

Online Sparse Reinforcement Learning

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Abstract

We investigate the hardness of online reinforcement learning in sparse linear Markov decision process (MDP), with a special focus on the high-dimensional regime where the ambient dimension is larger than the number of episodes. Our contribution is two-fold. First, we provide a lower bound showing that linear regret is generally unavoidable, even if there exists a policy that collects well-conditioned data. Second, we show that if the learner has oracle access to a policy that collects well-conditioned data, then a variant of Lasso fitted Q-iteration enjoys a regret of $\tilde{O}(N^{2/3})$ where N is the number of episodes.

1 Introduction

Sparse models in classical statistics often yield the best of both worlds: high representation power is achieved by including many features while sparsity leads to efficient estimation. There is a growing interest in applying the tools developed by statisticians to sequential settings such as contextual bandits and reinforcement learning (RL). As we now explore, in online RL this leads to a number of delicate trade-offs between assumptions and sample complexity. The use of sparsity in reinforcement learning (RL) has been explored before in the context of policy evaluation or policy optimization in the batch setting [Kolter and Ng, 2009, Geist and Scherrer, 2011, Hoffman et al., 2011, Painter-Wakefield and Parr, 2012, Ghavamzadeh et al., 2011]. As far as we know, there has been no work on the role of sparsity in online RL. In batch RL, the dataset is given a priori and the focus is typically on evaluating a given target policy or learning a near-optimal policy. By contrast, the central question in online RL is how to sequentially interact with the environment to balance the trade-off between exploration and exploitation, measured here by the cumulative regret. We ask the following question:

Under what circumstances does sparsity help when minimizing regret in online RL?

In sparse linear regression, the optimal estimation error rate generally scales polynomially with the sparsity s and only logarithmically in the ambient dimension d [Wainwright, 2019]. This is guaranteed under the sufficient and almost necessary condition that the data covariance matrix is well-conditioned, usually referred to restricted eigenvalue condition [Bickel et al., 2009] or compatibility condition [Van De Geer et al., 2009].

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The ‘almost necessary’ nature of the conditions for efficient estimation with sparsity leads to an unpleasant situation when minimizing regret. Even in sparse linear bandits, the worst-case regret is known to depend polynomially on the ambient dimension [Lattimore and Szepesvári, 2020, §24.3]. The reason is simple. By definition, a learner with small regret must play mostly the optimal action, which automatically leads to poorly conditioned data. Hence, making the right assumptions is essential in the high-dimensional regime where d is large relative to the time horizon. A number of authors have considered the contextual setting, where the regret can be made dimension free by making judicious assumptions on the context distribution [Bastani and Bayati, 2020, Wang et al., 2018, Kim and Paik, 2019].

When lifting assumptions from the bandit literature to RL it is essential to ensure that (a) the assumptions still help and (b) the assumptions remain reasonable. In some sense, our lower bound shows that a typical assumption that helps in linear bandits is by itself insufficient in RL. Specifically, in linear bandits the existence of a policy that collects well conditioned data is sufficient for dimension free regret. In RL this is no longer true because finding this policy may not be possible without first learning the transition structure, which cannot be done efficiently without well conditioned data, which yields an irresolvable chicken-and-egg problem.

Contribution We study online RL in episodic linear MDPs with ambient dimension d , sparsity s , episode length H and number of episodes N . Our contribution is two-fold:

- Our first result is a lower bound showing that $\Omega(Hd)$ regret is unavoidable in the worse-case when the dimension is large, even if the MDP transition kernel can be exactly represented by a sparse linear model and there exists an exploratory policy that collects well-conditioned data. The technical contribution is to craft a new class of hard-to-learn episodic MDPs. To overcome the difficulties caused by deterministic transitions from the constructed MDPs, we develop a novel stopping-time argument when calculating the KL-divergence.
- Our second result shows that if the learner has oracle access to an exploratory policy that collects well-conditioned data, then online Lasso fitted-Q-iteration in combination with the explore-then-commit template achieves a regret upper bound of $\tilde{O}(H^{4/3}s^{2/3}N^{2/3})$. The proof requires a non-trivial extension of high-dimensional statistics to Markov dependent data. As far as we know, this is the first regret bound that has no polynomial dependency on the feature dimension d in online RL.

1.1 Related work

Regret guarantees for online RL have received considerable attention in recent years. In episodic tabular MDPs with a homogeneous transition kernel, Azar et al. [2017] proved a minimax optimal regret of $O(\sqrt{H^2|S||A|N})$ achieved by a model-based algorithm. Jin et al. [2018] showed an $O(\sqrt{H^4|S||A|N})$ regret bound for Q-learning with inhomogeneous transition kernel. Under a linear MDP assumption, Jin et al. [2019] showed an $O(\sqrt{d^3H^4N})$ regret bound for an optimistic version of least-squares value iteration. Under a linear kernel MDP assumption [Zhou et al., 2020], Yang and Wang [2020] obtained an $O(dH^{5/2}\sqrt{N})$ regret bound by a model-based algorithm while Cai et al. [2019] obtained an $O(dH^2\sqrt{N})$ regret bound using an optimistic version of least-squares policy iteration. Zanette et al. [2020] derived an $O(d^2H^{5/2}\sqrt{N})$

regret bound for randomized least-squares value iteration. None of these works considered sparsity, and consequently the aforementioned regret bounds all have polynomial dependency on d .

There are several previous works focusing on sparse linear/contextual bandits that can be viewed as a simplified online RL problem. Abbasi-Yadkori et al. [2012] proposed an online-to-confidence-set conversion approach that achieves an $O(\sqrt{sdN})$ regret upper bound, where s is a known upper bound on the sparsity. The algorithm is not computationally-efficient, a deficit that is widely believed to be unavoidable. A matching lower bound is also known, which means polynomial dependence on d is generally unavoidable without additional assumptions [Lattimore and Szepesvári, 2020, §24.3]. In the contextual setting, where the action set changes from round to round, several works imposed various of careful assumptions on the context distribution such that polynomial dependency on d can be removed [Bastani and Bayati, 2020, Wang et al., 2018, Kim and Paik, 2019]. As far as well can tell, however, these assumptions are not easily extended to the MDP setting, where the contextual information available to the learner is not independent and identically distributed.

The use of feature selection in offline RL has also been investigated in a number of prior works. Kolter and Ng [2009], Geist and Scherrer [2011], Hoffman et al. [2011], Painter-Wakefield and Parr [2012], Liu et al. [2012] studied on-policy/off-policy evaluation with ℓ_1 -regularization for temporal-difference (TD) learning. Ghavamzadeh et al. [2011] and Geist et al. [2012] proposed Lasso-TD to estimate the value function in Markov reward processes and derived finite-sample statistical analysis. However, the aforementioned results can not be extended to online setting directly. One exception by Ibrahimi et al. [2012], who derived an $O(p\sqrt{N})$ regret bound in high-dimensional sparse linear quadratic systems where p is the dimension of the state space.

2 Preliminary

Notation. Denote by $\sigma_{\min}(X)$ and $\sigma_{\max}(X)$ the smallest and largest eigenvalues of a symmetric matrix X . Let $[n] = \{1, 2, \dots, n\}$. The relations \lesssim and \gtrsim stand for “approximately less/greater than” and are used to omit constant and poly-logarithmic factors. We use $\tilde{O}(\cdot)$ to omit polylog factors. For a finite set \mathcal{S} , let $\Delta_{\mathcal{S}}$ be the set of probability distributions over \mathcal{S} .

2.1 Problem definition

Episodic MDP. A finite episodic Markov decision process (MDP) is a tuple $(\mathcal{X}, \mathcal{A}, H, P, r)$ with \mathcal{X} the state-space, \mathcal{A} the action space, H the episode length, $P : \mathcal{X} \times \mathcal{A} \rightarrow \Delta_{\mathcal{X}}$ the transition kernel and $r : \mathcal{X} \times \mathcal{A} \rightarrow [0, 1]$ the reward function. As is standard, we assume that \mathcal{X} and \mathcal{A} are finite and that the reward function is known. We write $P(x'|x, a)$ for the probability of transitioning to state x' when taking action a in state x . A learner interacts with an episodic MDP as follows. In each episode, an initial state x_1 is sampled from an initial distribution $\xi_0 \in \Delta_{\mathcal{X}}$. Then, in each step $h \in [H]$, the learner observes a state $x_h \in \mathcal{X}$, takes an action $a_h \in \mathcal{A}$, and receives a deterministic reward $r(x_h, a_h)$. Then, the system evolves to a random next state x_{h+1} according to distribution $P(\cdot|x_h, a_h)$. The episode terminates when x_{H+1} is reached.

We define a (stationary) policy as a function $\pi : \mathcal{X} \rightarrow \Delta_{\mathcal{A}}$, that maps states to distributions over actions. A nonstationary policy is a sequence of maps from histories to probability distributions over actions. For

each $h \in [H]$ and policy π , the value function $V_h^\pi : \mathcal{X} \rightarrow \mathbb{R}$ is defined as the expected value of cumulative rewards received under policy π when starting from an arbitrary state at h th step; that is,

$$V_h^\pi(x) := \mathbb{E}^\pi \left[\sum_{h'=h}^H r(x_{h'}, a_{h'}) \middle| x_h = x \right],$$

where $a_{h'} \sim \pi(\cdot | x_{h'})$, $x_{h'+1} \sim P(\cdot | x_{h'}, a_{h'})$, and \mathbb{E}^π denotes the expectation over the sample path generated under policy π . Accordingly, we also define the action-value function $Q_h^\pi : \mathcal{X} \times \mathcal{A} \rightarrow \mathbb{R}$ which gives the expected cumulative reward when the learner starts from an arbitrary state-action pair at the h th step and follows policy π afterwards:

$$Q_h^\pi(x, a) := r(x, a) + \mathbb{E}^\pi \left[\sum_{h'=h+1}^H r(x_{h'}, a_{h'}) \middle| x_h = x, a_h = a \right].$$

Note, the conditioning in the above definitions is not quite innocent. In this form the value function is not well defined for states x that are not reachable by a given policy. This is easily rectified by defining the value function in terms of the Bellman equation or by being more rigorous about the probability space. The above definitions are standard in the literature and are left as is for reader's convenience.

Bellman equation. Since the action space and episode length are both finite, there always exists an optimal policy π^* which gives the optimal value $V_h^*(x) = \sup_\pi V_h^\pi(x)$ for all $x \in \mathcal{X}$ and $h \in [H]$ [Puterman, 2014, Szepesvári, 2010]. We denote the Bellman operator as

$$[\mathcal{T}V](x, a) := r(x, a) + \mathbb{E}_{x' \sim P(\cdot | x, a)}[V(x')],$$

and the Bellman equation for policy π becomes

$$\begin{aligned} Q_h^\pi(x, a) &= [\mathcal{T}V_{h+1}^\pi](x, a), \\ V_h^\pi(x) &= \mathbb{E}_{a \sim \pi(\cdot | x)}[Q_h^\pi(x, a)], V_{H+1}^\pi(x) = 0, \end{aligned} \tag{2.1}$$

which holds for all $(x, a) \in \mathcal{X} \times \mathcal{A}$. Similarly, the Bellman optimality equation is

$$\begin{aligned} Q_h^*(x, a) &= [\mathcal{T}V_{h+1}^*](x, a), \\ V_h^*(x) &= \max_{a \in \mathcal{A}} Q_h^*(x, a), V_{H+1}^*(x) = 0. \end{aligned} \tag{2.2}$$

Cumulative regret. In the online setting, the learner aims to minimize the cumulative regret by interacting with the environment over a number of episodes. At the beginning of the n th episode, an initial state x_1^n is sampled from ξ_0 and the agent executes policy π_n . We measure the performance of the learner over N episodes by the cumulative regret:

$$R_N = \sum_{n=1}^N (V_1^*(x_1^n) - V_1^{\pi_n}(x_1^n)). \tag{2.3}$$

The cumulative regret measures the expected loss of following the policy produced by the learner instead of the optimal policy. Therefore, the learner aims to follow a sequence of policies π_1, \dots, π_N such that the cumulative regret is minimized.

2.2 Sparse linear MDPs

Before we introduce sparse linear MDPs, we need to settle on a definition of a linear MDP. Let $\phi : \mathcal{X} \times \mathcal{A} \rightarrow \mathbb{R}^d$ be a feature map which assigns to each state-action pair a d -dimensional feature vector. A feature map combined with a parameter vector $w \in \mathbb{R}^d$ gives rise to the linear function $g_w : \mathcal{X} \times \mathcal{A} \rightarrow \mathbb{R}$ defined by $g_w(x, a) = \phi(x, a)^\top w$ and the subspace $\mathcal{G}_\phi = \{g_w : w \in \mathbb{R}^d\} \subseteq \mathbb{R}^{\mathcal{X} \times \mathcal{A}}$. Given a policy π and function $f : \mathcal{X} \times \mathcal{A} \rightarrow \mathbb{R}$, let $\tilde{\mathcal{T}}_\pi f : \mathcal{X} \times \mathcal{A} \rightarrow \mathbb{R}$ be the function defined by

$$[\tilde{\mathcal{T}}_\pi f](x, a) = r(x, a) + \mathbb{E}_{x' \sim P(\cdot|x, a), a \sim \pi(a|x')} [f(x', a)].$$

We call an MDP *linear* if \mathcal{G}_ϕ is closed under $\tilde{\mathcal{T}}_\pi$ for all policies π .¹ Yang and Wang [2019] and Jin et al. [2019] have shown that this is equivalent to assuming

$$P(x'|x, a) = \sum_{k \in [d]} \phi_k(x, a) \psi_k(x'),$$

for some functions $\psi_1, \dots, \psi_d : \mathcal{X} \rightarrow \mathbb{R}$ and all pairs of (x, a) . Note, the feature map ϕ is always assumed to be known to the learner. As far as we know, this notion of linearity was introduced by Bellman et al. [1963], Schweitzer and Seidmann [1985], who were motivated by the problem of efficiently computing the optimal policy for a known MDP with a large state-space.

When little priori information is available on how to choose the features, agnostic choices often lead to dimensions which can be as large as the number of episodes N . Without further assumptions, no procedure can achieve nontrivial performance guarantees, even when just considering simple prediction problems (e.g., predicting immediate rewards). However, effective learning with many more features than the sample-size is possible when only $s \ll d$ features are relevant. This motivates our definition of a sparse linear MDP.

Definition 2.1 (Sparse linear MDP). Fix a feature map $\phi : \mathcal{X} \times \mathcal{A} \rightarrow \mathbb{R}^d$ and assume the episodic MDP \mathcal{M} is linear in ϕ . We say \mathcal{M} is (s, ϕ) -sparse if there exists an active set $\mathcal{K} \subseteq [d]$ with $|\mathcal{K}| \leq s$ and some functions $\psi(\cdot) = (\psi_k(\cdot))_{k \in \mathcal{K}}$ such that for all pairs of (x, a) :

$$P(x'|x, a) = \sum_{k \in \mathcal{K}} \phi_k(x, a) \psi_k(x').$$

3 Hardness of online sparse RL

In this section we illustrate the fundamental hardness of online sparse RL in the high-dimensional regime by proving a minimax regret lower bound. The high-dimensional regime is referred to $N \leq d$. We first introduce a notion of an exploratory policy.

Definition 3.1 (Exploratory policy). Let Σ^π be the expected uncentered covariance matrix induced by policy π and feature map ϕ , given by

$$\Sigma^\pi := \mathbb{E}^\pi \left[\frac{1}{H} \sum_{h=1}^H \phi(x_h, a_h) \phi(x_h, a_h)^\top \right], \quad (3.1)$$

¹A different definition is called linear kernel MDP that the MDP transition kernel can be parameterized by a small number of parameters [Yang and Wang, 2020, Cai et al., 2019, Zanette et al., 2020, Zhou et al., 2020]

where $x_1 \sim \xi_0, a_h \sim \pi(\cdot|x_h), x_{h+1} \sim P(\cdot|x_h, a_h)$ and \mathbb{E}^π denotes expectation over the sample path generated under policy π . We call a policy π *exploratory* if $\sigma_{\min}(\Sigma^\pi) > 0$.

Remark 3.2. Intuitively, $\sigma_{\min}(\Sigma^\pi)$ characterizes how well the policy π explores in the feature space. Similar quantities also appear in the assumptions in the literature to ensure the success of policy evaluation/optimization with linear function approximation [Abbasi-Yadkori et al., 2019a, assumption A.4], [Duan and Wang, 2020, theorem 2], [Lazic et al., 2020, assumption A.3], [Abbasi-Yadkori et al., 2019b, assumption A.3] and [Agarwal et al., 2020, assumption 6.2].

Remark 3.3. In the tabular case, we can choose $\phi(x, a)$ as a basis vector in $\mathbb{R}^{|\mathcal{X}| \times |\mathcal{A}|}$. Let $\mu^\pi(x, a)$ be the frequency of visitation for state-action pair (x, a) under policy π and initial distribution ξ_0 :

$$\mu^\pi(x, a) = \frac{1}{H} \sum_{h=1}^H \mathbb{E}^\pi [\mathbb{I}((x_h, a_h) = (x, a))].$$

Then $\sigma_{\min}(\Sigma^\pi) > 0$ implies $\min_{x,a} \mu^\pi(x, a) > 0$. This means an exploratory policy in the tabular case will have positive visitation probability of each state-action pair.

The next theorem is a kind of minimax lower bound for online sparse RL. The key steps of the proof follow, with details and technical lemmas deferred to the appendix.

Theorem 3.4 (Minimax lower bound in high-dimensional regime). For any algorithm π , there exists a sparse linear MDP \mathcal{M} and associated exploratory policy π_e for which $\sigma_{\min}(\Sigma^{\pi_e})$ is a strictly positive universal constant independent of N and d , such that for any $N \leq d$,

$$R_N \geq \frac{1}{128} H d.$$

This theorem states that even if the MDP transition kernel can be exactly represented by a sparse linear model and there exists an exploratory policy, the learner could still suffer linear regret in the high-dimensional regime. This is in stark contrast to linear bandits, where the existence of an exploratory policy is sufficient for dimension-free regret. The problem in RL is that *finding* the exploratory policy can be very hard.

Proof of Theorem 3.4. The proof uses the standard information theoretic machinery, but with a novel hard-to-learn MDP construction and KL divergence calculation based on a stopping time argument. The intuition is to construct an informative state with only one of a large set of actions leading to the informative state deterministically. And the exploratory policy has to visit that informative state to produce well-conditioned data. In order to find this informative state, the learner should take a large number of trials that will suffer high regret.

Step 1: Construct a set of hard MDP instances. Let the state space \mathcal{X} consists of $\{x_0, x_i, x_u, x_g, x_b\}$. Here, x_0 is the initial state, x_i and x_u refer to informative and uninformative states, x_g and x_b refer to high-reward and low-reward states. Construct d different hard MDP instances: $\{\mathcal{M}_1, \dots, \mathcal{M}_d\}$ and they only differ at which action brings the learner to x_i . For each MDP $\mathcal{M}_k, k \in [d]$,

$$\theta = \left(\underbrace{\varepsilon, \dots, \varepsilon}_{s-1}, 0, \dots, 0, -1 \right)^\top, \quad (3.2)$$

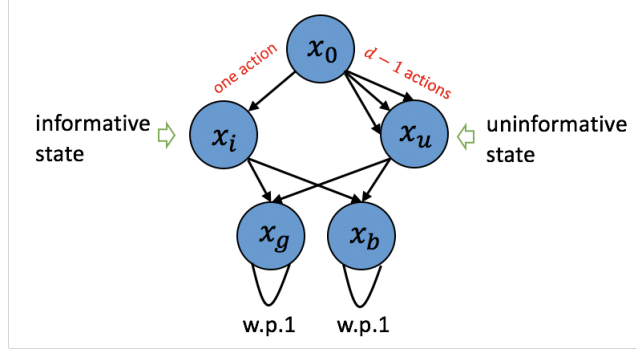


Figure 1: A hard-to-learn MDP instance that includes an informative state and an uninformative state.

where $\varepsilon > 0$ is a small constant to be tuned later, and $\bar{\theta}^{(k)} \in \mathbb{R}^{2d+2}$ as

$$\bar{\theta}^{(k)} = (\theta^\top, 1, 1, \underbrace{0, \dots, 0}_{k-1}, 1, \underbrace{0, \dots, 0}_{d-k})^\top. \quad (3.3)$$

We specify the transition probability of \mathcal{M}_k in the following steps:

1. Let the initial state x_0 associated with d actions as $\mathcal{A}_1 = \{a_1^0, \dots, a_d^0\}$. The transitions from x_0 to either x_i or x_u are *deterministic*. In MDP \mathcal{M}_k , only taking action a_k^0 brings the learner to x_i , and taking any other action *except* a_k^0 brings the learner to x_u . This information is parameterized into the last d coordinates of $\bar{\theta}^{(k)}$.
2. We construct a feature set \mathcal{S} associated to x_u and a feature set \mathcal{H} associated to x_i :

$$\begin{aligned} \mathcal{S} &= \left\{ z \in \mathbb{R}^d \mid z_d = 0, z_j \in \{-1, 0, 1\} \right. \\ &\quad \left. \text{for } j \in [d-1], \|z\|_1 = s-1 \right\}, \\ \mathcal{H} &= \left\{ z \in \mathbb{R}^d \mid z_j \in \{-1, 1\} \text{ for } j \in [d-1], z_d = 1 \right\}. \end{aligned}$$

Let $\mathcal{A}_2 = \{a_1^u, \dots, a_{|\mathcal{S}|}^u\}$ be the action set associated with x_u and $\mathcal{A}_3 = \{a_1^i, \dots, a_{|\mathcal{H}|}^i\}$ be the action set associated with x_i . We write $\varphi(x_u, a_j^u)$ as the j th element in \mathcal{S} and $\varphi(x_i, a_j^i)$ as the j th element in \mathcal{H} . At informative state x_i , the learner can take action $a_j^i \in \mathcal{A}_3$ and transits to either x_g or x_b according to

$$\begin{aligned} P(x_g | x_i, a_j^i) &= \varphi(x_i, a_j^i)^\top \theta = \phi(x_i, a_j^i)^\top \bar{\theta}^{(k)}, \\ P(x_b | x_i, a_j^i) &= 1 - \varphi(x_i, a_j^i)^\top \theta = 1 - \phi(x_i, a_j^i)^\top \bar{\theta}^{(k)}. \end{aligned}$$

where $\phi(x_i, a_j^i) = (\varphi(x_i, a_j^i), 0, \dots, 0)^\top \in \mathbb{R}^{2d+2}$. We define similarly when the learner at x_u .

3. At x_g or x_b , the learner will stay the current state for the rest of current episode no matter what actions to take.

One can verify that the above construction so far satisfies the sparse linear MDP assumption in Definition 2.1. In the end, the reward function is set to be $r(x, a) = 1$ if $x = x_g$ and $r(x, a) = 0$ otherwise. We now finish the construction of all the essential ingredients of $\{\mathcal{M}_1, \dots, \mathcal{M}_d\}$.

Remark 3.5. For \mathcal{M}_k , the overall action set will be $\mathcal{A}_1 \cup \mathcal{A}_2 \cup \mathcal{A}_3$. Now we specify the transitions that we have not mentioned so far. At x_0 , all the actions from \mathcal{A}_2 and \mathcal{A}_3 bring the learner to x_u . At x_i , all the actions from \mathcal{A}_1 and \mathcal{A}_2 bring the learner to either x_g or x_b with the same probability as a_1^i . At x_u , actions from \mathcal{A}_1 and \mathcal{A}_3 bring the learner to either x_g or x_b with the same probability as a_1^u .

Step 2: Construct an alternative set of MDPs. For each $k \in [d]$, the second step is to construct an alternative MDP $\widetilde{\mathcal{M}}_k$ that is hard to distinguish from \mathcal{M}_k and for which the optimal policy for \mathcal{M}_k is suboptimal for $\widetilde{\mathcal{M}}_k$ and vice versa. Fix a sequence of policies $\{\pi_1, \dots, \pi_N\}$. Let $\mathcal{D}_n = (S_1^n, A_1^n, \dots, S_H^n, A_H^n)$ be the sequence of state-action pairs in n th episode produced by π_n . Define $\mathcal{F}_h^n = \sigma(\mathcal{D}_1, \dots, \mathcal{D}_{h-1}, S_1^n, A_1^n, \dots, S_{h-1}^n, A_{h-1}^n, S_h^n)$. Let $\mathbb{F} = (\mathcal{F}_h^n)_{h \in [H], n \in [N]}$ be a filtration. Define the stopping time with respect to \mathbb{F} :

$$\tau_k = N \wedge \min \left\{ n : A_1^n = a_k^0 \right\}$$

that is the first episode the learner reaches the informative state. In other words, for $n \leq \tau_k - 1$, the learner always transits to x_u from x_0 . At x_u , the learner acts similarly as facing linear bandits where the number of arms is $|\mathcal{S}|$.

For $k \in [d]$, let $\mathbb{P}_k, \widetilde{\mathbb{P}}_k$ be the laws of $\mathcal{D}_1, \dots, \mathcal{D}_{\tau_k-1}$ induced by the interaction of $\{\pi_1, \dots, \pi_N\}$ and $\mathcal{M}_k, \widetilde{\mathcal{M}}_k$ accordingly. Let $\mathbb{E}_k, \widetilde{\mathbb{E}}_k$ be the corresponding expectation operators. In addition, denote a set \mathcal{S}' as

$$\begin{aligned} \mathcal{S}' = & \left\{ z \in \mathbb{R}^d \mid \|z\|_1 = s - 1, z_j \in \{-1, 0, 1\} \text{ for } j \in \{s, s+1, \dots, d-1\}, \right. \\ & \left. z_j = 0 \text{ for } j = \{1, \dots, s-1, d\} \right\}. \end{aligned} \quad (3.4)$$

Then we let

$$\widetilde{z}^{(k)} = \operatorname{argmin}_{z \in \mathcal{S}'} \mathbb{E}_k \left[\sum_{n=1}^{\tau_k} \langle \varphi(S_2^n, A_2^n), z \rangle^2 \right], \quad (3.5)$$

and construct the alternative $\widetilde{\theta}^{(k)} = \theta + 2\varepsilon \widetilde{z}^{(k)}$ where ε appears in Eq. (3.2). This is in contrast with $\bar{\theta}^{(k)}$ in Eq. (3.3) that specifies the original MDP \mathcal{M}_k . All the other ingredients of $\widetilde{\mathcal{M}}_k$ are the same as \mathcal{M}_k . Thus, we have constructed an alternative set of MDPs.

Step 3: Regret decomposition. Let $R_N(\mathcal{M}_k)$ be the cumulative regret of a sequence of policies $\{\pi_1, \dots, \pi_N\}$ interacting with MDP \mathcal{M}_k for N episodes. Recall that from the definition in Eq. (2.3), we have

$$R_N(\mathcal{M}_k) = \sum_{n=1}^N \left(V_1^{\pi^*}(x_1^n) - V_1^{\pi_n}(x_1^n) \right).$$

Denote $a^* = \operatorname{argmax}_{a_j^u \in \mathcal{A}_2} \varphi(x_u, a_j^u)^\top \theta$ be the optimal action when the learner is at x_u . The optimal policy π^* of MDP \mathcal{M}_k behaves in the following way for each episode:

- At state x_0 , the optimal policy takes an arbitrary action except a_k^0 to state x_u . There is no reward collected so far.
- At state x_u , the optimal policy takes action a^* and transits to good state x_g with probability $\varphi(x_u, a^*)^\top \theta$ or bad state x_b with probability $1 - \varphi(x_u, a^*)^\top \theta$.
- The learner stays at the current state for the rest of current episode.

Then the value function of π^* at n th episode is

$$\begin{aligned} V_1^{\pi^*}(x_1^n) &= (H-1)\mathbb{P}(A_2^n = a^*) \\ &= (H-1)\varphi(x_u, a^*)^\top \theta = (H-1)(s-1)\varepsilon. \end{aligned}$$

We decompose $R_N(\mathcal{M}_k)$ according to the stopping time τ_k :

$$\begin{aligned} R_N(\mathcal{M}_k) &\geq \sum_{n=1}^{\tau_k-1} \left(V_1^{\pi^*}(x_1^n) - V_1^{\pi_n}(x_1^n) \right) \\ &\geq \frac{H}{8} \mathbb{E}_k \left[\tau_k s \varepsilon - \sum_{n=1}^{\tau_k-1} \langle \varphi(S_2^n, A_2^n), \theta \rangle \right] \\ &= \frac{H}{8} \mathbb{E}_k \left[\tau_k s \varepsilon - \sum_{n=1}^{\tau_k-1} \sum_{j=1}^{s-1} \varphi_j(x_u, A_2^n) \varepsilon \right], \end{aligned} \tag{3.6}$$

where the last equation is due to S_2^n is always x_u until the stopping time τ_k .

Define an event

$$\mathcal{D}_k = \left\{ \sum_{n=1}^{\tau_k-1} \sum_{j=1}^{s-1} \varphi_j(x_u, A_2^n) \leq \frac{\tau_k s}{2} \right\}.$$

The next claim shows that when \mathcal{D}_k occurs, the regret is large in MDP \mathcal{M}_k , while if it does not occur, then the regret is large in MDP $\widetilde{\mathcal{M}}_k$. The detailed proof is deferred to Appendix B.1.

Claim 3.6. Regret lower bounds with respect to event \mathcal{D}_k :

$$R_N(\mathcal{M}_k) + R_N(\widetilde{\mathcal{M}}_k) \geq \frac{Hs\varepsilon}{8} \left(\mathbb{E}_k[\tau_k] + \widetilde{\mathbb{E}}_k[\tau_k \mathbb{I}(\mathcal{D}_k^c)] - \mathbb{E}_k[\tau_k \mathbb{I}(\mathcal{D}_k^c)] \right).$$

We construct an additional MDP \mathcal{M}_0 such that when the learner is at x_0 , no matter what actions to take, the learner will always transit to the uninformative state x_u . All the other structures remain the same with $\{\mathcal{M}_1, \dots, \mathcal{M}_d\}$. Let \mathbb{P}_0 be the laws of $\mathcal{D}_1, \dots, \mathcal{D}_{\tau_k-1}$ induced by the interaction of π and \mathcal{M}_0 and let \mathbb{E}_0 be the corresponding expectation operators. Then from Pinsker's inequality (Lemma C.5 in the Appendix), for any $k \in [d]$,

$$\begin{aligned} \left| \mathbb{E}_0[\tau_k] - \mathbb{E}_k[\tau_k] \right| &\leq N \sqrt{\frac{1}{2} \text{KL}(\mathbb{P}_0 \| \mathbb{P}_k)}, \\ \left| \widetilde{\mathbb{E}}_k[\tau_k \mathbb{I}(\mathcal{D}_k^c)] - \mathbb{E}_k[\tau_k \mathbb{I}(\mathcal{D}_k^c)] \right| &\leq N \sqrt{\frac{1}{2} \text{KL}(\widetilde{\mathbb{P}}_k \| \mathbb{P}_k)}, \end{aligned}$$

where $\text{KL}(\mathbb{P}, \mathbb{P}')$ is the KL divergence between probability measures \mathbb{P} and \mathbb{P}' . Combining with Claim 3.6, we have

$$R_N(\mathcal{M}_k) + R_N(\widetilde{\mathcal{M}}_k) \geq \frac{Hs\varepsilon}{8} \left(\mathbb{E}_0[\tau_k] - d\sqrt{\frac{1}{2}\text{KL}(\mathbb{P}_0\|\mathbb{P}_k)} - d\sqrt{\frac{1}{2}\text{KL}(\widetilde{\mathbb{P}}_k\|\mathbb{P}_k)} \right), \quad (3.7)$$

where we consider the high-dimensional regime such that $N \leq d$.

Step 4: Calculating the KL divergence. We make use of the following bound on the KL divergence between $\widetilde{\mathbb{P}}_k$ and \mathbb{P}_k , \mathbb{P}_0 and \mathbb{P}_k , which formalises the intuitive notion of information. When the KL divergence is small, the algorithm is unable to distinguish the two environments. The detailed proof is deferred to Appendix B.2.

Claim 3.7. The KL divergences between $\widetilde{\mathbb{P}}_k$ and \mathbb{P}_k , \mathbb{P}_0 and \mathbb{P}_k are upper bounded by the following when $N \leq d$:

$$\text{KL}(\widetilde{\mathbb{P}}_k\|\mathbb{P}_k) \leq 8\varepsilon^2(s-1)^2, \text{KL}(\mathbb{P}_0\|\mathbb{P}_k) = 0. \quad (3.8)$$

Combining with Eq. (B.11) and summing over the set of MDPs $\{\mathcal{M}_1, \dots, \mathcal{M}_d\}$,

$$\sum_{k=1}^d \left(R_N(\mathcal{M}_k) + R_N(\widetilde{\mathcal{M}}_k) \right) \geq \frac{Hs\varepsilon}{8} \left(\sum_{k=1}^d \mathbb{E}_0[\tau_k] - d^2\sqrt{8\varepsilon^2s^2} \right).$$

Picking $\varepsilon = 1/(8s)$, we have

$$\sum_{k=1}^d \left(R_N(\mathcal{M}_k) + R_N(\widetilde{\mathcal{M}}_k) \right) \geq \frac{H}{32} \left(\sum_{k=1}^d \mathbb{E}_0[\tau_k] - \frac{d^2}{4} \right).$$

Step 5: Summary. From the definition of the stopping time, one can see $\sum_{k=1}^d \mathbb{E}_0(\tau_k) \geq \sum_{k=1}^d k \geq d^2/2$. Therefore,

$$\sum_{k=1}^d \left(R_N(\mathcal{M}_k) + R_N(\widetilde{\mathcal{M}}_k) \right) \geq \frac{1}{128} H d^2.$$

Among two sets of MDPs $\{\mathcal{M}_k\}_{k=1}^d$ and $\{\widetilde{\mathcal{M}}_k\}_{k=1}^d$, for any sequence of policies $\{\pi_1, \dots, \pi_N\}$, there must exist a MDP \mathcal{M}_k such that

$$R_N(\mathcal{M}_k) \geq \frac{1}{128} H d.$$

This finishes the proof. □

4 Online Lasso Fitted-Q-iteration

In this section we prove that if the learner has oracle access to an exploratory policy, the online Lasso fitted-Q-iteration (Lasso-FQI) algorithm can have a dimension-free $\tilde{O}(N^{2/3})$ regret upper bound. We first introduce the online Lasso-FQI. Suppose the learner has the oracle access to an exploratory policy π_e (defined in Definition 3.1). The algorithm uses the explore-then-commit template and includes the following three phases:

- **Exploration phase.** The exploration phase includes N_1 episodes where N_1 will be chosen later based on regret bound and can be factorized as $N_1 = RH$, where $R > 1$ is an integer. At the beginning of each episode, the agent receives an initial state drawn from ξ_0 and executes the rest steps following the exploratory policy π_e . Let the dataset collected in the exploration stage as \mathcal{D} .
- **Learning phase.** Split \mathcal{D} into H folds: $\{\mathcal{D}_1, \dots, \mathcal{D}_H\}$ and each fold consists of R episodes. Based on the exploratory dataset \mathcal{D} , the agent executes an extension of fitted-Q-iteration [Ernst et al., 2005, Antos et al., 2008] combining with Lasso [Tibshirani, 1996] for feature selection. To define the algorithm, it is useful to introduce $Q_w(x, a) = \phi(x, a)^\top w$. For $a < b$, we also define the operator $\Pi_{[a,b]} : \mathbb{R} \rightarrow [a, b]$ that projects its input to $[a, b]$, i.e., $\Pi_{[a,b]}(x) = \max(\min(x, b), a)$. Initialize $\hat{w}_{H+1} = 0$. At each step $h \in [H]$, we fit \hat{w}_h through Lasso:

$$\hat{w}_h = \underset{w}{\operatorname{argmin}} \frac{1}{|\mathcal{D}_h|} \sum_{(x_i, a_i, x'_i) \in \mathcal{D}_h} (y_i - \phi(x_i, a_i)^\top w)^2 + \lambda_1 \|w\|_1, \quad (4.1)$$

where $y_i = \Pi_{[0,H]} \max_{a \in \mathcal{A}} Q_{\hat{w}_{h+1}}(x'_i, a)$ and λ_1 is a regularization parameter.

- **Exploitation phase.** For the rest $N - N_1$ episodes, the agent commits to the greedy policy with respect to the estimated Q-value $\{Q_{\hat{w}_h}\}_{h=1}^H$.

The full algorithm of online Lasso-FQI is summarized in Algorithm 1.

Remark 4.1. A key observation of Algorithm 1 is that the expected covariance matrix of data collected in the exploration phase could be well-conditioned due to the use of exploratory policy, e.g.,

$$\sigma_{\min} \left(\mathbb{E}^{\pi_e} \left[\frac{1}{N_1 H} \sum_{n=1}^{N_1} \sum_{h=1}^H \phi(x_h^n, a_h^n) \phi(x_h^n, a_h^n)^\top \right] \right) > 0.$$

This is the key condition to ensure the success of fast sparse feature selection in the learning and exploitation phases and eliminate the polynomial dependency of d in the cumulative regret.

Next we derive the regret guarantee for the online Lasso-FQI under the sparse linear MDP model. The proof is deferred to Appendix A. We need a notion of restricted eigenvalue that is common in high-dimensional statistics [Bickel et al., 2009, Bühlmann and Van De Geer, 2011].

Definition 4.2 (Restricted eigenvalue). Given a positive semi-definite matrix $Z \in \mathbb{R}^{d \times d}$ and integer $s \geq 1$, define the restricted minimum eigenvalue of Z as $C_{\min}(Z, s) :=$

$$\min_{S \subset [d], |S| \leq s} \min_{\beta \in \mathbb{R}^d} \left\{ \frac{\langle \beta, Z \beta \rangle}{\|\beta_S\|_2^2} : \|\beta_{S^c}\|_1 \leq 3 \|\beta_S\|_1 \right\}.$$

Algorithm 1 Online Lasso-FQI

1: **Input:** An episodic MDP $\mathcal{M} = (\mathcal{X}, \mathcal{A}, P, r, H)$, an exploratory policy π_e , exploration length N_1 , regularization parameter λ_1 ;
 # exploration phase
2: **Initialize.** $\mathcal{D} = \emptyset$.
3: **for** $n = 1, \dots, N_1$ **do**
4: Receive an initial state x_1^n .
5: **for** $h = 1, \dots, H$ **do**
6: Take the action $a_h^n = \pi_e(\cdot | x_h^n)$ and observe x_{h+1}^n .
7: Let $\mathcal{D} = \mathcal{D} \cup \{x_h^n, a_h^n, x_{h+1}^n\}$.
8: **end for**
9: **end for**
 # learning phase: Lasso fitted-Q-iteration
10: Partition the dataset \mathcal{D} into H folds such that each fold \mathcal{D}_h has R different episodes.
11: Initialize $Q_{\hat{w}_{H+1}}(x, a) = 0$.
12: **for** $h = H, \dots, 1$ **do**
13: Calculate regression targets for each $(x_i, a_i, x'_i) \in \mathcal{D}_h$:
$$y_i = \Pi_{[0, H]} \max_{a \in \mathcal{A}} Q_{\hat{w}_{h+1}}(x'_i, a).$$

14: Build training set $\{(x_i, a_i), y_i\}_{i \in \mathcal{D}_h}$ and fit \hat{w}_h through sparse linear regression in Eq. (4.1).
15: **end for**
 # exploitation phase
16: **for** $n = N_1 + 1$ to N **do**
17: Receive an initial state x_1^n .
18: **for** $h = 1, \dots, H$ **do**
19: Take greedy action $a_h^n = \operatorname{argmax}_a Q_{\hat{w}_h}(x_h^n, a)$ and transit to x_{h+1}^n .
20: **end for**
21: **end for**

Theorem 4.3 (Regret bound for online Lasso-FQI). Suppose the episodic MDP is (s, ϕ) -sparse as defined in Definition 2.1 and $\|\phi(x, a)\|_\infty \leq 1$ for any $(x, a) \in \mathcal{X} \times \mathcal{A}$. Assume the learner has oracle access to an exploratory policy π_e defined in Definition 3.1 and $C_{\min}(\Sigma^{\pi_e}, s)$ is a strictly positive universal constant independent of N and d . Choose the regularization parameter $\lambda_1 = H \sqrt{\log(2d)/N}$ and the number of episodes in the exploration phase N_1 as

$$N_1 = \left(\frac{2048 s^2 H^4 N^2}{C_{\min}(\Sigma^{\pi_e}, s)^2} \log(2dH/\delta) \right)^{\frac{1}{3}}.$$

With probability $1 - \delta$, the cumulative regret of online Lasso-FQI satisfies:

$$R_N \leq 2 \left(\frac{2048 \log(2dH/\delta)}{C_{\min}(\Sigma^{\pi_e}, s)^2} \right)^{\frac{1}{3}} H^{\frac{4}{3}} s^{\frac{2}{3}} N^{\frac{2}{3}}. \quad (4.2)$$

Remark 4.4. The condition that requires $C_{\min}(\Sigma^{\pi_e}, s)$ being dimension-free is weaker than requiring $\sigma_{\min}(\Sigma^{\pi_e})$ being dimension-free since $C_{\min}(\Sigma^{\pi_e}, s) \geq \sigma_{\min}(\Sigma^{\pi_e}) > 0$.

With oracle access of an exploratory policy, we obtain a dimension-free sub-linear regret bound. Without oracle access of such an exploratory policy, Theorem 3.4 implies a linear regret lower bound. On the other hand, without considering the sparsity, solving the MDP will suffer linear regret in the high-dimensional regime due to the well-known $\Omega(d\sqrt{N})$ lower bound. In summary, we emphasize that in high-dimensional regime, exploiting the sparsity to reduce the regret needs an exploratory policy but finding the exploratory policy is as hard as solving the MDP itself - an irresolvable “chicken and egg” problem.

5 Comparision with contextual bandits

In this section we investigate the difference between online RL and linear contextual bandits. When the planning horizon $H = 1$, the episodic MDP becomes to a contextual bandit. Specifically, consider a sparse linear contextual bandit. At n th episode, the environment generates a context x_n i.i.d from a distribution ξ_0 . The learner chooses an action $a_n \in \mathcal{A}$ and receives a reward:

$$Y_n = \phi(x_n, a_n)^\top \theta + \eta_n,$$

where $(\eta_n)_{n=1}^N$ is a sequence of independent standard Gaussian random variables and $\theta \in \mathbb{R}^d$ is a s -sparse unknown parameter vector.

We define an analogous exploratory policy as in Definition 3.1: for an exploratory policy π_e in a linear contextual bandit, it will satisfy

$$\sigma_{\min}(\Sigma^{\pi_e}) = \sigma_{\min}\left(\mathbb{E}^{\pi_e}\left[\phi(x_n, a_n)\phi(x_n, a_n)^\top\right]\right) > 0,$$

where $x_n \sim \xi_0$ and $a_n \sim \pi_e(\cdot|x_n)$. In episodic MDPs, since the MDP transition kernel is unknown, we can not find the exploratory policy without solving the MDP. However, in linear contextual bandits, as long as there exists an exploratory policy and the context distribution is known, we can obtain the exploratory policy by solving the following optimization problem:

$$\max_{\pi} \sigma_{\min}\left(\mathbb{E}_{x \sim \xi_0, a \sim \pi(\cdot|x)}\left[\phi(x, a)\phi(x, a)^\top\right]\right).$$

Thus, there is no additional cost of the regret to obtain the exploratory policy. Note that assuming known context distribution is much weaker than assuming known MDP transition kernel since we can learn the context distribution very quickly online. Following the rest step of online Lasso-FQI in Algorithm 1, we can replicate the $\tilde{O}(s^{2/3}N^{2/3})$ regret upper bound without oracle access of the exploratory policy.

6 Discussion

In this paper, we provide the first investigation of online sparse RL in the high-dimensional regime. In general, exploiting the sparsity to minimize the regret is hard without further assumptions. This also highlights some fundamental differences of sparse learning between online RL and supervised learning or contextual bandits.

A Proof of Theorem 4.3

Proof. In this section, we prove the regret bound of online Lasso fitted-Q-iteration. Recall that π_e is an exploratory policy that satisfies Definition 3.1, e.g.,

$$\sigma_{\min} \left(\mathbb{E}^{\pi_e} \left[\frac{1}{H} \sum_{h=1}^H \phi(x_h, a_h) \phi(x_h, a_h)^\top \right] \right) > 0,$$

where $x_1 \sim \xi_0, a_h \sim \pi(\cdot|x_h), x_{h+1} \sim P(\cdot|x_h, a_h)$ and \mathbb{E}^{π_e} denotes expectation over the sample path generated under policy π_e . Recall that N_1 is the number of episodes in exploration phase that will be specified later. Denote π_{N_1} as the greedy policy with respect to the estimated Q-value calculated from the Lasso fitted-Q-iteration in Algorithm 1. According to the design of Algorithm 1, we keep using π_{N_1} for the remaining $N - N_1$ episodes after exploration phase. From the definition of the cumulative regret in Eq. (2.3), we decompose R_N according to the exploration phase and exploitation phase:

$$R_N = \sum_{n=1}^N \left(V_1^*(x_1^n) - V_1^{\pi_n}(x_1^n) \right) = \underbrace{\sum_{n=1}^{N_1} \left(V_1^*(x_1^n) - V_1^{\pi_e}(x_1^n) \right)}_{I_1: \text{regret during exploring}} + \underbrace{\sum_{n=N_1+1}^N \left(V_1^*(x_1^n) - V_1^{\pi_{N_1}}(x_1^n) \right)}_{I_2: \text{regret during exploiting}}.$$

Since we assume $r \in [0, 1]$, from the definition of value functions, it is easy to see $0 \leq V_1^*(x), V_1^{\pi_e}(x) \leq H$ for any $x \in \mathcal{X}$. Thus, we can upper bound I_1 by

$$I_1 \leq N_1 H. \quad (\text{A.1})$$

To bound I_2 , we will bound $\|V_1^* - V_1^{\pi_{N_1}}\|_\infty$ first using the following lemma. The detailed proof is deferred to Lemma B.3. Recall that $C_{\min}(\Sigma^{\pi_e}, s)$ is the restricted eigenvalue in Definition 4.2 and we split the exploratory dataset into H folds with R episodes per fold.

Lemma A.1. Suppose the number of episodes in the exploration phase satisfies

$$N_1 \geq \frac{C_1 s^2 H \log(3d^2/\delta)}{C_{\min}(\Sigma^{\pi_e}, s)},$$

for some sufficiently large constant C_1 and $\lambda_1 = H \sqrt{\log(2d/\delta)/(RH)}$. Then we have with probability at least $1 - \delta$,

$$\|V_1^{\hat{\pi}_{N_1}} - V_1^*\|_\infty \leq \frac{32\sqrt{2}sH^3}{C_{\min}(\Sigma^{\pi_e}, s)} \sqrt{\frac{\log(2dH/\delta)}{N_1}}.$$

According to Lemma A.1, we have

$$I_2 \leq N \|V_1^{\hat{\pi}_{N_1}} - V_1^*\|_\infty \leq N \frac{32\sqrt{2}sH^3}{C_{\min}(\Sigma^{\pi_e}, s)} \sqrt{\frac{\log(2dH/\delta)}{N_1}}. \quad (\text{A.2})$$

Putting the regret bound during exploring (Eq. (A.1)) and the regret bound during exploiting (Eq. (A.2)), we have

$$R_N \leq N_1 H + N \frac{32\sqrt{2}sH^3}{C_{\min}(\Sigma^{\pi_e}, s)} \sqrt{\frac{\log(2dH/\delta)}{N_1}}.$$

We optimize N_1 by letting

$$N_1 H = N \frac{32\sqrt{2}sH^3}{C_{\min}(\Sigma^{\pi_e}, s)} \sqrt{\frac{\log(2dH/\delta)}{N_1}} \Rightarrow N_1 = \left(\frac{2048s^2H^4N^2}{C_{\min}(\Sigma^{\pi_e}, s)^2} \log(2dH/\delta) \right)^{1/3}. \quad (\text{A.3})$$

With this choice of N_1 , we have with probability at least $1 - \delta$

$$R_N \leq 2H \left(\frac{2048s^2H^4N^2}{C_{\min}(\Sigma^{\pi_e}, s)^2} \log(2dH/\delta) \right)^{1/3}.$$

□

Remark A.2. The optimal choice of N_1 in Eq. (A.3) requires the knowledge of s and $C_{\min}(\Sigma, s)$ that is typically not available in practice. Thus, we can choose a relatively conservative N_1 as

$$N_1 = (512H^4N^2 \log(2dH/\delta))^{1/3},$$

such that

$$R_N \leq 4 \frac{s}{C_{\min}(\Sigma^{\pi_e}, s)} H (512s^2H^4N^2 \log(2dH/\delta))^{1/3}.$$

B Additional proofs

B.1 Proof of Claim 3.6

Proof. We prove the first part. To simplify the notation, we write φ_{nj} short for $\varphi_j(x_u, A_2^n)$. From Eq. (3.6), we have

$$\begin{aligned} R_N(\mathcal{M}_k) &\geq (H-1) \mathbb{E}_k \left[\left((\tau_k - 1)(s-1)\varepsilon - \sum_{n=1}^{\tau_k} \sum_{j=1}^{s-1} \varphi_{nj} \varepsilon \right) \mathbb{I}(\mathcal{D}_k) \right] \\ &\geq \frac{Hs\varepsilon}{8} \mathbb{E}_k \left[\frac{\tau_k(s-1)\varepsilon}{2} \mathbb{I}(\mathcal{D}_k) \right]. \end{aligned}$$

Second, we derive a regret lower bound of alternative MDP $\widetilde{\mathcal{M}}_k$. Define $\widetilde{a}^* = \operatorname{argmax}_{a_j^u \in \mathcal{A}_2} \varphi(x_u, a_j^u)^\top \widetilde{\theta}^{(k)}$ as the optimal action when the learner is at state x_u in MDP \mathcal{M}_k . By a similar decomposition in Eq. (3.6),

$$\begin{aligned} R_N(\widetilde{\mathcal{M}}_k) &\geq (H-1) \left(\widetilde{\mathbb{E}}_k \left[\sum_{n=1}^{\tau_k-1} \langle \varphi(x_u, \widetilde{a}^*), \widetilde{\theta}^{(k)} \rangle \right] - \widetilde{\mathbb{E}}_k \left[\sum_{n=1}^{\tau_k-1} \langle \varphi_n, \widetilde{\theta}^{(k)} \rangle \right] \right) \\ &= (H-1) \widetilde{\mathbb{E}}_k \left[2\tau_k(s-1)\varepsilon - \sum_{n=1}^{\tau_k} \langle \varphi_n, \widetilde{\theta}^{(k)} \rangle \right]. \end{aligned} \quad (\text{B.1})$$

Next, we will find an upper bound for $\sum_{n=1}^{\tau_k-1} \langle \varphi_n, \tilde{\theta}^{(k)} \rangle$. From the definition of $\tilde{\theta}^{(k)}$ in Eq. (3.5),

$$\begin{aligned} \sum_{n=1}^{\tau_k-1} \langle \varphi_n, \tilde{\theta}^{(k)} \rangle &= \sum_{n=1}^{\tau_k} \langle \varphi_n, \theta + 2\varepsilon \tilde{z}^{(k)} \rangle \\ &= \sum_{n=1}^{\tau_k-1} \langle \varphi_n, \theta \rangle + 2\varepsilon \sum_{n=1}^{\tau_k-1} \langle \varphi_n, \tilde{z}^{(k)} \rangle \\ &\leq \sum_{n=1}^{\tau_k-1} \langle \varphi_n, \theta \rangle + 2\varepsilon \sum_{n=1}^{\tau_k-1} \sum_{j \in \text{supp}(\tilde{z}^{(k)})} |\varphi_{nj}|, \end{aligned} \quad (\text{B.2})$$

where the last inequality is from the definition of $\tilde{z}^{(k)}$ in Eq. (3.5). To bound the first term, we have

$$\begin{aligned} \sum_{n=1}^{\tau_k-1} \langle \varphi_n, \theta \rangle &= \sum_{n=1}^{\tau_k-1} \sum_{j=1}^{s-1} \varphi_{nj} \varepsilon \\ &\leq \varepsilon \sum_{n=1}^{\tau_k-1} \sum_{j=1}^{s-1} |\varphi_{nj}|. \end{aligned} \quad (\text{B.3})$$

Since all the φ_n come from \mathcal{S} which is a $(s-1)$ -sparse set, we have

$$\sum_{n=1}^{\tau_k-1} \sum_{j=1}^d |\varphi_{nj}| = (s-1)\tau_k,$$

which implies

$$\begin{aligned} \sum_{n=1}^{\tau_k-1} \left(\sum_{j=1}^{s-1} |\varphi_{nj}| + \sum_{j \in \text{supp}(\tilde{x})} |\varphi_{nj}| \right) &\leq \sum_{n=1}^{\tau_k-1} \sum_{j=1}^d |\varphi_{nj}| = (s-1)(\tau_k-1), \\ \sum_{n=1}^{\tau_k-1} \sum_{j=1}^{s-1} |\varphi_{nj}| &\leq (s-1)(\tau_k-1) - \sum_{n=1}^{\tau_k-1} \sum_{j \in \text{supp}(\tilde{x})} |\varphi_{nj}|. \end{aligned} \quad (\text{B.4})$$

Combining with Eq. (B.3),

$$\sum_{n=1}^{\tau_k-1} \langle \varphi_n, \theta \rangle \leq \varepsilon \left((s-1)(\tau_k-1) - \sum_{n=1}^{\tau_k-1} \sum_{j \in \text{supp}(\tilde{x})} |\varphi_{nj}| \right)$$

Plugging the above bound into Eq. (B.2), it holds that

$$\sum_{n=1}^{\tau_k-1} \langle \varphi_n, \tilde{\theta} \rangle \leq \varepsilon(s-1)(\tau_k-1) + \varepsilon \sum_{n=1}^{\tau_k} \sum_{j \in \text{supp}(\tilde{x})} |\varphi_{nj}|. \quad (\text{B.5})$$

When the event \mathcal{D}_k^c (the complement event of \mathcal{D}_k) happen, we have

$$\sum_{n=1}^{\tau_k-1} \sum_{j=1}^{s-1} |\varphi_{nj}| \geq \sum_{n=1}^{\tau_k-1} \sum_{j=1}^{s-1} \varphi_{nj} \geq \frac{(\tau_k-1)(s-1)}{2}.$$

Combining with Eq. (B.4), we have under event \mathcal{D}_k^c ,

$$\sum_{n=1}^{\tau_k-1} \sum_{j \in \text{supp}(\tilde{x})} |\varphi_{nj}| \leq \frac{(\tau_k-1)(s-1)}{2}. \quad (\text{B.6})$$

Putting Eqs. (B.1), (B.5), (B.6) together, it holds that

$$R_N(\widetilde{\mathcal{M}}_k) \geq (H-1) \widetilde{\mathbb{E}}_k \left[\frac{(\tau_k-1)(s-1)\varepsilon}{2} \mathbb{I}(\mathcal{D}_k^c) \right]. \quad (\text{B.7})$$

Putting the lower bounds of $R_N(\mathcal{M}_k)$ and $R_N(\widetilde{\mathcal{M}}_k)$ together, we have

$$\begin{aligned} R_N(\mathcal{M}_k) + R_N(\widetilde{\mathcal{M}}_k) &\geq (H-1) \left(\mathbb{E}_k \left[\frac{(\tau_k-1)(s-1)\varepsilon}{2} \mathbb{I}(\mathcal{D}_k) \right] + \widetilde{\mathbb{E}}_k \left[\frac{(\tau_k-1)(s-1)\varepsilon}{2} \mathbb{I}(\mathcal{D}_k^c) \right] \right) \\ &= \frac{Hs\varepsilon}{8} \left(\mathbb{E}_k \left[\tau_k \left(\mathbb{I}(\mathcal{D}_k) + \mathbb{I}(\mathcal{D}_k^c) \right) \right] + \widetilde{\mathbb{E}}_k [\tau_k \mathbb{I}(\mathcal{D}_k^c)] - \mathbb{E}_k [\tau_k \mathbb{I}(\mathcal{D}_k^c)] \right) \\ &= \frac{Hs\varepsilon}{8} \left(\mathbb{E}_k [\tau_k] + \widetilde{\mathbb{E}}_k [\tau_k \mathbb{I}(\mathcal{D}_k^c)] - \mathbb{E}_k [\tau_k \mathbb{I}(\mathcal{D}_k^c)] \right). \end{aligned}$$

This ends the proof. \square

B.2 Proof of Claim 3.7

Proof. The KL-calculation is inspired by Jaksch et al. [2010], but with novel stopping time argument. Denote the state-sequence up to n th episode, h th step as $\mathbb{S}_h^n = \{S_1^1, \dots, S_H^1, \dots, S_1^n, \dots, S_H^n\}$ and write $\mathcal{X}_h^n = \{x_0, x_i, x_u, x_g, x_b\}^{(n-1)H+h}$. For a fixed policy π interacting with the environment for n episodes, we denote $\mathbb{P}_k(\cdot)$ as the distribution over \mathbb{S}^n , where $S_1^n = x_0$, $A_h^n \sim \pi(\cdot | S_h^n)$, $S_{h+1}^n \sim \mathbb{P}_k(\cdot | S_h^n, A_h^n)$. Let \mathbb{E}_k denote the expectation w.r.t. distribution \mathbb{P}_k . By the chain rule, we can decompose the KL divergence as follows:

$$\text{KL}(\widetilde{\mathbb{P}}_k \| \mathbb{P}_k) = \mathbb{E} \left[\sum_{n=1}^{\tau_k-1} \sum_{h=1}^H \text{KL} \left[\widetilde{\mathbb{P}}_k(S_{h+1}^n | \mathbb{S}_h^n) \middle| \mathbb{P}_k(S_{h+1}^n | \mathbb{S}_h^n) \right] \right]. \quad (\text{B.8})$$

Given a random variable x , the KL divergence over two conditional probability distributions is defined as

$$\text{KL}(p(y|x), q(y|x)) = \sum_x \sum_y p(x, y) \log \left(\frac{p(y|x)}{q(y|x)} \right).$$

Then the KL divergence between $\tilde{\mathbb{P}}_k(S_{h+1}^n | \mathbb{S}_h^n)$ and $\mathbb{P}_k(S_{h+1}^n | \mathbb{S}_h^n)$ can be calculated as follows:

$$\begin{aligned}
& \text{KL} \left[\tilde{\mathbb{P}}_k(S_{h+1}^n | \mathbb{S}_h^n) \middle| \middle| \mathbb{P}_k(S_{h+1}^n | \mathbb{S}_h^n) \right] \\
&= \sum_{\mathbb{S}_h^n \in \mathcal{X}_h^n} \sum_{x \in \mathcal{X}} \tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_h^n) \log \left(\frac{\tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_h^n)}{\mathbb{P}_k(S_{h+1}^n = x | \mathbb{S}_h^n)} \right) \\
&= \sum_{\mathbb{S}_h^n \in \mathcal{X}_h^n} \sum_{x \in \mathcal{X}} \tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_h^n) \tilde{\mathbb{P}}_k(\mathbb{S}_h^n) \log \left(\frac{\tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_h^n)}{\mathbb{P}_k(S_{h+1}^n = x | \mathbb{S}_h^n)} \right) \\
&= \sum_{\mathbb{S}_{h-1}^n \in \mathcal{X}_{h-1}^n} \tilde{\mathbb{P}}_k(\mathbb{S}_{h-1}^n) \sum_{x' \in \mathcal{X}, a \in \mathcal{A}} \tilde{\mathbb{P}}_k(S_h^n = x', A_h^n = a | \mathbb{S}_{h-1}^n) \\
&\quad \cdot \sum_{x \in \mathcal{X}} \tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a) \log \left(\frac{\tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a)}{\mathbb{P}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a)} \right).
\end{aligned} \tag{B.9}$$

According to the construction of \mathcal{M}_k and $\widetilde{\mathcal{M}}_k$, the learner will remain staying at the current state when $x' = x_g$ or x_b , that implies

$$\tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a) = \mathbb{P}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a).$$

In addition, from the definition of stopping time τ_k , the learner will never transit to the informative state x_i . Therefore,

$$\begin{aligned}
& \text{KL} \left[\tilde{\mathbb{P}}_k(S_{h+1}^n | \mathbb{S}_h^n) \middle| \middle| \mathbb{P}_k(S_{h+1}^n | \mathbb{S}_h^n) \right] \\
&= \sum_{\mathbb{S}_{h-1}^n \in \mathcal{X}^{t-1}} \tilde{\mathbb{P}}_k(\mathbb{S}_{h-1}^n) \sum_{x' = x_0, x_i, x_u} \sum_{a \in \mathcal{A}} \tilde{\mathbb{P}}_k(S_h^n = x', A_h^n = a | \mathbb{S}_{h-1}^n) \\
&\quad \cdot \sum_{x \in \mathcal{X}} \tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a) \log \left(\frac{\tilde{\mathbb{P}}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a)}{\mathbb{P}_k(S_{h+1}^n = x | \mathbb{S}_{h-1}^n, S_h^n = x', A_h^n = a)} \right) \\
&= \sum_{a \in \mathcal{A}_2} \tilde{\mathbb{P}}_k(S_h^n = x_u, A_h^n = a) \sum_{x = x_g, x_b} \tilde{\mathbb{P}}_k(S_{h+1}^n = x | S_h^n = x_u, A_h^n = a) \log \left(\frac{\tilde{\mathbb{P}}_k(S_{h+1}^n = x | S_h^n = x_u, A_h^n = a)}{\mathbb{P}_k(S_{h+1}^n = x | S_h^n = x_u, A_h^n = a)} \right) \\
&= \sum_{a \in \mathcal{A}_2} \tilde{\mathbb{P}}_k(S_h^n = x_u, A_h^n = a) \left(\langle \varphi(x_u, a), \tilde{\theta}^{(k)} \rangle \log \left(\frac{\langle \varphi(x_u, a), \tilde{\theta}^{(k)} \rangle}{\langle \varphi(x_u, a), \theta \rangle} \right) + (1 - \langle \varphi(x_u, a), \tilde{\theta}^{(k)} \rangle) \log \left(\frac{1 - \langle \varphi(x_u, a), \tilde{\theta}^{(k)} \rangle}{1 - \langle \varphi(x_u, a), \theta \rangle} \right) \right)
\end{aligned}$$

where \mathcal{A}_2 is the action set associated to state x_u . Moreover, we will use Lemma C.4 to bound the above last term. Letting $q = \langle \varphi(x_u, a), \tilde{\theta}^{(k)} \rangle$ and $\epsilon = \langle \varphi(x_u, a), \theta - \tilde{\theta}^{(k)} \rangle$, it is easy to verify the conditions in Lemma C.4 as long as $\epsilon \leq (10(s-1))^{-1}$. Then we have

$$\begin{aligned}
\text{KL} \left[\tilde{\mathbb{P}}_k(S_{h+1}^n | \mathbb{S}_h^n) \middle| \middle| \mathbb{P}_k(S_{h+1}^n | \mathbb{S}_h^n) \right] &\leq \sum_{a \in \mathcal{A}_2} \tilde{\mathbb{P}}_k(S_h^n = x_u, A_h^n = a) \frac{2 \langle \tilde{\theta}^{(k)} - \theta, \varphi(x_u, a) \rangle^2}{\langle \tilde{\theta}^{(k)}, \varphi(x_u, a) \rangle} \\
&= \sum_{a \in \mathcal{A}_2} \tilde{\mathbb{P}}_k(S_h^n = x_u, A_h^n = a) \frac{8\epsilon^2 \langle \tilde{z}^{(k)}, \varphi(x_u, a) \rangle^2}{\langle \tilde{\theta}, \varphi(x_u, a) \rangle}.
\end{aligned}$$

Back to the KL-decomposition in Eq. (B.8), we have

$$\text{KL}(\tilde{\mathbb{P}}_k \| \mathbb{P}_k) \leq 8\varepsilon^2 \tilde{\mathbb{E}}_k \left[\sum_{n=1}^{\tau_k-1} \langle \varphi(x_u, A_2^n), \tilde{z} \rangle^2 \right].$$

To simplify the notations, we let $\varphi_n = \varphi(x_u, A_2^n)$.

Next, we use a simple argument “minimum is always smaller than the average”. We decompose the following summation over action set \mathcal{S}' defined in Eq. (3.4),

$$\begin{aligned} \sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} \langle \varphi_n, z \rangle^2 &= \sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} \left(\sum_{j=1}^d z_j \varphi_{nj} \right)^2 \\ &= \sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} \left(\sum_{j=1}^d (z_j \varphi_{nj})^2 + 2 \sum_{i < j} z_i z_j \varphi_{ni} \varphi_{nj} \right). \end{aligned}$$

We bound the above two terms separately. To bound the first term, we observe that

$$\sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} \sum_{j=1}^d (z_j \varphi_{nj})^2 = \sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} \sum_{j=1}^d |z_j \varphi_{nj}|, \quad (\text{B.10})$$

since both z_j, φ_{nj} can only take $-1, 0, +1$. In addition, $\sum_{t=1}^{\tau_k-1} \sum_{j=1}^d |\varphi_{nj}| = (s-1)\tau_k$. Since $z \in \mathcal{S}'$ that is $(s-1)$ -sparse, we have $\sum_{j=1}^d |z_j \varphi_{nj}| \leq s-1$. Therefore, we have

$$\sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} \sum_{j=1}^d |z_j \varphi_{nj}| \leq (s-1)(\tau_k-1) \binom{d-s-1}{s-2}. \quad (\text{B.11})$$

Putting Eqs. (B.10) and (B.11) together,

$$\sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} \sum_{j=1}^d (z_j \varphi_{nj})^2 \leq (s-1)(\tau_k-1) \binom{d-s-1}{s-2}. \quad (\text{B.12})$$

To bound the second term, we observe

$$\sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} 2 \sum_{i < j} z_i z_j \varphi_{ni} \varphi_{nj} = 2 \sum_{n=1}^{\tau_k-1} \sum_{i < j} \sum_{z \in \mathcal{S}'} z_i z_j \varphi_{ni} \varphi_{nj}.$$

From the definition of \mathcal{S}' , $z_i z_j$ can only take values of $\{1 * 1, 1 * -1, -1 * 1, -1 * -1, 0\}$. This symmetry implies

$$\sum_{z \in \mathcal{S}'} z_i z_j \varphi_{ni} \varphi_{nj} = 0,$$

which implies

$$\sum_{z \in \mathcal{S}'} \sum_{n=1}^{\tau_k-1} 2 \sum_{i < j} z_i z_j \varphi_{ni} \varphi_{nj} = 0. \quad (\text{B.13})$$

Combining Eqs. (B.12) and (B.13) together, we have

$$\sum_{z \in S'} \sum_{n=1}^{\tau_k-1} \langle \varphi_n, z \rangle^2 = \sum_{z \in S'} \sum_{n=1}^{\tau_k-1} \sum_{j=1}^d |z_j \varphi_{nj}| \leq (s-1)(\tau_k-1) \binom{d-s-1}{s-2}.$$

In the end, we use the fact that the minimum of $\tau_k - 1$ points is always smaller than its average,

$$\begin{aligned} \tilde{\mathbb{E}}_k \left[\sum_{n=1}^{\tau_k-1} \langle \varphi_n, \tilde{z} \rangle^2 \right] &= \min_{z \in S'} \tilde{\mathbb{E}}_k \left[\sum_{n=1}^{\tau_k-1} \langle \varphi_n, z \rangle^2 \right] \\ &\leq \frac{1}{|S'|} \sum_{z \in S'} \tilde{\mathbb{E}}_k \left[\sum_{n=1}^{\tau_k-1} \langle \varphi_n, z \rangle^2 \right] \\ &= \tilde{\mathbb{E}}_k \left[\frac{1}{|S'|} \sum_{z \in S'} \sum_{n=1}^{\tau_k-1} \langle \varphi_n, z \rangle^2 \right] \\ &\leq \frac{(s-1) \tilde{\mathbb{E}}_k[\tau_k-1] \binom{d-s-1}{s-2}}{\binom{d-s}{s-1}} \\ &\leq \frac{(s-1)^2 \tilde{\mathbb{E}}_k[\tau_k-1]}{d}. \end{aligned}$$

Therefore, we reach

$$\text{KL}(\tilde{\mathbb{P}}_k \| \mathbb{P}_k) \leq \frac{8\varepsilon^2(s-1)^2 \tilde{\mathbb{E}}_k[\tau_k-1]}{d} \leq \frac{8\varepsilon^2(s-1)^2 N}{d} \leq 8\varepsilon^2(s-1)^2,$$

since we consider the data-poor regime that $N \leq d$. It is obvious to see $\text{KL}(\mathbb{P}_0 \| \mathbb{P}_k) = 0$ from Eq. (B.9). This ends the proof. \square

B.3 Proof of Lemma A.1

Proof. Recall that in the learning phase, we split the data collected in the exploration phase into H folds and each fold consists of R episodes or RH sample transitions. For the update of each step h , we use a fresh fold of samples.

Step 1. We verify that the execution of Lasso fitted-Q-iteration is equivalent to the approximate value iteration. Recall that a generic Lasso estimator with respect to a function V at step h is defined in Eq. (4.1) as

$$\hat{w}_h(V) = \underset{w \in \mathbb{R}^d}{\text{argmin}} \left(\frac{1}{RH} \sum_{i=1}^{RH} \left(\Pi_{[0,H]} V(x_i^{(h)'}) - \phi(x_i^{(h)}, a_i^{(h)})^\top w \right)^2 + \lambda_1 \|w\|_1 \right).$$

Denote $V_w(x) = \max_{a \in \mathcal{A}} (r(x, a) + \phi(x, a)^\top w)$. For simplicity, we write $\hat{w}_h := \hat{w}_h(V_{\hat{w}_{h+1}})$ for short. Define an approximate Bellman optimality operator $\hat{\mathcal{T}}^{(h)} : \mathcal{X} \rightarrow \mathcal{X}$ as:

$$[\hat{\mathcal{T}}^{(h)} V](x) := \max_a \left[r(x, a) + \phi(x, a)^\top \hat{w}_h(V) \right]. \quad (\text{B.14})$$

Note this $\widehat{\mathcal{T}}^{(h)}$ is a randomized operator that only depends data from h th fold. The Lasso fitted-Q-iteration in learning phase of Algorithm 1 is equivalent to the following approximate value iteration:

$$[\widehat{\mathcal{T}}^{(h)}\Pi_{[0,H]}V_{\widehat{w}_{h+1}}](x) = \max_a \left[r(x, a) + \phi(x, a)^\top \widehat{w}_h \right] = \max_a Q_{\widehat{w}_h}(x, a) = V_{\widehat{w}_h}(x). \quad (\text{B.15})$$

Recall that the true Bellman optimality operator in state space $\mathcal{T} : \mathcal{X} \rightarrow \mathcal{X}$ is defined as

$$[\mathcal{T}V](x) := \max_a \left[r(x, a) + \sum_{x'} P(x'|x, a)V(x') \right]. \quad (\text{B.16})$$

Step 2. We verify that the true Bellman operator on $\Pi_{[0,H]}V_{\widehat{w}_{h+1}}$ can also be written as a linear form. From Definition 2.1, there exists some functions $\psi(\cdot) = (\psi_k(\cdot))_{k \in \mathcal{K}}$ such that for every x, a, x' , the transition function can be represented as

$$P(x'|x, a) = \sum_{k \in \mathcal{K}} \phi_k(x, a)\psi_k(x'), \quad (\text{B.17})$$

where $\mathcal{K} \subseteq [d]$ and $|\mathcal{K}| \leq s$. For a vector $\bar{w}_h \in \mathbb{R}^d$, we define its k th coordinate as

$$\bar{w}_{h,k} = \sum_{x'} \Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x')\psi_k(x'), \text{ if } k \in \mathcal{K}, \quad (\text{B.18})$$

and $\bar{w}_{h,k} = 0$ if $k \notin \mathcal{K}$. By the definition of true Bellman optimality operator in Eq. (B.16) and Eq. (B.17),

$$\begin{aligned} [\mathcal{T}\Pi_{[0,H]}V_{\widehat{w}_{h+1}}](x) &= \max_a \left[r(x, a) + \sum_{x'} P(x'|x, a)\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x') \right] \\ &= \max_a \left[r(x, a) + \sum_{x'} \phi(x, a)^\top \psi(x')\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x') \right] \\ &= \max_a \left[r(x, a) + \sum_{x'} \sum_{k \in \mathcal{K}} \phi_k(x, a)\psi_k(x')\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x') \right] \\ &= \max_a \left[r(x, a) + \sum_{k \in \mathcal{K}} \phi_k(x, a) \sum_{x'} \psi_k(x')\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x') \right] \\ &= \max_a \left[r(x, a) + \phi(x, a)^\top \bar{w}_h \right]. \end{aligned} \quad (\text{B.19})$$

We interpret \bar{w}_h as the ground truth of the Lasso estimator in Eq. (4.1) at step h in terms of the following sparse linear regression:

$$\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x'_i) = \phi(x_i, a_i)^\top \bar{w}_h + \varepsilon_i, i = 1 \dots, RH, \quad (\text{B.20})$$

where $\varepsilon_i = \Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x'_i) - \phi(x_i, a_i)^\top \bar{w}_h$. Define the filtration \mathcal{F}_i generated by $\{(x_1, a_1), \dots, (x_i, a_i)\}$ and also the data in folds $h+1$ to H . By the definition of $V_{\widehat{w}_{h+1}}$ and \bar{w}_h , we have

$$\begin{aligned} \mathbb{E}[\varepsilon_i | \mathcal{F}_i] &= \mathbb{E}[\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x'_i) | \mathcal{F}_i] - \phi(x_i, a_i)^\top \bar{w}_h \\ &= \sum_{x'} [\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x')P(x'|x_i, a_i) - \phi(x_i, a_i)^\top \bar{w}_h] \\ &= \sum_{k \in \mathcal{K}} \phi_k(x_i, a_i) \sum_{x'} [\Pi_{[0,H]}V_{\widehat{w}_{h+1}}(x')\psi_k(x') - \phi(x_i, a_i)^\top \bar{w}_h] = 0. \end{aligned}$$

Therefore, $\{\varepsilon_i\}_{i=1}^{RH}$ is a sequence of martingale difference noises and $|\varepsilon_i| \leq H$ due to the truncation operator $\Pi_{[0,H]}$. The next lemma bounds the difference between \hat{w}_h and \bar{w}_h within ℓ_1 -norm. The proof is deferred to Appendix B.4.

Lemma B.1. Consider the sparse linear regression described in Eq. (B.20). Suppose the number of episodes used in step h satisfies

$$R \geq \frac{C_1 \log(3d^2/\delta)s^2}{C_{\min}(\Sigma^{\pi_e}, s)},$$

for some absolute constant $C_1 > 0$. With the choice of $\lambda_1 = H\sqrt{\log(2d/\delta)/(RH)}$, the following holds with probability at least $1 - \delta$,

$$\|\hat{w}_h - \bar{w}_h\|_1 \leq \frac{16\sqrt{2}s}{C_{\min}(\Sigma^{\pi_e}, s)} H \sqrt{\frac{\log(2d/\delta)}{RH}}. \quad (\text{B.21})$$

Step 3. We start to bound $\|V_{\hat{w}_h} - V_h^*\|_\infty$ for each step h . By the approximate value iteration form Eq. (B.15) and the definition of optimal value function,

$$\begin{aligned} \|V_{\hat{w}_h} - V_h^*\|_\infty &= \|\hat{\mathcal{T}}^{(h)}\Pi_{[0,H]}V_{\hat{w}_{h+1}} - \mathcal{T}V_{h+1}^*\|_\infty \\ &= \|\hat{\mathcal{T}}^{(h)}\Pi_{[0,H]}V_{\hat{w}_{h+1}} - \mathcal{T}\Pi_{[0,H]}V_{\hat{w}_{h+1}}\|_\infty + \|\mathcal{T}\Pi_{[0,H]}V_{\hat{w}_{h+1}} - \mathcal{T}V_{h+1}^*\|_\infty. \end{aligned} \quad (\text{B.22})$$

The first term mainly captures the error between approximate Bellman optimality operator and true Bellman optimality operator. From linear forms Eqs. (B.15) and (B.19), it holds for any $x \in \mathcal{X}$,

$$\begin{aligned} &[\hat{\mathcal{T}}^{(h)}\Pi_{[0,H]}V_{\hat{w}_{h+1}}](x) - [\mathcal{T}\Pi_{[0,H]}V_{\hat{w}_{h+1}}](x) \\ &= \max_a \left[r(x, a) + \phi(x, a)^\top \hat{w}_h \right] - \max_a \left[r(x, a) + \phi(x, a)^\top \bar{w}_h \right] \\ &\leq \max_a |\phi(x, a)^\top (\hat{w}_h - \bar{w}_h)| \\ &\leq \max_{a,x} \|\phi(x, a)\|_\infty \|\hat{w}_h - \bar{w}_h\|_1. \end{aligned} \quad (\text{B.23})$$

Applying Lemma B.1, the following error bound holds with probability at least $1 - \delta$,

$$\|\hat{w}_h - \bar{w}_h\|_1 \leq \frac{16\sqrt{2}s}{C_{\min}(\Sigma^{\pi_e}, s)} H \sqrt{\frac{\log(2d/\delta)}{RH}}, \quad (\text{B.24})$$

where R satisfies $R \geq C_1 \log(3d^2/\delta)s^2/C_{\min}(\Sigma^{\pi_e}, s)$.

Note that the samples we use between phases are mutually independent. Thus Eq. (B.24) uniformly holds for all $h \in [H]$ with probability at least $1 - H\delta$. Plugging it into Eq. (B.23), we have for any stage $h \in [H]$,

$$\|\hat{\mathcal{T}}^{(h)}\Pi_{[0,H]}V_{\hat{w}_{h+1}} - \mathcal{T}\Pi_{[0,H]}V_{\hat{w}_{h+1}}\|_\infty \leq \frac{16\sqrt{2}s}{C_{\min}(\Sigma^{\pi_e}, s)} H \sqrt{\frac{\log(2dH/\delta)}{RH}}, \quad (\text{B.25})$$

holds with probability at least $1 - \delta$.

To bound the second term in Eq. (B.22), we observe that

$$\begin{aligned}
\|\mathcal{T}\Pi_{[0,H]}V_{\hat{w}_{h+1}} - \mathcal{T}V_{h+1}^*\|_\infty &= \max_x |\mathcal{T}\Pi_{[0,H]}V_{\hat{w}_{h+1}}(x) - \mathcal{T}V_{h+1}^*(x)| \\
&\leq \max_x \max_a \left| \sum_{x'} P(x'|x, a) \Pi_{[0,H]}V_{\hat{w}_{h+1}}(x') - \sum_{x'} P(x'|x, a) \Pi_{[0,H]}V_{h+1}^*(x') \right| \\
&\leq \|\Pi_{[0,H]}V_{\hat{w}_{h+1}} - V_{h+1}^*\|_\infty.
\end{aligned} \tag{B.26}$$

Plugging Eqs. (B.25) and (B.26) into Eq. (B.22), it holds that

$$\|V_{\hat{w}_h} - V_h^*\|_\infty \leq \frac{16\sqrt{2}s}{C_{\min}(\Sigma^{\pi_e}, s)} H \sqrt{\frac{\log(2dH/\delta)}{RH}} + \|\Pi_{[0,H]}V_{\hat{w}_{h+1}} - V_{h+1}^*\|_\infty, \tag{B.27}$$

with probability at least $1 - \delta$. Recursively using Eq. (B.27), the following holds with probability $1 - \delta$,

$$\begin{aligned}
\|\Pi_{[0,H]}V_{\hat{w}_1} - V_1^*\|_\infty &\leq \|V_{\hat{w}_1} - V_1^*\|_\infty \\
&= \frac{16\sqrt{2}s}{C_{\min}(\Sigma^{\pi_e}, s)} H \sqrt{\frac{\log(2dH/\delta)}{RH}} + \|\Pi_{[0,H]}V_{\hat{w}_2} - V_2^*\|_\infty \\
&\leq \|\Pi_{[0,H]}V_{\hat{w}_{H+1}} - V_{H+1}^*\|_\infty + H^2 \frac{16\sqrt{2}s}{C_{\min}(\Sigma^{\pi_e}, s)} \sqrt{\frac{\log(2dH/\delta)}{RH}} \\
&= H^2 \frac{16\sqrt{2}s}{C_{\min}(\Sigma^{\pi_e}, s)} \sqrt{\frac{\log(2dH/\delta)}{RH}},
\end{aligned}$$

where the first inequality is due to that $\Pi_{[0,H]}$ can only make error smaller and the last inequality is due to $V_{\hat{w}_{H+1}} = V_{H+1}^* = 0$. From Proposition 2.14 in Bertsekas [1995],

$$\|V_1^{\hat{\pi}_{N_1}} - V_1^*\|_\infty \leq H \|Q_{\hat{w}_1} - Q_1^*\|_\infty \leq 2H \|\Pi_{[0,H]}V_{\hat{w}_1} - V_1^*\|_\infty. \tag{B.28}$$

Putting the above together, we have with probability at least $1 - \delta$,

$$\|V_1^{\hat{\pi}_{N_1}} - V_1^*\|_\infty \leq \frac{32\sqrt{2}sH^3}{C_{\min}(\Sigma^{\pi_e}, s)} \sqrt{\frac{\log(2dH/\delta)}{N_1}},$$

when the number of episodes in the exploration phase has to satisfy

$$N_1 \geq \frac{C_1 s^2 H \log(3d^2/\delta)}{C_{\min}(\Sigma^{\pi_e}, s)},$$

for some sufficiently large constant C_1 . This ends the proof. \square

B.4 Proof of Lemma B.1

Proof. Denote the empirical covariance matrix induced by the exploratory policy π_e and feature map ϕ as

$$\hat{\Sigma}^{\pi_e} := \frac{1}{R} \sum_{r=1}^R \frac{1}{H} \sum_{h=1}^H \phi(x_h^r, a_h^r) \phi(x_h^r, a_h^r)^\top.$$

Recall that Σ^{π_e} is the population covariance matrix induced by the exploratory policy π_e defined in Eq. (3.1) and feature map ϕ with $\sigma_{\min}(\Sigma^{\pi_e}) > 0$. From the definition of restricted eigenvalue in (4.2) it is easy to verify $C_{\min}(\Sigma^{\pi_e}, s) \geq \sigma_{\min}(\Sigma^{\pi_e}) > 0$. For any $i, j \in [d]$, denote

$$v_{ij}^r = \frac{1}{H} \sum_{h=1}^H \phi_i(x_h^r, a_h^r) \phi_j(x_h^r, a_h^r) - \Sigma_{ij}^{\pi_e}.$$

It is easy to verify $\mathbb{E}[v_{ij}^r] = 0$ and $|v_{ij}^r| \leq 1$ since we assume $\|\phi(x, a)\|_{\infty} \leq 1$. Note that samples between different episodes are independent. This implies $v_{ij}^1, \dots, v_{ij}^R$ are independent. By standard Hoeffding's inequality (Proposition 5.10 in Vershynin [2010]), we have

$$\mathbb{P}\left(\left|\sum_{r=1}^R v_{ij}^r\right| \geq \delta\right) \leq 3 \exp\left(-\frac{C_0 \delta^2}{R}\right),$$

for some absolute constant $C_0 > 0$. Applying an union bound over $i, j \in [d]$, we have

$$\begin{aligned} \mathbb{P}\left(\max_{i,j} \left|\sum_{r=1}^R v_{ij}^r\right| \geq \delta\right) &\leq 3d^2 \exp\left(-\frac{C_0 \delta^2}{R}\right) \\ \Rightarrow \mathbb{P}\left(\|\hat{\Sigma}^{\pi_e} - \Sigma^{\pi_e}\|_{\infty} \geq \delta\right) &\leq 3d^2 \exp\left(-\frac{C_0 \delta^2}{R}\right). \end{aligned}$$

It implies the following holds with probability $1 - \delta$,

$$\|\hat{\Sigma}^{\pi_e} - \Sigma^{\pi_e}\|_{\infty} \leq \sqrt{\frac{\log(3d^2/\delta)}{R}}.$$

When the number of episodes $R \geq 32^2 \log(3d^2/\delta) s^2 / C_{\min}(\Sigma^{\pi_e}, s)^2$, the following holds with probability at least $1 - \delta$,

$$\|\hat{\Sigma}^{\pi_e} - \Sigma^{\pi_e}\|_{\infty} \leq \frac{C_{\min}(\Sigma^{\pi_e}, s)}{32s}.$$

Next lemma shows that if the restricted eigenvalue condition holds for one positive semi-definite matrix Σ_0 , then it holds with high probability for another positive semi-definite matrix Σ_1 as long as Σ_0 and Σ_1 are close enough in terms of entry-wise max norm.

Lemma B.2 (Corollary 6.8 in [Bühlmann and Van De Geer, 2011]). Let Σ_0 and Σ_1 be two positive semi-definite block diagonal matrices. Suppose that the restricted eigenvalue of Σ_0 satisfies $C_{\min}(\Sigma_0, s) > 0$ and $\|\Sigma_1 - \Sigma_0\|_{\infty} \leq C_{\min}(\Sigma_0, s)/(32s)$. Then the restricted eigenvalue of Σ_1 satisfies $C_{\min}(\Sigma_1, s) > C_{\min}(\Sigma_0, s)/2$.

Applying Lemma B.2 with $\hat{\Sigma}^{\pi_e}$ and Σ^{π_e} , we have the restricted eigenvalue of $\hat{\Sigma}^{\pi_e}$ satisfies $C_{\min}(\hat{\Sigma}^{\pi_e}, s) > C_{\min}(\Sigma^{\pi_e}, s)/2$ with high probability.

Note that $\{\varepsilon_i \phi_j(x_i, a_i)\}_{i=1}^{RH}$ is also a martingale difference sequence and $|\varepsilon_i \phi_j(x_i, a_i)| \leq H$. By Azuma-Hoeffding inequality,

$$\mathbb{P}\left(\max_{j \in [d]} \left|\frac{1}{RH} \sum_{i=1}^{RH} \varepsilon_i \phi_j(x_i, a_i)\right| \leq H \sqrt{\frac{\log(2d/\delta)}{RH}}\right) \geq 1 - \delta.$$

Denote event \mathcal{E} as

$$\mathcal{E} = \left\{ \max_{j \in [d]} \left| \frac{1}{RH} \sum_{i=1}^{RH} \varepsilon_i \phi_j(x_i, a_i) \right| \leq \lambda_1 \right\}.$$

Then $\mathbb{P}(\mathcal{E}) \geq 1 - \delta$. Under event \mathcal{E} , applying (B.31) in [Bickel et al. \[2009\]](#), we have

$$\|\hat{w}_h - \bar{w}_h\|_1 \leq \frac{16\sqrt{2}s\lambda_1}{C_{\min}(\Sigma^{\pi_e}, s)},$$

holds with probability at least $1 - 2\delta$. This ends the proof. \square

C Supporting lemmas

Lemma C.1 (Pinsker's inequality). Denote $\mathbf{x} = \{x_1, \dots, x_T\} \in \mathcal{X}^T$ as the observed states from step 1 to T . Then for any two distributions P_1 and P_2 over \mathcal{X}^\top and any bounded function $f : \mathcal{X}^\top \rightarrow [0, B]$, we have

$$\mathbb{E}_1 f(\mathbf{x}) - \mathbb{E}_2 f(\mathbf{x}) \leq \sqrt{\log 2/2B} \sqrt{\text{KL}(P_2 \| P_1)},$$

where \mathbb{E}_1 and \mathbb{E}_2 are expectations with respect to P_1 and P_2 .

Lemma C.2 (Bretagnolle-Huber inequality). Let \mathbb{P} and $\tilde{\mathbb{P}}$ be two probability measures on the same measurable space (Ω, \mathcal{F}) . Then for any event $\mathcal{D} \in \mathcal{F}$,

$$\mathbb{P}(\mathcal{D}) + \tilde{\mathbb{P}}(\mathcal{D}^c) \geq \frac{1}{2} \exp \left(-\text{KL}(\mathbb{P}, \tilde{\mathbb{P}}) \right), \quad (\text{C.1})$$

where \mathcal{D}^c is the complement event of \mathcal{D} ($\mathcal{D}^c = \Omega \setminus \mathcal{D}$) and $\text{KL}(\mathbb{P}, \tilde{\mathbb{P}})$ is the KL divergence between \mathbb{P} and $\tilde{\mathbb{P}}$, which is defined as $+\infty$, if \mathbb{P} is not absolutely continuous with respect to $\tilde{\mathbb{P}}$, and is $\int_{\Omega} d\mathbb{P}(\omega) \log \frac{d\mathbb{P}}{d\tilde{\mathbb{P}}}(\omega)$ otherwise.

The proof can be found in the book of [Tsybakov \[2008\]](#). When $\text{KL}(\mathbb{P}, \tilde{\mathbb{P}})$ is small, we may expect the probability measure \mathbb{P} is close to the probability measure $\tilde{\mathbb{P}}$. Note that $\mathbb{P}(\mathcal{D}) + \mathbb{P}(\mathcal{D}^c) = 1$. If $\tilde{\mathbb{P}}$ is close to \mathbb{P} , we may expect $\mathbb{P}(\mathcal{D}) + \tilde{\mathbb{P}}(\mathcal{D}^c)$ to be large.

Lemma C.3 (Divergence decomposition). Let \mathbb{P} and $\tilde{\mathbb{P}}$ be two probability measures on the sequence $(A_1, Y_1, \dots, A_n, Y_n)$ for a fixed bandit policy π interacting with a linear contextual bandit with standard Gaussian noise and parameters θ and $\tilde{\theta}$ respectively. Then the KL divergence of \mathbb{P} and $\tilde{\mathbb{P}}$ can be computed exactly and is given by

$$\text{KL}(\mathbb{P}, \tilde{\mathbb{P}}) = \frac{1}{2} \sum_{x \in \mathcal{A}} \mathbb{E}[T_x(n)] \langle x, \theta - \tilde{\theta} \rangle^2, \quad (\text{C.2})$$

where \mathbb{E} is the expectation operator induced by \mathbb{P} .

This lemma appeared as Lemma 15.1 in the book of [Lattimore and Szepesvári \[2020\]](#), where the reader can also find the proof.

Lemma C.4 (Lemma 20 in Jaksch et al. [2010]). Suppose $0 \leq q \leq 1/2$ and $\epsilon \leq 1 - 2q$, then

$$q \log \left(\frac{q}{q + \epsilon} \right) + (1 - q) \log \left(\frac{1 - q}{1 - q - \epsilon} \right) \leq \frac{2\epsilon^2}{q}.$$

Lemma C.5 (Pinsker’s inequality). For measures P and Q on the same probability space (Ω, \mathcal{F}) , we have

$$\delta(P, Q) = \sup_{A \in \mathcal{F}} (P(A) - Q(A)) \leq \sqrt{\frac{1}{2} \text{KL}(P, Q)}.$$

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