

## **SROR556 Group Project: Prophet**

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### **Introduction**

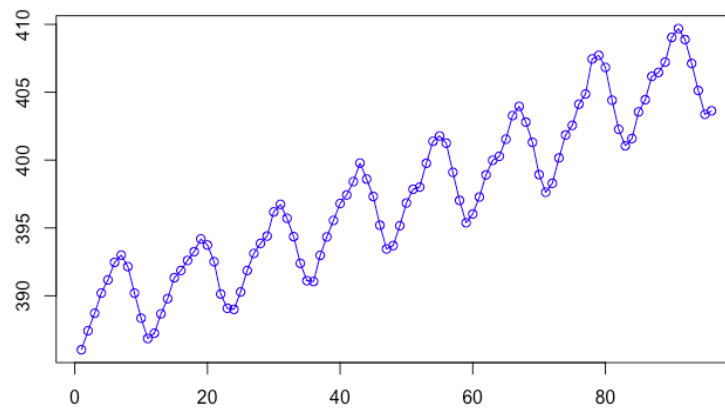
The American scientist Charles D. Keeling monitored carbon dioxide at Mauna Loa Observatory to first alert the world to the possibility of anthropogenic contribution to the "greenhouse effect" and global warming. "Greenhouse effect" has become a hot topic. Our group tried to predict the monthly average of CO<sub>2</sub> levels from Mauna Loa Observatory in Nov. 2019 based on historical data. We used the dataset of monthly mean carbon dioxide from Nov. 2009 to Oct. 2019 measured at Mauna Loa Observatory, Hawaii. The carbon dioxide data measured as the mole fraction in dry air with unit parts per million.

### **Heuristic of Approach and Assumptions of Model**

There are conspicuous seasonality and trend observed in the plot of data. Therefore, the method of classical decomposition is considered as a good choice. The model for the time series is assumed as  $X_t = m_t + s_t + Y_t$ . Where  $\{m_t\}$  is the trend,  $\{s_t\}$  is seasonal component and with known period  $d$  (i.e.,  $s_{t+d} = s_t$ ). Then, estimate the seasonal pattern and then form deseasonalized data:  $d_t = X_t - s_t$ . Next, use deseasonalized data to get an estimate of trend  $\{m_t\}$ . Then, remove trend and seasonality of the  $X_t$  based on estimation and form  $R_t$ . Next, analyze the residuals of:  $R_t = X_t - m_t - s_t$ . If the residuals are close to be iid, accept the estimated trend and seasonal pattern. Next, fit the  $X_t$  with models and do residuals analysis. The models with the iid residuals and lowest AICC standard will be selected for prediction. To check the accuracy of our data, we decided to use the method of cross validation, dividing the available data into 80% train set and

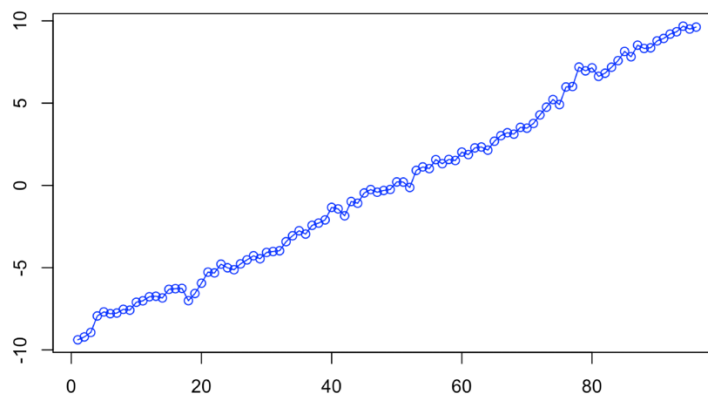
20% test set, using the model obtained from training set to predict testing set and get the residuals to test our model's accuracy.

### Description and Estimation for the model



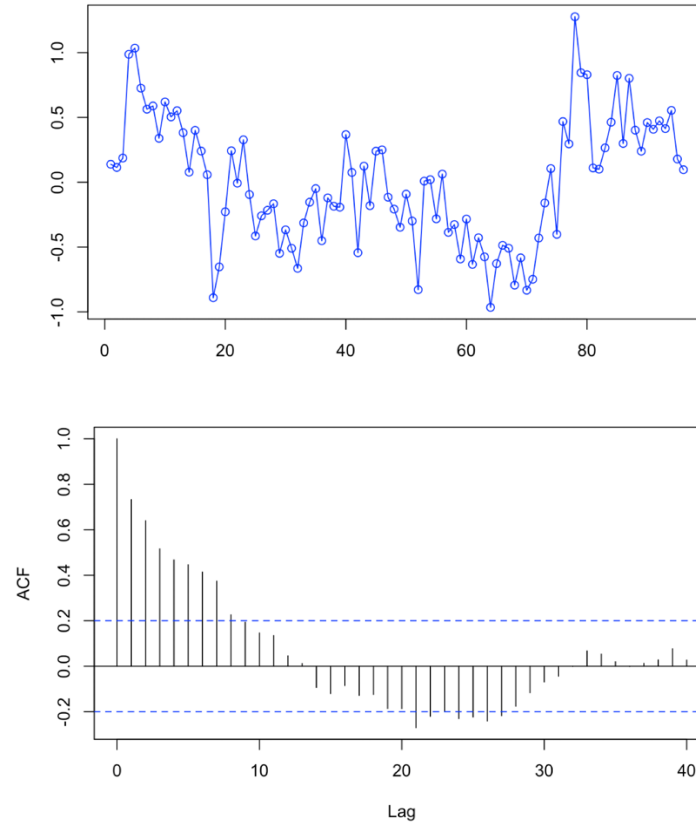
(Figure 1)

At first, we plotted the training set of the original time series (Figure 1), we can see clearly that there exist a trend component and a seasonal component, so our goal is to use classical elimination to remove the trend and seasonality component, and then fit residuals into a model to ideally obtain an iid noise. First, we counted there are roughly twelve data entries in each period, so we believe the period is twelve months (one year). Then, we obtain the deseasonalized data which is the original series with the estimated seasonal component removed.



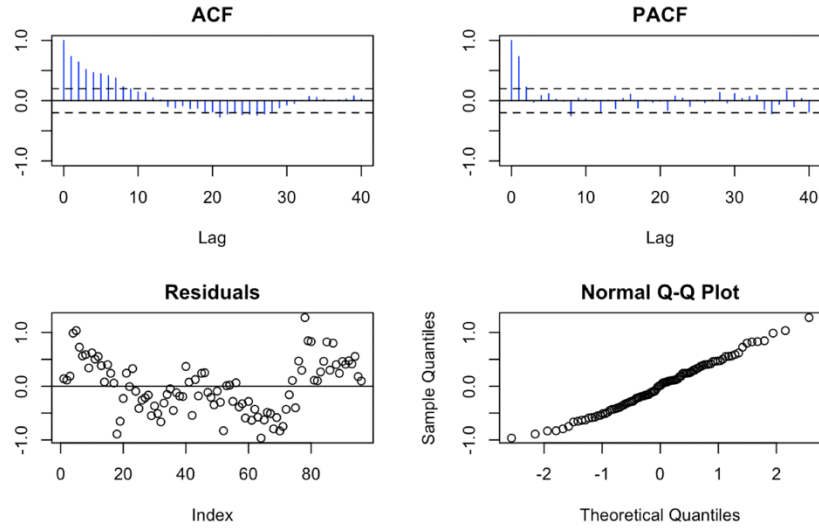
(Figure 2)

The graph of the deseasonalized data (Figure 2) suggests the presence of an additional linear trend function. In order to fit such a trend to the deseasonalized data, we removed the linear trend and obtained residuals called  $R_t$ .



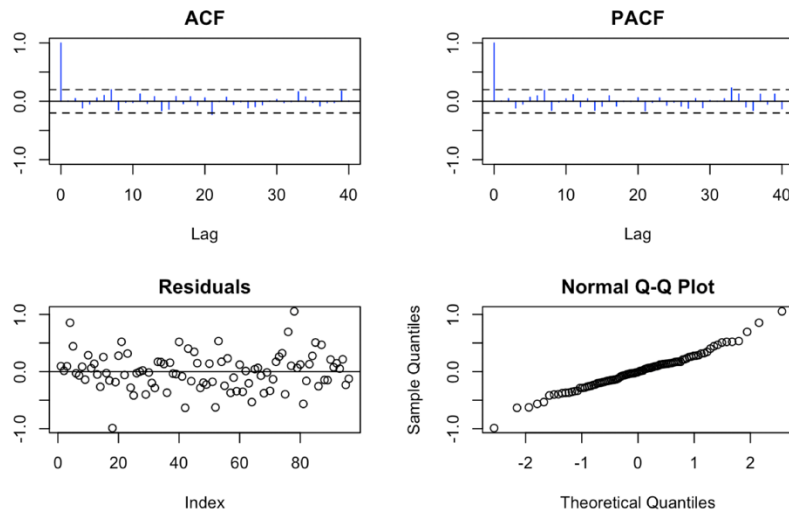
(Figure 3)

Yet based on the ACF plot (Figure 3), the residuals are not iid so we decided to further fit  $R_t$  into a model to finally obtain the expected iid noise. Having tried the ARMA(0,0) model, we can clearly see an AR(1) signature in PACF plot (Figure 4), so we decided to try fitting  $R_t$  into AR(1) and ARMA(1,1) model.



(Figure 4)

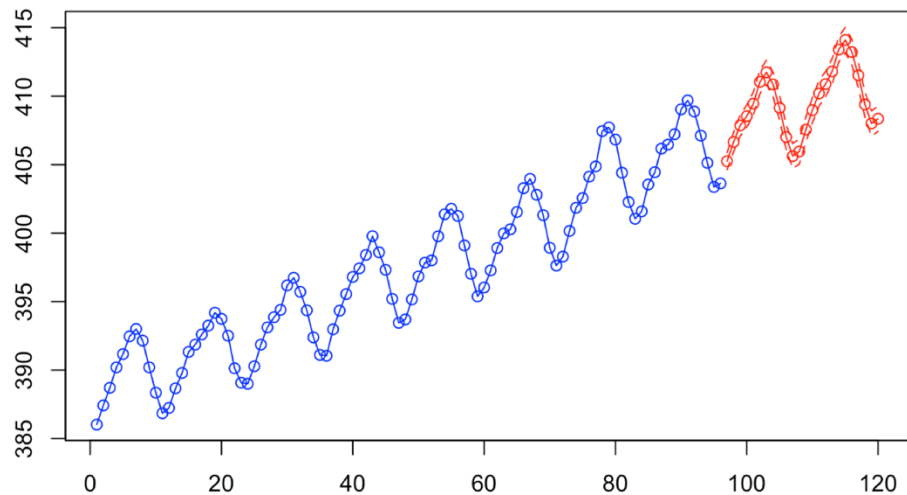
Fitting both AR(1) and ARMA(1,1) generated good ACF and PACF plot suggesting iid noise, they also pass all the tests of randomness. However, ARMA(1,1) has an AICC (59.88862) that's smaller than AR(1)'s AICC (62.42509) which indicates that ARMA(1,1) is a better model for  $R_t$ . For the final residuals obtained by subtracting the predicted value based on ARMA(1,1) model from  $R_t$ , seen from the output (Figure 5), the null hypothesis which suggests that our final residuals are iid noise is not rejected by all tests since the p-values are all greater than 0.05, including Ljung-Box test, McLeod-Li test, Turning points test, Different sign test, and Rank test.



Null hypothesis: Residuals are iid noise.			
Test	Distribution	Statistic	p-value
Ljung-Box Q	$Q \sim \text{chisq}(20)$	20.21	0.4448
McLeod-Li Q	$Q \sim \text{chisq}(20)$	13.32	0.8631
Turning points T	$(T-62.7)/4.1 \sim N(0,1)$	65	0.5685
Diff signs S	$(S-47.5)/2.8 \sim N(0,1)$	45	0.3792
Rank P	$(P-2280)/158 \sim N(0,1)$	2266	0.9294

(Figure 5)

We intended to use cross validation for the model we selected to predict the latest 20% of the data (i.e., data from November 2017 to October 2019). The reason why we included cross validation is that we want to test our model with the existing dataset to make sure that the predicted values are feasible. Using the forecast function in R with the first 80% of the data (i.e., data from November 2009 to October 2017) to predict the next 24 months average of CO<sub>2</sub> (Figure 6).



(Figure 6)

By subtracting the original values of the test dataset from our predicted values for test dataset, we got the residuals. As shown in the following output, the one sample t-test has a p-value of 0.1652, which is greater than 0.05. Therefore, we do not reject the null hypothesis that the true mean is equal to 0. Consequently, the cross-validation step affirms that our model selection is a good fit.

### One Sample t-test

```
data: resid_final
t = -1.4335, df = 23, p-value = 0.1652
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval:
 -0.28909571  0.05243445
sample estimates:
mean of x
-0.1183306
```

### Forecast

After selecting our model for the data, we used the forecast function in R with the whole dataset to predict the monthly average of CO<sub>2</sub> in November. The predicted number we have is 410.23 (more precisely, 410.2343) parts per million. The 95% confidence interval is from 409.61 to 410.8587 while the 99% confidence interval is from 409.4138 to 411.0549. These numbers seem quite reasonable given the newly released daily CO<sub>2</sub> in the first few days of November.

Step	Prediction	sqrt(MSE)	Lower Bound	Upper Bound
1	410.2343	0.3185491	409.4138	411.0549
Step	Prediction	sqrt(MSE)	Lower Bound	Upper Bound
1	410.2343	0.3185491	409.61	410.8587

### Scientific Justification

The data shows an obvious trend pattern for the CO<sub>2</sub> level at Mauna Loa. The average annual CO<sub>2</sub> level continues to increase in the past few decades. This pattern is consistent with the pattern of the average global CO<sub>2</sub> level. According to the Intergovernmental Panel on Climate Change (IPCC), studies have shown that “Anthropogenic greenhouse gas emissions have increased since the pre-industrial era, driven largely by economic and population growth”, which leads to atmospheric concentrations of carbon dioxide and the effects are “extremely likely to have been the dominant cause of the observed warming since the mid-20th century”. Many

researchers also support the theory that human actions like industrial production are the major factor for the increase of CO<sub>2</sub> level (Zhang and Mi. 2016, Baum et al. 2012, etc.).

There is an obvious seasonality pattern for the CO<sub>2</sub> level at Mauna Loa. The level is usually at its maximum in May and at its minimum around September or October. Researchers have tried to link the pattern with plant growth (Cleveland et al. 1983). One possible explanation is that in spring, the maxima were linked to an increase in rainfall during spring, contributing to vegetation growth at the site, thus increasing respiration rates at night-time. When respiration is larger than photosynthetic, plants consume oxygen and eliminate carbon dioxide. When summer arrives, the rapid growth of plants somewhat causes more intense photosynthetic activity and make the CO<sub>2</sub> level reaches its minimum around September (Fernández-Duque et al, 2019). Furthermore, a study based on the Atlantic Ocean points out that the influence of fossil fuel consumption is greater during the cold months, which may lead to the increase of CO<sub>2</sub> level during winter and cold spring. By contrast, clean air masses approaching the monitoring station are more frequent in summer, which may lead to a decrease of CO<sub>2</sub> level from May to September (García et al. 2016).

## Reference

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Cleveland, William S., Anne E. Freeny, and T. E. Graedel. "The seasonal component of atmospheric CO<sub>2</sub>: Information from new approaches to the decomposition of seasonal time series." *Journal of Geophysical Research: Oceans* 88.C15 (1983): 10934-10946.

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