# 10. Probability distributions

These notes consider the Chapter 10 of the handbook on various probability distributions.

## Discrete distributions

For random variables with countable number of possible values.

#### Binomial distribution

See page 200 for proper definition.

Useful when a random variable X has exactly two exclusive possible outcomes (e.g. success/fail) with known probabilities  $p \in [0, 1]$  (success) and q = 1 - p (fail). For  $n \in \mathbb{N}$  trials with  $k = 0, 1, 2, \ldots, n$  successes

$$X \sim Bin(n, p)$$

Probability Mass Function (PMF)

$$P(X = k) = \binom{n}{k} p^k q^{n-k}$$

Cumulative Distribution Function (CDF)

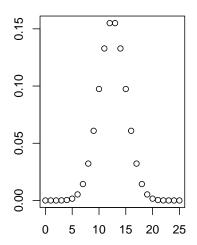
$$F(k) = P(X \leqslant k) = \sum_{i=0}^{k} \binom{n}{i} p^{i} q^{n-i}$$

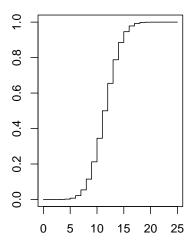
In the following code examples, equal probabilities are assumed, n = size, p = prob = 0.5.

```
# Sequence for visualization
binomial_seq <- seq(0, 25, by = 1)

# Functions
binomial_pmf <- dbinom(x = binomial_seq, size = 25, prob = 0.5)
binomial_cdf <- pbinom(q = binomial_seq, size = 25, prob = 0.5)

# Plot
par(mfrow = c(1, 2))
plot(binomial_seq, binomial_pmf, ann = FALSE)
plot(binomial_seq, binomial_cdf, type = "S", ann = FALSE)</pre>
```





```
# Probability for exactly 3 successes out of 10 trials
dbinom(x = 3, size = 10, prob = 0.5)
```

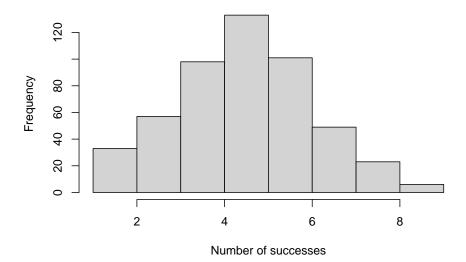
```
# Probability for up to 3 successes out of 10 trials
pbinom(q = 3, size = 10, prob = 0.5)
```

## [1] 0.171875

```
# Simulate 10 times how many successes there is using random numbers rbinom(n = 10, size = 10, prob = 0.5)
```

## [1] 2 6 6 6 6 6 1 4 2 7

```
# With large enough n, expected value (np = 5) should become visible
rbinom500 <- rbinom(n = 500, size = 10, prob = 0.5)
hist(rbinom500, xlab = "Number of successes", ylab = "Frequency", main = NULL)</pre>
```



## summary(rbinom500)

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 1.000 4.000 5.000 4.954 6.000 9.000
```

## Poisson distribution

See page 204.

Useful when estimating amounts in random processes where the expected value  $(\lambda)$  is known

$$X \sim Poi(\lambda)$$

PMF

$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}, \ k = 0, 1, 2, ..., n$$

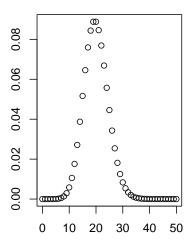
Note that the Poisson distribution can be used to approximate the binomial distribution when n is large and p is small. In this case,  $\lambda = np$  and

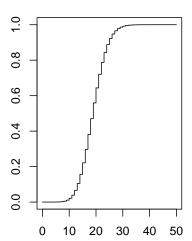
$$P(X = k) = \frac{(np)^k e^{-np}}{k!}$$

```
# Sequence for visualization
poisson_seq <- seq(0, 50, by = 1)

# Functions
poisson_pmf <- dpois(x = poisson_seq, lambda = 20)
poisson_cdf <- ppois(q = poisson_seq, lambda = 20)</pre>
```

```
# Plot
par(mfrow = c(1, 2))
plot(poisson_seq, poisson_pmf, ann = FALSE)
plot(poisson_seq, poisson_cdf, type = "S", ann = FALSE)
```





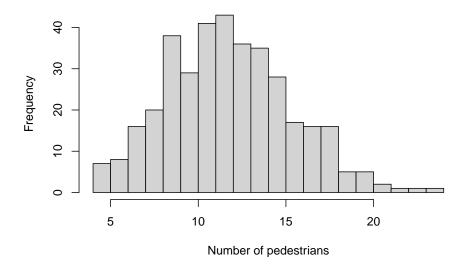
```
# A bridge is crossed by 12 people per minute, on average.
# What is the probability that exactly 16 people crosses it in a minute?
dpois(x = 16, lambda = 12)
```

```
# A bridge is crossed by 12 people per minute, on average.
# What is the probability that up to 16 people crosses it in a minute?
ppois(q = 16, lambda = 12)
```

## [1] 0.898709

```
# Simulate amount of people per minute 10 times
rpois(n = 10, lambda = 12)
```

## [1] 16 14 6 12 13 15 11 15 15 15



## summary(poisson\_pedestrians)

## Min. 1st Qu. Median Mean 3rd Qu. Max. ## 4.00 10.00 12.00 12.24 15.00 24.00

# Continuous distributions

For random variables with infinite number of possible values.

## Generic definitions

See page 208.

The probability that an event happens between an interval [a, b] can be calculated from the Probability Density Function (PDF) f(x) as an integral

$$P(a \leqslant X \leqslant b) = \int_{a}^{b} f(x) \, dx$$

Note that the point probabilities for continuous random variables are intrinsically zero

$$\int_{a}^{a} f(x) \, \mathrm{d}x = 0$$

CDF

$$P(X \leqslant t) = \int_{-\infty}^{t} f(x) \, dx$$

#### Normal distribution

See page 210.

Many phenomena in nature follow the normal distribution. Additionally, so called central limit theorem states that the averages of idenpendet random samples drawn from a population approximately follow the normal distribution - even if the population is better described by some other distribution!

Normal distribution is defined by expected value (and median and mode due to symmetry)  $\mu$  and (standard) deviation  $\sigma$ .

$$X \sim N(\mu, \sigma^2)$$

PDF

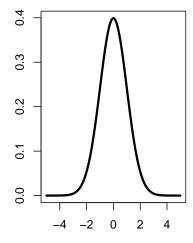
$$f(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

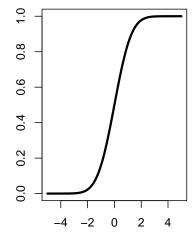
In the following code examples,  $\mu = \text{mean}$ ,  $\sigma = \text{sd}$ .

```
# Sequence for visualization
normal_seq <- seq(-5, 5, by = 0.01)

# Functions
normal_pmf <- dnorm(x = normal_seq, mean = 0, sd = 1)
normal_cdf <- pnorm(q = normal_seq, mean = 0, sd = 1)

# Plot
par(mfrow = c(1, 2))
plot(normal_seq, normal_pmf, type = "l", lwd = 3, ann = FALSE)
plot(normal_seq, normal_cdf, type = "l", lwd = 3, ann = FALSE)</pre>
```





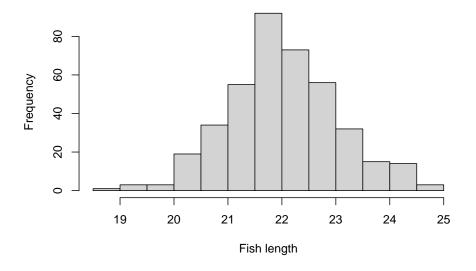
```
# The average length of a zebrafish is 22 mm with a standard diviation of 1 mm. # How likely is a fish 23 mm long? dnorm(x = 23, mean = 22, sd = 1)
```

```
# What fraction of the fish are up to 23 mm long?
pnorm(q = 23, mean = 22, sd = 1)
```

## [1] 0.8413447

```
# Simulate length of five random fish
rnorm(n = 5, mean = 22, sd = 1)
```

## [1] 23.35507 21.42742 23.97557 21.72851 23.58276



#### summary(normal\_zebrafish)

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 18.78 21.35 21.96 22.01 22.64 24.93
```

#### Student's t-distribution

See page 215.

Similar to the normal distribution, Student's t distribution can be used to estimate the expected value of a population using large enough amout of sample averages.

The Student's t distribution is defined by just one parameter, namely the degrees of freedom (DoF)  $\nu$ . Compared to the normal distribution, the advantage here is that prior knowledge of parameters  $\mu$  and  $\sigma$  is not needed - which oftentimes is the case. Additionally, with a large  $\nu$ , the Student's t distribution approaches the standard normal distribution N(0,1).

PDF is defined using the Gamma function  $\Gamma(n) = (n-1)!$ .

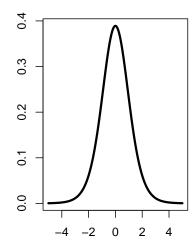
$$f(t) = \frac{\Gamma(\frac{\nu+1}{2})}{\sqrt{\pi\nu}\Gamma(\frac{\nu}{2})} \left(1 + \frac{t^2}{\nu}\right)^{-(\nu+1)/2}$$

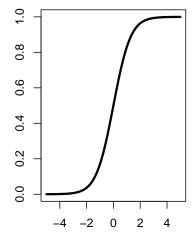
In the following code examples,  $\nu = \mathbf{df}$ .

```
# Sequence for visualization
students_t_seq <- seq(-5, 5, by = 0.01)

# Functions
students_t_pmf <- dt(x = students_t_seq, df = 10)
students_t_cdf <- pt(q = students_t_seq, df = 10)

# Plot
par(mfrow = c(1, 2))
plot(normal_seq, students_t_pmf, type = "l", lwd = 3, ann = FALSE)
plot(normal_seq, students_t_cdf, type = "l", lwd = 3, ann = FALSE)</pre>
```





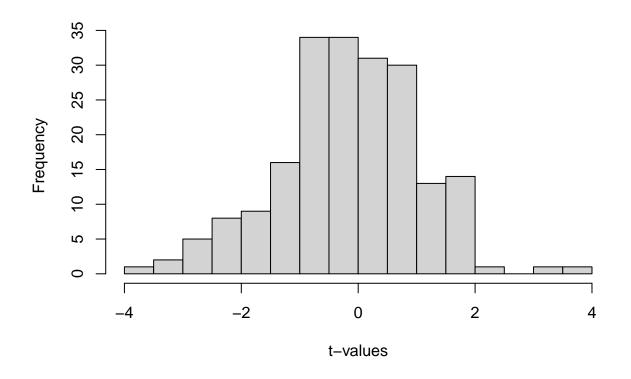
```
# Same as in the earlier examples
dt(x = 2, df = 20)
```

```
pt(q = 2, df = 20)
```

## [1] 0.9703672

```
rt(n = 5, df = 20)
```

**##** [1] 1.8902778 -0.9411542 0.2514571 0.2731861 -0.2830450



```
summary(student_values)
```

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## -3.9990 -0.8367 -0.1295 -0.1720 0.6384 3.7014
```

## Chi-squared distribution

See page 217.

The  $\chi^2$  distribution can be used to estimate the distribution of population variances. Suppose we have independent standard normal (N(0,1)) random variables  $X_1, X_2, \ldots, X_k$ , then the random variable

$$Q = \sum_{i=1}^{k} \chi_i^2$$

follows the  $\chi^2$  distribution  $(Q \sim \chi^2(k))$  with k-1 DoF.

PDF

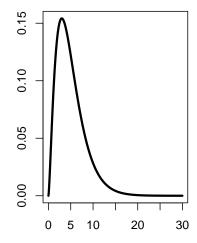
$$f(x; k) = \begin{cases} \frac{x^{k/2-1}e^{-x/2}}{2^{k/2}\Gamma\left(\frac{k}{2}\right)}, & x > 0; \\ 0, & \text{otherwise.} \end{cases}$$

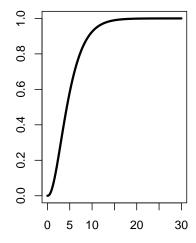
In the following code examples, **df** is the DoF.

```
# Sequence for visualization
chisq_seq <- seq(0, 30, by = 0.02)

# Functions
chisq_pmf <- dchisq(x = chisq_seq, df = 5)
chisq_cdf <- pchisq(q = chisq_seq, df = 5)

# Plot
par(mfrow = c(1, 2))
plot(chisq_seq, chisq_pmf, type = "l", lwd = 3, ann = FALSE)
plot(chisq_seq, chisq_cdf, type = "l", lwd = 3, ann = FALSE)</pre>
```





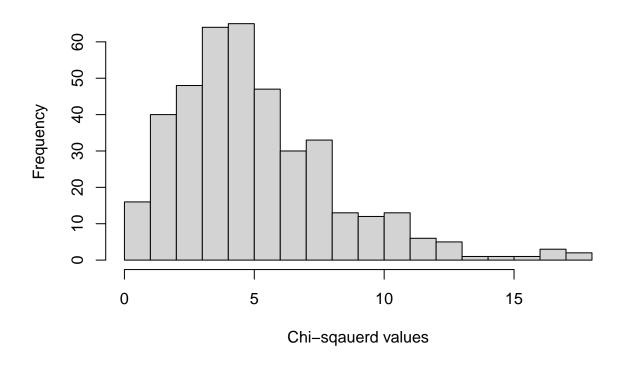
```
# Same as in the earlier examples
dchisq(x = 2, df = 4)

## [1] 0.1839397

pchisq(q = 2, df = 4)
```

```
rchisq(n = 5, df = 4)
```

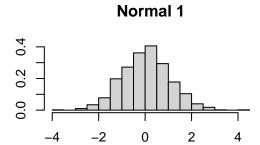
## [1] 2.1620175 1.6149888 4.3746883 1.4707055 0.5914959

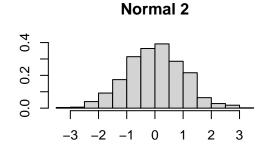


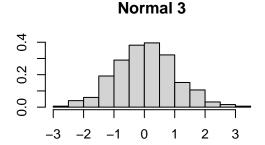
## summary(chisq\_values)

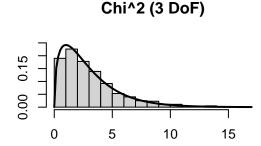
```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.05959 2.88340 4.49128 5.04503 6.56786 17.87608
```

```
# Let's also visualize the connection to the standard normal distribution
# First, generate some distributions
standard_normal_1 <- rnorm(n = 1000, mean = 0, sd = 1)
standard_normal_2 \leftarrow rnorm(n = 1000, mean = 0, sd = 1)
standard_normal_3 \leftarrow rnorm(n = 1000, mean = 0, sd = 1)
# Calculate the sum of squares (see random variable Q in the definition)
sum_of_squares = standard_normal_1^2 + standard_normal_2^2 + standard_normal_3^2
# Plots
par(mfrow = c(2, 2))
hist(standard_normal_1, breaks = 20,
     xlab = "", ylab = "", main = "Normal 1", prob = TRUE)
hist(standard_normal_2, breaks = 20,
     xlab = "", ylab = "", main = "Normal 2", prob = TRUE)
hist(standard_normal_3, breaks = 20,
     xlab = "", ylab = "", main = "Normal 3", prob = TRUE)
hist(sum_of_squares, breaks = 20, ylim = c(0, 0.25),
     xlab = "", ylab = "", main = "Chi^2 (3 DoF)", prob = TRUE)
# Overlap an ideal distribution with 3 DoF
curve(dchisq(x, df = 3), lwd = 2, add = TRUE)
```









Intuitively this should make sense: square of a standard normal distribution "flips" the negative values over the mean creating a skewed distribution and summing these enhances the effect.