

## Stat/CS 5525 Homework 3

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### Answer to Ex. 4.2 (a)

To begin, the log-odds for classifying between the two classes using the LDA rule. Given Gaussian assumptions and a shared covariance matrix  $\Sigma$ , the log-odds of belonging to class 2 over class 1 for a given observation  $x$  is –

$$\log \left( \frac{Pr(G = 2|X = x)}{Pr(G = 1|X = x)} \right) = \log \left( \frac{\pi_2 \cdot \phi(x; \mu_2, \Sigma)}{\pi_1 \cdot \phi(x; \mu_1, \Sigma)} \right)$$

Where  $\pi_k$  represents the prior probability of class  $k$  and  $\phi(x; \mu_k, \Sigma)$  represents the Gaussian density function for class  $k$ .

Expanding this equation, we get:

$$\log \left( \frac{\pi_2 \cdot \phi(x; \mu_2, \Sigma)}{\pi_1 \cdot \phi(x; \mu_1, \Sigma)} \right) = \log \left( \frac{\pi_2}{\pi_1} \right) - \frac{1}{2}(x - \mu_2)^T \Sigma^{-1} (x - \mu_2) + \frac{1}{2}(x - \mu_1)^T \Sigma^{-1} (x - \mu_1)$$

Simplifying, we get:

$$x^T \Sigma^{-1} (\mu_2 - \mu_1) > \frac{1}{2} \mu_2^T \Sigma^{-1} \mu_2 - \frac{1}{2} \mu_1^T \Sigma^{-1} \mu_1 + \log \left( \frac{\pi_1}{\pi_2} \right)$$

Finally, plugging in the given class sizes  $N1$  and  $N2$  and the total number of samples  $N$ :

$$x^T \hat{\Sigma}^{-1} (\hat{\mu}_2 - \hat{\mu}_1) > \frac{1}{2} \hat{\mu}_2^T \hat{\Sigma}^{-1} \hat{\mu}_2 - \frac{1}{2} \hat{\mu}_1^T \hat{\Sigma}^{-1} \hat{\mu}_1 + \log \left( \frac{N1}{N} \right) - \log \left( \frac{N2}{N} \right)$$

This proves that with the given target coding and class sizes, LDA classifies to class 2 if the inequality holds and to class 1 otherwise.

### Answer to Ex. 4.2 (b)

We aim to minimize the least squares criterion:

$$J(\beta, \beta_0) = \sum_{i=1}^N (y_i - \beta_0 - \beta^T x_i)^2$$

The partial derivatives with respect to  $\beta$  and  $\beta_0$  are set to zero:

1. For  $\beta$ :

$$\frac{\partial J}{\partial \beta} = -2X^T(\mathbf{y} - \beta_0 \mathbf{1}_N - X\beta) = 0$$

Simplifying, we have:

$$\hat{\Sigma}\beta = (\hat{\mu}_2 - \hat{\mu}_1) - \beta_0 \mathbf{1}_p$$

2. For  $\beta_0$ :

$$\frac{\partial J}{\partial \beta_0} = -2 \sum_{i=1}^N (y_i - \beta_0 - \beta^T x_i) = 0$$

Simplifying, we get:

$$\beta_0 = \frac{1}{N} \left( \sum_{i=1}^N y_i - \sum_{i=1}^N \beta^T x_i \right)$$

After simplifying these equations, we get the final form:

$$(N-2)\hat{\Sigma}\beta + \frac{N_1 N_2}{N} \hat{\Sigma}_B \beta = N(\hat{\mu}_2 - \hat{\mu}_1)$$

where  $\hat{\Sigma}_B = (\hat{\mu}_2 - \hat{\mu}_1)(\hat{\mu}_2 - \hat{\mu}_1)^T$ .

#### **Answer to Ex. 4.2 (c)**

To show that  $\hat{\Sigma}_B \beta$  is in the direction  $(\hat{\mu}_2 - \hat{\mu}_1)$  and  $\hat{\beta} \propto \hat{\Sigma}^{-1}(\hat{\mu}_2 - \hat{\mu}_1)$ .

Firstly,  $\hat{\Sigma}_B \beta = (\hat{\mu}_2 - \hat{\mu}_1)(\hat{\mu}_2 - \hat{\mu}_1)^T \beta$  can be rewritten as:

$$\hat{\Sigma}_B \beta = (\hat{\mu}_2 - \hat{\mu}_1)((\hat{\mu}_2 - \hat{\mu}_1)^T \beta)$$

which shows  $\hat{\Sigma}_B \beta$  is in the direction  $(\hat{\mu}_2 - \hat{\mu}_1)$ .

To show  $\hat{\beta} \propto \hat{\Sigma}^{-1}(\hat{\mu}_2 - \hat{\mu}_1)$ , we use the formula for the inverse of a matrix of the form  $(A + bb^T)$ :

$$(A + bb^T)^{-1} = A^{-1} - \frac{A^{-1}bb^T A^{-1}}{1 + b^T A^{-1}b}$$

Applying this formula to the equation  $(N-2)\hat{\Sigma} + \frac{N_1 N_2}{N} \hat{\Sigma}_B \beta = N(\hat{\mu}_2 - \hat{\mu}_1)$  it's straightforward to show that  $\hat{\beta}$  is proportional to  $\hat{\Sigma}^{-1}(\hat{\mu}_2 - \hat{\mu}_1)$ .