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Regression Modelling using Priors with Fisher Information Covariance Kernels (I-priors)



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### Objectives

- Outline an efficient computational method for estimating the parameters of an I-prior model in the continuous responses case.
- Extend the I-prior methodology to categorical responses for classification and inference.
- Explore the usefulness of I-priors towards variable selection.

### Introduction

Consider the following regression model for i = 1, ..., n:

$$y_i = f(x_i) + \epsilon_i$$

$$(\epsilon_1, \dots, \epsilon_n)^{\top} \sim N_n(0, \Psi^{-1})$$
(1)

where  $y_i \in \mathbb{R}$ ,  $x \in \mathcal{X}$ , and  $f \in \mathcal{F}$ . Let  $\mathcal{F}$  be a reproducing kernel Hilbert space (RKHS) with kernel  $h_{\lambda} : \mathcal{X} \times \mathcal{X} \to \mathbb{R}$ . The Fisher information for f is given by

$$\mathcal{I}(f(x), f(x')) = \sum_{k=1}^{n} \sum_{l=1}^{n} \Psi_{k,l} h_{\lambda}(x, x_k) h_{\lambda}(x', x_l).$$
 (2)

## The I-prior

The entropy maximising prior distribution for f, subject to constraints, is

$$\mathbf{f} = (f(x_1), \dots, f(x_n)) \sim N(\mathbf{f}_0, \mathcal{I}[f]).$$

#### Of interest are the

Posterior distribution for the regression function

$$p(\mathbf{f}|\mathbf{y}) = rac{p(\mathbf{y}|\mathbf{f})p(\mathbf{f})}{\int p(\mathbf{y}|\mathbf{f})p(\mathbf{f})\,\mathrm{d}\mathbf{y}};$$
 and

Posterior predictive distribution given new data

$$p(y_{\text{new}}|\mathbf{y}) = \int p(y_{\text{new}}|f_{\text{new}}, \mathbf{y}) p(f_{\text{new}}|\mathbf{y}) d\mathbf{y}.$$

### **Estimation**

Model parameters (error precision  $\Psi$ , RKHS scale parameters  $\lambda$ , and any others) may be estimated via

- Maximum (marginal) likelihood, a.k.a. empirical-Bayes;
- Expectation-maximisation (EM) algorithm; or
- Markov chain Monte Carlo (MCMC) methods.

Under the normal model (1), the posterior for y, given some x and model parameters, is normal with mean

$$\hat{y}(x) = f_0(x) + \mathbf{h}_{\lambda}^{\mathsf{T}}(x)\Psi H_{\lambda}(H_{\lambda}\Psi H_{\lambda} + \Psi^{-1})^{-1} (y - f_0(x))$$

and variance

$$\hat{\sigma}^2(x) = \mathbf{h}_{\lambda}^{\top}(x)(H_{\lambda}\Psi H_{\lambda} + \Psi^{-1})^{-1}\mathbf{h}_{\lambda}(x).$$

## Placeholder

# Image

Figure 1: Figure caption

### **Computational Hurdles**

For models with many scale parameters,

- Newton methods are problematic due to the presence of multiple local optima; while
- Gibbs sampling suffer from severe autocorrelation in the posterior samples.

The EM provides a reliable and straightforward framework for estimating the parameters of I-prior models.

Regardless, computational complexity is dominated by the  $n \times n$  matrix inversion in (X), which is  $O(n^3)$ . Suppose that  $H_\lambda \Psi H_\lambda = QQ^\top$ , where Q is a  $n \times q$  matrix, is a valid low-rank decomposition. Then

$$(H_{\lambda}\Psi H_{\lambda} + \Psi^{-1})^{-1} = \Psi - \Psi Q(I_q + Q^{\top}\Psi Q)^{-1}Q^{\top}\Psi$$

is a much cheaper  $O(nq^2)$  operation, especially if  $q \ll n$ .

We explore the Nyström method for low-rank matrix approximations, which we find works well for the fractional Brownian motion RKHS.

## **Categorical Responses**

Suppose now that each  $y_i \in \{1, \ldots, m\}$  and that

$$y_i \sim \operatorname{Cat}(p_{i1}, \dots, p_{im})$$

with probability mass function

$$p(y_i) = \prod_{i=1}^m p_{ij}^{y_{ij}}, \qquad y_{ij} = [y_i = j],$$

satisfying  $p_{ij} > 0$ ,  $\sum_{j} p_{ij} = 1$  for  $j \in \{1, ..., m\}$ . In the spirit of generalised linear models, take

$$E[y_{ij}] = p_{ij} = g^{-1}(f_j(x_i))$$

with some link function  $g: \mathbb{R} \to [0,1]$  and an I-prior on  $f_j$ .

Now, the marginal (on which the posterior depends),

$$p(\mathbf{y}) = \int \prod_{i=1}^{n} \prod_{j=1}^{m} \left[ \left\{ g^{-1}(f_j(x_i)) \right\}^{[y_i=j]} \cdot N_n(\mathbf{f}_{0j}, \mathcal{I}[f_j]) \, \mathrm{d}\mathbf{f}_j \right],$$

cannot be found in closed form.

## Variational Approximation

An approximation  $q(\mathbf{f})$  to the true posterior density  $p(\mathbf{f}|\mathbf{y})$  is considered, with q chosen to minimise the Kullback-Leibler divergence (under certain restrictions):

$$KL(q||p) = -\int \log \frac{p(\mathbf{f}|\mathbf{y})}{q(\mathbf{f})} q(\mathbf{f}) d\mathbf{f}.$$

By working in a fully Bayes setting, we append the parameters to  ${\bf f}$  and employ the variational approximation. The result is an iterative algorithm similar to the EM.

As this variational-EM works harmoniously with exponential family distributions, the probit link is preferred.

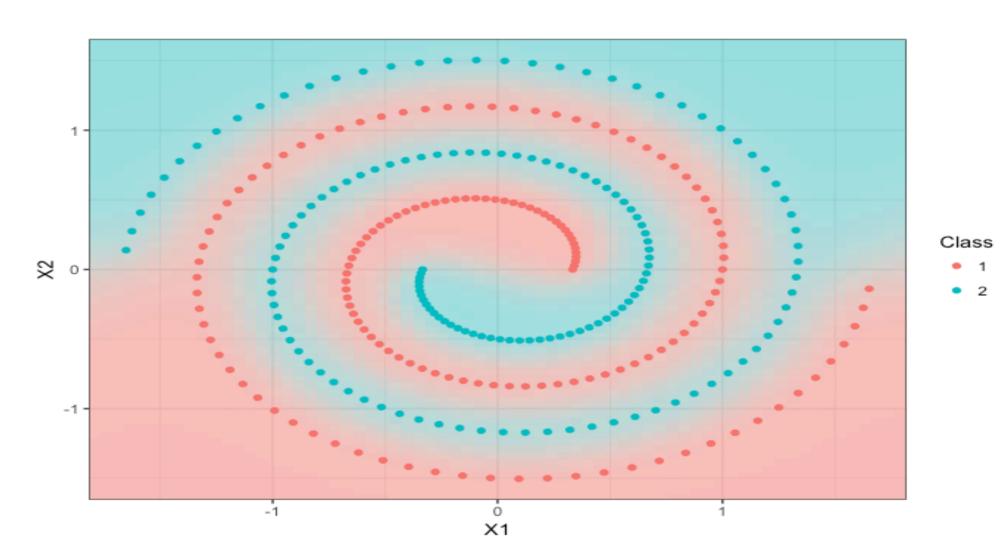


Figure 2: A toy example of binary classification using I-priors and the fBm-0.5 kernel over a two-dimensional predictor. Points indicate realisations, while background colours denote predicted classes.

### Variable Selection for Linear Models

In the original I-prior model, model selection can be done by comparing likelihoods (empirical Bayes factors). However, with p variables to select, the  $2^p$  comparisons will prove intractable.

For linear models of the form

$$(y_1, \dots, y_n)^{\top} \sim \mathrm{N}_n \left( \beta_0 \mathbf{1}_n + \sum_{j=1}^p \beta_j X_j, \Psi^{-1} \right),$$

the prior

$$(\beta_1, \dots, \beta_p)^{\top} \sim \mathrm{N}_p(0, \Lambda X^{\top} \Psi X \Lambda)$$

is an equivalent I-prior representation of (1) in the feature space of  $\beta$  under the linear kernel.

We employ a fully Bayesian treatment of the model in order to estimate posterior model probabilities

$$p(M|\mathbf{y}) \propto \int p(\mathbf{y}|M,\theta)p(\theta|M)p(M) d\theta$$

where M is the model index and  $\theta$  are model parameters.

Simulation studies together with real-data applications show promising results, outperforming methods such as greedy selection, g-priors, and regularisation (ridge and Lasso) under multicollinearity.

	Signal-to-noise Ratio (SNR)				
False choices	90%	<b>75</b> %	50%	25%	10%
0-2	0.84	0.92	0.81	0.79	0.20
3-5	0.15	0.08	0.18	0.20	0.47
>5	0.01	0.00	0.01	0.01	0.33

Table 1: Simulation results of choosing 100 pairwise correlated variables using I-priors. Under differing SNR, proportions of false choices were recorded.

### Conclusions

- The EM algorithm provides a straightforward and efficient method of estimating I-prior models.
- The dominating  $O(n^3)$  step in the algorithm can be reduced to  $O(nq^2)$ ,  $q \ll n$ , via low-rank matrix approximations.
- Advantages of I-priors in the continuous case extend well to binary and multinomial responses.
- I-priors work well for linear variable selection under multicollinearity.

### References

[1] J. M. Smith and A. B. Jones. Book Title.

[2] A. B. Jones and J. M. Smith. Article Title.

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