

# STAT430 Homework #7: Due Friday, April 15, 2022.

Name: **KEY**

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## Question 1

Let  $Y_1, \dots, Y_n$  be i.i.d.  $\text{Uniform}(0, \theta)$ . Show that  $Y_{(n)} = \max(Y_1, \dots, Y_n)$  is sufficient for  $\theta$ . Hint: The  $\text{Uniform}(0, \theta)$  density can be written  $f(y) = \theta^{-1} \cdot \mathbb{1}_{\{0 < y < \theta\}}$ , where  $\mathbb{1}_{\{A\}}$  is an indicator function which equals 1 if  $A$  is true and equals 0 otherwise.

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**Answer:**

The likelihood function is

$$\mathcal{L}(\theta \mid \mathbf{y}) = \prod_{i=1}^n \frac{1}{\theta} \mathbb{1}_{\{0 < y_i < \theta\}} = \frac{1}{\theta^n} \mathbb{1}_{\{y_{(n)} < \theta\}} \mathbb{1}_{\{y_{(1)} > 0\}}$$

We see that  $\mathcal{L}(\theta \mid \mathbf{y})$  can be written as a product of  $g(\theta, Y_{(n)}) = \frac{1}{\theta^n} \mathbb{1}_{\{y_{(n)} < \theta\}}$  and  $h(\mathbf{y}) = \mathbb{1}_{\{y_{(1)} > 0\}}$ . So by the factorization theorem,  $Y_{(n)}$  is a sufficient statistic for  $\theta$ .

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## Question 2

Let  $Y_1, \dots, Y_n$  be i.i.d. random variables with density

$$f(y) = \frac{\alpha y^{\alpha-1}}{\beta^\alpha} \quad 0 < y < \beta.$$

If  $\beta$  is known, find a one-dimensional sufficient statistic for  $\alpha$ .

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**Answer:**

The likelihood function is

$$\mathcal{L}(\alpha, \beta \mid \mathbf{y}) = \prod_{i=1}^n \frac{\alpha y_i^{\alpha-1}}{\beta^\alpha} \mathbb{1}_{\{0 < y_i < \beta\}} = \left(\frac{\alpha}{\beta^\alpha}\right)^n \left(\prod_{i=1}^n y_i\right)^{\alpha-1} \mathbb{1}_{\{y_{(1)} > 0\}} \mathbb{1}_{\{y_{(n)} < \beta\}}.$$

which can be factored into a product of  $g(\alpha, T(\mathbf{y})) = \left(\frac{\alpha}{\beta^\alpha}\right)^n (\prod_{i=1}^n y_i)^{\alpha-1}$  and  $h(\mathbf{y}) = \mathbb{1}_{\{y_{(1)} > 0\}} \mathbb{1}_{\{y_{(n)} < \beta\}}$ , so  $T(\mathbf{Y}) = \prod_{i=1}^n Y_i$  is a sufficient statistic for  $\alpha$ .

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### Question 3

Let  $X_1, \dots, X_n$  be an observation from the pdf

$$P(X = x; \theta) = \left(\frac{\theta}{2}\right)^{|x|} (1 - \theta)^{1-|x|}, \quad x = -1, 0, 1 \quad 0 \leq \theta \leq 1.$$

a) Show  $T_1 = T(X_1)$  is an unbiased estimator of  $\theta$ , where  $T(X)$  is defined

$$T(X) = \begin{cases} 2 & \text{if } x = 1 \\ 0 & \text{otherwise} \end{cases}$$

Show  $T_n = \frac{1}{n} \sum_{i=1}^n T(X_i)$  is unbiased for  $\theta$ .

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**Answer:**

$$E(T_1) = T_1(-1)P(X_1 = -1) + T_1(0)P(X_1 = 0) + T_1(1)P(X_1 = 1) = 2 \left(\frac{\theta}{2}\right) = \theta$$

Since  $ET_1 = \theta$ ,  $T_1$  is an unbiased estimator for  $\theta$ . Similarly, we can show that  $T_n$  is an unbiased estimator of  $\theta$

$$ET_n = \frac{1}{n} \sum_{i=1}^n ET(X_i) = \frac{n\theta}{n} = \theta$$

Since  $ET_n = \theta$ ,  $T_n$  is an unbiased estimator for  $\theta$ .

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b) Find the MVUE of  $\theta$  and show this estimator is better than  $T_n$ .

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**Answer:**

This was fun to figure out, but definitely a bit tricky. The UMVUE for  $\theta$  can be found by taking the expectation of an unbiased estimator of  $\theta$  conditioned on a sufficient (and complete) statistic. A sufficient statistic for  $\theta$  is  $S = \sum_{i=1}^n |X_i|$ . I am just going to assume it is complete as well, so the UMVUE for  $\theta$  is

$$E(T_n | S = s) = \frac{1}{n} \sum_{i=1}^n E(T(X_i) | S = s)$$

The expectation  $E(T(X_i) | S)$  is

$$\begin{aligned} E(T(X_i) | S = s) &= T(-1)P(X_i = -1 | S = s) + T(0)P(X_i = 0 | S = s) + T(1)P(X_i = 1 | S = s) \\ &= 2P(X_i = 1 | S = s) \\ &= \frac{2P(X_i = 1, S = s)}{P(S = s)} \\ &= \frac{2P(X_i = 1, \sum_{j \neq i} |X_j| = s - 1)}{P(S = s)} \\ &= \frac{2P(X_i = 1)P(\sum_{j \neq i} |X_j| = s - 1)}{P(S = s)} \\ &= \frac{2(\theta/2) \binom{n-1}{s-1} \theta^{s-1} (1-\theta)^{n-s}}{\binom{n}{s} \theta^s (1-\theta)^{n-s}} \\ &= \frac{\binom{n-1}{s-1}}{\binom{n}{s}} = \frac{(n-1)!}{(s-1)!(n-s)!} \frac{s!(n-s)!}{n!} = \frac{s}{n} \end{aligned}$$

Try figuring my reasoning out on that.

So the expectation  $E(T_n | S) = \frac{S}{n} = \frac{1}{n} \sum_{i=1}^n |X_i|$  is the UMVUE for  $\theta$ . To show that this is a "better" estimator than  $T_n$ , we have to show that the variance of the UMVUE is smaller than the variance of  $T_n$ . The variance of  $T_n$  is

$$\begin{aligned} Var T_n &= \frac{1}{n^2} \sum_{i=1}^n Var T(X_i) \\ &= \frac{1}{n^2} \sum_{i=1}^n (E T(X_i)^2 - (E T(X_i))^2) \\ &= \frac{n(4(\theta/2) - (\theta)^2)}{n^2} \\ &= \frac{\theta(2 - \theta)}{n} \end{aligned}$$

and the variance of  $\frac{1}{n}S$  is

$$\begin{aligned} \text{Var} \frac{1}{n}S &= \frac{1}{n^2} \sum_{i=1}^n \text{Var}|X_i| \\ &= \frac{1}{n^2} \sum_{i=1}^n (E|X_i|^2 - (E|X_i|)^2) \\ &= \frac{n(\theta/2 + \theta/2 - (\theta/2 + \theta/2)^2)}{n^2} \\ &= \frac{\theta(1-\theta)}{n} \end{aligned}$$

and it should be easy to see that  $\frac{\theta(2-\theta)}{n} > \frac{\theta(1-\theta)}{n}$ , so  $S/n$  is better than  $T_n$ , as it should be.

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#### Question 4

Let  $X_1, \dots, X_n$  be i.i.d. Geometric( $p$ ).

- a) Find the MoM estimator of  $p$ .
  - b) Prove that your estimator in part (a) is consistent for  $p$ .
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**Answer:**

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#### Question 5

Suppose  $Y_1, \dots, Y_n$  are i.i.d. according to density

$$f(y) = e^{-(y-\theta)}, \quad y \geq \theta, \quad \theta > 0.$$

- a) Find the MoM estimator of  $\theta$ .
  - b) Is the MoM estimator of  $\theta$  unbiased? If no, compute the bias.
  - c) Find the variance of the MoM estimator of  $\theta$ .
  - d) Find a sufficient statistic for  $\theta$ .
  - e) Find the MVUE of  $\theta$ .
  - f) Compare the mean squared error of the MoM estimator and the MVUE. Which one has the smallest MSE?
  - g) Find the MoM estimator of  $(\log(\theta))^{1/4}$ .
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**Answer:**

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### Question 6

Let  $X_1, \dots, X_n$  be i.i.d.  $\text{Uniform}(-\theta, \theta)$ .

- a) Find the MoM estimator of  $\theta$ .
- b) Find a one-dimensional sufficient statistic for  $\theta$ .
- c) Is the MoM estimator you found in part (a) the MVUE? Explain.

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**Answer:**

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### Question 7

For each of the following, state whether the statement is TRUE or FALSE. (*Note: Although I do not ask you to explain WHY each statement is true/false, for an exam you should understand the material well enough that you can WHY.*)

- 1) Suppose  $Y_1, \dots, Y_n$  are i.i.d. from distribution with parameter  $\theta$ , which is a function of  $\mu = E[Y_i]$ , and  $V[Y_i] < \infty$ . The MoM estimator of  $\theta^{-2/5}$  is a consistent estimator.
- 2) Suppose  $X$  is a random variable and  $g$  is a function. Then  $E[g(X)] = g(E[X])$ .
- 3) The MVUE and MLE are always functions of a minimal sufficient statistic.
- 4) MoM estimators are always unbiased.
- 5) MLEs are always unbiased.
- 6) MVUEs are always unbiased.

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**Answer:**

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