#### Applications and Examples

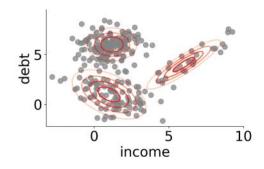
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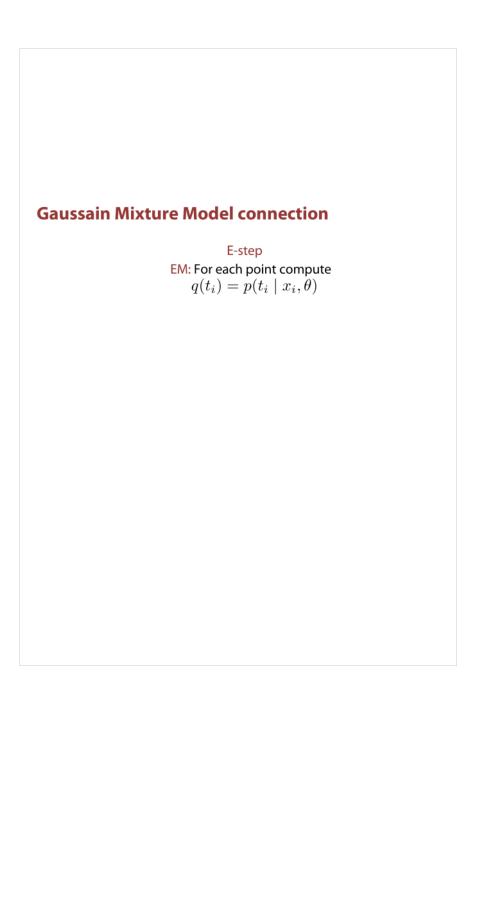
em for gmm

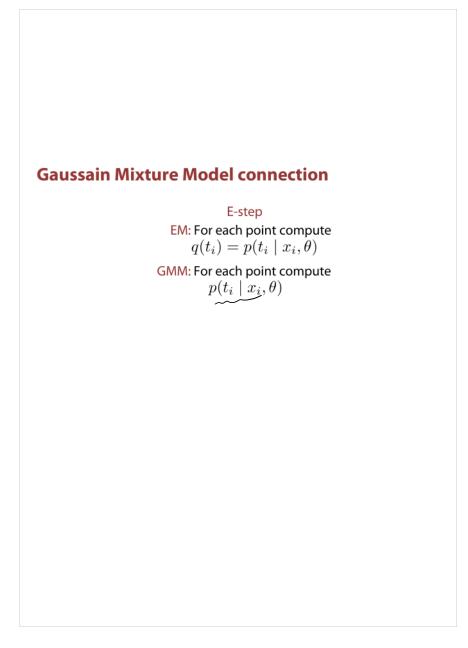
# **Applications of EM**

# **Gaussian Mixture Model revisited**



$$p(x \mid \theta) = \pi_1 \mathcal{N}(x \mid \mu_1, \Sigma_1) + \pi_2 \mathcal{N}(x \mid \mu_2, \Sigma_2) + \pi_3 \mathcal{N}(x \mid \mu_3, \Sigma_3)$$
$$\theta = \{\pi_1, \pi_2, \pi_3, \mu_1, \mu_2, \mu_3, \Sigma_1, \Sigma_2, \Sigma_3\}$$





#### **Gaussain Mixture Model connection**

#### E-step

EM: For each point compute  $q(t_i) = p(t_i \mid x_i, \theta)$ 

GMM: For each point compute  $p(t_i \mid x_i, \theta)$ 

#### M-step

EM: Update parameters to maximize  $\max_{\theta} \mathbb{E}_q \log p(X, T \mid \theta)$ 

GMM: Update Gaussian parameters to fit points assigned to them

$$\mu_1 = \frac{\sum_{i} p(t_i = 1 \mid x_i, \theta) x_i}{\sum_{i} p(t_i = 1 \mid x_i, \theta)}$$

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$$\mu_1 = \frac{\sum_i p(t_i = 1 \mid x_i, \theta) x_i}{\sum_i p(t_i = 1 \mid x_i, \theta)}$$

M-step derivation

$$\begin{aligned}
& \text{Max} & \sum_{i=1}^{N} \mathbb{E}_{q(ti)} \log P(x_i, ti|\theta) \\
& = \sum_{i=1}^{N} \sum_{c=1}^{2} q(ti=c) \log \left(\frac{1}{2} \exp\left(-\frac{(x_i - \mu_u)^2}{2\sigma_c^2}\right) \pi c\right) \\
& = \sum_{i=1}^{N} \sum_{c=1}^{2} q(ti=c) \left(\log \frac{\pi c}{Z} - \frac{(x_i - \mu_u)^2}{2\sigma_c^2}\right) \\
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$$= \frac{H}{\sum_{i=1}^{N} q(t_i x_i) \pi_i} - \left(\frac{\sum_{i=1}^{N} q(t_i z_i)}{\sum_{i=1}^{N} q(t_i z_i)}\right) \mu_i = 0.$$

$$\mu_1 = \frac{\sum_{i=1}^{N} q(t_i z_i) \pi_i}{\sum_{i=1}^{N} q(t_i z_i)}$$
Came for  $\mu_2$ ,  $\mu_3$ .

$$\sigma_{\nu}^{2} = \frac{\sum (x_{i} - \mu_{\nu})^{2} q_{i}(t_{i} = 0)}{\sum q_{i}(t_{i} = c)}$$

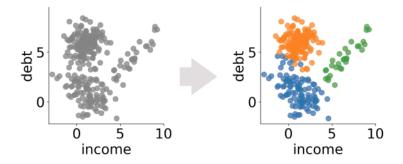
$$\pi_c = \frac{\sum_{i=0}^{\infty} q(t_i = c)}{N}$$

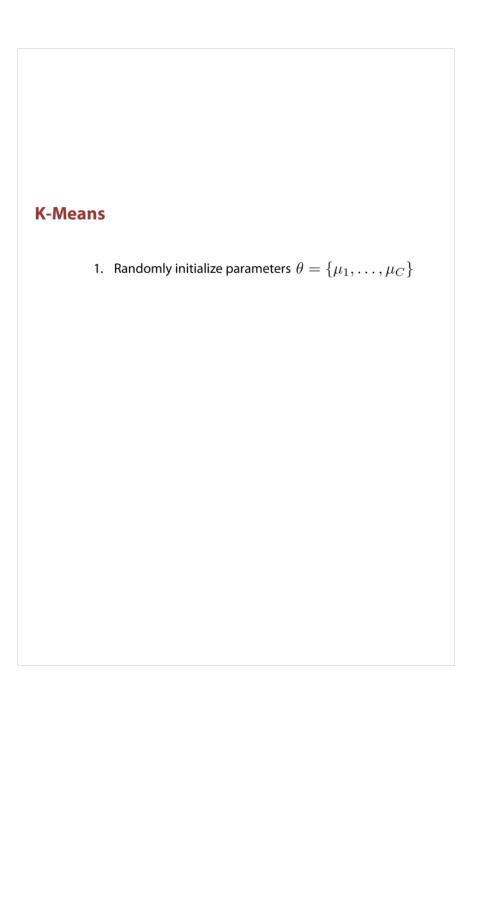
K-means from probabilitie perspective



📆 k-means1

# **K-Means connection**





### **K-Means**

- 1. Randomly initialize parameters  $\theta = \{\mu_1, \dots, \mu_C\}$
- 2. Until convergence repeat:
  - a) For each point compute closest centroid

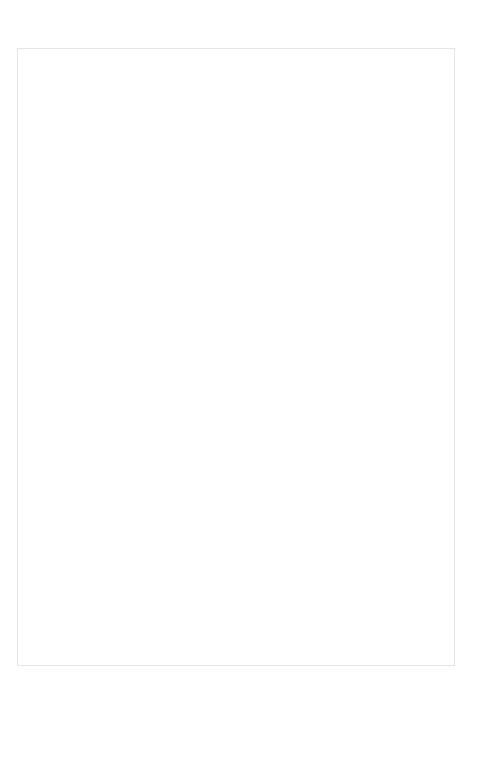
$$c_i = \operatorname*{arg\,min}_c \|x_i - \mu_c\|^2$$

#### **K-Means**

- 1. Randomly initialize parameters  $\theta = \{\mu_1, \dots, \mu_C\}$
- 2. Until convergence repeat:
  - a) For each point compute closest centroid

$$c_i = \mathop{\arg\min}_{c} \|x_i - \mu_c\|^2$$
 b) Update centroids

$$\mu_c = \frac{\sum_{i:c_i = c} x_i}{\#\{i:c_i = c\}}$$



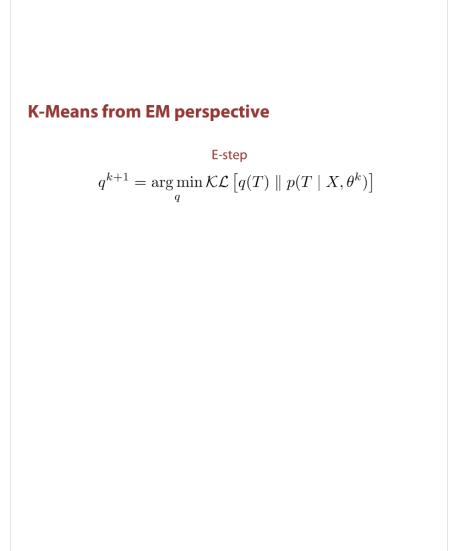


#### From GMM to K-means:

- Fix covariances to be identical  $\, \Sigma_c = I \,$
- Fix weights to be uniform  $\ \pi_c = \frac{1}{\# \ {
  m of \ Guassians}}$

$$p(x_i \mid t_i = c, \theta) = \frac{1}{Z} \exp\left(-0.5 \|x_i - \mu_c\|^2\right)$$

$$\text{The probability of } \text{Tenture } \text{Tenture$$



E-step

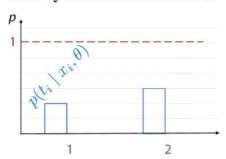
$$q^{k+1} = \underset{q \in Q}{\operatorname{arg \, min}} \, \mathcal{KL} \left[ q(T) \parallel p(T \mid X, \theta^k) \right]$$

Where  $\ensuremath{Q}$  is the set of delta-functions

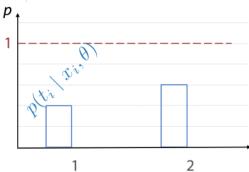
#### E-step

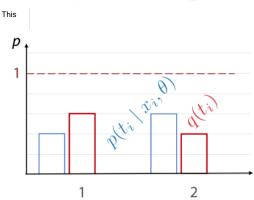
$$q^{k+1} = \operatorname*{arg\,min}_{q \in \textcolor{red}{Q}} \mathcal{KL} \left[ q(T) \parallel p(T \mid X, \theta^k) \right]$$

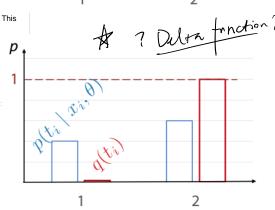
Where Q is the set of delta-functions



What is the closest approximation of the following distribution in the family of delta functions?







### K-Means from EM perspective

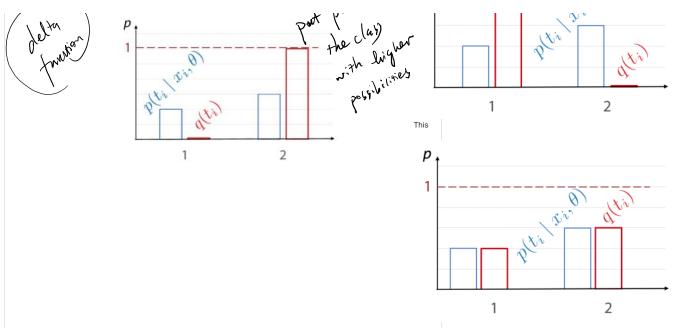
#### E-step

$$q^{k+1} = \operatorname*{arg\,min}_{q \in \textit{\textbf{Q}}} \textit{\textbf{KL}}\left[q(T) \parallel p(T \mid X, \theta^k)\right]$$

Where Q is the set of delta-functions P and P are P and P and P are P are P and P are P and P are P and P are P are P and P are P are P and P are P and P are P are P and P are P are P and P are P and P are P are P are P and P are P are P and P are P are P and P are P are P are P and P are P and P are P are P and P are P are P are P are P and P are P and P are P are P are P and P are P a

The distribution  $_{\rm qq}$  is indeed a delta function (unlike in some other answers) and the KL divergence between this  $_{\rm qq}$  and  $_{\rm pp}$  is lower than the corresponding KL divergence in other answers. In this case, KL divergence equals

 $\label{eq:local_kkl} $$ \mathcal{KL}(q \ | p) = 0 \cdot dot \cdot + 1 \cdot dot \cdot \int_{0.3} \operatorname{approx} 0.52 KL(q \cdot p) = 0$ 



#### From <a href="https://www.coursera.org/learn/bayesian-methods-in-machine-learning/lecture/qvh0F/k-means-from-probabilistic

# K-Means from EM perspective

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

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$$c_i = \operatorname*{max}_{c} p(t_i = c \mid x_i, \theta)$$

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$$c_i = \operatorname*{arg\,max}_{c} p(t_i = c \mid x_i, \theta)$$

$$p(t_i \mid x_i, \theta) = \frac{1}{Z} p(x_i \mid t_i, \theta) p(t_i \mid \theta)$$

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

$$c_i = \underset{c}{\text{arg max}} p(t_i = c \mid x_i, \theta)$$

$$p(t_i \mid x_i, \theta) = \frac{1}{Z} p(x_i \mid t_i, \theta) p(t_i \mid \theta)$$

$$= \frac{1}{Z} \exp\left(-0.5||x_i - \mu_c||^2\right) \pi_c$$

E-step

q is a function

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

$$c_i = \arg \max_{c} p(t_i = c \mid x_i, \theta) = \arg \min_{c} ||x_i - \mu_c||^2$$

$$c_{i} = \underset{c}{\operatorname{arg \, max}} p(t_{i} = c \mid x_{i}, \theta) = \underset{c}{\operatorname{arg \, min}} \|x_{i} - \mu_{c}\|^{2}$$

$$p(t_{i} \mid x_{i}, \theta) = \frac{1}{Z} p(x_{i} \mid t_{i}, \theta) p(t_{i} \mid \theta)$$

$$= \frac{1}{Z} \exp\left(-0.5\|x_{i} - \mu_{c}\|^{2}\right) \pi_{c}$$

$$\underset{can \quad \sigma \, min}{\operatorname{min}} t$$

$$do \, \text{not} \, depend \, \sigma m \, \overset{c}{\underline{c}}$$

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

$$c_i = \operatorname*{arg\,min}_{c} \|x_i - \mu_c\|^2$$

E-step

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$
$$c_i = \underset{c}{\arg\min} \|x_i - \mu_c\|^2$$

Exactly like in K-Means!

To muximize

اما

max 
$$\sum_{i=1}^{N} \mathbb{E}_{q(t_i)} \log P(x_i, t_i | M)$$
 $M_c = \frac{\sum_{i=1}^{N} q(t_i = c) \cdot x_i}{\sum_{i=1}^{N} q(t_i = c)} = \frac{\sum_{i=c,t=c}^{N} x_i}{\# i : c : = c}$ 
 $q(t_i) = \begin{cases} 1 & \text{if } t_i = C^* \\ 0 & \text{if } t_i \neq C^* \end{cases}$ 

, IX

k-means

### K-Means from EM perspective

M-step

$$\theta^{k+1} = \operatorname*{arg\,max}_{\theta} \mathbb{E}_{q^{k+1}} \log p(X, T \mid \theta)$$
 
$$\mu_c^{k+1} = \frac{\sum_{i: c_i = c} x_i}{\#\{i: c_i = c\}}$$

M-step

$$\theta^{k+1} = \underset{\theta}{\arg\max} \mathbb{E}_{q^{k+1}} \log p(X, T \mid \theta)$$
$$\mu_c^{k+1} = \frac{\sum_{i:c_i=c} x_i}{\#\{i:c_i=c\}}$$

Exactly like in K-Means!



Summary

• With fixed covariance matrices  $\underline{\Sigma_c=I}$  (?) restricting to  $\overline{t}$  step to a specific distribution.
• Simplified E-step (approximate  $p(t_i\mid x_i,\theta)$  with delta function)

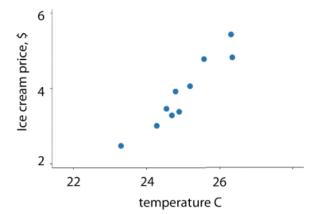
Thus K-Means is faster but less flexible than GMM

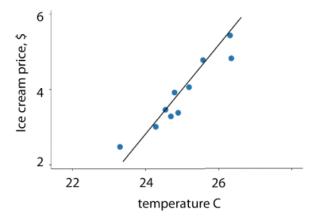


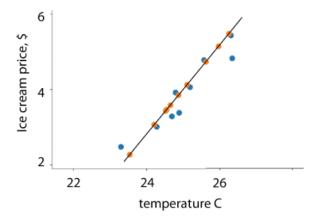
probabilistic PCA

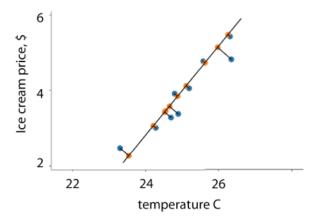
# **Ice Cream conspiracy**

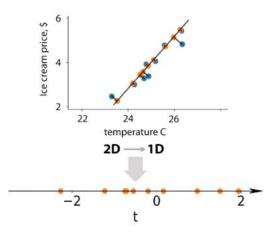










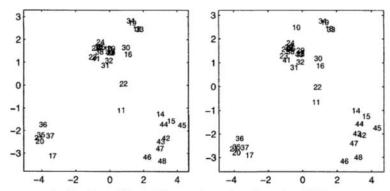


Dimensional reductions

Fast

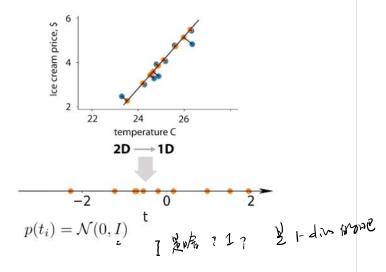
Want: formulate per in probabilistic term?

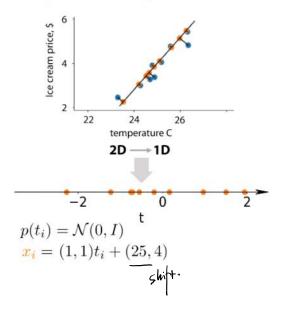
Why? Account for missing data

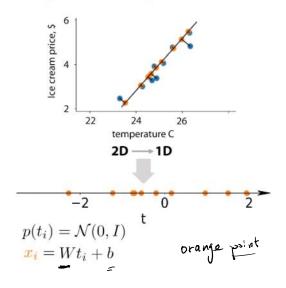


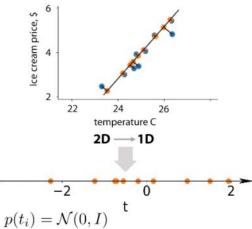
Projection of the Tobamovirus data by using PCA on the full data set and PPCA with 136 missing values

[source: Tipping, M. E., & Bishop, C. M. (1999), Probabilistic principal component analysis]

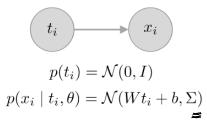


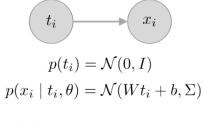






 $p(t_i) = \mathcal{N}(0, I)$  $x_i = Wt_i + b + \varepsilon_i, \quad \varepsilon_i \sim \mathcal{N}(0, \Sigma)$  blue points = or angle points + some noise 1-dim => 21 plus some ganssian noise





 $\max_{\theta} p(X \mid \theta)$ 

$$p(t_i) = \mathcal{N}(0, I)$$

$$p(x_i \mid t_i, \theta) = \mathcal{N}(Wt_i + b, \Sigma)$$

$$\max_{\theta} p(X \mid \theta) = \prod_{i=1}^{N} p(x_i \mid \theta)$$

$$t_i$$

$$p(t_i) = \mathcal{N}(0, I)$$
$$p(x_i \mid t_i, \theta) = \mathcal{N}(Wt_i + b, \Sigma)$$

$$\max_{\theta} p(X \mid \theta) = \prod_{i=1}^{N} p(x_i \mid \theta)$$
$$= \prod_{i} \int p(x_i \mid t_i, \theta) p(t_i) dt_i$$

How is it even possible for Expectation Maximization algorithm to maximize a function (  $f(\theta)=p(X\mid\theta)$ ) without being able to compute the value of the function, let alone the gradient  $\nabla_{\theta}\ f(\theta)$ ?

- We cannot compute the value of the function (because of intractable integrals), but we can approximate it and so we can build an approximate EM algorithm which will not guarantee as the exact solution (hence no magic), but will usually work on practice.
  - It's actually impossible.
  - We don't need to be able to compute the value of the function to build its lower bound and to optimize this bound.

#### This should not be selected

It's true that we can compute some lower bound even if we know very little about the function (e.g.  $p(X\mid\theta)\geq 0$  for any distribution p), but the lower bound will not necessarily be useful. When applying EM-algorithm, on the E-step we minimize the gap to 0 and the lower bound becomes exact at the current point (  $\mathcal{L}(\theta^k,q^{k+1})=\log p(X\mid\theta^k)=f(\theta^k)$ ), which means that if we can compute the lower bound  $\mathcal L$  at any given point  $\theta$ , we can compute the original function at any point as well.

 $=\prod_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{gap to 0 and the lower bound} \\ \mathcal{L}(\theta^{k},q^{k+1}) = \log p(X\mid\theta^{k}) \\ \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$   $=\bigcap_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$   $=\bigcap_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$   $=\bigcap_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$   $=\bigcap_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$   $=\bigcap_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$   $=\bigcap_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$   $=\bigcap_{i}\int p(x_{i}\mid t_{i},\theta)p(t_{i})dt_{i} \qquad \text{lower bound $\mathcal{L}$ at any given practice of the same because as well.}$ 

$$p(t_i) = \mathcal{N}(0, I)$$

$$p(x_i \mid t_i, \theta) = \mathcal{N}(Wt_i + b, \Sigma)$$

$$\max_{\theta} p(X \mid \theta) = \prod_{i=1}^{N} p(x_i \mid \theta)$$

$$= \prod_{i} \int p(x_i \mid t_i, \theta) p(t_i) dt_i$$

$$\text{conjugacy, } \mathcal{N}(\mu_i, \Sigma_i) \longrightarrow \text{conjugate...}$$
The MLE of this formula is the same as pea-
$$\text{graste of fine? The poblarity interpretation:}$$

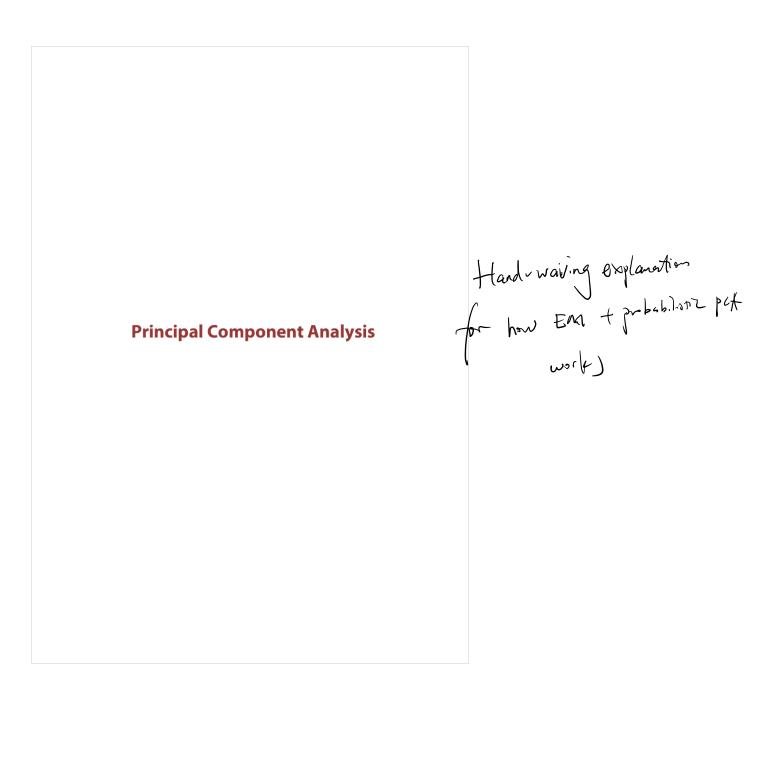
$$\text{graste of fine? The poblarity interpretation:}$$

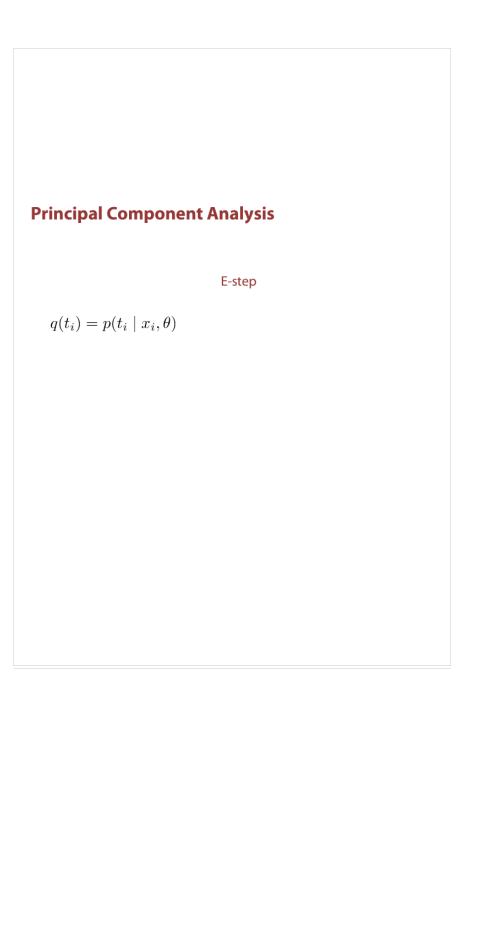
$$\text{still with!}$$

$$\text{there it is conjugate...}$$

$$\text{problem with analytical solution. No need for EM
$$\text{there it is conjugate...}$$$$

probabilistic PCA2



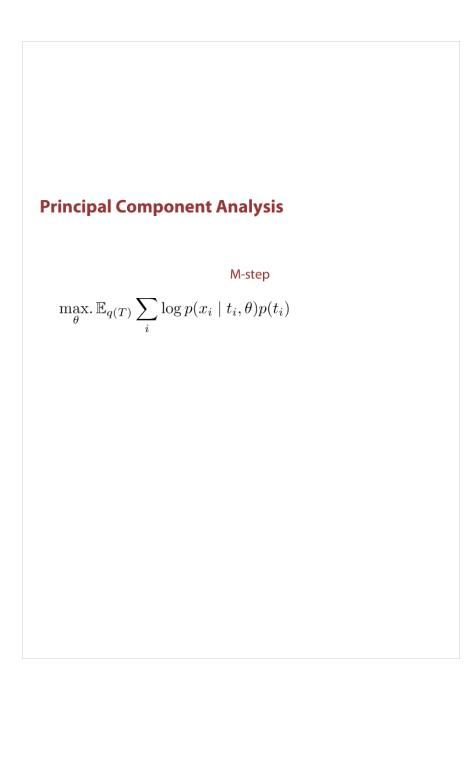


E-step

$$q(t_i) = p(t_i \mid x_i, \theta) = \frac{p(x_i \mid t_i, \theta)p(t_i)}{Z}$$

E-step

$$q(t_i) = p(t_i \mid x_i, \theta) = \frac{p(x_i \mid t_i, \theta)p(t_i)}{Z}$$
$$= \mathcal{N}(\widetilde{\mu}_i, \widetilde{\Sigma}_i)$$



?

$$\max_{\theta} \mathbb{E}_{q(T)} \sum_{i} \log p(x_i \mid t_i, \theta) p(t_i)$$
$$= \sum_{i} \mathbb{E}_{q(t_i)} \log \left( \frac{1}{Z} \exp (\dots) \exp (\dots) \right)$$

$$\max_{\theta} \mathbb{E}_{q(T)} \sum_{i} \log p(x_i \mid t_i, \theta) p(t_i)$$

$$= \sum_{i} \log \frac{1}{Z}$$

$$+ \sum_{i} \mathbb{E}_{q(t_i)} \log (\exp (\dots) \exp (\dots))$$

$$\max_{\theta} \mathbb{E}_{q(T)} \sum_{i} \log p(x_i \mid t_i, \theta) p(t_i)$$

$$= \sum_{i} \log \frac{1}{Z}$$

$$+ \sum_{i} \mathbb{E}_{q(t_i)} \log \left( \exp(\ldots) \exp\left(-\frac{t_i^2}{2}\right) \right)$$

$$\begin{aligned} & \max_{\theta} \mathbb{E}_{q(T)} \sum_{i} \log p(x_i \mid t_i, \theta) p(t_i) \\ & = \sum_{i} \log \frac{1}{Z} \\ & + \sum_{i} \mathbb{E}_{q(t_i)} \log \left( \exp \left( -\frac{(x - Wt_i - b)^2}{2\sigma^2} \right) \exp \left( -\frac{t_i^2}{2} \right) \right) \end{aligned}$$

$$\begin{aligned} &\max_{\theta} \mathbb{E}_{q(T)} \sum_{i} \log p(x_{i} \mid t_{i}, \theta) p(t_{i}) \\ &= \sum_{i} \log \frac{1}{Z} \\ &+ \sum_{i} \mathbb{E}_{q(t_{i})} \log \left( \exp \left( -\frac{(x - Wt_{i} - b)^{2}}{2\sigma^{2}} \right) \exp \left( -\frac{t_{i}^{2}}{2} \right) \right) \\ &\xrightarrow{at_{i}^{2} + ct_{i} + d} \\ &\text{can piece into the probability of the proba$$

### **Summary**

#### Probabilistic formulation of PCA

- Allows for missing values
- Straightforward iterative scheme for large dimensionalities

Don't need linear Algebra. 1

- Can do mixture of PPCA
- Hyperparameter tuning (number of components or choose between diagonal and full covariance)

| 3. | Select real-world problems which can be modeled using Gaussian Mixture Model (GMM)                                  |
|----|---|
|    | Amount of time till the next bus arrival  |
|    | Un-selected is correct  |
|    | Blood type distribution of people of different ethnicities  |
|    | This should not be selected  Blood type is discrete variable, so it can not be modeled using Gaussian distribution. |
|    | Height distribution of people of different ethnicities  |
|    | Correct For each ethnicity we can model height using Gaussian distribution.   |
|    | Rainfall measurement within 4 different seasons   |
|    | Correct For each season rainfall measurement can be modeled using Gaussian distribution.                            |
|    |   |
| 5  | Select correct statements about Probabilistic Principle Component Analysis (PPCA)                                   |
|    | PPCA can be computationally more efficient than naive version of its deterministic analog (PCA)                     |
|    | This should be selected   |
|    | After training the model we can sample new data from the resulting distribution                                     |
|    | Correct Revise Probabilistic PCA video  |
|    | PPCA is a linear dimensionality reduction   |
|    | This should be selected   |
|    | PPCA can be used to visialize multidimensional data   |
|    | Correct Revise Prohabilistic PCA video  |

| 4. | Choose reasonable criteriums for stopping EM iterations |  |  |
|----|---|--|--|
|    |   | Parameter values stabilized (changed less than the predefines epsilon in the last iteration)   |  |
|    | Corr  | ect  |  |
|    |   | Log-likelihood lower bound stabilized (changed less than the predefines epsilon in the last iteration)   |  |
|    | Corr  | ect  |  |
|    |   | Constraints of the original optimization problem (e.g. the prior probability weights in GMM should be non-negative and sum up to one) become satisfied |  |
|    | Un-s  | elected is correct   |  |
|    |   | Log-likelihood lower bound reached the predefined constant value   |  |
|    | Un-s  | elected is correct   |  |