Queueing Theory and Simulation, lecture 8

Nicky van Foreest

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1 Overview

1.1 Past

- Finish approximations, G/G/1 queue in tandem
- Start with exact work: renewal reward theorem
- application to queueing: fraction of time the servers are busy is $\lambda E[S]$.

1.2 Now

Very elegant use of sample paths (ideas based on simuation) to obtain crucial work horses of queueing theory:

- Level crossing
- PASTA

1.3 Future

- Little's law
- All is set in place to analyze exact queueing models

2 Level crossing argument

2.1 Notation and concepts

$$A(n,t) = \sum_{k=1}^{\infty} I_{A_k \le t} I_{L(A_k - t) = n} = \sum_{k=1}^{A(t)} I_{L(A_k - t) = n}$$
(1)

$$Y(n,t) = \int_0^t I_{L(s)=n} \, \mathrm{d}s,$$
 (2)

$$p(n,t) = \frac{Y(n,t)}{t},\tag{3}$$

$$D(n,t) = \sum_{k=1}^{\infty} I_{D_k \le t} I_{L(D_k)=n} = \sum_{k=1}^{D(t)} I_{L(D_k)=n}$$
(4)

2.2 Useful limits

$$\lambda(n) = \lim_{t \to \infty} \frac{A(n,t)}{Y(n,t)} \tag{5}$$

$$p(n) = \lim_{t \to \infty} p(n, t) \tag{6}$$

$$\mu(n+1) = \lim_{t \to \infty} \frac{D(n,t)}{Y(n+1,t)}.$$
(7)

2.3 Level crossing

$$1 \ge |A(n,t) - D(n,t)|,\tag{8}$$

$$\lim_{t \to \infty} \frac{A(n,t)}{t} = \lim_{t \to \infty} \frac{D(n,t)}{t},\tag{9}$$

$$\lim_{t \to \infty} \frac{A(n,t)}{t} = \lim_{t \to \infty} \frac{A(n,t)}{Y(n,t)} \frac{Y(n,t)}{t} = \lambda(n)p(n), \tag{10}$$

$$\lim_{t \to \infty} \frac{D(n,t)}{t} = \lim_{t \to \infty} \frac{D(n,t)}{Y(n+1,t)} \frac{Y(n+1,t)}{t}$$
(11)

$$= \mu(n+1)p(n+1), \tag{12}$$

$$\implies$$
 (13)

$$\lambda(n)p(n) = \mu(n+1)p(n+1). \tag{14}$$

We have a recursion!

2.4 Specific queueing systems

By making proper choices for $\lambda(n)$ and $\mu(n)$ we can model many different queueing systems

- M/M/1: $\lambda(n) = \lambda, \mu(n) = \mu$.
- $M/M/\infty$: $\lambda(n) = \lambda, \mu(n) = \mu n$.
- M/M/c: $\lambda(n) = \lambda, \mu(n) = \mu \min\{n, c\}$.
- M/M/c/c: $\lambda(n) = \lambda 1\{n < c\}, \mu(n) = \mu n$.

2.5 Implications

$$p(n+1) = \frac{\lambda(n)}{\mu(n+1)}p(n),\tag{15}$$

$$p(n+1) = \frac{\lambda(n)\lambda(n-1)\cdots\lambda(0)}{\mu(n+1)\mu(n)\cdots\mu(1)}p(0)$$
(16)

$$1 = p(0) \left(1 + \sum_{n=0}^{\infty} \frac{\lambda(n)\lambda(n-1)\cdots\lambda(0)}{\mu(n+1)\mu(n)\cdots\mu(1)} \right)$$
(17)

(18)

2.6 KPIs

$$\mathsf{E}\left[L\right] = \sum_{n=0}^{\infty} np(n),\tag{19}$$

$$P[L \ge n] = \sum_{i=n}^{\infty} p(i). \tag{20}$$

3 PASTA

3.1 Notation and concepts

$$\frac{A(n,t)}{t} = \frac{A(n,t)}{Y(n,t)} \frac{Y(n,t)}{t} \to \lambda(n)p(n), \quad \text{as } t \to \infty,$$
(21)

$$\frac{A(n,t)}{t} = \frac{A(t)}{t} \frac{A(n,t)}{A(t)},\tag{22}$$

$$\frac{A(n,t)}{A(t)} = \frac{1}{A(t)} \sum_{k=1}^{A(t)} I_{L(A_k-)=n},$$
(23)

$$\pi(n) = \lim_{t \to \infty} \frac{1}{A(t)} \sum_{k=1}^{A(t)} I_{L(A_k -) = n},$$
(24)

$$\Rightarrow$$
 (25)

$$\lambda(n)p(n) = \lambda\pi(n) \tag{26}$$

3.2 PASTA (Poisson arrivals see time averages)

- $\lambda(n) = \lambda \iff p(n) = \pi(n)$.
- Jobs arrive as a Poisson process $\implies \lambda(n) = \lambda$. (Hard to prove)

 \Longrightarrow

• Jobs arrive as a Poisson process $\implies p(n) = \pi(n)$

In other words, when you sample a system at times $\{T_k\}$ such that $T_k - T_{k-1}$ are exp. distributed, then sample averages converge to the time averages!