#### **Advanced Statistical Computing**

Fall 2018

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#### Review

- Gibbs sampler
- Grouping and collapsing
- Convergence check
- Sequential Monte Carlo
  - Acceptance rejection method
  - Importance sampling

#### Importance sampling

• Importance sampling:

to evaluate 
$$E_f[h(X)] = \int h(x)f(x)dx$$
  
based on generating a sample  $X_1, \dots, X_n$  from a given distribution  $g$  and approximating

$$E_f[h(X)] \approx \frac{1}{m} \sum_{j=1}^m \frac{f(X_j)}{g(X_j)} h(X_j)$$

which is based on

$$E_f[h(X)] = \int_{\aleph} h(x) \frac{f(x)}{g(x)} g(x) dx$$

- For high dimensional problem, how to design trial distribution is challenging.
- Suppose the target density of  $\mathbf{x} = (x_1, x_2, ..., x_d)$ can be decomposed as

$$\pi(\mathbf{x}) = \pi(x_1)\pi(x_2 \mid x_1)\cdots\pi(x_d \mid x_1,...,x_{d-1})$$
 then constructed trial density as

$$g(\mathbf{x}) = g_1(x_1)g_2(x_2 \mid x_1) \cdots g_d(x_d \mid x_1,...,x_{d-1})$$

$$w(\mathbf{x}) = \frac{\pi(x_1)\pi(x_2 \mid x_1) \cdots \pi(x_d \mid x_1, ..., x_{d-1})}{g_1(x_1)g_2(x_2 \mid x_1) \cdots g_d(x_d \mid x_1, ..., x_{d-1})}$$

Suggest a recursive way of computing and monitoring importance weight. Denote

$$\mathbf{x_t} = (x_1, x_2, \dots, x_t)$$

then we have

$$w_t(\mathbf{x_t}) = w_{t-1}(\mathbf{x_{t-1}}) \frac{\pi(x_t \mid \mathbf{x_{t-1}})}{g_t(x_t \mid \mathbf{x_{t-1}})}$$

- Advantages of the recursion scheme
  - Can stop generating further components of x if the partial weight is too small.
  - Can take advantage of  $\pi(x_t | \mathbf{x_{t-1}})$  in designing  $g_t(x_t | \mathbf{x_{t-1}})$
- However, the scheme is impractical since it requires the knowledge of marginal distribution  $\pi(x_t)$ .

- Add another layer of complexity:
- Introduce a sequence of "auxiliary distributions"  $\pi_1(x_1)\pi_2(\mathbf{x}_2)\pi_d(\mathbf{x})$  such that  $\pi_t(\mathbf{x}_t)$  is a reasonable approximation of the marginal distribution  $\pi(\mathbf{x}_t)$ , for  $t=1,\ldots,d$ -1 and  $\pi_d=\pi$ .
- Note the  $\pi_d$  are only required to be known up to a normalizing constant.

### The SIS procedure

For 
$$t = 2, ..., d$$
,

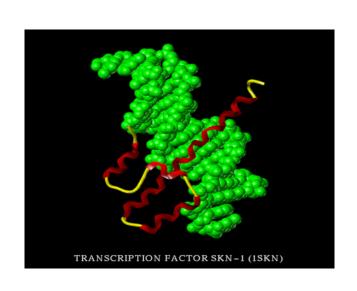
• Draw  $X_t = x_t$  from  $g_t(x_t | x_{t-1})$ , and let

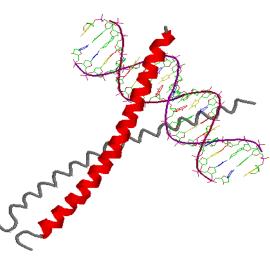
$$\mathbf{x}_{\mathbf{t}} = (\mathbf{x}_{\mathbf{t-1}}, x_t)$$

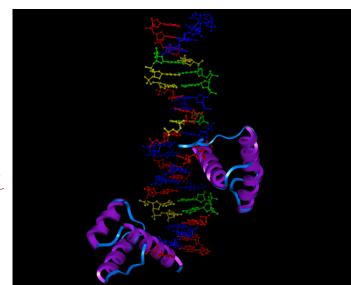
- Compute  $u_t = \frac{\pi_t(\mathbf{x_t})}{\pi_{t-1}(\mathbf{x_{t-1}})g_t(x_t \mid \mathbf{x_{t-1}})}$ and let  $w_t = w_{t-1}u_t$
- $u_t$ : incremental weight.
- The key idea is to breaks a difficult task into manageable pieces.
- If  $w_t$  is getting too small, reject.

# Applications of MCMC and SMC

# Appliation: Transcription Factor Binding Sites Discovery





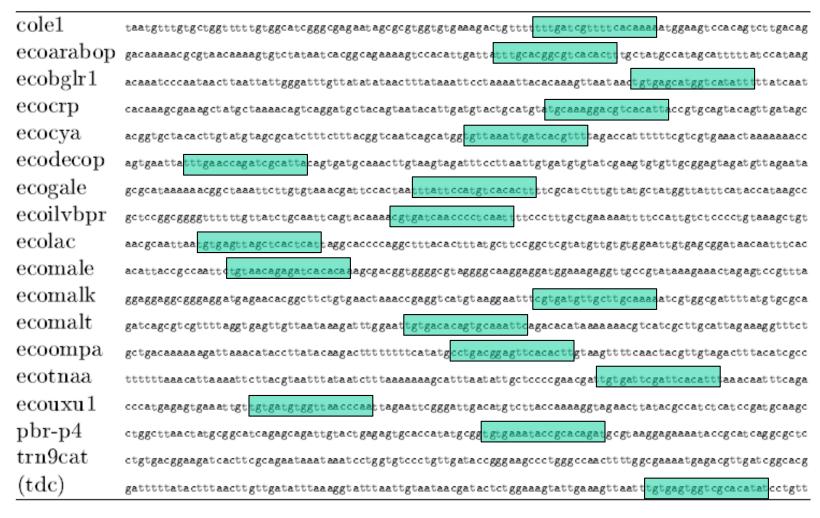


#### Example: cyclic receptor protein (CRP)

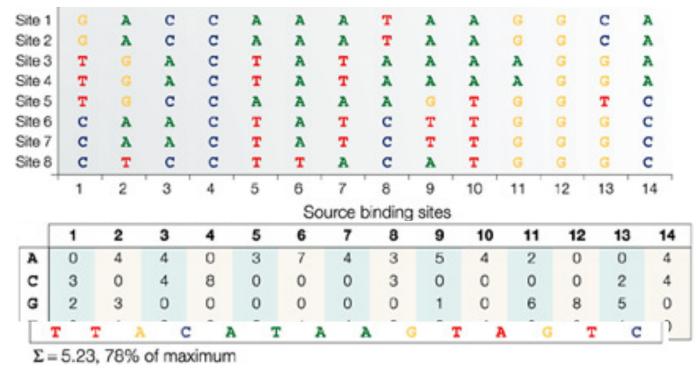
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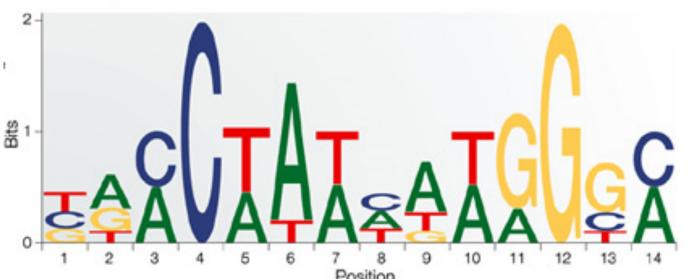
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#### Example: cyclic receptor protein (CRP)



#### Transcription factor binding site (TFBS)





# Existing *de novo* motif finding algorithms

• Consensus Hertz et al. 1990

• Gibbs Motif Sampler Lawrence et al. 1993

• MEME Bailey and Elkan 1994

• AlignACE Roth et al. 1998

• BioProspector Liu et al. 2001

• MDScan Liu et al. 2002

• Mobydick Bussemaker et al. 2000

• • •

Review

Tompa et al. 2005

#### Motif identification model

 $\begin{array}{c} a_{1}\\ a_{2}\\ a_{3}\\ a_{4}\\ c_{5}\\ c_$ 

Alignment variable  $A = \{a_1, a_2, ..., a_J\}$ 

#### Posterior distributions

• The posterior conditional distribution for alignment variable *A* 

$$p(a_{j} = l \mid \boldsymbol{\theta_{0}}, \boldsymbol{\Theta}, \boldsymbol{R_{j}}, \boldsymbol{A_{-j}}) \propto \prod_{k=1}^{4} \theta_{0k}^{h_{k}(\boldsymbol{R_{j}})} \prod_{i=1}^{w} \prod_{k=1}^{4} \left(\frac{\theta_{ik}}{\theta_{0k}}\right)^{h_{k}(r_{j,l+i-1})} \propto \prod_{i=1}^{w} \prod_{k=1}^{4} \left(\frac{\theta_{ik}}{\theta_{0k}}\right)^{h_{k}(r_{j,l+i-1})}$$

DNA sequence data

$$\boldsymbol{R} = (\boldsymbol{R}_1, ..., \boldsymbol{R}_J)$$

Lawrence et al. Science 1993, Liu et al. JASA 1995

### Motif Alignment Model



**The missing data:** Alignment variable:  $A = \{a_1, a_2, ..., a_k\}$ 

- Every **non-site positions** follows a common multinomial with  $p_0 = (p_{0,1}, ..., p_{0,20})$
- Every position i in the motif element follows probability distribution  $p_i = (p_{i,1}, ..., p_{i,20})$

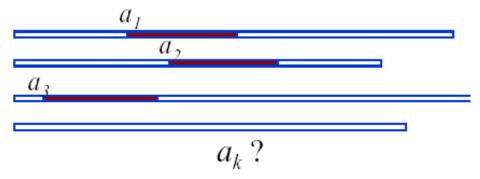
#### Statistical Model

- Objects:
  - Seq: sequence data to search for motif
  - $-\theta_0$ : non-motif (genome background) probability
  - $-\theta$ : motif probability matrix parameter
  - $-\pi$ : site locations
- Problem:  $P(\theta, \pi \mid \text{seq}, \theta_0)$
- Approach: alternately estimate
  - $-\pi$  by  $P(\pi \mid \theta, \text{ seq}, \theta_0)$
  - $-\theta$  by P( $\theta \mid \pi$ , seq,  $\theta_0$ )

#### The Algorithm

- Initialize by choosing random starting positions
- Iterate the following steps many times;
  - Randomly or systematically choose a sequence to exclude
  - Carry out the predictive-updating step to update the starting position
  - Stop when no more observable changes in likelihood.

## The Predictive Updating Step



• Compute predictive frequencies of each position *i* in motif

 $c_{ij}$  = count of amino acid type j at position i.  $c_{0j}$  = count of amino acid type j in all non-site positions.  $q_{ij} = (c_{ij} + b_j)/(K-1+B)$ ,  $B=b_1 + \cdots + b_K$  "pseudo-counts"

• Sample from the predictive distribution of  $a_k$ 

$$P(a_k = l+1) \propto \prod_{i=1}^{w} \frac{q_{i,R_k(l+i)}}{q_{0,R_k(l+i)}}$$

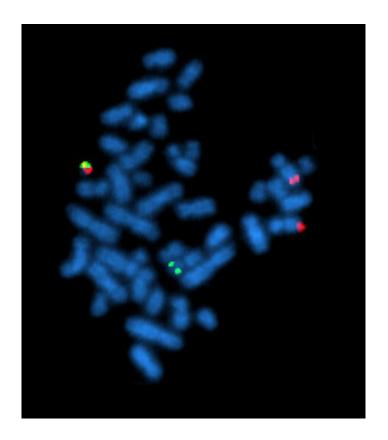
#### References

- Lawrence et al. (1993) Science.
- Liu, Neuwald and Lawrence (1995) *JASA*.
- Liu and Lawrence (1999) Bioinformatics.

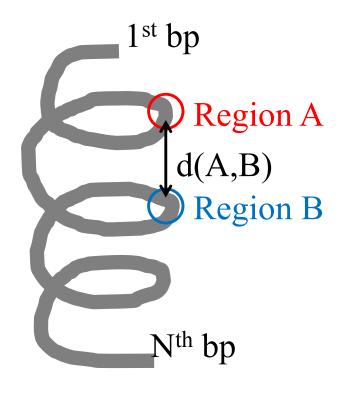
# Infer the 3D shape of chromosomes

### Microscopic Methods

• Fluorescent *in situ* hybridization (FISH)

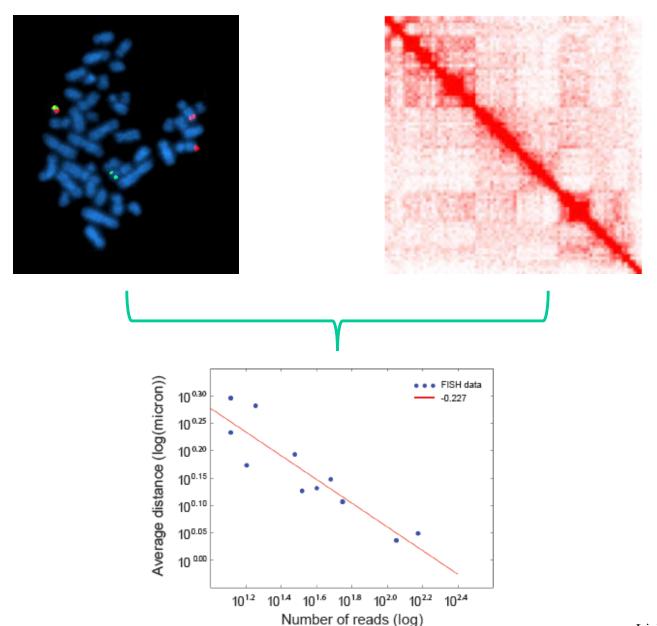


#### FISH Data Representation

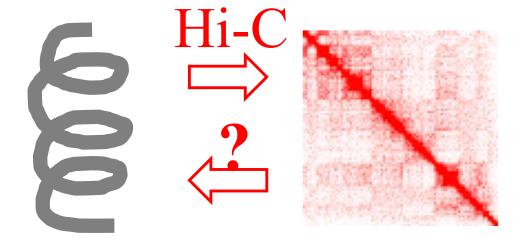


3D chromosomal structure

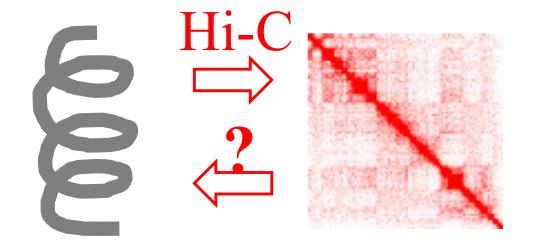
#### Contact Frequency vs. Spatial Distance



# Problem setting

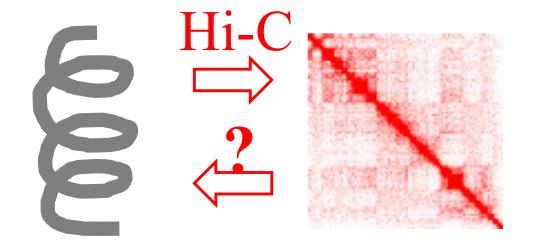


#### Problem setting



- Challenges:
- > Sequencing uncertainties
- ➤ Biases: enzyme, GC content, mappability

#### Problem setting

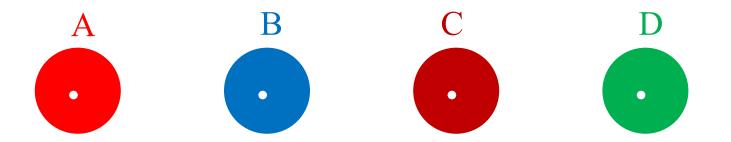


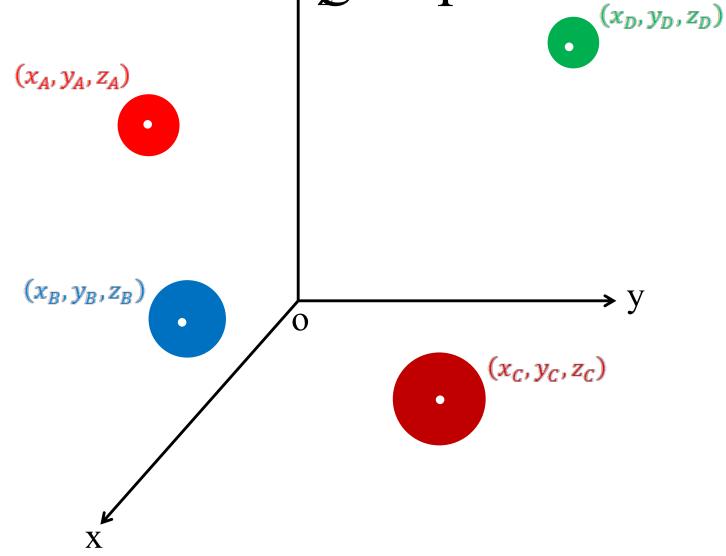
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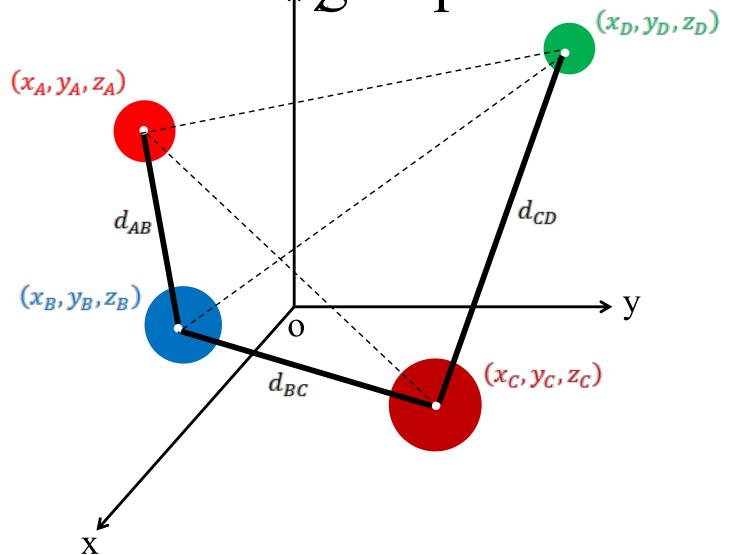
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ACGTAGCTAGATACTGTAGTGTAGTTTTGGAACCTGAGGG

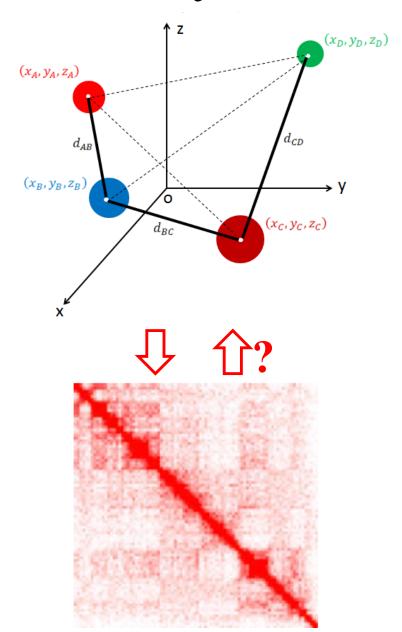
ACGTAGCTAG ATACTGTAGT GTAGTTTGGA ACCTGAGGG



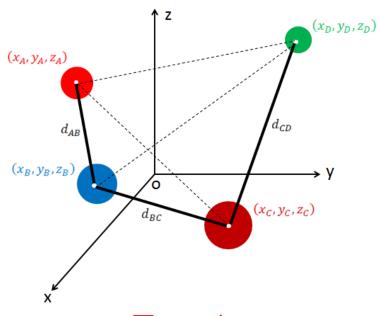




# Bayesian Statistical Model



### Bayesian Statistical Model



 $u_{ij}$ : # of reads between loci i and j $(x_i, y_i, z_i)$ : Euclidian coordinates of locus i

 $d_{ii}$ : spatial distance between loci i and j

 $e_i$ : # of enzyme cut site in locus i

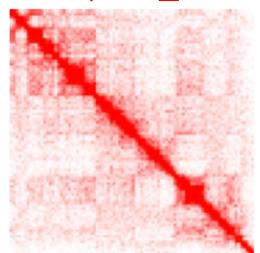
 $g_i$ : GC content of locus i

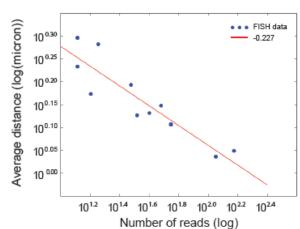
 $m_i$ : mappability of locus i



Hi-C read counts: population summation  $u_{ij} \sim Poisson(\theta_{ij})$ 

Hi-C read counts vs. spatial distance: log-log linear





$$\log(\theta_{ij}) = \beta_0 + \beta_1 \log(d_{ij})$$
  
+  $\beta_e \log(e_i e_j) + \beta_g \log(g_i g_j)$   
+  $\beta_m \log(m_i m_j)$ 

### Bayesian Statistical Model

#### Likelihood:

$$L(u_{ij}, 1 \le i < j \le N | x_i, y_i, z_i, 1 \le i \le N, \beta_0, \beta_1, \beta_e, \beta_g, \beta_m) = \prod_{1 \le i < j \le N} \frac{e^{-\theta_{ij}} \theta_{ij}^{u_{ij}}}{u_{ij}!}$$

$$\log(\theta_{ij}) = \beta_0 + \beta_1 \log \left( \sqrt{(x_i - x_j)^2 + (y_i - y_j)^2 + (z_i - z_j)^2} \right) + \beta_e \log(e_i e_j) + \beta_g \log(g_i g_j) + \beta_m \log(m_i m_j)$$

### Bayesian Statistical Model

• Likelihood:  $\binom{N}{2}$  data points, 3N + 5 parameters

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Posterior distribution

$$\pi(x_i, y_i, z_i, 1 \le i \le N, \beta_0, \beta_1, \beta_e, \beta_g, \beta_m | u_{ij}, 1 \le i < j \le N)$$

$$\propto L(u_{ij}, 1 \le i < j \le N | x_i, y_i, z_i, 1 \le i \le N, \beta_0, \beta_1, \beta_e, \beta_g, \beta_m) prior$$

• Algorithm: Bayesian 3D constructor for Hi-C data (BACH)

$$\pi(x_i, y_i, z_i, 1 \leq i \leq N, \beta_0, \beta_1, \beta_e, \beta_g, \beta_m \mid u_{ij}, 1 \leq i < j \leq N)$$

Algorithm: Bayesian 3D constructor for Hi-C data (BACH)

$$\pi(x_i, y_i, z_i, 1 \le i \le N, \beta_0, \beta_1, \beta_e, \beta_g, \beta_m | u_{ij}, 1 \le i < j \le N)$$

➤ Initialization 1: use Poisson regression to obtain the initial values for

$$eta_0$$
,  $eta_e$ ,  $eta_g$ ,  $eta_m$  . Set  $eta_1 = -1$  .

$$u_{ij} \sim Poisson(\theta_{ij}) \log(\theta_{ij}) = \beta_0 + \beta_e \log(e_i e_j) + \beta_g \log(g_i g_j) + \beta_m \log(m_i m_j)$$

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Initialization 2: use sequential important sampling to get the initial 3D chromosomal structure  $\{x_i, y_i, z_i, 1 \le i \le N\}$ .

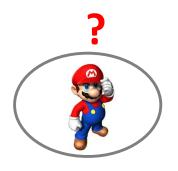
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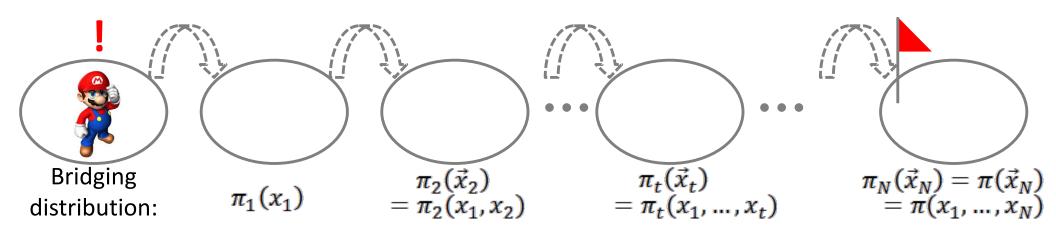
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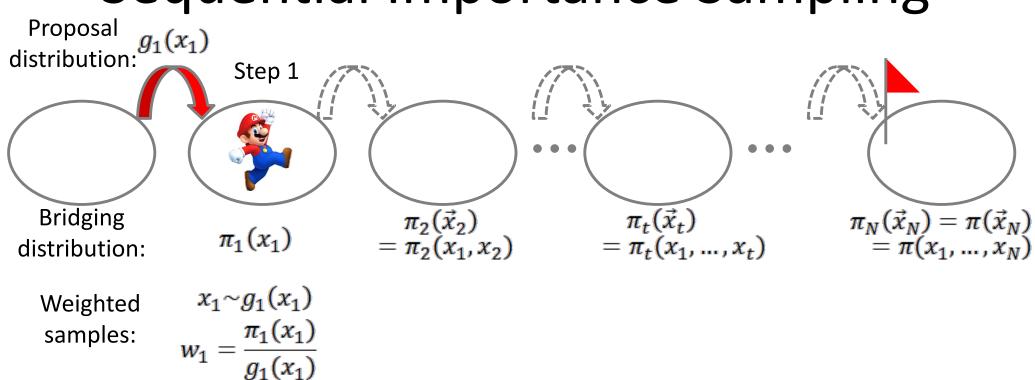
- Initialization 2: use sequential important sampling to get the initial 3D chromosomal structure  $\{x_i, y_i, z_i, 1 \le i \le N\}$ .
- ➤ Refinement: use Gibbs sampler with hybrid Monte Carlo to refine the initial values for parameters.



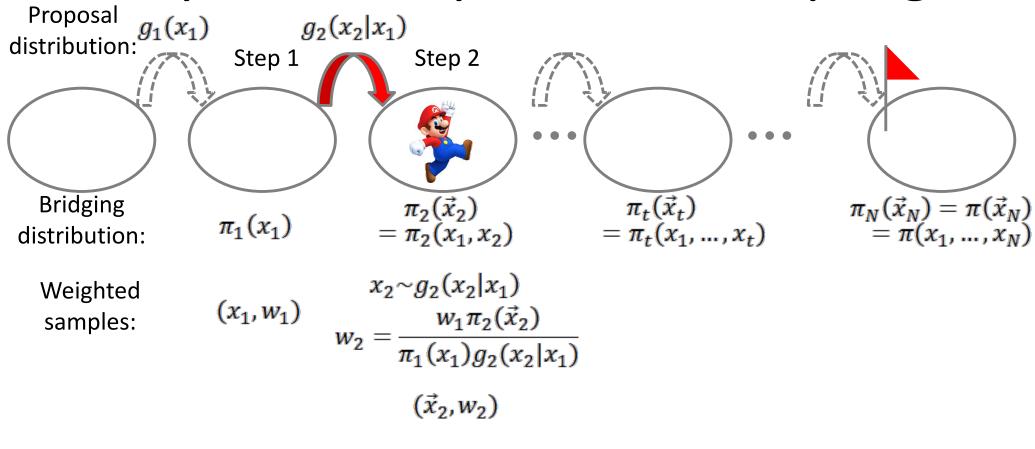


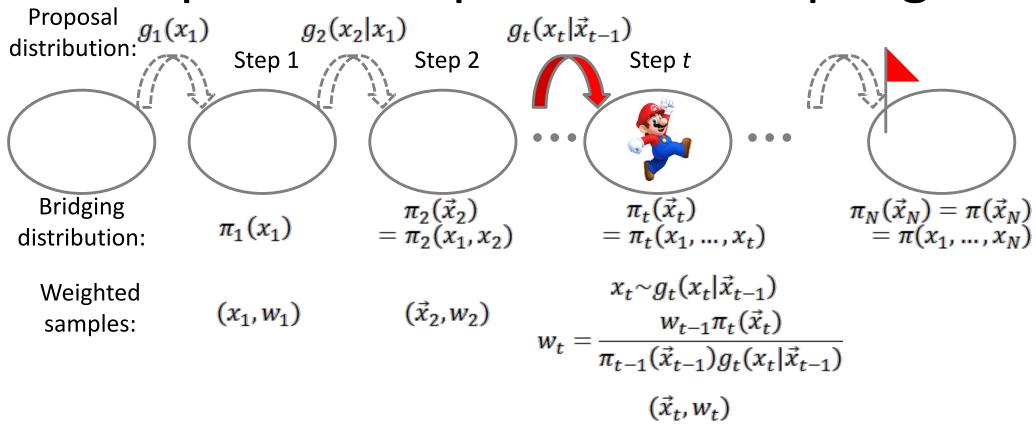
$$\pi(\vec{x}_N) = \pi(x_1, \dots, x_N)$$

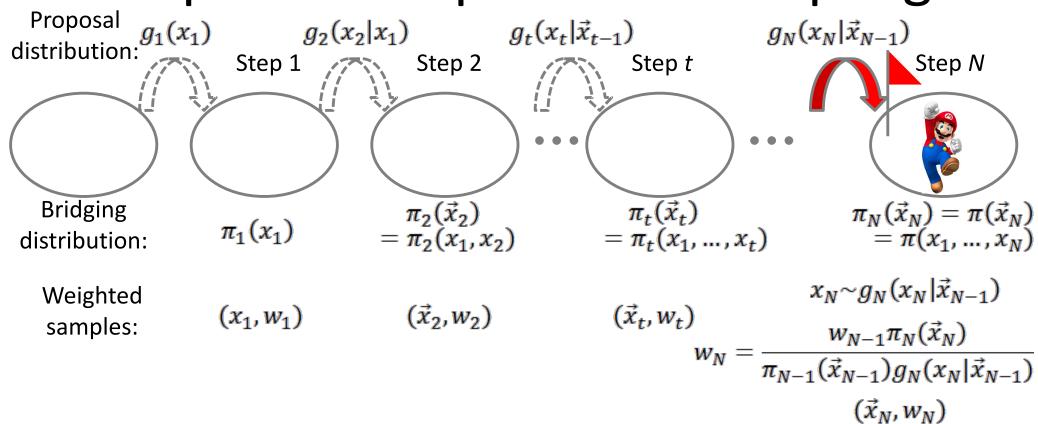


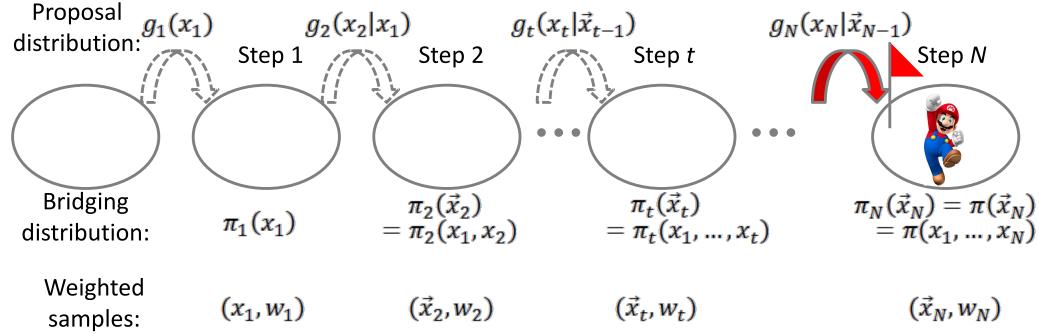


 $(x_1, w_1)$ 









Sequential Importance Sampling (SIS) Algorithm:

- (1) Design bridging distributions  $\pi_t(\vec{x}_t)$  and proposal distributions  $g_t(x_t|\vec{x}_{t-1})$
- (2) Sequentially draw weighted samples  $x_t \sim g_t(x_t | \vec{x}_{t-1})$ , and update weight

$$w_t = \frac{w_{t-1}\pi_t(\vec{x}_t)}{\pi_{t-1}(\vec{x}_{t-1})g_t(x_t|\vec{x}_{t-1})}$$

#### SIS in BACH: Outline

 Goal: use sequential importance sampling to sequentially put N loci into 3D space, i.e. sample from:

$$\pi(x_i, y_i, z_i, 1 \le i \le N | u_{ij}, 1 \le i < j \le N)$$

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 Goal: use sequential importance sampling to sequentially put N loci into 3D space, i.e. sample from:

$$\pi(x_i, y_i, z_i, 1 \le i \le N | u_{ij}, 1 \le i < j \le N)$$

Bridging distributions:

$$\pi_t(x_i, y_i, z_i, 1 \le i \le t | u_{ij}, 1 \le i < j \le t)$$

#### SIS in BACH: Outline

 Goal: use sequential importance sampling to sequentially put N loci into 3D space, i.e. sample from:

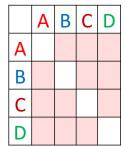
$$\pi(x_i, y_i, z_i, 1 \le i \le N | u_{ij}, 1 \le i < j \le N)$$

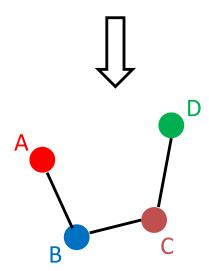
Bridging distributions:

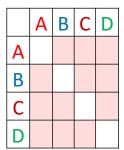
$$\pi_t(x_i, y_i, z_i, 1 \le i \le t | u_{ij}, 1 \le i < j \le t)$$

 Proposal distributions (given the first t-1 loci, put the t th locus in to 3D space):

$$g_t(x_t, y_t, z_t | x_i, y_i, z_i, 1 \le i \le t - 1, u_{ij}, 1 \le i < j \le t)$$





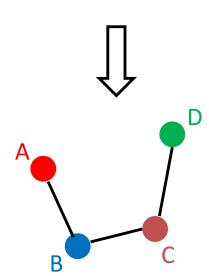


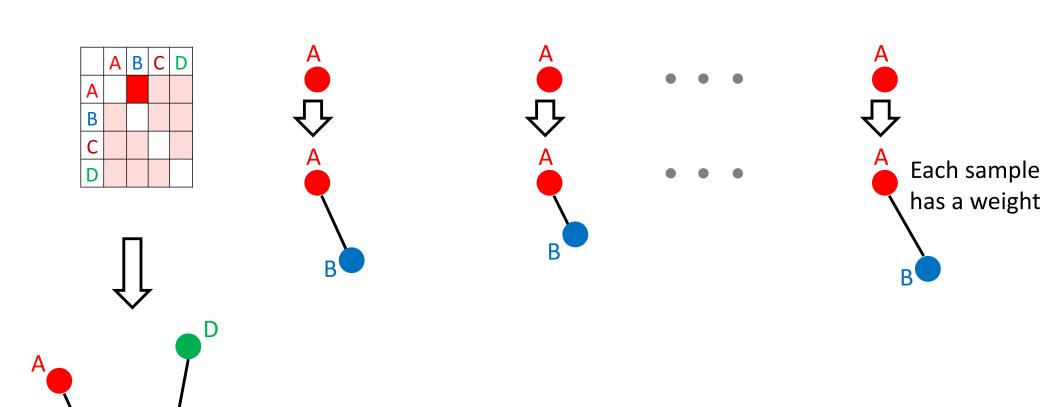


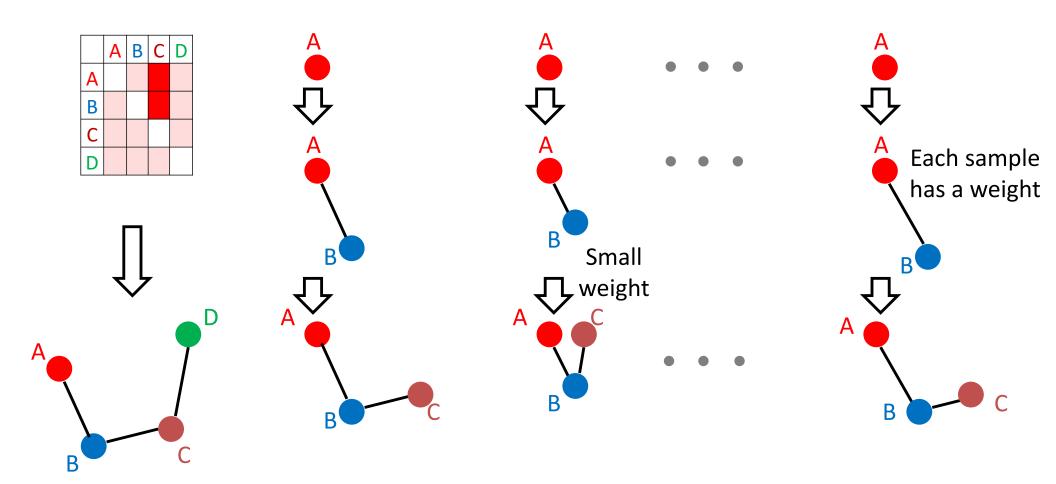


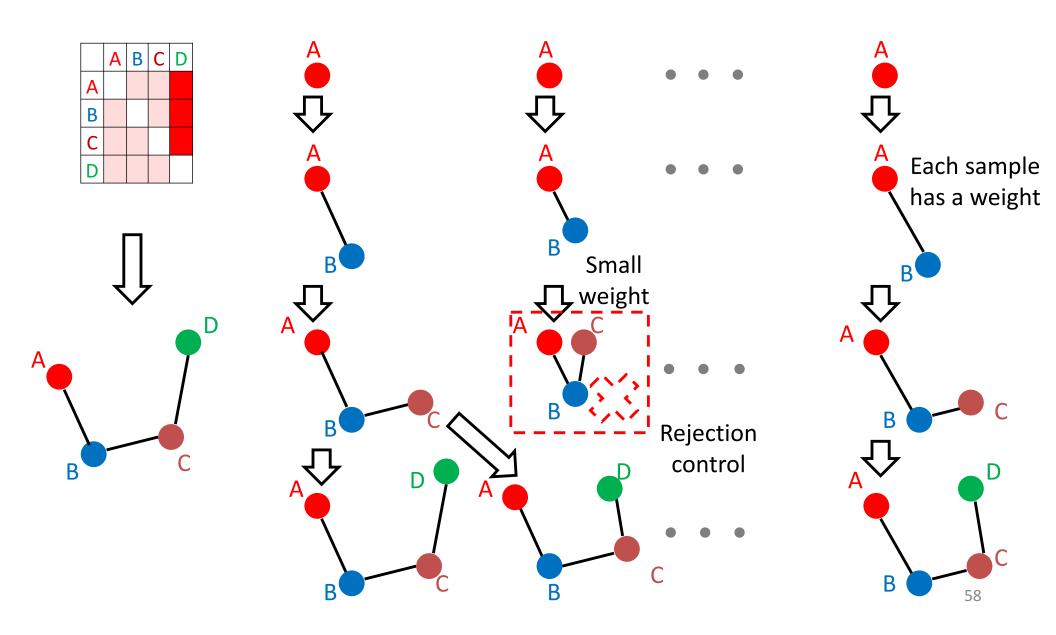












### Hybrid Monte Carlo

 Goal: do efficient group move to refine initial 3D chromosomal structure, since local 3D coordinates are highly correlated.

Combine molecular dynamics with Metropolis acceptance-rejection rule.

### Hybrid Monte Carlo in BACH

Goal: sampling from

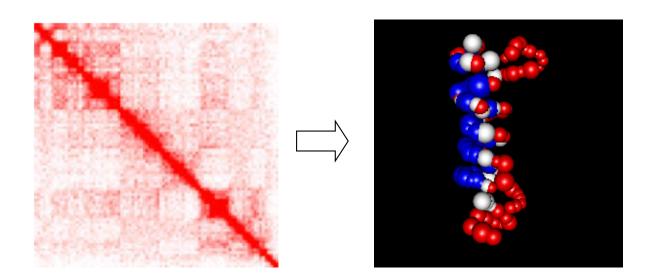
$$\pi(x_i, y_i, z_i, 1 \le i \le N | u_{ij}, 1 \le i < j \le N)$$

• Take partial derivate of log likelihood over 3D coordinates  $(x_i, y_i, z_i, 1 \le i \le N)$ .

 Run the leap-frog algorithm, adaptively tune the time interval to achieve acceptance rate ~ 90%.

### Conclusions

- BACH: reconstruct chromosome 3D structures from Hi-C data
- Remove systematic biases
- Predicted spatial distances are consistent with FISH data
- Elongation of chromatin is highly associated with genetic/epigenetic features.
- Separation of compartments of A and B can be visualized.



### References

Hu M, Deng K, Qin ZS, Dixon J, Selvaraj S, Fang J, Ren B, Liu JS.
 (2013) Bayesian inference of three-dimensional chromosomal organization. *PLoS Comput Biol.* 9 e1002893.

http://www.people.fas.harvard.edu/~junliu/BACH/

 Dixon JR, Selvaraj S, Yue F, Kim A, Li Y, Shen Y, Hu M, Liu JS and Ren B. (2012) Topological domains in mammalian genomes identified by analysis of chromatin interactions. *Nature*, 485, 376-380.

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