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# Assignment - 1: BUAN 6312 Harikrishna Dev HXD220000

## Answers

1. Use the data in CHARITY [obtained from Franses and Paap (2001)] to answer the following questions:

- What is the average gift in the sample of 4,268 people (in Dutch guilders)? What percentage of people gave no gift?

Average gift is **7.4444704**.

Percentage of people who have no gift is **0.60004687** (i.e. 60%)

```
. use "C:\Users\hxd220000\Desktop\Data Sets- STATA\charity.dta"

. egen gift_mean = mean(gift)

. display gift_mean
7.4444704

. egen count_total = count(respond)

. egen count_gift = count(gift), by (respond)

. display count_gift
2561

. display count_total
4268

. gen count_per = count_gift /count_total

. display count_per
.60004687
```

- What is the average mailings per year? What are the minimum and maximum values?

Average Mailings per year: 2.049555

Min Mailings per year: 0.25

Max mailings per year: 3.5

```
. summarize mailsyear
```

Variable	Obs	Mean	Std. dev.	Min	Max
-----+-----					
mailsyear	4,268	2.049555	.66758	.25	3.5

- Estimate the model.

$$gift = \beta_0 + \beta_1 \times mailsyear + u$$

by OLS and report the results in the usual way, including the sample size and R-squared.

$$gift = 2.01408 + 2.649546 \times mailsyear + u$$

Sample size = 4268 R-sqr = 0.0138

```
. reg gift mailsyear
```

Source	SS	df	MS	Number of obs	=	4,268
Model	13349.7251	1	13349.7251	F(1, 4266)	=	59.65
Residual	954750.114	4,266	223.804528	Prob > F	=	0.0000
				R-squared	=	0.0138
				Adj R-squared	=	0.0136
Total	968099.84	4,267	226.880675	Root MSE	=	14.96

gift	Coefficient	Std. err.	t	P> t	[95% conf. interval]
mailsyear	2.649546	.3430598	7.72	0.000	1.976971 3.322122
_cons	2.01408	.7394696	2.72	0.006	.5643347 3.463825

- Interpret the slope coefficient. If each mailing costs one guilder, is the charity expected to make a net gain on each mailing? Does this mean the charity makes a net gain on every mailing? Explain.

The slope coefficient estimates that one mailing might be resulting in 2.65 additional guilders per year. So, the expected profit per mailing is  $2.64 - 1 = 1.65$  *guilders*. This value could be lesser or higher based on different scenarios, but the overall average using this model would be the given values.

- What is the smallest predicted charitable contribution in the sample? Using this simple regression analysis, can you ever predict zero for gift?

We know that  $\min(mailsyear) = 0.25$ . This makes the smallest predicted contribution  $gift = 2.01408 + 2.649546 \times 0.25 = 2.6764665$ . As  $\min(gift) = 2.67$ , we can't get zero as a predict value.

2. The file CEOSAL2 contains data on 177 chief executive officers and can be used to examine the effects of firm performance on CEO salary.

- Estimate a model relating annual salary to firm sales and market value. Make the model of the constant elasticity variety for both independent variables. Report the results in the usual way.

As `lsalary`, `lmktval` and `lsales` are created in the dataset, we are able to generate the `lm` model.

$$\ln(\text{salary}) = 4.620917 + 0.106708 \times \ln(\text{mktval}) + 0.1621283 \times \ln(\text{sales}) + u$$

```
. use "CEOSAL2.dta"
```

```
. reg lsalary lsales lmktval
```

Source	SS	df	MS	Number of obs	=	177
Model	19.3365617	2	9.66828083	F(2, 174)	=	37.13
Residual	45.3096514	174	.260400295	Prob > F	=	0.0000
				R-squared	=	0.2991
				Adj R-squared	=	0.2911
Total	64.6462131	176	.367308029	Root MSE	=	.51029

lsalary	Coefficient	Std. err.	t	P> t	[95% conf. interval]
lsales	.1621283	.0396703	4.09	0.000	.0838315 .2404252
lmktval	.106708	.050124	2.13	0.035	.0077787 .2056372
_cons	4.620917	.2544083	18.16	0.000	4.118794 5.123041

- Add profits to the model from part (i), re-estimate the model and report the results in the usual way. Why can this variable not be included in logarithmic form? Would you say that these firm performance variables explain most of the variation in CEO salaries?

As profits can be negative, we cannot use the log function on the profits values.

```
. reg lsalary lsales lmktval profits
```

Source	SS	df	MS	Number of obs	=	177
Model	19.3509799	3	6.45032663	F(3, 173)	=	24.64
Residual	45.2952332	173	.261822157	Prob > F	=	0.0000
				R-squared	=	0.2993
				Adj R-squared	=	0.2872
Total	64.6462131	176	.367308029	Root MSE	=	.51169

lsalary	Coefficient	Std. err.	t	P> t	[95% conf. interval]
lsales	.1613683	.0399101	4.04	0.000	.0825949 .2401416
lmktval	.0975286	.0636886	1.53	0.128	-.0281782 .2232354
profits	.0000357	.000152	0.23	0.815	-.0002643 .0003356
_cons	4.686924	.3797294	12.34	0.000	3.937425 5.436423

From the regression results, we can that *profits* has a *p value* > 0.05. This helps to conclude that the assumption that co-efficient of *profits* = 0 cannot be rejected

- Add the variable *ceoten* to the model in part (ii), re-estimate the model and report the results in the usual way. What is the estimated percentage return for another year of CEO tenure, holding other factors fixed?

$$\ln(\text{salary}) = 4.620917 + 0.106708 \times \ln(\text{mktval}) + 0.1621283 \times \ln(\text{sales}) + 0.0116847 \times \text{ceoten} + u$$

```
reg lsalary lsales lmktval profits ceoten
```

Source	SS	df	MS	Number of obs	=	177
Model	20.5768102	4	5.14420254	F(4, 172)	=	20.08
Residual	44.0694029	172	.256217459	Prob > F	=	0.0000
				R-squared	=	0.3183
				Adj R-squared	=	0.3024
Total	64.6462131	176	.367308029	Root MSE	=	.50618

lsalary	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
lsales	.1622339	.0394826	4.11	0.000	.0843012	.2401667
lmktval	.1017598	.063033	1.61	0.108	-.022658	.2261775
profits	.0000291	.0001504	0.19	0.847	-.0002677	.0003258
ceoten	.0116847	.005342	2.19	0.030	.0011403	.022229
_cons	4.55778	.3802548	11.99	0.000	3.807213	5.308347

From the above regression equation, we can conclude

$$\frac{\frac{\Delta \text{salary}}{\text{salary}}}{\Delta \text{ceoten}} = 0.012$$

An addition in a year of tenure results in an increase in their salary by 1.2% on an average.

- Find the sample correlation coefficient between the variables  $\ln(\text{mktval})$  and  $\text{profits}$ . Are these variables highly correlated? What does this say about the OLS estimators? [Hint: You can use the stata command `correlate.`]

The covariance between  $\ln(\text{mktval})$  and  $\text{profits}$  is **0.78**. This implies there is a highly correlated variables. This doesn't affect the OLS estimators as we assume they are independent variables. Correlation doesn't mean causation.

```
. correlate lmktval profits
(obs=177)
```

	lmktval	profits
lmktval	1.0000	
profits	0.7769	1.0000

3. Refer to the example used in Lecture 4 to compare the returns to education at junior colleges and four-year colleges. The model after rearrangement is  $\log(wage) = \beta_0 + \theta_1 jc + \beta_2 totcoll + \beta_3 exper + u$ , where *totcoll* is total years of college. Use the data set *TWOYEAR*, which comes from Kane and Rouse (1995).

- Run the regression above and report the OLS estimates in the usual form, including the standard errors, sample size and R-squared. How do you interpret  $\theta_1$ ? Is it statistically significant?

$$\ln(wage) = 1.472326 - 0.0101795 \times jc + 0.0768762 \times totcoll + 0.0049442 \times exper + u$$

$\theta_1$  is the percentage increase in wage for unit increase in Junior college credit. As *p values* > 0.05, we can conclude the parameter *jc* is not significant.

```
. use "TWOYEAR.dta"
```

```
. reg lwage jc totcoll exper
```

Source	SS	df	MS	Number of obs	=	6,763
Model	357.752575	3	119.250858	F(3, 6759)	=	644.53
Residual	1250.54352	6,759	.185019014	Prob > F	=	0.0000
				R-squared	=	0.2224
				Adj R-squared	=	0.2221
Total	1608.29609	6,762	.237843255	Root MSE	=	.43014

lwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]
jc	-.0101795	.0069359	-1.47	0.142	-.0237761 .003417
totcoll	.0768762	.0023087	33.30	0.000	.0723504 .0814021
exper	.0049442	.0001575	31.40	0.000	.0046355 .0052529
_cons	1.472326	.0210602	69.91	0.000	1.431041 1.51361

- The variable *phsrank* is the person's high school percentile. (A higher number is better. For example, 90 means you are ranked better than 90 percent of your graduating class.) Find the smallest, largest, and average *phsrank* in the sample

Parameters of *phsrank* are as follows: Minimum: 0 Maximum: 100 Average: 56.15703

```
. summarize phsrank
```

Variable	Obs	Mean	Std. dev.	Min	Max
phsrank	6,763	56.15703	24.27296	0	99

- Add *phsrank* to the model and report the OLS estimates in the usual form. Is *phsrank* statistically significant? How much is 10 percentage points of high school rank worth in terms of wage?

$$\ln(wage) = 1.458747 - 0.0093108 \times jc + 0.0754756 \times totcoll + 0.0049396 \times exper + 0.0003032 \times phsrank -$$

We can see that *phsrank* has a *p value* > 0.05 . Therefore, we can conclude that *phsrank* is not significant.

From the regression model, we can conclude

$$\frac{\frac{\Delta wage}{wage}}{\Delta phsrank} = 0.0003032$$

This implies that a 10% increase in *phsrank* would result in a 0.3% increase in wage.

```
reg lwage jc totcoll exper phsrank
```

Source	SS	df	MS	Number of obs	=	6,763
Model	358.050568	4	89.5126419	F(4, 6758)	=	483.85
Residual	1250.24552	6,758	.185002297	Prob > F	=	0.0000
				R-squared	=	0.2226
				Adj R-squared	=	0.2222
Total	1608.29609	6,762	.237843255	Root MSE	=	.43012

lwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
jc	-.0093108	.0069693	-1.34	0.182	-.0229728	.0043512
totcoll	.0754756	.0025588	29.50	0.000	.0704595	.0804918
exper	.0049396	.0001575	31.36	0.000	.0046308	.0052483
phsrank	.0003032	.0002389	1.27	0.204	-.0001651	.0007716
_cons	1.458747	.0236211	61.76	0.000	1.412442	1.505052

- Compare regression results in (i) and (iii), does adding phsrank to the model substantively change the conclusions on the returns to two- and four-year colleges? Explain.

After adding an additional variable *phsrank*, we see not many changes in terms of magnitude of coefficients of the linear model generated. Therefore, the base point remains unchanged: the return to a junior college is estimated to be somewhat smaller, but the difference is not significant and standard significant levels.

- The data set contains a variable called *id*. Explain why if you add *id* to the model you expect it to be statistically insignificant. What is the two-sided p-value?

The variable *id* is a unique identifier for each employee. It doesn't have any relation with wage. The *p value* > 0.05 which shows that it is not significant.

*p values (t test on id) = 0.507*

```
. reg lwage jc totcoll exper phsrank id
```

Source	SS	df	MS	Number of obs	=	6,763
Model	358.132086	5	71.6264171	F(5, 6757)	=	387.13
Residual	1250.16401	6,757	.185017612	Prob > F	=	0.0000
				R-squared	=	0.2227
				Adj R-squared	=	0.2221
Total	1608.29609	6,762	.237843255	Root MSE	=	.43014

lwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
jc	-.0093159	.0069696	-1.34	0.181	-.0229784	.0043467
totcoll	.075414	.0025606	29.45	0.000	.0703943	.0804336
exper	.004941	.0001575	31.37	0.000	.0046322	.0052498
phsrank	.0003179	.00024	1.32	0.185	-.0001525	.0007883
id	1.40e-07	2.10e-07	0.66	0.507	-2.73e-07	5.52e-07
_cons	1.452215	.0255898	56.75	0.000	1.402051	1.502379

4. Use the data set GPA1 to answer this question.

- Run the regression  $\text{colGPA}$  on  $\text{PC}$ ,  $\text{hsGPA}$ , and  $\text{ACT}$  and obtain a 95% confidence interval for  $\beta_{\text{PC}}$ . Is the estimated coefficient statistically significant at the 5% level against a two-sided alternative?

The confidence interval of  $\beta_{\text{PC}}$  is  $(0.0440271, 0.2705913)$ .

The  $p$  value of  $\beta_{\text{PC}} < 0.05$ , which makes its statistically significant.

```
. use "GPA1.dta"
```

```
. reg colGPA PC hsGPA ACT
```

Source	SS	df	MS	Number of obs	=	141
Model	4.25741863	3	1.41913954	F(3, 137)	=	12.83
Residual	15.1486808	137	.110574313	Prob > F	=	0.0000
				R-squared	=	0.2194
				Adj R-squared	=	0.2023
Total	19.4060994	140	.138614996	Root MSE	=	.33253

colGPA	Coefficient	Std. err.	t	P> t	[95% conf. interval]
PC	.1573092	.0572875	2.75	0.007	.0440271 .2705913
hsGPA	.4472417	.0936475	4.78	0.000	.2620603 .632423
ACT	.008659	.0105342	0.82	0.413	-.0121717 .0294897
_cons	1.26352	.3331255	3.79	0.000	.6047871 1.922253

- Discuss the statistical significance of the estimates  $\beta_{\text{hsGPA}}$  and  $\beta_{\text{ACT}}$  in part (i). Is  $\text{hsGPA}$  or  $\text{ACT}$  the more important predictor of  $\text{colGPA}$ ? Explain.

$\beta_{\text{hsGPA}}$  has a  $p$  value  $< 0.05$ , which implies that the variable  $\text{hsGPA}$  is statistically significant.

$\beta_{\text{ACT}}$  has a  $p$  value  $> 0.05$ , which implies that the variable  $\text{hsGPA}$  is not statistically significant.

We can conclude that  $\text{hsGPA}$  is more important predictor of  $\text{colGPA}$ .

- Add the two indicators  $\text{fathcoll}$  and  $\text{mothcoll}$  to the regression in part (i). Is either individually significant? Are they jointly statistically significant at the 5% level?

From the new regression model, we can see that both  $\text{fathcoll}$  and  $\text{mothcoll}$  are not statistically significant.

```
. reg colGPA PC hsGPA ACT fathcoll mothcoll
```

Source	SS	df	MS	Number of obs	=	141
Model	4.31210399	5	.862420797	F(5, 135)	=	7.71
Residual	15.0939955	135	.111807374	Prob > F	=	0.0000
				R-squared	=	0.2222
				Adj R-squared	=	0.1934
Total	19.4060994	140	.138614996	Root MSE	=	.33438

colGPA	Coefficient	Std. err.	t	P> t	[95% conf. interval]
PC	.1518539	.0587161	2.59	0.011	.0357316 .2679763
hsGPA	.4502203	.0942798	4.78	0.000	.2637639 .6366767
ACT	.0077242	.0106776	0.72	0.471	-.0133929 .0288413
fathcoll	.0417999	.0612699	0.68	0.496	-.079373 .1629728
mothcoll	-.0037579	.0602701	-0.06	0.950	-.1229535 .1154377

_cons		1.255554	.3353918	3.74	0.000	.5922526	1.918856
-----							



5. Use the data in WAGE1 for this exercise.

- Estimate the equation  $wage = \beta_0 + \beta_1 educ + \beta_2 exper + \beta_3 tenure + u$  and report the OLS estimates in the usual form. Save the residuals and plot a histogram. [Hint: 1) You can obtain the residuals of each prediction by using the residuals command and storing these values in a variable named whatever you'd like, e.g., predict resid\_wage, residuals. 2) You can use the histogram command to plot a histogram, e.g., histogram resid\_wage.]

$$wage = -2.872735 + 0.5989651 \times educ + 0.0223395 \times exper + 0.1692687 \times tenure + u$$

```
. use "WAGE1.dta"

. reg wage educ exper tenure
```

Source	SS	df	MS	Number of obs	=	526
Model	2194.1116	3	731.370532	F(3, 522)	=	76.87
Residual	4966.30269	522	9.51398984	Prob > F	=	0.0000
Total	7160.41429	525	13.6388844	R-squared	=	0.3064
				Adj R-squared	=	0.3024
				Root MSE	=	3.0845

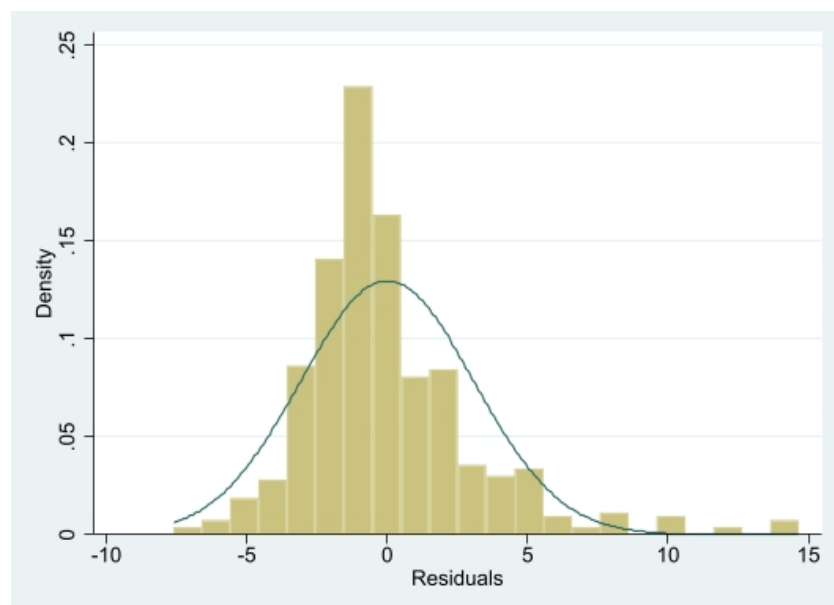
wage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
educ	.5989651	.0512835	11.68	0.000	.4982176	.6997126
exper	.0223395	.0120568	1.85	0.064	-.0013464	.0460254
tenure	.1692687	.0216446	7.82	0.000	.1267474	.2117899
_cons	-2.872735	.7289643	-3.94	0.000	-4.304799	-1.440671

The following code helps us plot a frequency histogram of the residuals

```
. predict resid_wage, residuals

. histogram resid_wage, normal
(bin=22, start=-7.6067705, width=1.011835)
```

Histogram plot of the residuals



- Repeat part (i), but with  $\log(wage)$  as the dependent variable.

```
. reg lwage educ exper tenure
```

Source	SS	df	MS	Number of obs	=	526
Model	46.8741776	3	15.6247259	F(3, 522)	=	80.39
Residual	101.455574	522	.194359337	Prob > F	=	0.0000
				R-squared	=	0.3160
				Adj R-squared	=	0.3121
				Root MSE	=	.44086
Total	148.329751	525	.28253286			

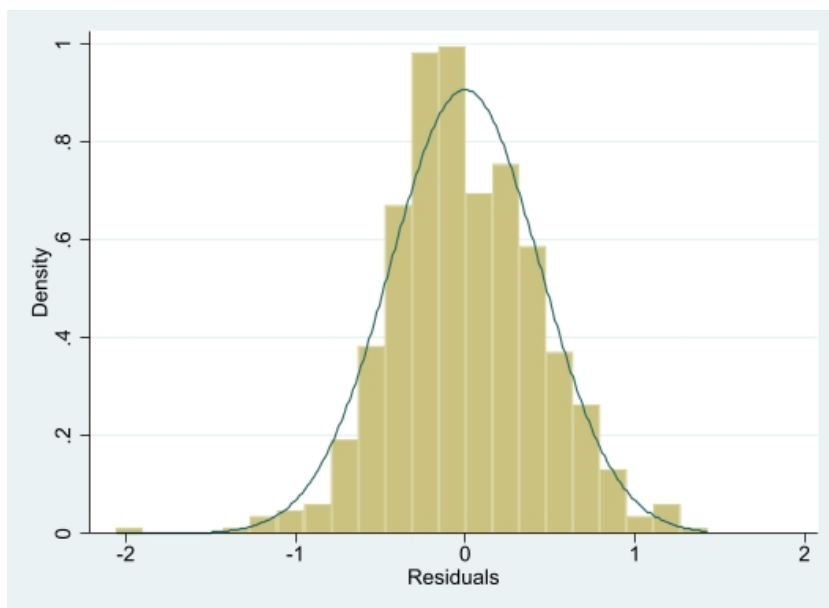
  

lwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]
educ	.092029	.0073299	12.56	0.000	.0776292 .1064288
exper	.0041211	.0017233	2.39	0.017	.0007357 .0075065
tenure	.0220672	.0030936	7.13	0.000	.0159897 .0281448
_cons	.2843595	.1041904	2.73	0.007	.0796756 .4890435

The following code helps us plot a frequency histogram of the residuals

```
. predict resid_lwage, residuals
. histogram resid_lwage, normal
(bin=22, start=-2.0580163, width=.15845945)
```

Histogram plot of the residuals



- Would you say that Assumption MLR.6 is closer to being satisfied for the level-level model or the log-level model? Explain.

The log-level model looks to be a normal distributed better than the level-level mode. The level-level model looks to be positively skewed as well. We can also see that  $sd(residual\_wage)$  is five time the  $sd(residual\_wage)$ . This signifies the homoscedasticity of the log-level model.

```
. summarize resid_lwage
```

Variable	Obs	Mean	Std. dev.	Min	Max
resid_lwage	526	1.27e-10	.4396006	-2.058016	1.428092

```
. summarize resid_wage
```

Variable	Obs	Mean	Std. dev.	Min	Max
resid_wage	526	1.90e-09	3.07565	-7.606771	14.6536

6. The model we used in class to explain the standardized outcome on a final exam ( $stndfnl$ ) in terms of percentage of classes attended, prior college grade point average, and ACT score is  $stndfnl = \beta_0 + \beta_1 \times atndrte + \beta_2 priGPA + \beta_3 ACT + \beta_4 priGPA^2 + \beta_5 \times ACT^2 + \beta_6 \times priGPA \times atndrte + u$

- Argue that  $\frac{\Delta stndfnl}{\Delta priGPA} = \beta_2 + 2\beta_4 \times priGPA + \beta_6 \times atndrte$

Assuming the regression equation is as follows:

$$stndfnl = \beta_0 + \beta_1 \times atndrte + \beta_2 priGPA + \beta_3 ACT + \beta_4 priGPA^2 + \beta_5 \times ACT^2 + \beta_6 \times priGPA \times atndrte + u$$

we partially differentiate wrto  $priGPA$  on both sides.

$$\frac{\delta stndfnl}{\delta priGPA} = \frac{\delta(\beta_0 + \beta_1 \times atndrte + \beta_2 priGPA + \beta_3 ACT + \beta_4 priGPA^2 + \beta_5 \times ACT^2 + \beta_6 \times priGPA \times atndrte + u)}{\delta priGPA}$$

$$= 0 + 0 + \beta_2 + 0 + 2 \times \beta_4 \times priGPA + 0 + \beta_6 \times priGPA \times 1$$

as

$$\frac{\delta(constant)}{\delta priGPA} = 0$$

and

$$\frac{\delta priGPA^2}{\delta priGPA} = 2 \times priGPA$$

$$\therefore \frac{\delta stndfnl}{\delta priGPA} = \beta_2 + 2 \times \beta_4 \times priGPA + \beta_6 \times atndrte$$

It can also be written as

$$\frac{\Delta stndfnl}{\Delta priGPA} = \beta_2 + 2 \times \beta_4 \times priGPA + \beta_6 \times atndrte$$

- Use the equation above to estimate the partial effect of priGPA on stndfnl when priGPA is at its mean value 2.59, and atndrte is also at its mean value 82. Interpret your estimate. [Hint: The estimated OLS equation can be found in Lecture 5.]

We know that

$$\beta_2 = -1.63, \beta_4 = 0.296, \beta_6 = 0.0056$$

and

$$\hat{priGPA} = 2.59, \hat{atndrte} = 82$$

we can compute that

$$\frac{\Delta stndfnl}{\Delta priGPA} = -1.63 + 2 \times 0.296 + 2.59 + 0.0056 \times 82 = 0.36248$$

$$\therefore \frac{\Delta stndfnl}{\Delta priGPA} = 0.36248$$

- Show that the equation can be re-written as  $stndfnl = \theta_0 + \beta_1 atndrte + \theta_2 priGPA + \beta_3 ACT + \beta_4 (priGPA - 2.59)^2 + \beta_5 ACT^2 + \beta_6 priGPA \cdot (atndrte - 82) + u$ , where  $\theta_2 = \beta_2 + 2\beta_4(2.59) + \beta_6(82)$ . How do you interpret  $\theta_2$ ?

We can solve the equation using the following method:

$$stndfnl = \theta_0 + \beta_1 atndrte + \theta_2 \times priGPA + \beta_3 \times ACT + \beta_4 \times (priGPA - 2.59)^2 + \beta_5 ACT^2 + \beta_6 priGPA \cdot (atndrte - 82) + u$$

$$= \theta_0 + \beta_1 \times atndrte + \beta_2 \times priGPA + \beta_3 \times ACT + \beta_3 \times ACT + \beta_4 \times (priGPA - 2.59)^2 + \beta_4 \times 2 \times 2.59 \times priGPA - \beta_4 \times (2.59)^2 + \beta_6 \times priGPA \times (atndrte - 82) + \beta_6 \times 82 \times priGPA + u$$

$$= [\beta_0 - \beta_4 \times (2.59)^2] + \beta_1 atndrte + [\beta_2 + 2 \times \beta_4 \times 2.59 + \beta_6 \times (0.82)] \times priGPA + \beta_3 \times ACT + \beta_4 \times (priGPA - 2.59)^2 + \beta_5 \times ACT^2 + \beta_6 \times priGPA \times (atndrte - 82) \times (atndrte - 82) + u$$

$$= \theta_0 + \beta_1 atndrte + \theta_2 priGPA + \beta_3 ACT + \beta_4 (priGPA - 2.59)^2 + \beta_5 ACT^2 + \beta_6 priGPA (atndrte - 82) + u$$

When we run the regression associated with this last model, we obtain  $\theta^2 \approx -0.091$  and  $se(\theta^2) \approx 0.363$ . This implies a very small t statistic for  $\theta^2$ .

- Following (iii), suppose that, in place of  $priGPA \cdot (atndrte - 82)$ , you put  $(priGPA - 2.59) \cdot (atndrte - 82)$ . Now how do you interpret the coefficients on  $atndrte$  and  $priGPA$ ?

Using  $(priGPA - 2.59) \cdot (atndrte - 0.82)$  in place of  $priGPA \cdot (atndrte - 0.82)$  gives:

$$stndfnl = \beta_0 + \beta_1 \times atndrte + \beta_2 priGPA + \beta_3 ACT + \beta_4 (priGPA - 2.59)^2 + \beta_5 \times ACT^2 + \beta_6 \times (priGPA - 2.59) \times (atndrte - 82) \times (atndrte - 82) + u$$

$$stndfnl = \beta_0 + \beta_1 \times atndrte + \beta_2 priGPA + \beta_3 ACT + \beta_4 (priGPA - 2.59)^2 + \beta_5 \times ACT^2 + \beta_6 \times (priGPA - 2.59) \times (atndrte - 82)^2 + u$$

## Code used in STATA

```
-----
      name:  HW1
      log:   C:\Users\hxd220000\Desktop\HW1.log
    log type: text
  opened on: 22 Feb 2023, 17:14:15

. display count_per
.60004687

. egen gift_mean = mean(gift)
variable gift_mean already defined
r(110);

. sysuse auto, clear
(1978 automobile data)

. use "C:\Users\hxd220000\Desktop\Data Sets- STATA\charity.dta"

. egen gift_mean = mean(gift)

. display gift_mean
7.4444704

. egen count_total = count(respond)

. egen count_gift = count(gift), by (respond)

. display count_gift
2561

. display count_total
4268
```

```
. gen count_per = count_gift /count_total
```

```
. display count_per
.60004687
```

```
. save "C:\Users\hxd220000\Desktop\Charity_v2.dta"
file C:\Users\hxd220000\Desktop\Charity_v2.dta saved
```

```
. summarize mailsyear
```

Variable	Obs	Mean	Std. dev.	Min	Max
-----+-----					
mailsyear	4,268	2.049555	.66758	.25	3.5

```
. reg gift mailsyear
```

Source	SS	df	MS	Number of obs =	4,268
-----+-----				F(1, 4266) =	59.65
Model	13349.7251	1	13349.7251	Prob > F =	0.0000
Residual	954750.114	4,266	223.804528	R-squared =	0.0138
-----+-----				Adj R-squared =	0.0136
Total	968099.84	4,267	226.880675	Root MSE =	14.96

gift	Coefficient	Std. err.	t	P> t	[95% conf. interval]
-----+-----					
mailsyear	2.649546	.3430598	7.72	0.000	1.976971 3.322122
_cons	2.01408	.7394696	2.72	0.006	.5643347 3.463825

```
. use "CEOSAL2.dta"
```

```
. reg lsalary lsales lmktval
```

Source	SS	df	MS	Number of obs =	177
-----+-----				F(2, 174) =	37.13
Model	19.3365617	2	9.66828083	Prob > F =	0.0000
Residual	45.3096514	174	.260400295	R-squared =	0.2991
-----+-----				Adj R-squared =	0.2911
Total	64.6462131	176	.367308029	Root MSE =	.51029

lsalary	Coefficient	Std. err.	t	P> t	[95% conf. interval]
-----+-----					
lsales	.1621283	.0396703	4.09	0.000	.0838315 .2404252
lmktval	.106708	.050124	2.13	0.035	.0077787 .2056372
_cons	4.620917	.2544083	18.16	0.000	4.118794 5.123041

```
. reg lsalary lsales lmktval profits
```

Source	SS	df	MS	Number of obs =	177
-----+-----				F(3, 173) =	24.64
Model	19.3509799	3	6.45032663	Prob > F =	0.0000
Residual	45.2952332	173	.261822157	R-squared =	0.2993
-----+-----				Adj R-squared =	0.2872
Total	64.6462131	176	.367308029	Root MSE =	.51169

lsalary	Coefficient	Std. err.	t	P> t	[95% conf. interval]
-----+-----					
lsales	.1613683	.0399101	4.04	0.000	.0825949 .2401416
lmktval	.0975286	.0636886	1.53	0.128	-.0281782 .2232354
profits	.0000357	.000152	0.23	0.815	-.0002643 .0003356
_cons	4.686924	.3797294	12.34	0.000	3.937425 5.436423

```
. reg lsalary lsales lmktval profits ceoten
```

Source	SS	df	MS	Number of obs =	177
-----+-----				F(4, 172) =	20.08
Model	20.5768102	4	5.14420254	Prob > F =	0.0000

Residual		44.0694029	172	.256217459	R-squared	=	0.3183
	+				Adj R-squared	=	0.3024
Total		64.6462131	176	.367308029	Root MSE	=	.50618

lsalary		Coefficient	Std. err.	t	P> t	[95% conf. interval]
lsales		.1622339	.0394826	4.11	0.000	.0843012 .2401667
lmktval		.1017598	.063033	1.61	0.108	-.022658 .2261775
profits		.0000291	.0001504	0.19	0.847	-.0002677 .0003258
ceoten		.0116847	.005342	2.19	0.030	.0011403 .022229
_cons		4.55778	.3802548	11.99	0.000	3.807213 5.308347

```
. correlate lmktval profits
(obs=177)
```

		lmktval	profits
lmktval		1.0000	
profits		0.7769	1.0000

```
. use "TWOYEAR.dta"
```

```
. reg lwage jc totcoll exper
```

Source		SS	df	MS	Number of obs	=	6,763
	+				F(3, 6759)	=	644.53
Model		357.752575	3	119.250858	Prob > F	=	0.0000
Residual		1250.54352	6,759	.185019014	R-squared	=	0.2224
	+				Adj R-squared	=	0.2221
Total		1608.29609	6,762	.237843255	Root MSE	=	.43014

lwage		Coefficient	Std. err.	t	P> t	[95% conf. interval]
jc		-.0101795	.0069359	-1.47	0.142	-.0237761 .003417
totcoll		.0768762	.0023087	33.30	0.000	.0723504 .0814021
exper		.0049442	.0001575	31.40	0.000	.0046355 .0052529
_cons		1.472326	.0210602	69.91	0.000	1.431041 1.51361

```
. display jc
0
```

```
. summarize jc
```

Variable		Obs	Mean	Std. dev.	Min	Max
jc		6,763	.3388946	.7721268	0	3.833333

```
. reg lwage jc totcoll exper phsrnk
```

Source		SS	df	MS	Number of obs	=	6,763
	+				F(4, 6758)	=	483.85
Model		358.050568	4	89.5126419	Prob > F	=	0.0000
Residual		1250.24552	6,758	.185002297	R-squared	=	0.2226
	+				Adj R-squared	=	0.2222
Total		1608.29609	6,762	.237843255	Root MSE	=	.43012

lwage		Coefficient	Std. err.	t	P> t	[95% conf. interval]
jc		-.0093108	.0069693	-1.34	0.182	-.0229728 .0043512
totcoll		.0754756	.0025588	29.50	0.000	.0704595 .0804918
exper		.0049396	.0001575	31.36	0.000	.0046308 .0052483
phsrnk		.0003032	.0002389	1.27	0.204	-.0001651 .0007716
_cons		1.458747	.0236211	61.76	0.000	1.412442 1.505052

```
. summarize phsrnk
```

Variable	Obs	Mean	Std. dev.	Min	Max
phsrank	6,763	56.15703	24.27296	0	99

```
. reg lwage jc totcoll exper phsrank id
```

Source	SS	df	MS	Number of obs	=	6,763
Model	358.132086	5	71.6264171	F(5, 6757)	=	387.13
Residual	1250.16401	6,757	.185017612	Prob > F	=	0.0000
				R-squared	=	0.2227
				Adj R-squared	=	0.2221
Total	1608.29609	6,762	.237843255	Root MSE	=	.43014

lwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]
jc	-.0093159	.0069696	-1.34	0.181	-.0229784 .0043467
totcoll	.075414	.0025606	29.45	0.000	.0703943 .0804336
exper	.004941	.0001575	31.37	0.000	.0046322 .0052498
phsrank	.0003179	.00024	1.32	0.185	-.0001525 .0007883
id	1.40e-07	2.10e-07	0.66	0.507	-2.73e-07 5.52e-07
_cons	1.452215	.0255898	56.75	0.000	1.402051 1.502379

```
. use "GPA1.dta"
```

```
. reg colGPA PC hsGPA ACT
```

Source	SS	df	MS	Number of obs	=	141
Model	4.25741863	3	1.41913954	F(3, 137)	=	12.83
Residual	15.1486808	137	.110574313	Prob > F	=	0.0000
				R-squared	=	0.2194
				Adj R-squared	=	0.2023
Total	19.4060994	140	.138614996	Root MSE	=	.33253

colGPA	Coefficient	Std. err.	t	P> t	[95% conf. interval]
PC	.1573092	.0572875	2.75	0.007	.0440271 .2705913
hsGPA	.4472417	.0936475	4.78	0.000	.2620603 .632423
ACT	.008659	.0105342	0.82	0.413	-.0121717 .0294897
_cons	1.26352	.3331255	3.79	0.000	.6047871 1.922253

```
. reg colGPA PC hsGPA ACT fathcoll mothcoll
```

Source	SS	df	MS	Number of obs	=	141
Model	4.31210399	5	.862420797	F(5, 135)	=	7.71
Residual	15.0939955	135	.111807374	Prob > F	=	0.0000
				R-squared	=	0.2222
				Adj R-squared	=	0.1934
Total	19.4060994	140	.138614996	Root MSE	=	.33438

colGPA	Coefficient	Std. err.	t	P> t	[95% conf. interval]
PC	.1518539	.0587161	2.59	0.011	.0357316 .2679763
hsGPA	.4502203	.0942798	4.78	0.000	.2637639 .6366767
ACT	.0077242	.0106776	0.72	0.471	-.0133929 .0288413
fathcoll	.0417999	.0612699	0.68	0.496	-.079373 .1629728
mothcoll	-.0037579	.0602701	-0.06	0.950	-.1229535 .1154377
_cons	1.255554	.3353918	3.74	0.000	.5922526 1.918856

```
. use "WAGE1.dta"
```

```
. reg wage educ exper tenure
```

Source	SS	df	MS	Number of obs	=	526
Model	2194.1116	3	731.370532	F(3, 522)	=	76.87
Residual	4966.30269	522	9.51398984	Prob > F	=	0.0000
				R-squared	=	0.3064
				Adj R-squared	=	0.3024



Total | 7160.41429      525   13.6388844   Root MSE      =      3.0845

wage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
educ	.5989651	.0512835	11.68	0.000	.4982176	.6997126
exper	.0223395	.0120568	1.85	0.064	-.0013464	.0460254
tenure	.1692687	.0216446	7.82	0.000	.1267474	.2117899
_cons	-2.872735	.7289643	-3.94	0.000	-4.304799	-1.440671

```
. predict resid_wage, residuals
```

```
. histogram resid_wage
(bin=22, start=-7.6067705, width=1.011835)
```

```
. graph save "Graph" "C:\Users\hxd220000\Desktop\Graph.gph"
file C:\Users\hxd220000\Desktop\Graph.gph saved
```

```
. graph export "C:\Users\hxd220000\Desktop\Graph.jpg", as(jpg) name("Graph") qua
> lity(90)
file C:\Users\hxd220000\Desktop\Graph.jpg written in JPEG format
```

```
. reg lwage educ exper tenure
```

Source	SS	df	MS	Number of obs	=	526
Model	46.8741776	3	15.6247259	F(3, 522)	=	80.39
Residual	101.455574	522	.194359337	Prob > F	=	0.0000
Total	148.329751	525	.28253286	R-squared	=	0.3160
				Adj R-squared	=	0.3121
				Root MSE	=	.44086

lwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
educ	.092029	.0073299	12.56	0.000	.0776292	.1064288
exper	.0041211	.0017233	2.39	0.017	.0007357	.0075065
tenure	.0220672	.0030936	7.13	0.000	.0159897	.0281448
_cons	.2843595	.1041904	2.73	0.007	.0796756	.4890435

```
. predict resid_wage, residuals
variable resid_wage already defined
r(110);
```

```
. predict resid_lwage, residuals
```

```
. histogram resid_lwage
(bin=22, start=-2.0580163, width=.15845945)
```

```
. graph export "C:\Users\hxd220000\Desktop\Graph-lwage.jpg", as(jpg) name("Graph
> ") quality(100)
file C:\Users\hxd220000\Desktop\Graph-lwage.jpg written in JPEG format
```

```
. graph save "Graph" "C:\Users\hxd220000\Desktop\Graph-lwage.gph"
file C:\Users\hxd220000\Desktop\Graph-lwage.gph saved
histogram resid_wage, normal
(bin=22, start=-7.6067705, width=1.011835)
```

```
. graph save "Graph" "C:\Users\hxd220000\Desktop\Graph.gph", replace
file C:\Users\hxd220000\Desktop\Graph.gph saved
```

```
. graph export "C:\Users\hxd220000\Desktop\Graph.jpg", as(jpg) name("Graph") qua
> lity(100) replace
file C:\Users\hxd220000\Desktop\Graph.jpg written in JPEG format
```

```
. histogram resid_lwage, normal
(bin=22, start=-2.0580163, width=.15845945)
```

```
. graph save "Graph" "C:\Users\hxd220000\Desktop\Graph-lwage.gph", replace
file C:\Users\hxd220000\Desktop\Graph-lwage.gph saved
```

```
. graph export "C:\Users\hxd220000\Desktop\Graph-lwage.jpg", as(jpg) name("Graph
```

```
> ") quality(100) replace
file C:\Users\hxd220000\Desktop\Graph-lwage.jpg written in JPEG format
```

```
. summarize resid_lwage
```

Variable	Obs	Mean	Std. dev.	Min	Max
resid_lwage	526	1.27e-10	.4396006	-2.058016	1.428092

```
. summarize resid_wage
```

Variable	Obs	Mean	Std. dev.	Min	Max
resid_wage	526	1.90e-09	3.07565	-7.606771	14.6536