## Lecture 11

## Professor Virginia R. Young Transcribed by Hao Chen

September 28, 2022

## **Expectations**

Recap:

$$\mathbb{E}X = \int_{-\infty}^{\infty} x \ dF(x)$$

E.g.: 
$$X \sim U(a,b)$$

$$\mathbb{E}X = \int_{a}^{b} x f(x) \ dx = \frac{b+a}{2}$$

**E.g.:**  $X \sim \text{Gamma}(\alpha, \lambda)$ 

$$\begin{split} \mathbb{E}X &= \int_0^\infty x f(x) \; dx \\ &= \int_0^\infty x \cdot \frac{x^{\alpha - 1} e^{-\lambda x}}{\Gamma(\alpha)} \lambda^\alpha \; dx \\ &= \frac{\lambda^\alpha}{\Gamma(\alpha)} \int_0^\infty x^{(\alpha + 1) - 1} e^{-\lambda x} \; dx \\ &= \frac{\lambda^\alpha}{\Gamma(\alpha)} \cdot \frac{\Gamma(\alpha + 1)}{\lambda^{\alpha + 1}} \int_0^\infty \frac{\lambda^{\alpha + 1} x^{(\alpha + 1) - 1} e^{-\lambda x}}{\Gamma(\alpha + 1)} \; dx \end{split}$$

where we have

$$\int_0^\infty \frac{\lambda^{\alpha+1} x^{(\alpha+1)-1} e^{-\lambda x}}{\Gamma(\alpha+1)} \ dx = 1$$

Thus,

$$\mathbb{E}X = \frac{\Gamma(\alpha+1)}{\Gamma(\alpha)} \frac{1}{\lambda} = \frac{\alpha}{\lambda}$$

**E.g.:** Special case of the Gamma $(\alpha, \lambda)$  with  $\alpha = 1$  is  $\text{Exp}(\lambda)$  Another way to compute  $\mathbb{E}$ . Let  $S(x) = 1 - F(x) = \mathbb{P}(X > x)$ , or survival function of X.

$$\mathbb{E}X = \int_0^\infty (1 - F(x)) dx$$
$$= \int_0^\infty S(x) dx$$
$$= \int_0^\infty e^{-\lambda x} dx$$
$$= -\frac{1}{\lambda} e^{-\lambda x} \Big|_0^\infty = \frac{1}{\lambda}$$

Therefore,  $\mathbb{E}X = \frac{1}{\lambda}$ .

**E.g.:**  $X \sim \text{Pareto}(\alpha, \theta)$ 

$$S(x) = \begin{cases} 1 & x < 0 \\ \left(\frac{\theta}{\theta + x}\right)^{\alpha} & x \ge 0 \end{cases}$$
$$f(x) = F'(x) = -S'(x)$$

$$\mathbb{E}X = \int_0^\infty \left(\frac{\theta}{\theta + x}\right)^\alpha dx$$
$$= \theta^\alpha \frac{(\theta + x)^{-\alpha + 1}}{-\alpha + 1} \Big|_{x=0}^\infty$$

$$\mathbb{E}X = \left\{ \begin{array}{ll} \infty & \alpha \leq 1 \\ \frac{\theta}{\alpha - 1} & \alpha > 1 \end{array} \right.$$

E.g.: 
$$X \sim N(\mu, \sigma^2)$$

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

f(x) symmetric with respect to  $x = \mu \implies \mathbb{E}X = \mu$ .

$$\mathbb{E}X = \int_{-\infty}^{\infty} x \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx$$

Let  $z = \frac{x-\mu}{\sigma}$ , then  $dz = \frac{dx}{\sigma}$ ,

$$\begin{split} \mathbb{E}X &= \int_{-\infty}^{\infty} (\sigma z + \mu) \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2} \, dz \\ &= \frac{\sigma}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z e^{-\frac{1}{2}z^2} \, dz + \mu \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2} \, dz \\ &= \left( -e^{-\frac{1}{2}z^2} \right) \bigg|_{z=-\infty}^{\infty} + \mu \\ &= \mu \end{split}$$

**Defn:** Expectation of a function g of X

$$\mathbb{E}(g(X)) = \int_{-\infty}^{\infty} g(x) dF(x)$$
$$= \lim_{\|\pi\| \to 0} \sum_{i=-\infty}^{\infty} g(x_i^*) (F(x_i) - F(x_{i-1}))$$

where  $x_{-\infty} = -\infty$  and  $x_{\infty} = \infty$ .

If X is discrete with pmf p,

$$\mathbb{E}(g(X)) = \sum_{x_k} g(x_k) p(x_k)$$

If X is continuous with pdf f,

$$\mathbb{E}(g(X)) = \int_{-\infty}^{\infty} g(x)f(x) \ dx$$

 $\mathbb{E}$  is a linear operator

$$g(x) = ax + b, \qquad a, b \in \mathbb{R}$$

$$\begin{split} \mathbb{E}(aX+b) &= \int_{-\infty}^{\infty} (ax+b) \; F(x) \\ &= a \int_{-\infty}^{\infty} x \; dF(x) + b \int_{-\infty}^{\infty} 1 \; dF(x) \\ &= a \mathbb{E} X + b \end{split}$$

**Defn:** Variance of X:  $g(x) = (x - \mathbb{E}X)^2$ 

$$Var X = \mathbb{E}[(X - \mathbb{E}X)^2]$$

$$= \mathbb{E}[X^2 - 2\mathbb{E}X \cdot X + (\mathbb{E}X)^2]$$

$$= \mathbb{E}(X^2) - 2\mathbb{E}X \cdot \mathbb{E}X + (\mathbb{E}X)^2$$

$$= \mathbb{E}(X^2) - (\mathbb{E}X)^2$$

Typically, we call  $\mathbb{E}(X^k)$  as  $k^{\underline{th}}$  moment, so  $\mathbb{E}(X^2)$  is second moment. Similarly, we call  $\mathbb{E}((X - \mathbb{E}X)^k)$  as  $k^{\underline{th}}$  central moment, so  $\operatorname{Var} X$  is  $2^{nd}$  central moment.

For  $a, b \in \mathbb{R}$ ,

$$Var(aX + b)$$

$$= \mathbb{E}((aX + b - \mathbb{E}(aX + b))^{2})$$

$$= \mathbb{E}((aX + b - (a\mathbb{E}X + b))^{2})$$

$$= \mathbb{E}((aX - a\mathbb{E}X)^{2})$$

$$= \mathbb{E}(a^{2}(X - \mathbb{E}X)^{2})$$

$$= a^{2}\mathbb{E}((X - \mathbb{E}X^{2}))$$

$$= a^{2}Var X$$

**E.g.:**  $X \sim N(\mu, \sigma^2)$  Given that  $\mathbb{E}X = \mu$ ,

$$\operatorname{Var}(X) = \int_{-\infty}^{\infty} (x - \mu)^2 \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2} \left(\frac{x - \mu}{\sigma}\right)^2} dx$$

Let  $z = \frac{x-\mu}{\sigma}$ , then  $dz = \frac{dx}{\sigma}$ ,

$$\operatorname{Var} X = \int_{-\infty}^{\infty} \sigma^2 z^2 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2} dz$$
$$= \frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z \left( z e^{-\frac{1}{2}z^2} dz \right)$$

Consider that u = z and  $dt = ze^{-\frac{1}{2}z^2} dz$ ,

$$Var X = \sigma^{2} \frac{1}{\sqrt{2\pi}} \left[ -ze^{-\frac{1}{2}z^{2}} \Big|_{-\infty}^{\infty} + \int_{-\infty}^{\infty} e^{-\frac{1}{2}z^{2}} dz \right]$$
$$= \sigma^{2} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}z^{2}} dz = \sigma^{2}$$