

# ON BAYESIAN G-COMPUTATION AND INSTRUMENTAL VARIABLES FOR DETERMINING EFFICACY FROM OBSERVATIONAL DATA

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## 1. INTRODUCTION

Clinical trials are often not practical for long-term outcomes. As such, researchers must consider observational data for determining the effect (or lack of effect) of a given intervention,  $X$ , on a given outcome,  $Y$ . Two main obstacles with the analysis of time-varying observational data are:

- **Confounding by indication:** if exposure,  $X$ , is related to  $Y$  (e.g. if patients are more likely to receive a new therapy because they aren't doing well).
- **Confounding by trends in time:** if outcomes,  $Y$ , are coincidentally changing over the same time period as changes in exposure to the intervention,  $X$ , (e.g. the overall .... )

One strategy is to apply  $g$ -computation, as described in Borsi (2012) [7]. Another, seemingly cruder strategy, would involve treating calendar period of reaching eligibility for treatment as an instrumental variable, see Greenland (2000) [9]. The theoretical properties of the  $g$ -computation approach have been previously examined, see Johnston et al. (2008) [2], and applications exist, such as in the context of assessing the effect of antiretroviral therapy on incident AIDS, see Cainn et al. (2009) [6].

Either approach requires delicate assumptions. For  $g$ -computation, one assumes that all time-varying confounders are measured in the data. For IV, one assumes that a chosen observed variable is indeed an “instrumental variable”. Each approach involves very strong, yet different assumptions, see Table 1. This paper investigates the merits of each approach in terms of power to detect efficacy (or lack thereof) and robustness to violations of assumptions.

**Application-** Since they were approved 20 years ago, the proportion of MS patients taking beta-interferons has gone up dramatically, but outcomes, as measured by time from disease onset to disease progression, have not improved substantially over calendar time. Tremlett

et al. (2008) [3] and Tremlett et al. (2010) [4] review recent advances in the understanding of the natural history of MS. In order to determine whether or not beta-interferons are an effective treatment, one can look over extensive subject-level observational data, as in the analysis of Karim et al. (2014) [1]. If calendar time can be considered an “instrumental variable”, other methods of analysis may prove advantageous.

	<b>IV <math>g</math>-comp with H   <math>g</math>-comp without H</b>		
<b>Assumption</b>			
H is an Instrumental Variable, $\theta_6 = 0$	✓	✗	✓
H has positive causal effect on X, $\theta_2 > 0$	✓	✗	✗
No unknown confounders, $\theta_7 = \theta_8 = 0$	✗	✓	✓
<b>Required Data</b>			
H	✓	✓	✗
X	✗	✓	✓
C	✗	✓	✓

TABLE 1. Assumptions and required data for both approaches.

The objective is to determine the significance (superiority or non-inferiority) of the target value  $T_{(1,1)-(0,0)}$ :

$$(1) \quad T_{(1,1)-(0,0)} = Pr(Y_2 = 1 | do(X_1 = 1, X_2 = 1)) - Pr(Y_2 = 1 | do(X_1 = 0, X_2 = 0))$$

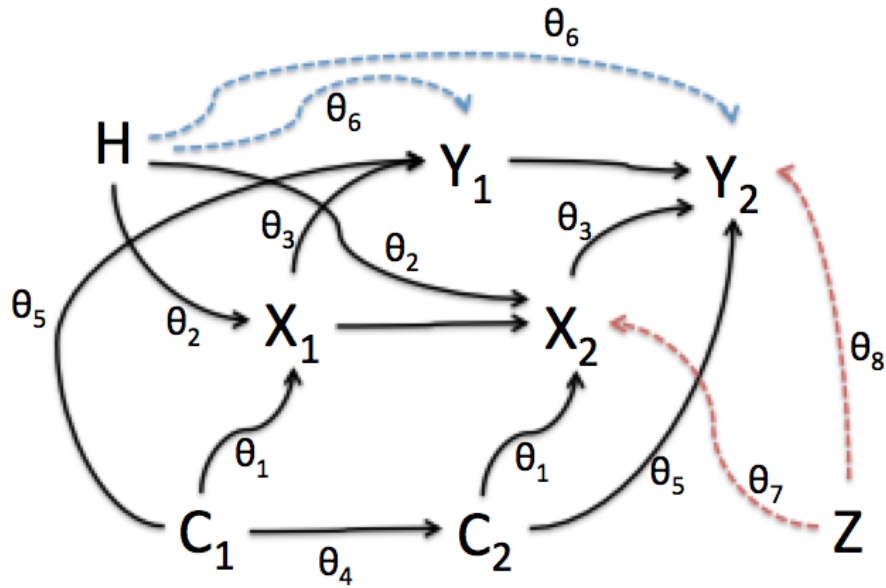


FIGURE 1. The two time-point model considered. We do not suggest that this model will be adequate for all applications. However, it is sufficiently rich to provide a vehicle for investigating the consequences of preferential sampling, and for the application that is described in Section X of the paper.

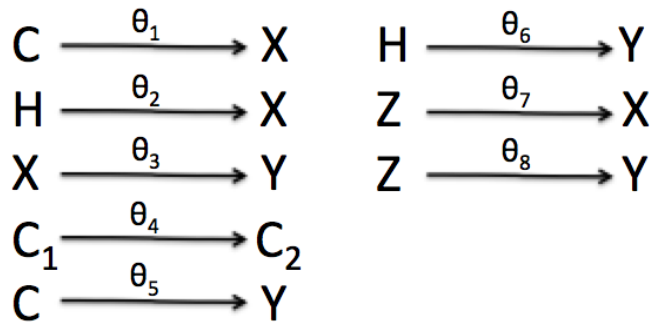


FIGURE 2. Parameters considered.

Variable	Representation		Interpretation	
	if = 0	if = 1		
H	early	late	Instrumental Variable	Calendar time
C	?	?	Known Confounder	?
Y	ill	healthy	Outcome of interest	Disease progression
Z	?	?	Unknown additional confounder	
X	no treatment	treatment	Intervention of interest	Treatment
Parameter	if = 0	if > 0		
$\theta_1$	<b>X</b>	<b>X</b>	<b>X</b>	
$\theta_2$	<b>X</b>	<b>X</b>	<b>X</b>	
$\theta_3$	Null is true	Alt. is true	intervention effect	treatment effect
$\theta_4$	<b>X</b>	<b>X</b>	<b>X</b>	
$\theta_5$	<b>X</b>	<b>X</b>	<b>X</b>	
$\theta_6$	<b>X</b>	<b>X</b>	<b>X</b>	
$\theta_7$	<b>X</b>	<b>X</b>	<b>X</b>	
$\theta_8$	<b>X</b>	<b>X</b>	<b>X</b>	

TABLE 2. Model Components.

Given an outcome measure, a treatment variable, a set of unmeasured confounders, and, possibly, a set of measured confounders, a variable satisfies the conditions of an instrumental variable (also referred to as an instrument) if the following conditions hold:

- (1) it is correlated with the treatment variable
- (2) given the treatment variable and all confounders (measured and unmeasured), it is conditionally independent of outcome; and
- (3) it is independent of the entire set of unmeasured confounders [9].

**With  $Y_1$ , without  $H$ :**

$$\begin{aligned}
 Pr(Y_2 = 1 | do(X_1 = x_1, X_2 = x_2)) &= \sum_{y_1=0}^1 \sum_{c_1=0}^1 \sum_{c_2=0}^1 Pr(C_1 = c_1) \\
 &\quad \cdot Pr(C_2 = c_2 | C_1 = c_1, X_1 = x_1) \\
 &\quad \cdot Pr(Y_1 = y_1 | C_1 = c_1, X_1 = x_1) \\
 &\quad \cdot Pr(Y_2 = 1 | C_2 = c_2, X_2 = x_2, Y_1 = y_1)
 \end{aligned}$$

**With  $Y_1$ , and with  $H$ :**

$$\begin{aligned}
Pr(Y_2 = 1|do(X_1 = x_1, X_2 = x_2)) &= \sum_{y_1=0}^1 \sum_{c_1=0}^1 \sum_{c_2=0}^1 \sum_{h=0}^1 Pr(H = h) \cdot Pr(C_1 = c_1|H = h) \\
&\cdot Pr(C_2 = c_2|C_1 = c_1, X_1 = x_1, H = h) \\
&\cdot Pr(Y_1 = y_1|C_1 = c_1, X_1 = x_1, H = h) \\
&\cdot Pr(Y_2 = 1|C_2 = c_2, X_2 = x_2, Y_1 = y_1, H = h)
\end{aligned}$$

**Without  $Y_1$ , without  $H$ :**

$$\begin{aligned}
Pr(Y_2 = 1|do(X_1 = x_1, X_2 = x_2)) &= \sum_{c_1=0}^1 \sum_{c_2=0}^1 Pr(C_1 = c_1) \\
&\cdot Pr(C_2 = c_2|C_1 = c_1, X_1 = x_1) \\
&\cdot Pr(Y_2 = 1|C_2 = c_2, X_2 = x_2)
\end{aligned}$$

**Without  $Y_1$ , with  $H$ :**

$$\begin{aligned}
Pr(Y_2 = 1|do(X_1 = x_1, X_2 = x_2)) &= \sum_{c_1=0}^1 \sum_{c_2=0}^1 \sum_{h=0}^1 Pr(H = h) \cdot Pr(C_1 = c_1|H = h) \\
&\cdot Pr(C_2 = c_2|C_1 = c_1, X_1 = x_1, H = h) \\
&\cdot Pr(Y_2 = 1|C_2 = c_2, X_2 = x_2, H = h)
\end{aligned}$$

## 2. METHODS

**Frequentist g-comp with logistic model:**  $Pr(A = 1|B = b, C = c) = predict(logistic(Pr(A = 1|B = b, C = c)))$

**Bayesian g-comp with Uniform priors:**  $Pr(A = 1|B = b, C = c) \sim Beta(1 + \sum_{i=1}^n 1_{(A=1, B=b, C=c)}(x_i), 1 + \sum_{i=1}^n 1_{(A=0, B=b, C=c)}(x_i))$

- run down of the simple sandbox of two time-points, everything binary - g-computation (perhaps motivated several different ways) - IV idea

Comparison in a focussed scenario - confounding by indication present (C encourages treatment start) - different possible forms of treatment effect (through C and/or around C) - both sets of assumptions met

Comparison across a broad range of scenarios

- set up distribution giving rise to a wide array of states of the world - look for patterns, e.g., is the power of the g-computation driven only by  $\theta_{11}$ - $\theta_{00}$ , or are there other important factors

Violations of assumptions

- any defensible sense of one procedure being more robust than the other - how quickly is power lost as we relax assumptions

### 3. APPLICATION

### 4. DISCUSSION

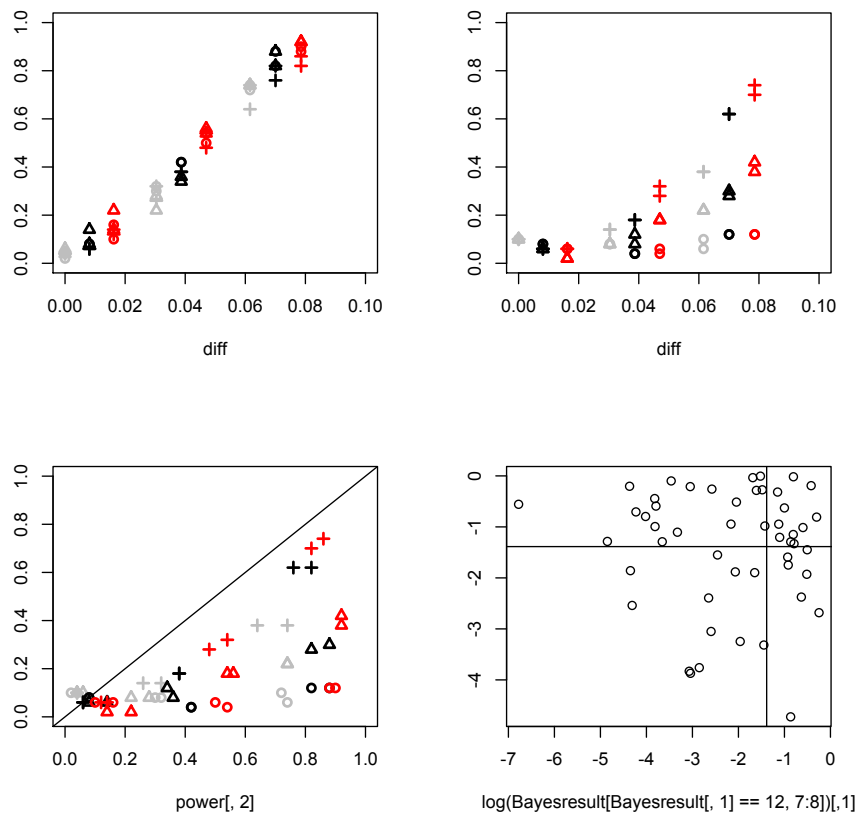


FIGURE 3. Results of Simulation study.

## REFERENCES

- [1] M. E. Karim, P. Gustafson, J. Petkau, Y. Zhao, A. Shirani, E. Kingwell, C. Evans, M. van der Kop, J. Oger, and H. Tremlett, “Marginal structural cox models for estimating the association between  $\beta$ -interferon exposure and disease progression in a multiple sclerosis cohort,” *American journal of epidemiology*, vol. 180, no. 2, pp. 160–171, 2014.
- [2] K. Johnston, P. Gustafson, A. Levy, and P. Grootendorst, “Use of instrumental variables in the analysis of generalized linear models in the presence of unmeasured confounding with applications to epidemiological research,” *Statistics in medicine*, vol. 27, no. 9, pp. 1539–1556, 2008.
- [3] H. Tremlett, Y. Zhao, and V. Devonshire, “Natural history of secondary-progressive multiple sclerosis,” *Multiple sclerosis*, vol. 14, no. 3, pp. 314–324, 2008.
- [4] H. Tremlett, Y. Zhao, P. Rieckmann, and M. Hutchinson, “New perspectives in the natural history of multiple sclerosis,” *Neurology*, vol. 74, no. 24, pp. 2004–2015, 2010.
- [5] O. Saarela, D. A. Stephens, E. E. Moodie, and M. B. Klein, “On bayesian estimation of marginal structural models,” *Biometrics*, 2015.

- [6] L. E. Cain, S. R. Cole, S. Greenland, T. T. Brown, J. S. Chmiel, L. Kingsley, and R. Detels, "Effect of highly active antiretroviral therapy on incident aids using calendar period as an instrumental variable," *American journal of epidemiology*, p. kwp002, 2009.
- [7] L. Borsi, *Estimating the Causal Effect of Switching to Second-line Antiretroviral HIV Treatment Using G-computation*. PhD thesis, ETH, Department of Mathematics, 2012.
- [8] M. Kuroki and Z. Cai, "Instrumental variable tests for directed acyclic graph models,"
- [9] S. Greenland, "An introduction to instrumental variables for epidemiologists," *International journal of epidemiology*, vol. 29, no. 4, pp. 722–729, 2000.
- [10] K. Peter, *Marginal structural models and causal inference*. PhD thesis, Master thesis, ETH, 2011.
- [11] B. Chakraborty and E. E. Moodie, "G-computation: Parametric estimation of optimal dtrs," in *Statistical Methods for Dynamic Treatment Regimes*, pp. 101–112, Springer, 2013.
- [12] S. L. Taubman, J. M. Robins, M. A. Mittleman, and M. A. Hernán, "Intervening on risk factors for coronary heart disease: an application of the parametric g-formula," *International Journal of Epidemiology*, vol. 38, no. 6, pp. 1599–1611, 2009.
- [13] J. M. Snowden, S. Rose, and K. M. Mortimer, "Implementation of g-computation on a simulated data set: demonstration of a causal inference technique," *American Journal of Epidemiology*, p. kwq472, 2011.