

$$1.a. \quad \int_{-4}^5 c(x+4)dx = 1$$

$$t = x + 4$$

$$\Rightarrow c \int_0^9 t dt = 1$$

$$\Rightarrow c \cdot \frac{t^2}{2} \Big|_0^9 = c \cdot \frac{81}{2} = 1$$

$$\Rightarrow c = \frac{2}{81}$$

1 mark

$$y = x|x| = \begin{cases} -x^2, & x < 0 \\ x^2, & x \geq 0 \end{cases} \quad \uparrow \text{on } \mathbb{R}$$

$$S_y = (-16, 25) = (-16, 0) \cup (0, 25)$$

- For $-16 < y < 0$, we have

$$P(X|X| \leq y) = P(X < 0, X^2 \geq -y) +$$

$$P(X > 0, X^2 \leq y)$$

$$= P(X \leq -\sqrt{-y}) + 0$$

$$= \frac{2}{81} \int_{-4}^{-\sqrt{-y}} (x+4) dx$$

$$= \frac{(4 - \sqrt{-y})^2}{81}$$

1 mark

- For $0 < y < 25$, we have

$$\begin{aligned} P(X|X_1 \leq y) &= P(X < 0, X^2 \geq -y) \\ &\quad + P(X > 0, X^2 \leq y) \\ &= P(X < 0) + \\ &\quad P(0 < X \leq \sqrt{y}) \\ &= P(X \leq \sqrt{y}) \\ &= \frac{2}{81} \int_{-4}^{\sqrt{y}} (x+4) dx \\ &= \frac{(4+\sqrt{y})^2}{81}. \end{aligned}$$

1 mark

So,

$$F_Y(y) = \begin{cases} 0 & , y < -16 \\ \frac{(4-\sqrt{y})^2}{81} & , -16 \leq y < 0, \\ \frac{(4+\sqrt{y})^2}{81} & , 0 \leq y < 25, \\ 1 & , y \geq 25. \end{cases}$$

1 mark

F_Y is differentiable (except at a finite number of points).

So, the pdf is

$$f_Y(y) = F'_Y(y)$$

$$= \begin{cases} \frac{(4 - \sqrt{-y})}{81 \cdot \sqrt{-y}}, & -16 \leq y < 0 \\ \frac{(4 + \sqrt{y})}{81 \cdot \sqrt{y}}, & 0 \leq y < 25 \\ 0, & \text{otherwise.} \end{cases}$$

1 mark

1. b. (i) Let $x > 0$.

$$F(x) = \int_0^x 2\beta t e^{-\beta t^2} dt$$
$$z = \beta t^2$$
$$= \int_0^{\beta x^2} e^{-z} dz$$
$$= 1 - e^{-\beta x^2}$$

$$F(x) = \begin{cases} 0 & , x < 0 \\ 1 - e^{-\beta x^2} & , x \geq 0. \end{cases}$$

1 mark

Let the median be m_0 .

$$F(m_0) = \frac{1}{2}$$

$$\Rightarrow 1 - e^{-\beta m_0^2} = \frac{1}{2}$$

$$\Rightarrow e^{-\beta m_0^2} = \frac{1}{2}$$

$$\Rightarrow \beta m_0^2 = \log_e 2$$

$$\Rightarrow m_0 = \sqrt{\frac{\log_e 2}{\beta}} .$$

1 mark

Now,

$$\log_e f(t) = \log_e 2\beta + \log t - \beta t^2$$

$$\Rightarrow f'(t) = f(t) \left[\frac{1}{t} - 2\beta t \right] = 0$$

$$\Rightarrow t_0 = \frac{1}{\sqrt{2\beta}} \quad [1 \text{ mark}]$$

$$f''(t) = f(t) \left[-\frac{1}{t^2} - 2\beta \right] +$$

$$f(t) \left[\frac{1}{t} - 2\beta t \right]^2$$

$$= f(t) \left[-4\beta + 4\beta^2 t^2 \right]$$

$$f''(t_0) = f(t_0) \cdot \left[-4\beta + 2\beta \right]$$

$$= -2\beta f(t_0) < 0 \dots$$

$$[\because \beta > 0 \text{ & } f > 0]$$

[1 mark]

1. b . (ii) From $M_X(t)$, it is clear
that

$$p(x) = \begin{cases} \frac{1}{8} & x = -1, \\ \frac{1}{4} & x = 0, \\ \frac{5}{8} & x = 2. \end{cases}$$

1 mark

$$E(X) = -\frac{1}{8} + \frac{10}{8} = \frac{9}{8}$$

$$E(X^2) = \frac{1}{8} + \frac{20}{8} = \frac{21}{8}$$

$$\text{Var}(X) = \frac{21}{8} - \frac{81}{64} = \frac{87}{64}$$

1 mark

2.a.i) $X_1, \dots, X_n \stackrel{iid}{\sim} N(0,1)$

$\Rightarrow \bar{X}_n$ is independent of
[Student's theorem] $S_n^2 \quad \forall n \geq 1.$

$\Rightarrow \bar{X}_n$ is independent of
 $\frac{1}{S_n} \quad \forall n \geq 1$

Fix $n=5$.

$$E\left[\frac{\bar{X}_5}{S_5}\right] = E\left[\bar{X}_5\right] \cdot E\left[\frac{1}{S_5}\right]$$
$$= 0 \quad \boxed{1 \text{ mark}}$$

$$\text{Var}\left[\frac{\bar{X}_5}{S_5}\right] = E\left[\frac{\bar{X}_5^2}{S_5^2}\right] - 0$$
$$= E\left[\bar{X}_5^2\right] \cdot E\left[\frac{1}{S_5^2}\right]$$

$$E\left[\bar{X}_5^2\right] = \frac{1}{5} \quad \boxed{1 \text{ mark}}$$

$$4S_5^2 \sim \chi_4^2$$

$$= \frac{4}{5} \cdot E\left[\frac{1}{4S_5^2}\right]$$

$$= \frac{4}{5} \cdot \frac{2^{2/2} \cdot \sqrt{2/2}}{2^{4/2} \cdot \sqrt{4/2}}$$

$$= \frac{4}{5} \cdot \frac{2 \cdot 1}{4 \cdot 1}$$

$$\text{So, } \text{Var} \left[\frac{\bar{X}_5}{S_5} \right] = \frac{2}{5} .$$

1 mark

2.a.ii) X_1, \dots, X_n iid $\text{Unif}(0,1)$

Claim : $X_{(1)} \xrightarrow{P} 0$ as $n \rightarrow \infty$
 $X_{(n)} \xrightarrow{P} 1$ as $n \rightarrow \infty$.

Proof : $E[X_{(1)}^2] = n \int_0^1 y^2 (1-y)^{n-1} dy$

$$= n \cdot B(3, n)$$

$$= n \cdot \frac{2! (n-1)!}{(n+2)!}$$

$$= \frac{2}{(n+2)(n+1)} \rightarrow 0$$

$\Rightarrow X_{(1)} \xrightarrow{L_2} 0$ as $n \rightarrow \infty$.

1 mark

Similarly,

$$E[1 - X_{(n)}]^2 = n \int_0^1 y^{n-1} (1-y)^2 dy$$

$$= n \cdot B(n, 3)$$

$\rightarrow 0$ as $n \rightarrow \infty$.

$$\Rightarrow X(n) \xrightarrow{L_2} 1 \text{ as } n \rightarrow \infty.$$

1 mark

Combining these two statements,
we have

$$X_D X(n) \xrightarrow{P} 0 \text{ as } n \rightarrow \infty.$$

1 mark

$$2.b. \quad \begin{pmatrix} z_1 \\ z_2 \end{pmatrix} \sim N_2 \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{bmatrix} 1 & \gamma_2 \\ \gamma_2 & 1 \end{bmatrix} \right)$$

i). $\max\{z_1, z_2\} = \frac{z_1 + z_2 + |z_1 - z_2|}{2}$

$$E[\max\{z_1, z_2\}] = \frac{0 + E|z_1 - z_2|}{2}$$

$$z_1 - z_2 \sim N(0, 1).$$

$$\text{So, } E|z_1 - z_2| = \sqrt{\frac{2}{\pi}}.$$

$$\Rightarrow E[\max\{z_1, z_2\}] = \frac{1}{\sqrt{2\pi}}.$$

ii). Recall $z_2 | z_1 \sim N\left(\frac{z_1}{2}, \frac{3}{4}\right)$.

$$E[e^{t z_1 z_2}]$$

1 mark

$$= E_{z_1} \left\{ E[e^{t z_1 z_2} | z_1] \right\}$$

$$= E_{z_1} \left\{ e^{\frac{t z_1^2}{2} + \frac{3 t^2 z_1^2}{8}} \right\}$$

$$= E_{Z_1} \left(e^{sZ_1^2} \right), \quad s = \frac{t}{2} + \frac{3t^2}{8}$$

$$= \frac{1}{\sqrt{1-2s}} \quad \text{for } s < \frac{1}{2} \quad [1 \text{ mark}]$$

$$[\because Z_1^2 \sim \chi_1^2]$$

$$= \frac{1}{\sqrt{1-(t+\frac{3t^2}{4})}} \quad \text{for } -2 < t < \frac{2}{3}. \quad [1 \text{ mark}]$$

3. a. $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Unif}(\theta-a, \theta+a)$

i> $L(\theta, a) = \prod_{i=1}^n f(X_i)$

$$= \frac{1}{(2a)^n}, \quad \theta-a \leq X_{(1)} < X_{(n)} \\ \leq \theta+a.$$

\downarrow
free of θ !

Now,

$$X_{(n)} - a \leq \theta \leq X_{(1)} + a.$$

Given a , L is maximized if

$$\theta \in [X_{(n)} - a, X_{(1)} + a].$$

In particular, let us consider

$$\hat{\theta}_{MLE} = \frac{X_{(n)} - a + X_{(1)} + a}{2}$$

$$= \frac{X_{(1)} + X_{(n)}}{2}.$$

1 mark

Note that L is \downarrow in a .

So, L is maximized if a is minimized. Note that

$$\underline{\theta - X_{(1)} \leq a \leq X_{(n)} - \theta}$$

$$\Rightarrow \hat{a}_{MUE} = \hat{\theta}_{MUE} - \bar{X}_{(1)}$$

$$= \frac{X_{(n)} - X_{(1)}}{2} . \boxed{1 \text{ mark}}$$

ii) $E[X_{(1)}] = n \int_{\theta-a}^{\theta+a} x \cdot \frac{1}{2a} \cdot \left[1 - \frac{x-\theta+a}{2a} \right]^{n-1} dx$

$$z = \frac{x-\theta+a}{2a}, dz = \frac{dx}{2a}$$

$$= n \int_0^1 (2az + \theta - a) \cdot (1-z)^{n-1} dz$$

$$= 2an \int_0^1 z (1-z)^{n-1} dz +$$

$$(\theta - a)^n \int_0^1 (1-z)^{n-1} dz$$

$$= 2an \text{Beta}(2, n) + \frac{(\theta-a)}{n} \cdot n$$

$$= 2an \cdot \frac{1! (n-1)!}{(n+1)!} + (\theta-a)$$

mark

$$= \frac{2a}{(n+1)} + (\theta-a).$$

$$E[X_{(n)}] = n \int_{\theta-a}^{\theta+a} x \cdot \frac{1}{2a} \cdot \left(\frac{x-\theta+a}{2a}\right)^{n-1} dx$$

$$z = \frac{x-\theta+a}{2a}, dz = \frac{dx}{2a}$$

$$= n \int_0^1 (2az + \theta - a) \cdot z^{n-1} dz$$

$$= 2an \int_0^1 z^n dz + (\theta-a)n \int_0^1 z^{n-1} dz$$

$$= \frac{2a}{n+1} + (\theta - a)$$

$$= 2a - \frac{2a}{n+1} + (\theta - a)$$

$$= (\theta + a) - \frac{2a}{n+1}$$

Nomr,

$$E[\hat{\theta}_{MLE}] = E\left[\frac{X_{(1)} + X_{(n)}}{2}\right]$$

$$= \frac{1}{2} \left[\cancel{\frac{2a}{n+1}} + \theta - a + \theta + a - \cancel{\frac{2a}{n+1}} \right]$$

$$= \theta \text{ (Yes)}$$

[mark]

$$E[\hat{a}_{MLE}] = E\left[\frac{X_{(n)} - X_{(1)}}{2}\right]$$

$$= \frac{1}{2} \left[(\theta + a) - \frac{2a}{n+1} - \frac{2a}{n+1} - \theta + a \right]$$

$$= \frac{1}{2} \left[2a - \frac{4a}{n+1} \right] = a - \frac{2a}{n+1}$$

$$= \frac{(n-1)}{(n+1)} \cdot a$$

$$\neq a \quad (\text{No})$$

1 mark

3. b. Using the NP lemma, we get

if $\prod_{i=1}^n \frac{2x_i}{1} > c$, then we reject H_0 .

Critical region: $\prod_{i=1}^n x_i > c/2^n$. 1 mark

$$P_{H_0} \left(\prod_{i=1}^n x_i > c_1 \right)$$

$$= P_{H_0} \left(-2 \sum_{i=1}^n \log x_i < -2 \log c_1 \right)$$

Under H_0 , $X \sim U(0,1)$

$$\begin{aligned} -2 \log X &\sim \chi^2 \\ \Rightarrow -2 \sum_{i=1}^n \log x_i &\sim \chi^2_{2n} \quad (\text{Additivity of } \chi^2) \end{aligned}$$

$$\text{Now, } P(\chi^2_{2n} < k) = \alpha$$

$$\Rightarrow k = \chi^2_{2n} (1-\alpha)$$

1 mark

For $\alpha = 0.10$ and $n = 10$, we

get $K = \chi^2_{20}(0.90) = 12.443.$

1 mark