

- 1) Bernoulli Distribution
- 2) Binomial Distribution
- 3) Poisson Distribution.

# Let  $X_1, X_2, \dots, X_k$  be independent RVs where  $X_i \sim b(n_i, p)$ . Then

$$S_n = \sum_{k=1}^n X_k \sim b\left(\sum_{k=1}^n n_k, p\right)$$

$$M_{S_n}(t) = E[e^{t S_n}] = E[e^{t(X_1 + X_2 + \dots + X_n)}] \quad \left| \begin{array}{l} X \sim b(n, p) \\ M_X(t) = \underline{(1-p+pe^t)^n} \end{array} \right.$$

$$= E\left[\prod_{k=1}^n e^{t X_k}\right]$$

$$= \prod_{k=1}^n E[e^{t X_k}]$$

$$= \prod_{k=1}^n M_{X_k}(t)$$

$$= \prod_{k=1}^n (1-p+pe^t)^{n_k} \quad \because X_k \sim b(n_k, p)$$

$$= \prod_{k=1}^n (1-p+pe^t)^{n_k} \quad \because X_k \sim b(n_k, p)$$

$$= \underline{(1-p+pe^t)^{\sum_{k=1}^n n_k}} \longrightarrow \text{MGF of } b(\sum_{k=1}^n n_k, p)$$

$$= S_n \sim b(\sum_{k=1}^n n_k, p)$$

# Let  $X_1, X_2, \dots, X_k$  be independent Poisson RVs with  $X_k \sim P(\lambda_k)$ ,  $k=1, \dots, n$

$$\sum_{k=1}^n X_k \sim P(\sum_{k=1}^n \lambda_k)$$

Use MGF to prove this

$$\left\{ \begin{array}{l} M_X(t) \\ = e^{\lambda(e^t-1)} \\ \text{if } X \sim P(\lambda) \end{array} \right.$$

(exercise)

Exercise: Let  $X$  and  $Y$  be independent Poisson RVs with parameters  $\lambda_1 > 0$ ,  $\lambda_2 > 0$ , respectively. Show that

the conditional distribution of  $X$  given  $X+Y$  is a binomial distribution.

Exercise: If  $X \sim P(\lambda)$  and the conditional distribution of  $Y$  given  $X=x$  is  $b(x, p)$ , then  $Y \sim P(\lambda p)$ .

Negative Binomial distribution:

Consider a succession of trials of an experiment with success probability  $p$ .

Let us compute the probability of observing exactly  $r$  successes, where  $r \geq 1$  is fixed integer.

We do this by counting the number

of failures before  $r^{\text{th}}$  success.

$X \rightarrow$  no. of failures before  $r^{\text{th}}$  success.

To make this happen  $X+r$  trials are required.

$(X+r)^{\text{th}}$  trial is success.

and there are exactly  $X$  failures in  $X+r-1$  trials.

$$P\{X=k\} = \binom{k+r-1}{k} p^r (1-p)^k$$

$k=0, 1, 2, \dots$

A RV  $X$  with PMF given by  $\oplus$  is said to follow negative binomial dist.  $\oplus$

$$\binom{k+r-1}{k} = \frac{(k+r-1)!}{k! (r-1)!}$$

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$$= \frac{(k+r-1)(k+r-2)\dots(r+1)(r)(r-1)!}{k! \cancel{(r-1)!}}$$

$$= \frac{(-1)^k \cdot (-r)(-r-1)\dots(-r-k+1)}{k!} \quad \checkmark$$

$$\parallel \binom{-r}{k} \quad \checkmark$$

$$\binom{k+r-1}{k} = (-1)^k \binom{-r}{k}$$

$$\text{let } q = (1-p) < 1$$

$$\underline{(1-q)^{-r}} = \sum_{k=0}^{\infty} \binom{-r}{k} (-q)^k$$

$$= \sum_{k=0}^{\infty} (-1)^k \binom{-r}{k} \underline{q^k}$$

$$\boxed{1 = \sum_{k=0}^{\infty} \binom{k+r-1}{k} p^r (1-p)^k}$$

$$M_X(t) = E[e^{tx}]$$

$$= \sum_{k=0}^{\infty} \frac{(k+r-1) \dots r}{k!} t^k$$

$$= \sum_{k=0}^{\infty} \binom{k+r-1}{k} p^r (1-p)^k e^{tk}$$

$$= p^r \sum_{k=0}^{\infty} \binom{k+r-1}{k} ((1-p)e^t)^k$$

$$M_X(t) = p^r (1 - (1-p)e^t)^{-r} \quad \text{if } \underline{e^t(1-p) < 1}$$

$$e^t < \frac{1}{1-p}$$

$$t < \log\left(\frac{1}{1-p}\right)$$

$$E[X] = \frac{r(1-p)}{p}$$

$$\text{Var}(X) = \frac{r(1-p)}{p^2}$$

exercise.

Special case:

$$\underline{r=1}$$

$X \rightarrow$  number of failures before first success

$$P\{X=k\} = p \cdot (1-p)^k$$

$$k=0, 1, \dots, \infty$$

Here  $X$  is said to follow a geometric distribution.

$$E[X] = \frac{1-p}{p}$$

$$\text{Var}(X) = \frac{1-p}{p^2}$$

Example: An oil company conducts a geological study that indicates that an exploratory oil well should have 20% chance of striking oil.

- i) what is the probability that third strike comes in seventh drill.
- ii) what is the probability that first strike comes on third well drilled.

(Exercise) i)  $x=3$ ,  $P\{X=4\}$   
 $X \sim \text{NB}(p=0.2)$

ii)  $X \sim$  geometric distribution

$$P\{X=2\}, \quad p=0.2$$

Theorem: If  $X$  has a geometric distribution, then for any two non-negative integers  $m, n$ ,

$$\underline{P\{X > m+n \mid X > m\}} = \underline{P\{X \geq n\}}$$

Proof:

$$\begin{aligned} P\{X > m+n \mid X > m\} &= \frac{P\{X > m+n, X > m\}}{P\{X > m\}} \\ &= \frac{P\{X > m+n\}}{P\{X > m\}} = \frac{(1-p)^{m+n+1}}{(1-p)^{m+1}} \\ &= (1-p)^n = P\{X \geq n\} \end{aligned}$$

Theorem: Let  $X_1, X_2, \dots, X_n$  be



independent geometric RVs with parameters  $p_1, p_2, \dots, p_n$ , respectively. Then

$X_{(1)} = \min \{X_1, \dots, X_n\}$  is also a geometric RV with parameter

$$p^* = 1 - \prod_{i=1}^n (1 - p_i)$$

Proof :-  $X \equiv$  geometric RV with parameter  $p$

$$\underline{P\{X \leq k\} = 1 - P\{X > k\}}$$

$$\underline{F(k) = 1 - (1-p)^{k+1}} \quad \forall k.$$

$$\underline{P\{X_{(1)} \leq k\}} = 1 - P\{X_{(1)} > k\}$$

$$= 1 - P\{\underline{\min\{X_1, \dots, X_n\} > k}\}$$

$$= 1 - P\{X_1 > k, X_2 > k, \dots, X_n > k\}$$

$$= 1 - \prod_{i=1}^n \underline{P\{X_i > k\}} \quad |$$

$$\begin{aligned}
 & \prod_{i=1}^n \mathbb{P}\{X_i > k\} \\
 &= 1 - \prod_{i=1}^n (1 - p_i)^{k+1} \\
 &= 1 - \left( \prod_{i=1}^n (1 - p_i) \right)^{k+1} \\
 &= 1 - (1 - p^*)^{k+1}
 \end{aligned}$$

$$\begin{aligned}
 p^* &= 1 - \prod_{i=1}^n (1 - p_i) \\
 \Rightarrow \prod_{i=1}^n (1 - p_i) &= 1 - p^*
 \end{aligned}$$

$$F(k) = 1 - (1 - p^*)^{k+1}$$

$X_{(1)}$

$\Rightarrow X_{(1)}$  is geometric RV with parameter  $p^*$ .

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