α, β) . Second, it allows us to quantify the importance of each psychological component in explaining empirical choices. We

find that altruism is most important, reciprocity second most important, and inequity aversion least important in rationalizing our experimental results.

Recall that Table 1 predicts which strategies are permissible under each model by comparing giving behavior at the two specified decision nodes under the LHS and RHS columns. At these two decision nodes, subjects can choose among four strategies: (give, give), (give, keep), (keep, give), and (keep, keep). When the prediction is ~ 0 , the model predicts that giving rates at both decision nodes will be statistically indistinguishable from 0, so subjects should play (keep, keep). When the prediction is \sim , the model predicts equivalent actions at the two decision nodes: (give, give) or (keep, keep). Third, when the prediction is > 0, the model predicts that giving at the node under LHS would be strictly greater than giving at the node under RHS, which would be equivalent to 0. The only action that aligns with such a prediction is (give, keep). Lastly, when the prediction is >, the model predicts greater giving rates at the node under LHS than under RHS. This means subjects may give at both nodes, keep at both nodes, or give at the LHS node and keep at the RHS node. The only action inconsistent with the prediction of > is (keep, give).

The exceptions to this method are Propositions 5, 7, and 8. Proposition 5 predicts that P1's giving rate after P2 gives will not significantly differ between the exclusive and nonexclusive games. This does not restrict how P1's mixed strategy gets realized. Subjects who choose (give, keep) or (keep, give) do not definitively violate Proposition 5, since it is possible that they are indifferent between the two decisions and choose at random. Hence, (give, give), (keep, keep), (give, keep) and (keep, give) can all occur even when P1's giving rate is the same in the two games. Under any model incorporating inequity aversion, Propositions 7 and 8 depend on altruism and inequity aversion parameters (A, α, β) . Depending on the specific values (A, α, β) take on, (give, give), (keep, keep), (give, keep), and (keep, give) are all plausible strategies.

Table 4 tabulates the number of subjects whose strategies align with different model predictions. Columns (1)-(4) count the number of subjects by choice at the nodes under the LHS and RHS columns. For example, Proposition 1 generates predictions for P1's behavior in the nonexclusive game. 55 subjects chose to give if P2 gave and give if P2 kept. 146 subjects chose to keep if

P2 gave and keep if P2 kept. 163 subjects chose to give if P2 gave and keep if P2 kept. 39 subjects chose to keep if P2 gave and give if P2 kept. Subjects' strategies under Propositions 2-8 can be read in a similar fashion.

Columns (5)-(12) summarize model performance, measured in the proportion of subjects whose behaviors can be rationalized by model predictions. Model S (column 1), which assumes that subjects only care about material payoffs, can only explain 35-36% of decisions by P1s, since it predicts that P1 would always play (keep, keep).

Models that exclude altruism (I, R, and IR) fail to explain behavior for the majority of subjects. Only incorporating reciprocity (Model R, column 8), explains behavior for 19% subjects regarding Propositions 2 and 3, since it predicts that P1 must play (give, keep). If reciprocity alone motivated P1's behavior, she would always give after receiving a gift from P2 in the treatment games. She would never give in the control game, where she does not benefit from someone else. This prediction is inconsistent with the behaviors of the 80.6% of P1s who give in the control game or keep after receiving a gift from P2 in the treatment games. Models I and IR (columns 7 and 11) explain behavior for 31% of subjects with respect to Proposition 4, since they predict that subjects must play (give, keep): they give in the control game and keep in the nonexclusive game after P2 keeps. Without altruism, P1 would never give if P2 kept in the nonexclusive game. But, inequity aversion would lead P1s to give in the control game to equalize payoffs for all players. Models I and IR fail to explain the behaviors of the 61% of subjects who keep in the control game or give in the nonexclusive game after P2 keeps. Together, these results establish the importance of altruism in generalized reciprocity games.

Excluding reciprocity motives would also fail to explain behavior by a large fraction of subjects. The model with only altruism (Model A, column 6) can explain behavior for 50% of subjects with respect to Proposition 1, since it predicts that altruistic P1s would be equally likely to give independent of whether P2 gave or kept in the nonexclusive game. P1 would only play (give, give) or (keep, keep). Model A cannot explain behavior for the other 50% who choose to give at one node but to keep at the other. The model with altruism and inequity aversion (Model AI,

column 9) performs comparatively better, but only explains the decisions of 70% of subjects with respect to Propositions 2 and 3. It predicts that, absent reciprocity motives, giving rates by P1 in the control game and the treatment games after P2 gave should be equal. It cannot explain the behavior of the 30% of subjects that choose (give, keep) or (keep, give).

We are then left with the model with altruism and reciprocity (Model AR, column 10) and the model with altruism, reciprocity, and inequity aversion (Model AIR, column 12). The two models generate identical predictions for P1's behavior, but the AIR model performs slightly better in explaining 93% of P2's behavior. In contrast, the AR model explains 88-89% of P2's behavior with respect to Propositions 7 and 8. The AR model predicts that in the absence of inequity aversion, P2's giving in the treatment games should be greater than P1's giving in the control game, meaning subjects should only play (give, give), (keep, keep), and (give, keep). These predictions are at odds with the t-test results from Table 3, where aggregate giving rates are higher for P1 in the control game than P2 in the treatment games. They cannot explain the 10-12% of subjects that keep in the treatment games but give in the control game. Their behaviors are better rationalized under the AIR model, where inequity aversion can explain why they give in the control game but not in the treatment games given individual parameters (A, α, β) .

We now evaluate the relative importance of altruism, reciprocity, and inequity aversion in explaining our behavior. Taking into account all propositions, models without altruism can only explain 19.35-34.73% of subjects' behavior. Models without reciprocity can explain 30.77-69.73% of subjects' behavior, and models without inequity aversion can explain 19.35-87.59% of subjects' behavior. By comparison, the AIR model can explain behavior for 89.08% of subjects. This means that adding altruism to the IR model increases the proportion of subjects whose behavior can be rationalized by 89.08-30.77=58.31%, or 235 subjects. Adding reciprocity to the AI model increases predictive power from 69.73% to 89.08% of subjects, a gain of 19.35% or 78 subjects whose behavior is now rationalized. Adding inequity aversion to the AR model increases predictive power from 87.59% to 89.08%, a gain of 1.49% or 6 subjects. Altruism and reciprocity substantially increase model performance, while the improvement from incorporating inequity aversion appears

marginal.

Lastly, we note that different psychological components help us understand behavior for different players. Altruism and reciprocity are key for describing P1's behavior, and therefore explain why one might give to a third party after receiving a gift. Inequity aversion plays no role here, as the AR and AIR models perform equally well in rationalizing P1's behavior. Rather, inequity aversion helps explain P2's experimental behavior. One may anticipate that P1's payit-forward behavior would increase giving rates for P2 in the treatment games compared to P1 in the control game, since it would extend the impact of P2's generosity to benefit P0. However, we find the opposite, and inequity aversion helps explain why. Inequity aversion makes P2 less likely to give in the treatment games than P1 in the control game, since in the latter case P1 can equalize payoffs across all players while in the former case P2 cannot.

5 Conclusion

Our study evaluates the importance of reciprocity motives, altruism, and inequity aversion in explaining pay-it-forward behavior. We establish a psychological game-theoretic framework which formulates predictions for giving behavior under different models of prosocial behavior. We then test these predictions using a novel experiment that demonstrates the existence of generalized reciprocal exchange while controlling for alternate explanations such as income effects, relative payoffs, and social image considerations. The experimental design allows us to exploit within-subject variation when comparing across various game nodes. That is, by assuming that subjects have constant prosocial preferences across games, we isolate distinct patterns in giving behavior in different games and player roles. We find that reciprocity incentives are critical to explain the generalized reciprocity effect, where receiving a gift makes P1 more likely to give. However, this effect does not appear to encourage giving by P2, even though P2's generosity would have been magnified by P1's pay-it-forward behavior. Rather, inequity aversion makes P2 *less* likely to give, making the transmission of generosity unlikely to start in the first place.

Our findings address the question of how generosity spreads within communities, which has been documented by prior work (Fowler and Christakis, 2010) but is not well understood. We provide experimental evidence that people pay forward kind acts to unrelated others; namely, that kindness engenders further kindness. Our results speak to why people help strangers with no expectation of meeting them again, why people choose to cooperate with co-workers even at some cost to themselves, and why some communities are cooperative while others are hostile or competitive.

References

- Bartlett, M. Y. and D. DeSteno (2006). Gratitude and prosocial behavior: Helping when it costs you. *Psychological Science* 17(4), 319–325.
- Battigalli, P., R. Corrao, and M. Dufwenberg (2019). Incorporating belief-dependent motivation in games. *Journal of Economic Behavior and Organization* 167, 185–218.
- Battigalli, P. and M. Dufwenberg (2009). Dynamic psychological games. *Journal of Economic Theory 144*(1), 1–35.
- Battigalli, P. and M. Dufwenberg (2021). Belief-dependent motivations and psychological game theory. forthcoming, Journal of Economic Literature.
- Ben-Ner, A., L. Putterman, F. Kong, and D. Magan (2004). Reciprocity in a two-part dictator game. *Journal of Economic Behavior and Organization* 53(3), 333–352.
- Berger, U. (2011). Learning to cooperate via indirect reciprocity. *Games and Economic Behavior 72*(1), 30–37.
- Binmore, K. G. (1994). Game theory and the social contract: playing fair, Volume 1. MIT press.
- Bolton, G. E., E. Katok, and A. Ockenfels (2005). Cooperation among strangers with limited information about reputation. *Journal of Public Economics 89*(8), 1457–1468. The Experimental Approaches to Public Economics.
- Bolton, G. E. and A. Ockenfels (2000). ERC: A theory of equity, reciprocity, and competition.

- American Economic Review 90(1), 166–193.
- Charness, G., N. Du, and C.-L. Yang (2011). Trust and trustworthiness reputations in an investment game. *Games and Economic Behavior 72*(2), 361–375.
- Charness, G. and M. Rabin (2002). Understanding social preferences with simple tests. *Quarterly Journal of Economics* 117(3), 817–869.
- Cox, J. C., D. Friedman, and S. Gjerstad (2007). A tractable model of reciprocity and fairness. *Games and Economic Behavior 59*(1), 17–45.
- Cox, J. C., D. Friedman, and V. Sadiraj (2008). Revealed altruism. *Econometrica* 76(1), 31–69.
- Desteno, D., M. Bartlett, J. Wormwood, L. Williams, and L. Dickens (2010). Gratitude as moral sentiment: Emotion-guided cooperation in economic exchange. *Emotion 10*, 289–93.
- Dufwenberg, M. and G. Kirchsteiger (2004). A theory of sequential reciprocity. *Games and Economic Behavior* 47(2), 268–298.
- Ebrahimji, A. (2020, December). Over 900 cars paid for each other's meals at a Dairy Queen drive-thru in Minnesota. [Online; posted 09-December-2020].
- Engelmann, D. and U. Fischbacher (2009). Indirect reciprocity and strategic reputation building in an experimental helping game. *Games and Economic Behavior 67*, 399–407.
- Falk, A. and U. Fischbacher (2006). A theory of reciprocity. *Games and Economic Behavior 54*(2), 293–315.
- Fehr, E. and S. Gächter (2000). Fairness and retaliation: The economics of reciprocity. *Journal of Economic Perspectives 14*(3), 159–181.
- Fehr, E. and K. M. Schmidt (1999). A theory of fairness, competition, and cooperation. *Quarterly Journal of Economics* 114(3), 817–868.
- Fowler, J. H. and N. A. Christakis (2010). Cooperative behavior cascades in human social networks. *Proceedings of the National Academy of Sciences* 107(12), 5334–5338.
- Gaudeul, A., C. Keser, and S. Müller (2021). The evolution of morals under indirect reciprocity. *Games and Economic Behavior 126*, 251–277.
- Gong, B. and C.-L. Yang (2019). Cooperation through indirect reciprocity: The impact of higher-

- order history. Games and Economic Behavior 118, 316-341.
- Gray, K., A. F. Ward, and M. I. Norton (2014). Paying it forward: Generalized reciprocity and the limits of generosity. *Journal of Experimental Psychology: General* 143(1), 247.
- Gul, F. and W. Pesendorfer (2016). Interdependent preference models as a theory of intentions. *Journal of Economic Theory 165*, 179–208.
- Heller, Y. and E. Mohlin (2017). Observations on cooperation. *Review of Economic Studies 85*(4), 2253–2282.
- Herne, K., O. Lappalainen, and E. Kestilä-Kekkonen (2013). Experimental comparison of direct, general, and indirect reciprocity. *The Journal of Socio-Economics* 45, 38–46.
- Hu, Y., J. Ma, Z. Luan, J. S. Dubas, and J. Xi (2019). Adolescent indirect reciprocity: Evidence from incentivized economic paradigms. *Journal of Adolescence* 74, 221–228.
- Iwagami, A. and N. Masuda (2010). Upstream reciprocity in heterogeneous networks. *Journal of Theoretical Biology 265*(3), 297–305.
- Jiang, L. and J. Wu (2019). Belief-updating rule and sequential reciprocity. *Games and Economic Behavior 113*, 770–780.
- Khadjavi, M. (2017). Indirect reciprocity and charitable giving: Evidence from a field experiment. *Management Science 63*(11), 3708–3717.
- Levine, D. K. (1998). Modeling altruism and spitefulness in experiments. *Review of economic dynamics* 1(3), 593–622.
- Malmendier, U., V. L. te Velde, and R. A. Weber (2014). Rethinking reciprocity. *Annual Review of Economics* 6(1), 849–874.
- McCullough, M. E., M. B. Kimeldorf, and A. D. Cohen (2008). An adaptation for altruism: The social causes, social effects, and social evolution of gratitude. *Current Directions in Psychological Science* 17(4), 281–285.
- Melamed, D., B. Simpson, and J. Abernathy (2020). The robustness of reciprocity: Experimental evidence that each form of reciprocity is robust to the presence of other forms of reciprocity. *Science* 6(23), eaba0504.

- Mujcic, R. and A. Leibbrandt (2018). Indirect reciprocity and prosocial behaviour: Evidence from a natural field experiment. *The Economic Journal* 128(611), 1683–1699.
- Nava, E., E. Croci, and C. Turati (2019). 'I see you sharing, thus I share with you': Indirect reciprocity in toddlers but not infants. *Palgrave Communications* 5(66), 1–9.
- Nowak, M. A. and K. Sigmund (1998a). The dynamics of indirect reciprocity. *Journal of Theoretical Biology* 194(4), 561–574.
- Nowak, M. A. and K. Sigmund (1998b). Evolution of indirect reciprocity by image scoring. *Nature 393*, 573–577.
- Ohtsuki, H. and Y. Iwasa (2006). The leading eight: Social norms that can maintain cooperation by indirect reciprocity. *Journal of Theoretical Biology* 239(4), 435–444.
- Ong, D. and H. H. Lin (2011). Deserving altruism: An experiment in pure indirect reciprocity.

 Mimeo.
- Rabin, M. (1993). Incorporating fairness into game theory and economics. *American Economic Review 83*(5), 1281–1302.
- Seinen, I. and A. Schram (2006). Social status and group norms: Indirect reciprocity in a repeated helping experiment. *European Economic Review 50*(3), 581–602.
- Simpson, B., A. Harrell, D. Melamed, N. Heiserman, and D. V. Negraia (2018). The roots of reciprocity: Gratitude and reputation in generalized exchange systems. *American Sociological Review 83*(1), 88–110.
- Sobel, J. (2005). Interdependent preferences and reciprocity. *Journal of Economic Literature 43*(2), 392–436.
- Takahashi, N. (2000). The emergence of generalized exchange. *American Journal of Sociology* 105(4), 1105–1134.
- Tsvetkova, M. and M. W. Macy (2014). The social contagion of generosity. *PLOS One 9*(2), e87275. van Apeldoorn, J. and A. Schram (2016). Indirect reciprocity: A field experiment. *PLOS One 11*(4), e0152076.
- Wu, J. (2018). Indirect higher order beliefs and cooperation. Experimental Economics 21(4), 858-

876.

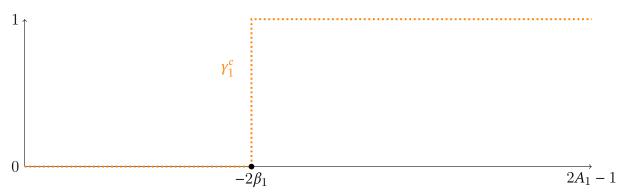
Yoeli, E., M. Hoffman, D. G. Rand, and M. A. Nowak (2013). Powering up with indirect reciprocity in a large-scale field experiment. *Proceedings of the National Academy of Sciences* 110(2), 10424–10429.

Zeckhauser, R., J. Swanson, and K. Lockwood (2006). The value of reputation on eBay: A controlled experiment. *Experimental Economics 9*, 79–101.

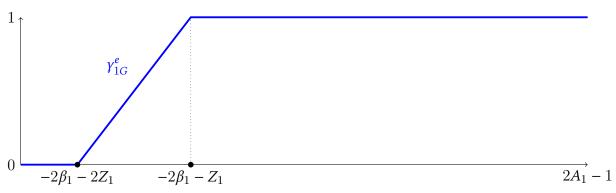
Figures & Tables

Figure 2: P1's equilibrium giving rates in different games

(a) P1's equilibrium giving rate γ_1^c in the control game



(b) P1's equilibrium giving rate γ_{1G}^e in the exclusive game



(c) P1's equilibrium giving rates γ_{1G}^n and γ_{1K}^n in the nonexclusive game

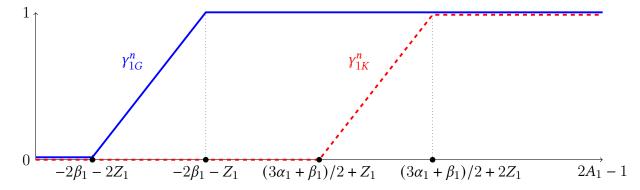


Figure 3: P1's equilibrium giving rates comparisons

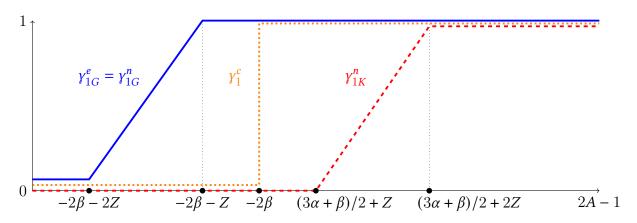


Figure 4: Initial movers' equilibrium giving rates

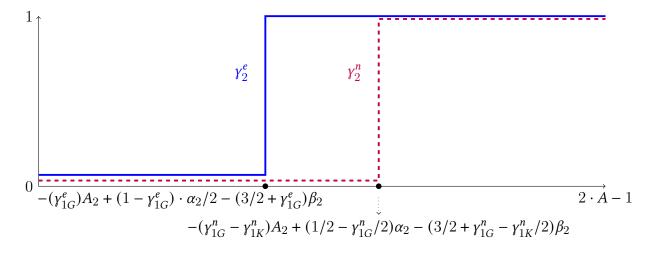


Table 1: Predicted comparisons of giving under different utilities

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)		
Mo	S	A	I	R	ΑI	AR	IR	AIR		
\overline{A}	= 0	> 0	= 0	= 0	> 0	> 0	= 0	> 0		
α_i ,	=0	=0	> 0	= 0	> 0	=0	> 0	> 0		
Z	=0	=0	=0	> 0	=0	> 0	> 0	> 0		
Proposition	LHS	RHS	Predictions							
1	γ_{1G}^n	γ_{1K}^n	~ 0	~	> 0	> 0	>	>	> 0	>
2	γ_{1G}^{e}	γ_1^c	~ 0	~	~	> 0	~	>	>	>
3	γ_{1G}^n	$\gamma_1^{\bar{c}}$	~ 0	~	~	> 0	~	>	>	>
4	γ_1^c	γ_{1K}^n	~ 0	~	> 0	~ 0	>	>	> 0	>
5	γ_{1G}^n	γ_{1G}^e	~ 0	~	~	~	~	~	~	~
6	γ_2^e	γ_2^n	~ 0	>	~	~ 0	>	>	~	>
7	$\gamma_2^{\overline{e}}$	$\gamma_1^{\overline{c}}$	~ 0	>	X	~ 0	X	>	X	X
8	$\gamma_2^{\bar{n}}$	γ_1^c	~ 0	~	X	~ 0	X	>	X	X

Note: Giving rate is denoted by γ . The superscript denotes game type, where c stands for control, e for exclusive, and n for nonexclusive. The subscript G stands for P1's decision after P2 gives, and K for P1's decision after P2 keeps. ~ 0 indicates that LHS \sim RHS \sim 0; \sim indicates that LHS \sim RHS \sim 0; and K indicates that the predictions are ambiguous and depend on the psychological parameters.

Table 2: Summary statistics of study subjects

	Full Sample	Accurate Responders
	Study Characteristics	
% saw exclusive game first	0.522	0.519
	(0.0249)	(0.0278)
study duration (minutes)	28.73	29.02
	(0.543)	(0.630)
median study duration (minutes)	26.75	26.77
bonus payment	2.553	2.731
	(0.0692)	(0.0768)
median bonus payment	3	3
wrong answers	2.157	1.324
	(0.117)	(0.0878)
median wrong answers	1	1
	Demographics	
% female	0.275	0.275
	(0.0223)	(0.0248)
% college graduate	0.829	0.802
	(0.0188)	(0.0222)
% employed	0.931	0.917
	(0.0127)	(0.0154))
Citizenship/residency/language flue		
% US citizen	0.684	0.673
	(0.0234)	(0.0263)
% native English speaker	0.763	0.755
	(0.0213)	(0.0240)
% US resident	0.727	0.701
	(0.0222)	(0.0255)
Race/ethnicity	, ,	,
% Black	0.129	0.0988
	(0.0167)	(0.0166)
% Asian	0.293	0.306
	(0.0227)	(0.0256)
% Hispanic	0.0496	0.0494
-	(0.0108)	(0.0121)
% White	0.501	0.515
	(0.0249)	(0.0278))
% Other race/ethnicity	0.0273	0.0309
•	(0.00813)	(0.00962)
Age	, ,	, ,
% 16-25 years old	0.159	0.160
·	(0.0182)	(0.0204)
% 26-35 years old	0.496	0.491
•	(0.0249)	(0.0278)
% 36-45 years old	0.223	0.219
•	(0.0208)	(0.0230)
% 46-55 years old	0.0670	0.0648
- y	(0.0125)	(0.0137)
% 56-65 years old	0.0422	0.0494
	(0.0100)	(0.0121)
% 65 or older	0.0124	0.0154
of the state of th	(0.00552)	(0.00686)
Observations	403	324

Notes: Standard errors in parentheses.

Table 3: Experimental results of giving rates comparisons

(a) Full sample, N = 403

Prop	Experimental	Consistent with predictions?								
	mean (standard error)	p-value (t-stat)	S	A	I	R	AI	AR	IR	AIR
1	$\widehat{\gamma}_{1G}^n = .541(0.025)$ > $\widehat{\gamma}_{1K}^n = .233(0.021)$	p < 0.0001 (9.6753)	No	No	No	No	Yes	Yes	No	Yes
2	$\widehat{\gamma}_{1G}^{e} = .531(0.025)$ > $\widehat{\gamma}_{1}^{c} = .447(0.025)$	p = 0.0010 (3.1112)	No	No	No	No	No	Yes	Yes	Yes
3	$\hat{\gamma}_{1G}^{n_1} = .541(0.025)$ > $\hat{\gamma}_{1G}^{c} = .447(0.025)$	p = 0.0002 (3.5481)	No	No	No	No	No	Yes	Yes	Yes
4	$\widehat{\gamma}_{1}^{c} = .447(0.025)$ > $\widehat{\gamma}_{1K}^{n} = .233(0.021)$	$p < 0.0001 \\ (-7.1665)$	No	No	No	No	Yes	Yes	No	Yes
5	$\widehat{\gamma}_{1G}^{nH} = .541(0.025)$ $\sim \widehat{\gamma}_{1G}^{e} = .531(0.025)$	p = 0.3402 (0.4121)	No	Yes						
6	$\widehat{\gamma}_2^e = .402(0.024)$ $\sim \widehat{\gamma}_2^n = .390(0.024)$	$p = 0.2542 \\ (-0.6618)$	No	D	Yes	D	D	D	Yes	D
7	$\widehat{\gamma}_2^e = .402(0.024)$ $< \widehat{\gamma}_1^c = .447(0.025)$	p = 0.0156 (-2.1612)	No	No	Yes	No	Yes	No	Yes	Yes
8	$\widehat{\gamma}_2^n = .390(0.024)$ $< \widehat{\gamma}_1^c = .447(0.025)$	$p = 0.0043 \\ (-2.6404)$	No	No	Yes	No	Yes	No	Yes	Yes

Note: D indicates that the experimental results are directionally but not significantly consistent with predictions.

(b) Accurate responders, N = 324

Prop	Experimental	Consistent with predictions?								
	mean (standard error)	p-value (t-stat)	S	A	I	R	ΑI	AR	IR	AIR
1	$\widehat{\gamma}_{1G}^n = .562(0.028)$ > $\widehat{\gamma}_{1K}^n = .198(0.022)$	p < 0.0001 (10.7992)	No	No	No	No	Yes	Yes	No	Yes
2	$\widehat{\gamma}_{1G}^e = .522(0.028) > \widehat{\gamma}_1^c = .454(0.028)$	p = 0.0057 (2.5448)	No	No	No	No	No	Yes	Yes	Yes
3	$ \widehat{\gamma}_{1G}^n = .562(0.028) $ $ > \widehat{\gamma}_1^c = .454(0.028) $	p = 0.0001 (3.9263)	No	No	No	No	No	Yes	Yes	Yes
4	$ \widehat{\gamma}_{1K}^n = .198(0.022) $ $ < \widehat{\gamma}_1^c = .454(0.028) $	$p < 0.0001 \\ (-8.0592)$	No	No	No	No	Yes	Yes	No	Yes
5	$\widehat{\gamma}_{1G}^n = .562(0.028)$ > $\widehat{\gamma}_{1G}^e = .522(0.028)$	p = 0.0562 (1.5920)	No	Yes						
6	$\widehat{\gamma}_2^e = .441(0.028)$ > $\widehat{\gamma}_2^n = .407(0.027)$	$p = 0.0429 \\ (-1.7231)$	No	Yes	No	No	Yes	Yes	No	Yes
7	$ \widehat{\gamma}_2^e = .441(0.028) $ $ \sim \widehat{\gamma}_1^c = .454(0.028) $	$p = 0.2781 \\ (-0.5892)$	No	No	Yes	No	Yes	No	Yes	Yes
8	$\widehat{\gamma}_2^n = .407(0.027)$ $< \widehat{\gamma}_1^c = .454(0.028)$	$p = 0.0160 \\ (-2.1549)$	No	No	Yes	No	Yes	No	Yes	Yes

Table 4: Comparing empirical strategies to model predictions (N = 403)

		T									
		(1) (2) (3) (4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	
Prop	LHS,RHS	GG,GK,KG,KK	S	A	I	R	AI	AR	IR	AIR	
		strategies	subjects with consistent behavior								
		(p-value)		(%)							
1	$\gamma_{1G}^n \ \gamma_{1K}^n$	55 163 39 146	146	201	163	163	364	364	63	364	
	10 11		(36.23%)	(49.88%)	(40.45%)	(40.45%)	(90.32%)	(90.32%)	(40.45%)	(90.32%)	
2	$\gamma_{1G}^e \gamma_1^c$	136 78 44 145	145	281	281	78	281	359	359	359	
			(35.98%)	(69.73%)	(69.73%)	(19.35%)	(69.73%)	(89.08%)	(89.08%)	(89.08%)	
3	$\gamma_{1G}^n \gamma_1^c$	140 78 40 145	145	285	285	78	285	363	363	363	
			(35.98%)	(70.72%)	(70.72%)	(19.35%)	(70.72%)	(90.07%)	(90.07%)	(90.07%)	
	c n						0.4				
4	$\gamma_1^c \gamma_{1K}^n$	56 124 38 185	185	241	124 %	185	365	365	124	365	
			(45.91%)	(59.80%)	(30.77 %)	(45.91%)	(90.57%)	(90.57%)	(30.77%)	(90.57%)	
5	ne	169 49 45 140	140	309	403	403^{a}	309	403^{a}	403^{a}	403^{a}	
3	$\gamma_{1G}^n \ \gamma_{1G}^e$	169 49 45 140									
			(34.73%)	(76.67%)	(76.67%)	(100%)	(76.67%)	(100%)	(100%)	(100%)	
6	$\gamma_2^e \gamma_2^n$	131 31 26 215	215	377	346	215	377	377	346	377	
O	12 12	151 51 20 215	(53.35%)	(93.55%)	(85.86%)	(53.35%)	(93.55%)	(93.55%)		(93.55%)	
			(33.3370)	(73.3370)	(03.0070)	(33.3370)	(73.3370)	(73.3370)	(03.0070)	(73.3370)	
7	$\gamma_2^e \gamma_1^c$	136 26 44 197	197	359	377^{b}	197	377^{b}	359^{b}	377^{b}	377^{b}	
	02 01		(48.88%)	(89.08%)	(93.55%)	(48.88%)	(93.55%)	(89.08%)	(93.55%)	(93.55%)	
			()	()	(212212)	()	(2122.0)	()	(2.22.0)	(2.22.3)	
8	$\gamma_2^n \gamma_1^c$	130 27 50 196	196	353	376 ^b	196	376^{b}	353	376 ^b	376 ^b	
	. 2 • 1		(48.64%)	(87.59%)	(93.30%)	(48.64%)	(93.30%)	(87.59%)	(93.30%)	(93.30%)	
	1									' 	

Columns (1)-(4) tabulate subjects that choose give or keep at the two nodes corresponding to (LHS, RHS). P-values from a one-sided Fisher's exact test for all propositions except Proposition 5, which reports results from a two-sided test given the AIR model predictions from Table 1. Columns (5)-(12) tabulate the number of subjects whose strategies are consistent with predictions under each model we consider.

^aPlayers can play mixed strategies, so all strategy combinations can comply with the predictions.

 $[^]b$ Violators play either GK or KG. We report the lower percentage of violators in the table.

Online Appendix (not for publication)

A Theory

A.1 Games

To simplify the exposition of giving rates, we define the following notation.

Definition 4. For any real number $x \in \mathbb{R}$, define $[\![x]\!]$ to be 1 if x is bigger than 1, x if x is between 0 and 1, and 0 if x is smaller than 0. Mathematically, $[\![x]\!] \equiv \max\{0, \min\{1, x\}\}$.

A.1.1 The control game

First, consider the control game (Figure 1a).

P2 is endowed with 2 chips, P1 with 3 chips, and P0 with 0 chips. P2 cannot decide on anything in this game, and exists to keep relative payoffs similar to the treatment games. P1 can either keep all 3 chips so that P0 has 0 chips, or pass 1 chip to P0 so that P0 has 2 chips. P0 cannot decide on anything, and can only receive chips from P1.

Lemma 1. In the control game, P1 prefers giving if and only if $2 \cdot A_1 + 2 \cdot \beta_1 \ge 1$.

P1 can either keep 1 chip (so that P0 gains no chips) or give up 1 chip (so that P0 gains 2 chips). When P1 gives up one chip, the material payoffs change from $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$ to $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$. By giving, she lowers her material payoff by 1 unit, but increases her altruistic payoff by $2 \cdot A_1$ units as P0's payoff increases from 0 to 2. Moreover, because inequity aversion gives P1 disutility from having more chips than other players, she gains $2 \cdot \beta$ from giving and equalizing payoffs across all three players. Overall, the psychological gain of giving by P1 is

$$2 \cdot A_1 + 2 \cdot \beta_1. \tag{1}$$

Figure 2a depicts P1's equilibrium giving rate as A_1 varies. In equilibrium, there is no mixed strategy except for when $2 \cdot A_1 + 2 \cdot \beta_1 = 1$, so the equilibrium giving rate can be represented

by an indicator function: $\gamma_1^c = 1_{2 \cdot A_1 + 2 \cdot \beta_1 \ge 1}$. P1 is more inclined to give the higher her altruistic factor A_1 and the higher her advantageous inequity aversion β_1 (that is, the more she dislikes having more than other players). Pure altruism and/or advantageous inequity aversion—but not disadvantageous inequity aversion or reciprocity—helps rationalize giving by P1 in the control game.

A.1.2 Exclusive game

Second, consider a three-player game in which P0's channel of receiving chips is exclusive (Figure 1b). P2 is endowed with 3 chips, P1 with 1 chip, and P0 with 0. P2 can either keep all 3 chips or give 1 chip to P1 so that P1's chip count increases from 1 to 3. Only upon receiving additional chips can P1 choose to give. If P1 gives 1 chip, P0 gets 2 chips. If P1 keeps, P2 gets 0.

Lemma 2. In the exclusive game, P2 prefers giving if $1 + (1 - \gamma_{1G}^e) \cdot \alpha_2/2 \le (2 + \gamma_{1G}^e) \cdot A_2 + (3/2 + \gamma_{1G}^e) \cdot \beta_2$, and P1 gives with probability $\gamma_{1G}^e = [(2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2]$.

Compared to keeping, giving lowers P2's material payoff but increases her utility from altruism and from having an equitable distribution of payoffs among all players in the group. The left-hand side of the inequality in Lemma 2 represents the two ways P2 loses from giving. Compared to keeping, giving lowers P2's material payoff by 1. P2 may also suffer utility loss from disadvantageous inequity, since P1 may keep after she gives, making her material payoff lower than P1's. More precisely, if P1 keeps after P2 gives, P2 incurs utility loss from disadvantageous inequity of $(3-2) \cdot \alpha_2/2 = \alpha_2/2$. Because P1 keeps with probability $1 - \gamma_{1G}^e$ after P2 gives, giving would lower P2's expected utility by $(1 - \gamma_{1G}^e) \cdot \alpha_2/2$.

The right-hand side of the inequality represents the two ways P2 gains from giving. Giving increases P2's altruistic payoff by $A_2 \cdot (2 + \gamma_{1G}^e)$ in expectation. Giving also rectifies inequality aversion, in that P2 is less likely to have more than other players. If P2 keeps, she suffers disutility of $[(3-1)+(3-0)] \cdot \beta_2/2 = 5 \cdot \beta_2/2$ from advantageous inequity, since she will have higher

¹⁰When $2 \cdot A_1 + 2 \cdot β_1 = 1$, P1 is indifferent between giving and keeping, since the change in material payoffs balances out the change in psychological payoffs. In equilibrium, P1 can choose to give with any probability $γ_1^c ∈ [0, 1]$. Without loss of generality, we assume that P1 chooses to give.

payoffs compared to P1 and P0. If P2 gives, two scenarios can happen. If P1 gives, P2 suffers no inequity aversion since all players will have 2 chips. If P1 keeps (which happens with probability $1 - \gamma_{1G}^e$), then P2 suffers $(2 - 0) \cdot \beta_2/2 = \beta_2$ from getting more than P0. Hence, by giving, P2 gains in expectation $5 \cdot \beta_2/2 - (1 - \gamma_{1G}^e) \cdot \beta_2 = (3/2 + \gamma_{1G}^e) \cdot \beta_2$. In summary, altruism and inequity aversion motivate P2 to give.

Upon receiving 2 chips from P2, P1 faces the following trade-off. If P1 gives, P1 loses one unit of material payoff, but gains in the three psychological components. A 2-chip gain for P0 gives P1 an altruistic payoff gain of $2 \cdot A_1$ and $2 \cdot \beta_1$ from equalizing payoffs. Furthermore, P1's kind action of giving earns P1 a generalized reciprocity payoff of $Z_1 \cdot (2 - \gamma_{1G}'')$, where γ_{1G}'' is P1's belief of P2's belief of P1's probability of giving. Altogether, P1's psychological gain from giving after P2 gives is

$$2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \cdot (2 - \gamma_{1G}^{"}). \tag{2}$$

In equilibrium, P1's second-order belief must equate with her strategy ($\gamma_{1G}'' = \gamma_{1G}^e$). Hence, if $2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \ge 1$, then $\gamma_{1G}^e = 1$; if $2 \cdot A_1 + 2 \cdot \beta_1 + 2 \cdot Z_1 \le 1$, then $\gamma_{1G}^e = 0$. Otherwise, P1 gives with a probability strictly between 0 and 1 that makes her indifferent between giving and keeping: $\gamma_{1G}^e = (2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2$. P1's inclination to give increases with altruism A_1 , advantageous inequity aversion β_1 , and generalized reciprocity Z_1 . Figure 2b depicts how P1's equilibrium giving rate in the exclusive game varies with altruism A_1 .

A.1.3 Nonexclusive game

Finally, consider a three-player game in which P0's channel of receiving chips is nonexclusive (Figure 1c). P2 is endowed with 3 chips, P1 with 1, and P0 with 0. P2 can either keep all the chips so that P1's chip count remains unchanged, or give away 1 chip so that P1's chip count increases by 2. Regardless of P2's decision, P1 can keep all the chips or give away 1 chip so that P0's chip count increases by 2.

Lemma 3. In the nonexclusive game, P2 prefers giving if and only if $1 + (1/2 - \gamma_{1G}^n/2) \cdot \alpha_2 \le (2 + \gamma_{1G}^n - \gamma_{1K}^n) \cdot A_2 + (3/2 + \gamma_{1G}^n - \gamma_{1K}^n/2) \cdot \beta_2$, P1 gives with probability $\gamma_{1G}^n = [(2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2]$

after P2 gives, and P1 gives with probability $\gamma_{1K}^n = [(2 \cdot A_1 - 3 \cdot \alpha_1/2 - \beta_1/2 - 1)/Z_1 - 1]$ after P2 keeps.

P2's gains from giving are similar to those in the exclusive game. The only difference is that after P2 keeps, P1 gives 1 chip with probability γ_{1K}^n , and 1 new chip gets created from P1's gift to P0. P2 then gains A_2 from altruism and $\beta_2/2$ since payoffs become more equal after P0 gains 2 chips. Since the psychological penalty to keeping is less severe for P2 in the nonexclusive game, the net benefit of giving is smaller in the nonexclusive game than the exclusive game by $(A_2 + \beta_2/2) \cdot \gamma_{1K}^n$.

If P2 chooses to give, then P1 faces the same trade-off as in the exclusive game. Therefore, the equilibrium giving rate γ_{1G}^n in the nonexclusive game is characterized in the same way as in the exclusive game. If P2 chooses to keep, reciprocity motives will make P1 more inclined to keep. If P1 gives instead, her reciprocity motives will generate a utility loss of $Z_1 \cdot (2 - \gamma_{1K}^n)$. In addition, by giving, P1 changes the material payoffs from $(\pi_2, \pi_1, \pi_0) = (3, 1, 0)$ to (3, 0, 2), which results in $3 \cdot \alpha_1/2$ units increase in disadvantageous inequity aversion and β_1 units of increase in advantageous inequity aversion. Overall, the psychological gain of giving by P1 after P2 keeps is

$$2 \cdot A_1 - 3 \cdot \alpha_1 / 2 - \beta_1 / 2 - Z_1 \cdot (2 - \gamma_{1G}''). \tag{3}$$

Figure 2c depicts P1's giving rates in the nonexclusive game. If P1 gives after P2 keeps, then her decision can only be justified by altruism. Note that there is no simultaneous mixing of both giving decisions in equilibrium under any combination of parameters. This is because the conditions for indifference are different at the two decision nodes, when P1 is indifferent between giving and keeping at one decision node, she is not indifferent at the other.

A.2 Predictions

In this subsection, we predict giving rates under the general AIR utility ($A_i > 0$, $\alpha_i > 0$, $\beta_i > 0$, and $Z_i > 0$). Table 1 summarizes the predictions under different utility functions. Our results

center on giving rates comparisons, and we say that one is more inclined to give than another in the following sense.

Definition 5. Player *i* is *more* (*less*) *inclined* to take action *s* at node *H* than player *j* to take action *s'* at node *H'*, $\sigma_i(s|H) > (\text{resp.}, \prec) \sigma_j(s'|H')$, if for $\alpha_i = \alpha_j$ and $Z_i = Z_j$, $\sigma_i(s|H) \ge (\text{resp.}, \le) \sigma_i(s'|H')$ for all parameters, and the inequality holds strictly for some parameters.

A.2.1 Giving by last mover P1

P1 is the last mover in all three games. Figure 3 shows P1's equilibrium giving rates for different altruistic factors. The comparisons are unambiguous: $\gamma_{1G}^e \sim \gamma_{1G}^n > \gamma_1^c > \gamma_{1K}^n$. We discuss the pairwise comparisons of these giving decisions in the following five propositions.¹¹

First, compare P1's two giving decisions in the nonexclusive game, after P2 gives versus after P2 keeps. Regardless of P2's choice, P1 incurs the same material loss (1 unit) and altruistic gain (2 units) from giving. However, if P1 chooses to give after P2 keeps, she incurs a larger inequity aversion loss and reciprocity loss than if she chooses to give after P2 gives. The material loss is the same: 1 unit. The psychological gains are different, as characterized in Equations (2) and (3). The altruistic gain is the same, $2 \cdot A_1$, but both the inequity aversion payoff and reciprocity payoff are worse—in fact, negative—when P1 gives after P2 keeps.

Proposition 1. In the nonexclusive game, P1 is more inclined to give after P2 gives than after P2 keeps. That is, $\gamma_{1G}^n > \gamma_{1K}^n$.

Since there is no simultaneous mixed strategy in equilibrium (as argued in the previous section), Proposition 1 implies that when P1 chooses to give after P2 keeps, P1 must also give after P2 gives. When P1 chooses to keep after P2 gives, P1 must also keep after P2 keeps. It is possible that P1 gives after P2 gives and keeps after P2 keeps, but never possible for P1 to keep after P2 gives and give after P2 keeps. Any subject who does so violates theoretical predictions as long as subjects exhibit reciprocity motives or inequity aversion.

¹¹There are $4 \times 3/2 = 6$ different pairwise comparisons for the four decisions. We do not directly compare γ_{1G}^e and γ_{1K}^n , but discuss the other five pairwise comparisons in the five propositions.

Now compare the psychological gain of giving in the control game to that in the exclusive game. The choice for P1 is the same in terms of material payoffs: either (2, 2, 2) by giving or (2, 3, 0) by keeping. If subjects have reciprocity motives, a gift from P2 increases P1's giving rate in the exclusive game.

Proposition 2. Relative to the control game, P1 is more inclined to give in the exclusive game after P2 gives. That is, $\gamma_{1G}^e > \gamma_1^c$.

Similarly, a gift from P2 increases P1's giving rate in the nonexclusive game relative to the control game.

Proposition 3. Relative to the control game, P1 is more inclined to give in the nonexclusive game after P2 gives. That is, $\gamma_{1G}^n > \gamma_1^c$.

However, both reciprocity motives and inequity aversion would decrease P1's giving inclination after P2 keeps in the nonexclusive game.

Proposition 4. Relative to the control game, P1 is less inclined to give in the nonexclusive game after P2 keeps. That is, $\gamma_{1K}^n < \gamma_1^c$.

Finally, P1's inclination to give is the same in the exclusive and nonexclusive games after P2 gives, and this result holds for all utility preferences we consider.

Proposition 5. P1 is equally inclined to give after P2 gives in the nonexclusive game and the exclusive game. That is, $\gamma_{1G}^n \sim \gamma_{1G}^e$.

Note that it is possible that a player is indifferent between giving and keeping in equilibrium in the exclusive and nonexclusive games, because she mixes between giving and keeping in equilibrium. Hence, when subjects are observed to give in one game and keep in another, the difference in observed behavior neither validates nor invalidates the prediction that they are equally likely to give. We discuss this prediction further in Section 4.5.

A.2.2 Giving by initial movers

We compare the first movers in these games: P1 in the control game and P2 in the treatment games. To summarize, the initial mover's giving rate is higher in the exclusive game than in the nonexclusive game, but it is unclear whether the initial mover is more inclined to give in the control game than in the treatment games, since altruism pushes for greater giving while inequity aversion pushes for lower giving in the treatment games.

First, P2's incentives to give are greater in the exclusive than the nonexclusive game. In the exclusive game, P2 knows that keeping will prevent P1 from giving, while in the nonexclusive game P2 knows that P1 can still give even if she kept. In particular, knowing that P1 can give even after P2 keeps in the nonexclusive game will increase P2's expected utility from keeping by $\gamma_{1K}^n \cdot A_2$ from altruism and $\gamma_{1K}^n \cdot \beta_2/2$ from having more equal payoffs. Figure 4 depicts the comparison of giving rates for initial movers.

Proposition 6. P2 in the exclusive game is more inclined to give than P2 in the nonexclusive game. That is, $\gamma_2^e > \gamma_2^n$.

Next, compare the giving rates of P1 in the control game and P2 in the exclusive game. Since we compare the giving decisions for the same subject, we can assume that all psychological parameters are the same for the subject across games and player roles: $A_1 = A_2 \equiv A$, $\alpha_1 = \alpha_2 \equiv \alpha$, $\beta_1 = \beta_2 \equiv \beta$, and $Z_1 = Z_2 \equiv Z$. Here, altruism and inequity aversion might work in opposite directions. By giving, P1 in the control game gets an altruistic payoff of $2 \cdot A$, and P2 in the exclusive game gets an altruistic payoff of $(2 + \gamma_{1G}^e) \cdot A$, because P1 in the exclusive game generates additional $\gamma_{1G}^e \cdot A$ units of altruistic payoff for P2 by passing to P0. Therefore, by altruism alone, P1 in the control game would be less inclined to give than P2 in the exclusive game.

Inequity aversion could create the opposite effect from altruism. P2 does not have a sure chance of equalizing payoffs in the exclusive game since P1 may choose to keep after P2 gives, while in the control game P1 will certainly equalize payoffs by choosing to give. This uncertainty decreases expected utility from giving of P2 in the exclusive game compared to P1 in the control

game. Furthermore, if P1 keeps, P2's inequity aversion causes her to suffer disutility $(3-2) \cdot \alpha/2 = \alpha/2$ from having a lower payoff than P1. Since this occurs with probability $1 - \gamma_{1G}^e$ and results in material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$, the loss in expectation is $(1 - \gamma_{1G}^e) \cdot \alpha/2$.

Proposition 7. P1 in the control game is more inclined to give than P2 in the exclusive game, i.e., $\gamma_1^c > \gamma_2^e$ if and only if

The comparison between P1 in the control game and P2 in the nonexclusive game is similar to the logic above. A higher altruistic payoff incentivizes P2 in the nonexclusive game to give more, since her gift could improve payoffs for both P1 and P0. But, P1 can equalize payoffs with certainty in the control game whereas P2 cannot in the nonexclusive game. In summary, altruism alone would drive greater giving by P2 in the nonexclusive game, but inequity aversion would drive greater giving by P1 in the control game.

Proposition 8. P1 in the control game is more inclined to give than P2 in the nonexclusive game, i.e., $\gamma_1^c > \gamma_2^n$ if and only if $(1/2 - \gamma_{1G}^n/2) \cdot \alpha \ge (\gamma_{1G}^n - \gamma_{1K}^n) \cdot A + (\gamma_{1G}^n - \gamma_{1K}^n - 1/2) \cdot \beta$.

A.3 Omitted proofs

A.3.1 Remark on efficient strategies

For technical reasons, when we define kindness we ignore Pareto-inefficient strategies and focus on Pareto-efficient ones. In our experiments, players do not have inefficient strategies. For the sake of completeness for the theory and to be consistent with previous theories, we keep this assumption. Intuitively, a strategy is inefficient if another strategy provides (i) no lower material payoff for any player for any history of play and the subsequent choices of others and (ii) a strictly higher payoff for some player for some history of play and subsequent choices by the

others. Formally, player i's set of efficient strategies is

$$\begin{split} \Sigma_i^e := & \bigg\{ \sigma_i \in \Sigma_i \bigg| \nexists \widehat{\sigma}_i \in \Sigma_i \text{ such that } \forall h \in H, \sigma_{-i} \in \Sigma_{-i}, k \in N, \\ & \pi_k(\widehat{\sigma}_i(h), \sigma_{-i}(h)) \geq \pi_k(\sigma_i(h), \sigma_{-i}(h)) \quad \text{with strict inequality for some } (h, \sigma_{-i}, k) \bigg\}. \end{split}$$

A.3.2 Proof of the theorem on equilibrium existence

Proof of Theorem 1. Let $\Sigma_i(h)$ denote i's set of (potentially random) choices at history $h \in H$. For any $s \in \Sigma_i(h)$, let $\sigma_i(h, s)$ denote player i's strategy that specifies the choice s at h, but is the same as $\sigma_i(h)$ otherwise—i.e., at every history in $H \setminus \{h\}$. Define correspondence $\beta_{i,h} : \Sigma \to \Sigma_i(h)$ by

$$\beta_{i,h}(\sigma) = \operatorname*{arg\,max}_{x \in X_i(h)} u_i(\sigma_i(h,s), (\sigma_j(h), (\sigma_k(h))_{k \neq j})_{j \neq i}),$$

and define correspondence $\beta: \Sigma \to \Pi_{(i,h) \in N \times H} \Sigma_i(h)$ by

$$\beta(\sigma) = \prod_{(i,h)\in N\times H} \beta_{i,h}(\sigma).$$

The set $\prod_{(i,h)\in N\times H} \Sigma_i(h)$ is topologically equivalent to the set Σ , so $\beta:\Sigma\to \Pi_{(i,h)\in N\times H}\Sigma_i(h)$ is equivalent to a correspondence $\gamma:\Sigma\to\Sigma$ (which is a direct redefinition of β). Every fixed point of γ is an equilibrium. To see this, note that a fixed point $\beta_{i,h}$ satisfies utility maximization under consistent beliefs. Here, because $\beta_{i,h}$ specifies the optimal choices at each $h\in H$, altogether, $\beta_{i,h}$ specifies the optimal strategies in $\Sigma_i(h,s)$. Hence, β and γ are combined best-response correspondences. Since γ is a correspondence from Σ to Σ , it is amenable to fixed-point analysis.

It remains to show that γ possesses a fixed point. Berge's maximum principle guarantees that $\beta_{i,h}$ is nonempty, closed-valued, and upper hemicontinuous, since $\Sigma_i(h)$ is nonempty and compact and u_i is continuous (since π_i , κ_{ij} , and λ_{ijk} are all continuous). In addition, $\beta_{i,h}$ is convex-valued, since $\Sigma_i(h)$ is convex and u_i is linear—and hence quasiconcave—in i's own choice. Hence, $\beta_{i,h}$ is nonempty, closed-valued, upper hemicontinuous, and convex-valued. These properties extend to β and γ . Hence, it follows by Kakutani's fixed-point theorem that γ admits a fixed point. \square

A.3.3 Proof of lemmas on equilibrium giving strategy

Proof of Lemma 1. P1's choice is between giving, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$, and keeping, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$. P1's utility from giving is $u_1(g_1) = 2 + 4 \cdot A_1$. P1's utility from keeping is $u_1(k_1) = 3 + A_1 \cdot 2 - \beta_1 \cdot (3 - 2)/2 - \beta_1 \cdot (3 - 0)/2 = 3 + 2 \cdot A_1 - 2 \cdot \beta_1$. P1 prefers giving if and only if $u_1(g_1) = 2 + 4 \cdot A_1 \ge u_1(k_1) = 3 + 2 \cdot A_1 - 2 \cdot \beta_1$, that is, $2 \cdot A_1 + 2 \cdot \beta_1 \ge 1$.

Proof of Lemma 2. Suppose P1 believes that P2 believes that P1 gives with probability $\gamma_{1G}^{"}$. The equitable payoff of P1 is $(1 + 3 - \gamma_{1G}'')/2 = 2 - \gamma_{1G}''/2$, so giving by P2 to P1 shows a kindness of $3 - \gamma_{1G}'' - (2 - \gamma_{1G}''/2) = 1 - \gamma_{1G}''/2$. P1's utility from giving, which results in material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$, is $u_1(g_{1G}, \gamma_{1G}'') = 2 + 4 \cdot A_1 + Z_1 \cdot (+1) \cdot (1 - \gamma_{1G}''/2)$, and P1's utility from keeping, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$, is $u_1(k_{1G}, \gamma_{1G}'') = 3 + 2$. $A_1 - 2 \cdot \beta_1 + Z_1 \cdot (-1) \cdot (1 - \gamma_{1G}''/2)$. Therefore, P1's utility from giving with probability γ_{1G} is $u_1(\gamma_{1G},\gamma_{1G}'') = \gamma_{1G} \cdot \left[-1 + 2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \cdot (2 - \gamma_{1G}'')\right] + 3 + 2 \cdot A_1 - 2 \cdot \beta_1 - Z_1 \cdot (1 - \gamma_{1G}''/2).$ If $2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \ge 1$, then $\gamma_{1G} = 1$. If $2 \cdot A_1 + 2 \cdot \beta_1 + 2 \cdot Z_1 \le 1$, then $\gamma_{1G} = 0$. If $2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 < 1 < 2 \cdot A_1 + 2 \cdot \beta_1 + 2 \cdot Z_1, \text{ then } -1 + 2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \cdot (2 - \gamma_{1G}) = 0, \text{ which rearranges}$ to $\gamma_{1G} = 2 - (1 - 2 \cdot A_1 - 2 \cdot \beta_1)/Z_1$. Therefore, in equilibrium, $\gamma_{1G}^e = [2 + (2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1]$. Suppose P2 believes that P1 gives with probability γ'_{1G} . P2's expected utility from giving, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$ with probability γ'_{1G} and $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$ (2,3,0) with probability $1 - \gamma'_{1G}$, is $\gamma'_{1G} \cdot (2 + 4 \cdot A_2) + (1 - \gamma'_{1G}) \cdot (2 + 3 \cdot A_2 - \alpha_2/2 - \beta_2) =$ $2 + 4 \cdot A_2 - (1 - \gamma'_{1G}) \cdot (A_2 + \alpha_2/2 + \beta_2)$. P2's utility from keeping, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (3, 1, 0)$, is $3 + 1 \cdot A_2 - \beta_2 \cdot (3 - 1)/2 - \beta_2 \cdot (3 - 0)/2 = 3 + A_2 - 5 \cdot \beta_2/2$. P2 prefers giving if $2 + 4 \cdot A_2 - (1 - \gamma_{1G}') \cdot (A_2 + \alpha_2/2 + \beta_2) \ge 3 + A_2 - 5 \cdot \beta_2/2$, which is simplified to $3 \cdot A_2 + 5 \cdot \beta_2/2 - (1 - \gamma_{1G}') \cdot (A_2 + \alpha_2/2 + \beta_2) \ge 1$. In equilibrium, $\gamma_{1G}' = \gamma_{1G}$, so the inequality is rearranged to $(2 + \gamma_{1G}) \cdot A_2 - (1 - \gamma_{1G}) \cdot \alpha_2 / 2 + (3/2 + \gamma_{1G}) \cdot \beta_2 \ge 1$.

Proof of Lemma 3. Suppose P1 believes that P2 believes that P1 gives with probability γ_{1G}'' when P2 gives, and gives with probability γ_{1K}'' when P2 keeps. First, suppose P2 keeps. P0's equitable

payoff is 1, and P1's equitable payoff is $[(1 - \gamma_{1K}'') + (3 - \gamma_{1G}'')]/2 = 2 - \gamma_{1K}''/2 - \gamma_{1G}''/2$.

First, consider when P2 keeps. P1's utility from giving, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (3, 0, 2)$, is $u_1(k_2, g_{1K}, \cdots) = 0 + 5 \cdot A_1 - \alpha_1 \cdot (3 - 0)/2 - \alpha_1 \cdot (2 - 0)/2 + Z_1 \cdot (+1) \cdot (\gamma_{1G}''/2 - 1 - \gamma_{1K}''/2)) = 5 \cdot A_1 - 5 \cdot \alpha_1/2 + Z_1 \cdot (\gamma_{1G}''/2 - 1 - \gamma_{1K}''/2)$, and P1's utility from keeping, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (3, 1, 0)$, is $u_1(k_2, k_{1K}, \cdots) = 1 + 3 \cdot A_1 - \alpha_1 \cdot (3 - 1)/2 - \beta_1 \cdot (1 - 0)/2 - Z_1 \cdot (\gamma_{1G}''/2 - 1 - \gamma_{1K}''/2)) = 1 + 3 \cdot A_1 - 3 \cdot \alpha_1/2 - \beta_1/2 - Z_1 \cdot (\gamma_{1G}''/2 - 1 - \gamma_{1K}''/2)$. Fixing γ_{1G}'' and γ_{1K}'' , we have $u_1(k_2, g_{1K}, \cdots) - u_1(k_2, k_{1K}, \cdots) = -1 + 2 \cdot A_1 - 3 \cdot \alpha_1/2 - \beta_1/2 + 2 \cdot Z_1 \cdot (\gamma_{1G}''/2 - 1 - \gamma_{1K}''/2) = -1 + 2 \cdot A_1 - 3 \cdot \alpha_1/2 - \beta_1/2 - Z_1 \cdot (2 - \gamma_{1G}'' + \gamma_{1K}'')$.

Second, consider when P2 gives. Regarding the reciprocity payoff, the only change is in the flip of the sign of λ_{121} . P1's utility of giving, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$, is $u_1(g_2, g_{1G}, \dots) = 2 + 4 \cdot A_1 + Z_1 \cdot (1 - \gamma_{1G}''/2 + \gamma_{1K}''/2)$. P1's utility of keeping, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$, is $u_1(g_2, k_{1G}, \dots) = 3 + 2 \cdot A_1 - \beta_1 \cdot (3 - 2)/2 - \beta_1 \cdot (3 - 0)/2 - Z_1 \cdot (1 - \gamma_{1G}''/2 + \gamma_{1K}''/2) = 3 + 2 \cdot A_1 - 2 \cdot \beta_1 - Z_1 \cdot (1 - \gamma_{1G}''/2 + \gamma_{1K}''/2)$. Hence, $u_1(g_2, g_{1G}, \dots) - u_1(g_2, k_{1G}, \dots) = -1 + 2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \cdot (2 - \gamma_{1G}'' + \gamma_{1K}'')$.

Comparing the net benefit of giving after P2 gives and that after P2 keeps, we have

$$u_1(q_2, q_{1G}, \dots) - u_1(q_2, k_{1G}, \dots) \ge u_1(k_2, q_{1G}, \dots) - u_1(k_2, k_{1G}, \dots).$$

Hence, whenever P1 decides to give after P2 keeps, she will also choose to give after P2 gives. In other words, P1 is more inclined to give after P2 gives than after P2 keeps: $\gamma_{1G} \geq \gamma_{1K}$. Given this inequality, there are five possible cases regarding γ_{1G} and γ_{1K} .

- 1. Strategies $\gamma_{1G} = 1$ and $\gamma_{1K} = 1$ are supported in equilibrium when and only when $2 \cdot A_1 3 \cdot \alpha_1/2 \beta_1/2 2 \cdot Z_1 \ge 1$.
- 2. Strategies $\gamma_{1G} = 1$ and $0 < \gamma_{1K} < 1$ are supported in equilibrium when and only when $2 \cdot A_1 3 \cdot \alpha_1/2 \beta_1/2 2 \cdot Z_1 \le 1 \le 2 \cdot A_1 3 \cdot \alpha_1/2 \beta_1/2 Z_1$. In this case, $2 \cdot A_1 3 \cdot \alpha_1/2 \beta_1/2 Z_1 \cdot (1 + \gamma_{1K}) = 1$, which is rearranged to $\gamma_{1K} = (2 \cdot A_1 3 \cdot \alpha_1/2 \beta_1/2 1 Z_1)/Z_1$.

- 3. Strategies $\gamma_{1G} = 1$ and $\gamma_{1K} = 0$ are supported in equilibrium when and only when $2 \cdot A_1 3 \cdot \alpha_1/2 \beta_1/2 Z_1 \le 1 \le 2 \cdot A_1 + 2 \cdot \beta_1 + Z_1$.
- 4. Strategies $0 < \gamma_{1G} < 1$ and $\gamma_{1K} = 0$ are supported in equilibrium when and only when $2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \le 1 \le 2 \cdot A_1 + 2 \cdot \beta_1 + 2 \cdot Z_1$. In this case, $2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 \cdot (2 \gamma_{1G}) = 1$, which is rearranged to $\gamma_{1G} = 2 + (2 \cdot A_1 + 2 \cdot \beta_1 1)/Z_1$.
- 5. Strategies $\gamma_{1G} = 0$ and $\gamma_{1K} = 0$ are supported in equilibrium when and only when $2 \cdot A_1 + 2 \cdot \beta_1 + 2 \cdot Z_1 \le 1$.

In summary, in the nonexclusive game, P1 gives with probability $\gamma_{1G}^n = [(2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2]$ after P2 gives, and gives with probability $\gamma_{1K}^n = [(2 \cdot A_1 - \alpha_1 - \beta_1/2 - 1)/Z_1 - 1]$ after P2 keeps.

Consider P2's action next. Suppose P2 believes that P1 gives with probability γ'_{1G} and γ'_{1K} when P2 gives and keeps, respectively. P2's expected utility from giving, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (2, 2, 2)$ with probability γ'_{1G} and $(\pi_2, \pi_1, \pi_0) = (2, 3, 0)$ with probability $1 - \gamma'_{1G}$, is $(2+4\cdot A_2) \cdot \gamma'_{1G} + (2+3\cdot A_2 - \alpha_2/2 - \beta_2) \cdot (1-\gamma'_{1G}) = 2 + (3+\gamma'_{1G}) \cdot A_2 - (1-\gamma'_{1G}) \cdot (\alpha_2/2 + \beta_2)$. P2's expected utility from keeping, which yields material payoffs $(\pi_2, \pi_1, \pi_0) = (3, 0, 2)$ with probability γ'_{1K} and $(\pi_2, \pi_1, \pi_0) = (3, 1, 0)$ with probability $1 - \gamma'_{1K}$, is $\gamma'_{1K} \cdot [3+2\cdot A_2 - \beta_2 \cdot (3-0)/2 - \beta_2 \cdot (3-2)/2] + (1-\gamma'_{1K}) \cdot [3+A_2-\beta_2 \cdot (3-0)/2 - \beta_2 \cdot (3-1)/2] = 3 + (1+\gamma'_{1K}) \cdot A_2 - (5-\gamma'_{1K}) \cdot \beta_2/2$. P2 prefers giving if and only if $2 + (3+\gamma'_{1G}) \cdot A_2 - (1-\gamma'_{1G}) \cdot (\alpha_2/2 + \beta_2) \ge 3 + (1+\gamma'_{1K}) \cdot A_2 - (5-\gamma'_{1K}) \cdot \beta_2/2$, which, as $\gamma_{1G} = \gamma'_{1G}$ and $\gamma_{1K} = \gamma'_{1K}$ in equilibrium, is rearranged to $(2+\gamma_{1G}-\gamma_{1K}) \cdot A_2 - (1/2-\gamma_{1G}/2) \cdot \alpha_2 + (3/2+\gamma_{1G}-\gamma_{1K}/2) \cdot \beta_2 \ge 1$.

A.3.4 Proofs of propositions on equilibrium giving comparisons

Proof of Proposition 1. In the nonexclusive game, P1 gives with probability $\gamma_{1G}^n = [(2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2]$ after P2 gives, and P1 gives with probability $\gamma_{1K}^n = [(2 \cdot A_1 - 3 \cdot \alpha_1/2 - \beta_1/2 - 1)/Z_1 - 1]$ after P2 keeps. Since $(2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2 > (2 \cdot A_1 - 3 \cdot \alpha_1/2 - \beta_1/2 - 1)/Z_1 - 1$ for any combination of nonnegative parameters A_1 , α_1 , β_1 , and A_2 , A_1 , and the inequality is strict as long as $A_1 \neq 0$.

Proof of Proposition 2. Explicitly, P1's equilibrium probability of giving in the control game is

$$\gamma_1^c = \begin{cases} 1 & \text{if } 2 \cdot A_1 + 2 \cdot \beta_1 > 1, \\ 0 & \text{if } 2 \cdot A_1 + 2 \cdot \beta_1 < 1, \end{cases}$$

and P1's equilibrium probability of giving in the exclusive game is

$$\gamma_{1G}^{e} = \begin{cases} 1 & \text{if } 2 \cdot A_1 + 2 \cdot \beta_1 + Z_1 > 1, \\ (2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2 & \text{if } 2 \cdot A_1 + 2 \cdot \beta_1 + 2 \cdot Z_1 \ge 1 \ge 2 \cdot A_1 + 2 \cdot \beta_1 + Z_1, \\ 0 & \text{if } 2 \cdot A_1 + 2 \cdot \beta_1 + 2 \cdot Z_1 < 1. \end{cases}$$

When $Z_1=0$, the condition for $\gamma_{1G}^e=1$ and the condition for $\gamma_1^c=1$ coincide, and the condition for $\gamma_{1G}^e=0$ and the condition for $\gamma_{1G}^c=0$ also coincide. When $Z_1>0$, the set of parameters for $\gamma_{1G}^e=1$ is a strict superset of that for $\gamma_1^c=1$, and the set of parameters for $\gamma_{1G}^e=0$ is a strict subset of that for $\gamma_1^c=0$. For the set range of parameters for $0<\gamma_{1G}^e<1$, $\gamma_1^c=0$. Hence, $\gamma_{1G}^e\geq\gamma_1^c$ for any combination of parameters. Hence, $\gamma_{1G}^e>\gamma_1^c$

Proof of Proposition 3. The proof mimics the proof of Proposition 2, with superscripts e replaced by superscripts n. Alternatively, by Proposition 5, $\gamma_{1G}^e \sim \gamma_{1G}^n$, so by transitivity of the inclination, $\gamma_{1G}^n > \gamma_1^c$.

Proof of Proposition 4. Explicitly, P1's equilibrium probability of giving in the control game is

$$\gamma_1^c = \begin{cases} 1 & \text{if } 2 \cdot A_1 + 2 \cdot \beta_1 > 1, \\ 0 & \text{if } 2 \cdot A_1 + 2 \cdot \beta_1 < 1, \end{cases}$$

and P1's equilibrium probability of giving after P2 keeps in the nonexclusive game is

$$\gamma_{1K}^{n} = \begin{cases} 1 & \text{if } 2 \cdot A_{1} - 3 \cdot \alpha_{1}/2 - \beta_{1}/2 - 2 \cdot Z_{1} > 1, \\ (2 \cdot A_{1} - 3 \cdot \alpha_{1}/2 - \beta_{1}/2 - 1)/Z_{1} - 1 & \text{if } 2 \cdot A_{1} - 3 \cdot \alpha_{1}/2 - \beta_{1}/2 - 2 \cdot Z_{1} \leq 1 \\ & \leq 2 \cdot A_{1} - 3 \cdot \alpha_{1}/2 - \beta_{1}/2 - Z_{1} \\ 0 & \text{if } 2 \cdot A_{1} - 3 \cdot \alpha_{1}/2 - \beta_{1}/2 - Z_{1} < 1. \end{cases}$$

When $\alpha_1 = \beta_1 = Z_1 = 0$, the two decisions coincide. When $\alpha_1 > 0$, $\beta_1 > 0$, and/or $Z_1 > 0$, the set of parameters for $\gamma_1^c = 1$ is a strict superset of that for $\gamma_1^{1K} = 1$, and the set of parameters for $\gamma_1^c = 0$ is a strict subset of that for $\gamma_1^{1K} = 0$. Hence, $\gamma_1^c > \gamma_{1K}^n$.

Proof of Proposition 5. P1 gives with probability $\gamma_{1G}^e = [(2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2]$ after P2 gives in the exclusive game. Equally, P1 gives with probability $\gamma_{1G}^e = [(2 \cdot A_1 + 2 \cdot \beta_1 - 1)/Z_1 + 2]$ after P2 gives in the nonexclusive game. Hence, P1 is equally inclined to give in the two treatment games after P2 gives.

Proof of Proposition 6. As shown by the inequality condition in Lemma 2 that characterizes P2's preference for giving, P2's net benefit of giving over keeping in the exclusive game is $B^e \equiv (2 + \gamma_{1G}^e) \cdot A_2 - (1/2 - \gamma_{1G}^e/2) \cdot \alpha_2 + (3/2 + \gamma_{1G}^e) \cdot \beta_2 - 1$. Similarly, by the inequality condition in Lemma 3 that characterizes P2's preference for giving, P2's net benefit of giving over keeping in the nonexclusive game is $B^n \equiv (2 + \gamma_{1G}^n - \gamma_{1K}^n) \cdot A_2 - (1/2 - \gamma_{1G}^n/2) \cdot \alpha_2 + (3/2 + \gamma_{1G}^n - \gamma_{1K}^n) \cdot \beta_2 - 1$. By Proposition 5, $\gamma_{1K}^n = \gamma_{1G}^n$. Then, $B^e - B^n = \gamma_{1K}^n \cdot A_2 + \gamma_{1K}^n \cdot \beta_2$. Since $A_1 \ge 0$ and $\beta_1 \ge 0$ in the general AIR utility function, and $\gamma_{1K}^n > 0$ in equilibrium, $B^e - B^n \ge 0$. The higher net benefit of giving over keeping in the exclusive game implies a higher inclination of giving in the exclusive game than the nonexclusive game.

Proof of Proposition 7. By the inequality condition in Lemma 1 that characterizes P1's preference for giving, P1's net benefit of giving over keeping in the control game is $B^c \equiv 2 \cdot A_1 + 2 \cdot \beta_1 - 1$. As shown by the inequality condition in Lemma 2 that characterizes P2's preference for giving,

P2's net benefit of giving over keeping in the exclusive game is $B^e \equiv (2 + \gamma_{1G}^e) \cdot A_2 - (1/2 - \gamma_{1G}^e/2) \cdot \alpha_2 + (3/2 + \gamma_{1G}^e) \cdot \beta_2 - 1$. For the same subject, i.e., $A_1 = A_2 \equiv A$, $\alpha_1 = \alpha_2 \equiv \alpha$, $\beta_1 = \beta_2 \equiv \beta$, $Z_1 = Z_2 \equiv Z$, the difference in the net benefits is $B^e - B^c = \gamma_{1G}^e \cdot A - (1 - \gamma_{1G}^e) \cdot \alpha/2 + (\gamma_{1G}^e - 1/2) \cdot \beta$. Therefore, P1 in the control game is more inclined to give than P2 in the exclusive game, if and only if $B^e - B^c \leq 0$, that is, $\gamma_{1G}^e \cdot A + (\gamma_{1G}^e - 1/2) \cdot \beta \leq (1 - \gamma_{1G}^e) \cdot \alpha/2$.

Proof of Proposition 8. By the inequality condition in Lemma 1 that characterizes P1's preference for giving, P1's net benefit of giving over keeping in the control game is $B^c \equiv 2 \cdot A_1 + 2 \cdot \beta_1 - 1$. By the inequality condition in Lemma 3 that characterizes P2's preference for giving, P2's net benefit of giving over keeping in the nonexclusive game is $B^n \equiv (2 + \gamma_{1G}^n - \gamma_{1K}^n) \cdot A_2 - (1/2 - \gamma_{1G}^n/2) \cdot \alpha_2 + (3/2 + \gamma_{1G}^n - \gamma_{1K}^n) \cdot \beta_2 - 1$. For the same subject, i.e., $A_1 = A_2 \equiv A$, $\alpha_1 = \alpha_2 \equiv \alpha$, $\beta_1 = \beta_2 \equiv \beta$, $Z_1 = Z_2 \equiv Z$, the difference in the net benefits is $B^n - B^c = (\gamma_{1G}^n - \gamma_{1K}^n) \cdot A - (1/2 - \gamma_{1G}^n/2) \cdot \alpha + (\gamma_{1G}^n - \gamma_{1K}^n - 1/2) \cdot \beta$. Therefore, P1 in the control game is more inclined to give than P2 in the nonexclusive game, if and only if $B^n - B^c \leq 0$, that is, $(\gamma_{1G}^n - \gamma_{1K}^n) \cdot A + (\gamma_{1G}^n - \gamma_{1K}^n - 1/2) \cdot \beta \leq (1/2 - \gamma_{1G}^n/2) \cdot \alpha$.

B Additional experimental results

Table B1: Experimental results of giving rates comparisons, robustness checks

(a) ≤ 1 questions answered incorrectly (N=227)

Prop	Experimental	result		Consistent with predictions?								
	mean (standard error)	p-value (t-stat)	S	A	I	R	AI	AR	IR	AIR		
1	$\widehat{\gamma}_{1G}^n = .617(0.032)$	<i>p</i> < 0.0001	No	No	No	No	Yes	Yes	No	Yes		
1	$> \widehat{\gamma}_{1K}^n = .176(0.025)$	(12.1047)	110	110	110	110	103	105		ics		
2	$\widehat{\gamma}^e_{1G} = .568(0.033)$	p = 0.0055	No	No	No	No	No	Yes	Yes	Yes		
2	$> \widehat{\gamma}_1^c = .502(0.033)$	(2.5665)	110	110	110	110	110	105	105	165		
3	$\widehat{\gamma}_{1G}^n = .617(0.032)$	p < 0.0001	No	No	No	No	No	Yes	Yes	Yes		
Ü	$> \widehat{\gamma}_1^c = .502(0.033)$	(4.1530)	1,0	110						165		
4	$\hat{\gamma}_1^c = .502(0.033)$	p < 0.0001	No	No	No	No	Yes	Yes	No	Yes		
	$> \widehat{\gamma}_{1K}^n = .176(0.025)$	(-8.8330)	1,0		140	110						
5	$\widehat{\gamma}_{1G}^n = .617(0.032)$	p=0.0352	No	No	No	No	No	No	No	No		
-	$> \widehat{\gamma}_{1G}^e = .568(0.033)$	(1.8175)								- 10		
6	$\hat{\gamma}_2^e = .480(0.033)$	p = 0.0635	No	Yes	No	No	Yes	Yes	No	Yes		
Ü	$<\widehat{\gamma}_2^n = .511(0.033)$	(-1.5321)	1,0	100	1,0	1,0	100	100	1,0	165		
7	$\hat{\gamma}_2^e = .511(0.033)$	p=0.3192	No	No	Yes	No	o Yes	No	Yes	Yes		
•	$\sim \widehat{\gamma}_1^c = .502(0.033)$	(0.4706)	1,0	NO	103	140		110	103	165		
8	$\hat{\gamma}_2^n = .480(0.033)$	p = 0.1381	No	Yes	Yes	No	Yes	No	Yes	Yes		
	$\sim \widehat{\gamma}_1^c = .502(0.033)$	(-1.0915)			200							

 $Note: D\ indicates\ that\ the\ experimental\ results\ are\ directionally\ but\ not\ significantly\ consistent\ with\ predictions.$

(b) \leq 6 incorrect answers (N = 378)

Prop	Experimental result			Consistent with predictions?								
	mean (standard error)	p-value (t-stat)	S	A	I	R	AI	AR	IR	AIR		
1	$\widehat{\gamma}_{1G}^n = .545(0.026)$	p < 0.0001	No	No	No	No	Yes	Yes	No	Yes		
1	$> \hat{\gamma}_{1K}^n = .222(0.021)$	(10.0619)	0		110			100		165		
2	$\widehat{\gamma}^e_{1G} = .513(0.026)$	p=0.0092	No	No	No	No	No	Yes	Yes	Yes		
2	$> \widehat{\gamma}_1^c = .450(0.026)$	(2.3677)	110	NO	110					165		
3	$\widehat{\gamma}_{1G}^n = .545(0.026)$	p = 0.0002	No	No	No) No	o No	Yes	Yes	Yes		
3	$> \widehat{\gamma}_1^c = .450(0.026)$	(3.5499)		110	110			105		200		

4	$\widehat{\gamma}_{1K}^n = .222(0.021)$	p < 0.0001	No	No	No	No	Yes	Yes	No	Yes
	$<\widehat{\gamma}_1^c = .450(0.026)$	(-7.4632)								
5	$\widehat{\gamma}_{1G}^n = .545(0.026)$	p = 0.0954	No	Yes						
	$> \widehat{\gamma}_{1G}^e = .513(0.026)$	(1.3105)								
6	$\hat{\gamma}_2^e = .423(0.025)$	p = 0.0829	No	Yes	No	No	Yes	Yes	No	Yes
	$> \hat{\gamma}_2^n = .397(0.025)$	(-1.3885)								
7	$\hat{\gamma}_2^e = .423(0.025)$	p = 0.1022	No	No	Yes	No	Yes	No	Yes	Yes
	$\sim \widehat{\gamma}_1^c = .450(0.026)$	(-1.2710)					Yes	No	Yes	
8	$\hat{\gamma}_2^n = .397(0.025)$	p = 0.0068	No	No	Yes	No				Yes
	$<\widehat{\gamma}_1^c = .450(0.026)$	(-2.4785)								

(c) Accurate responders, <45 min (N=298)

Prop	Experimental result		Consistent with predictions?									
_	mean (standard error)	p-value (t-stat)	S	A	I	R	AI	AR	IR	AIR		
1	$\widehat{\gamma}_{1G}^n = .577(0.029)$	<i>p</i> < 0.0001	No	No	No	No	o Yes	Yes	No	Yes		
1	$>\widehat{\gamma}^n_{1K}=.195(0.023)$	(10.7353)	110	110	110	110	103	105	140	103		
2	$\widehat{\gamma}^e_{1G} = .540(0.029)$	p = 0.0020	No	No	No	No	No	Yes	Yes	Yes		
2	$> \widehat{\gamma}_1^c = .460(0.029)$	(2.9041)	110	110	110	110	110	105	ies	103		
3	$\widehat{\gamma}_{1G}^n = .577(0.029)$	p < 0.0001	No	No	No	No	No	Yes	Yes	Yes		
3	$> \widehat{\gamma}_1^c = .460(0.029)$	(4.0376)	110	110			110			103		
4	$\widehat{\gamma}_{1K}^n = .195(0.023)$	p < 0.0001	No	No	No	No	Yes	Yes	No	Yes		
1	$<\widehat{\gamma}_1^c = .460(0.029)$	(-7.8845)	NO	NO	110	140	100		110	105		
5	$\widehat{\gamma}_{1G}^n = .577(0.029)$	p=0.0797	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes		
3	$> \widehat{\gamma}_{1G}^e = .540(0.029)$	(1.4107)	110	105	105	105	105	105	105	100		
6	$\widehat{\gamma}_2^e = .456(0.029)$	p=0.0524	No	Yes	No	No	Yes	Yes	No	Yes		
O	$> \widehat{\gamma}_2^n = .423(0.029)$	(-1.6267)	110	105	110	110	105	105	110	100		
7	$\widehat{\gamma}_2^e = .456(0.029)$	p=0.4395	No	No	Yes	No	Yes	No	Yes	Yes		
,	$\sim \widehat{\gamma}_1^c = .460(0.029)$	(-0.1522)	1,0	1,0	100	110	103	1.0	105	100		
8	$\widehat{\gamma}_2^n = .423(0.029)$	p = 0.0506	No	No	Yes	No	o Yes	No	Yes	Yes		
	$<\widehat{\gamma}_1^c = .460(0.029)$	(-1.6445)	110									

(d) Accurate responders, saw nonexclusive game first (N=156)

Prop	Experimental result			Consistent with predictions?								
	mean (standard error)	p-value (t-stat)	S	A	I	R	AI	AR	IR	AIR		
1	$\widehat{\gamma}_{1G}^n = .551(0.040)$	<i>p</i> < 0.0001	No	No	No	No	Yes	Yes	No	Yes		
1	$> \widehat{\gamma}_{1K}^n = .167(0.030)$	(7.7913)	No	NO	NO	NO	163	103		165		
2	$\widehat{\gamma}^e_{1G} = .487(0.040)$	p = 0.0506	No	No	No	No	No	Yes	Yes	Yes		
2	$> \widehat{\gamma}_1^c = .417(0.040)$	(1.6488)	110	110	110	110	110	103	168	103		
3	$\widehat{\gamma}_{1G}^n = .551(0.040)$	p = 0.0015	No	No	No	No	No	Yes	Yes	Yes		
3	$> \widehat{\gamma}_1^c = .417(0.040)$	(3.0159)	INO	110	110	110				103		
4	$\widehat{\gamma}_{1K}^n = .167(0.030)$	p < 0.0001	No	No	No	No	Yes	Yes	No	Yes		
1	$<\widehat{\gamma}_1^c = .417(0.040)$	(-5.5394)			NO	110	105			105		
5	$\widehat{\gamma}_{1G}^n = .551(0.040)$	p = 0.0339	No	No	No	No	No No	No	No	No		
3	$> \widehat{\gamma}_{1G}^e = .487(0.040)$	(1.8396)	110	110	110	110	110	110	110	110		
6	$\hat{\gamma}_2^e = .417(0.040)$	p = 0.0302	No	Yes	No	No	Yes	Yes	No	Yes		
O	$> \widehat{\gamma}_2^n = .358(0.039)$	(-1.8921)	110	103	110	110	103	103	110	105		
7	$\widehat{\gamma}_2^e = .417(0.040)$	p = 0.5000	No	No	Yes	No	Yes	No	Yes	Yes		
,	$\sim \widehat{\gamma}_1^c = .417(0.040)$	(0.000)	110	110	168	NO	168	INO	ies	105		
8	$\widehat{\gamma}_2^n = .359(0.039)$	p = 0.0359	No	No	Yes	No	Yes	No	Yes	Yes		
8	$<\widehat{\gamma}_1^c = .417(0.040)$	(-1.8131)	110	110	ies	INO						

(e) Accurate responders, saw exclusive game first (N = 168)

Prop	Experimental result			Consistent with predictions?								
	mean (standard error)	p-value (t-stat)	S	A	I	R	AI	AR	IR	AIR		
1	$\widehat{\gamma}_{1G}^n = .571(0.038)$ $> \widehat{\gamma}_{1K}^n = .226(0.032)$	p < 0.0001 (7.4669)	No	No	No	No	Yes	Yes	No	Yes		
2	$\widehat{\gamma}_{1G}^e = .554(0.038)$ $> \widehat{\gamma}_1^c = .488(0.039)$	p = 0.0239 (1.9931)	No	No	No	No	No	Yes	Yes	Yes		
3	$\hat{\gamma}_{1G}^n = .571(0.038)$ $> \hat{\gamma}_1^c = .488(0.039)$	p = 0.0064 (2.5137)	No	No	No	No	No	Yes	Yes	Yes		
4	$\widehat{\gamma}_{1K}^n = .226(0.032)$ $< \widehat{\gamma}_1^c = .488(0.039)$	p < 0.0001 (-5.8375)	No	No	No	No	Yes	Yes	No	Yes		

5	$\widehat{\gamma}_{1G}^n = .571(0.038)$	p = 0.3117	No	Yes						
	$\sim \widehat{\gamma}^e_{1G} = .554(0.038)$	(0.4921)								
6	$\widehat{\gamma}_{2}^{e} = .464(0.039)$	p = 0.3194	No	No	Yes	No	No	No	Yes	No
	$\sim \widehat{\gamma}_2^n = .452(0.039)$	(-0.4703)								
7	$\hat{\gamma}_2^e = .464(0.039)$	p = 0.1863	No	No	Yes	No	Yes	No	Yes	Yes
	$\sim \widehat{\gamma}_1^c = .488(0.039)$	(-0.8939)		110	200					100
8	$\hat{\gamma}_2^n = .452(0.039)$	p=0.1129	No	Yes	Yes	No	Yes	No	Yes	Yes
	$\sim \widehat{\gamma}_1^c = .488(0.039)$	(-1.2266)							165	