Stochastic Process and Forecasting Assignment

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Question 2

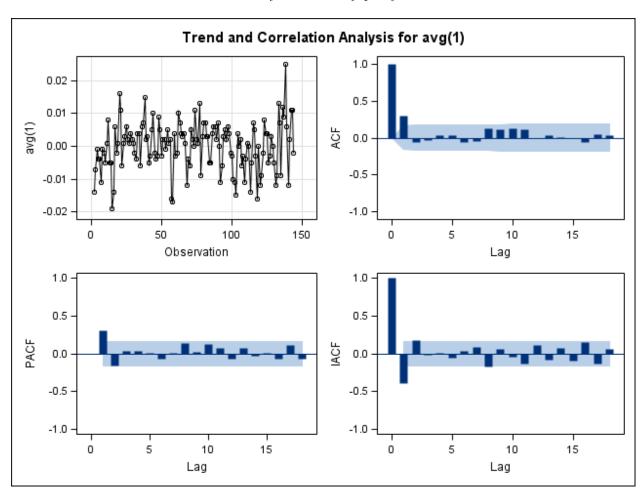
i)

The following SAS code was used to identify and estimate appropriate univariate ARIMA models for each of the avg and end exchange rate time series:

```
proc arima data=LatDol;
identify var=avg(1) nlag=18;
estimate q=1 noint;
identify var=end nlag=18;
estimate p=1;
run;
```

The diagnostics plots indicates that the first difference of avg has an ACF cut-off at lag 1, thereby leading us to believe the data follows an Arima(0,1,1) model with the following model equation:

$$\Delta Y_t = -0.36347\epsilon_{t-1} + \epsilon_t$$



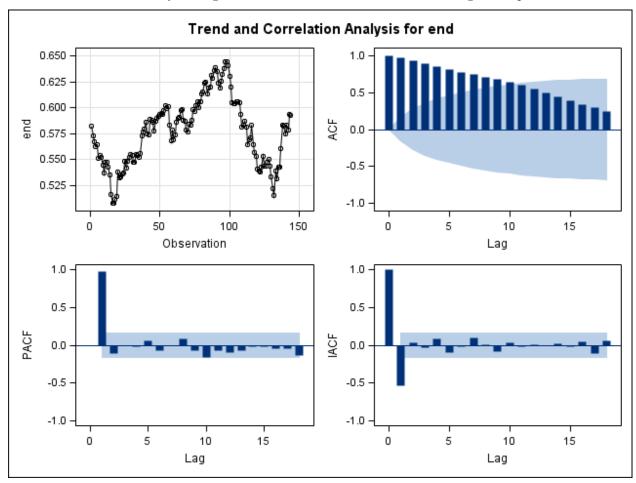
In addition the Ljung-Box chi-square statistics do not have any significant p-values, giving further weight to the chosen model:

To Lag	Chi- Square	DF	Pr > ChiSq			Autocorr	elations		
6	1.30	5	0.9346	-0.010	-0.041	-0.034	0.032	0.046	-0.051
12	8.50	11	0.6681	-0.071	0.133	0.046	0.079	0.104	-0.066
18	11.49	17	0.8299	0.059	-0.022	0.030	-0.087	0.075	0.015
24	16.13	23	0.8497	-0.014	-0.121	-0.044	-0.074	-0.025	-0.066

end is found to follow an AR(1) model with the following model equation:

$$Y_t = 0.58167 + 0.96945Y_{t-1} + \epsilon_t$$

This model was identified by the lag 1 cut-off in the PACF as shown in the diagnostics plots below:



In addition the Ljung-Box chi-square statistics do not have any significant p-values, giving further weight to the chosen model:

Autocorrelation Check of Residuals

То	Chi-		Pr >	
Lag	Square	DF	ChiSq	Autocorrelations

12 11.19 11 0.4279 -0.050 0.	31 0.153 0.092 0.081 0.076
18 13.25 17 0.7192 -0.035 -0. 24 18.91 23 0.7064 -0.024 -0.	

ii)

The following SAS code identifies and fits a VAR(p) model to the bivariate data:

```
proc varmax data=LatDol;
model end avg /
minic = (type=sbc p=(0:10) q=0) noint dif=(end(1) avg(1)) PRINT=DIAGNOSE;
output lead=3;
run;
```

Which gives the following truncated output:

Model Parameter Estimates

			Standard			
Equation	Parameter	Estimate	Error	t Value	Pr > t	Variable
end	AR1_1_1	0.03600	0.11185	0.32	0.7480	end(t-1)
	AR1_1_2	0.11050	0.12356	0.89	0.3727	avg(t-1)
avg	AR1_2_1	0.70995	0.07553	9.40	0.0001	end(t-1)
	AR1_2_2	-0.22236	0.08344	-2.66	0.0086	avg(t-1)

Covariances of Innovations

Variable	end	avg
end	0.00007	0.00004
avg	0.00004	0.00003

Based off which the fitted model is as follows:

$$\begin{pmatrix} x_t \\ y_t \end{pmatrix} = \begin{pmatrix} .03600 & 0.11050 \\ .70995 & 0.22236 \end{pmatrix} \begin{pmatrix} x_{t-1} \\ y_{t-1} \end{pmatrix} + \epsilon_t$$

The following Portmanteau Statistics shows significant p-values at all lags indicating that this VAR(1) model does not fit adequatly:

Portmanteau Test for Cross Correlations of Residuals

Up To			
Lag	DF	Chi-Square	Pr > ChiSq
2	4	23.29	0.0001
3	8	23.46	0.0028
4	12	26.37	0.0095
5	16	34.09	0.0053
6	20	35.06	0.0198
7	24	43.30	0.0092
8	28	44.79	0.0232
9	32	51.56	0.0157
10	36	58.37	0.0106
11	40	63.56	0.0103
12	44	64.87	0.0219

The following outputs show the forecasts of the average exchange rate for the next three months of 2006 are: $0.59273,\,0.59261,\,0.59261$

Forecasts

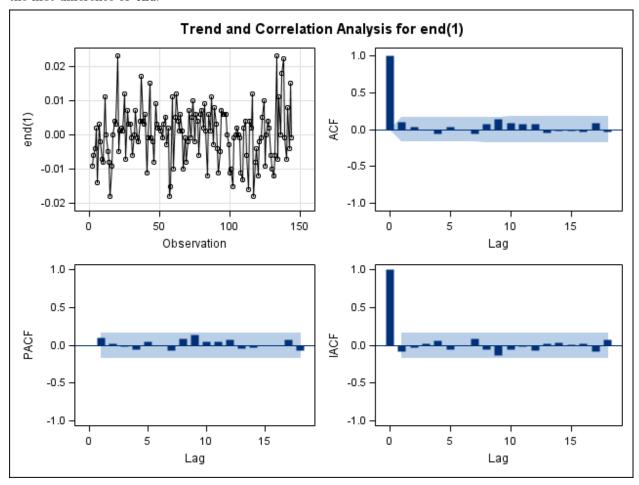
Variable	Obs	Forecast	Standard Error	95% Confiden	ce Limits
end	145	0.59274	0.00810	0.57686	0.60863
	146	0.59270	0.01204	0.56911	0.61630
	147	0.59269	0.01530	0.56270	0.62268
avg	145	0.59273	0.00547	0.58201	0.60346
	146	0.59261	0.01102	0.57101	0.61421
	147	0.59261	0.01435	0.56448	0.62074

iii)

The following SAS code was used to identify and estimate an appropriate transfer function model where avg was specified as the output variable and end the input variable:

```
proc arima data=LatDol;
identify var=end(1) nlag=18;
estimate q=1 noint;
identify var=avg(1) crosscor=end(1) nlag=18;
estimate q=1 input=(1 $ (1) / end) noint;
forecast lead=3 out=results;
run;
```

The following diagnostics plots show from the lag 1 cutoff that an ARIMA(0,1,1) model is appropriate for the first difference of end:



Hence the estimation procedure shows the following model for end:

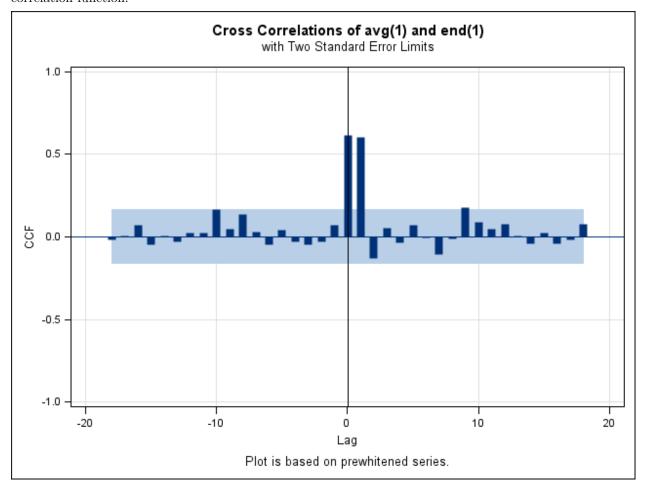
$$\Delta X_t = \eta_t + 0.09801 \eta_{t-1}$$

The following output shows the p-values from the portmanteau statistics indicate the model fits well:

Autocorrelation Check of Residuals

To Lag	Chi- Square	DF	Pr > ChiSq			Autocorr	elations		
6	0.95	5	0.9668	0.003	0.029	-0.003	-0.065	0.035	-0.003
12	7.14	11	0.7880	-0.069	0.066	0.134	0.066	0.062	0.069
18	9.75	17	0.9136	-0.045	-0.017	-0.018	-0.034	0.101	-0.044
24	14.42	23	0.9143	-0.015	-0.067	-0.102	-0.033	-0.013	-0.105

The next statement shows the output of the second identify statement which shows the following cross-correlation function:



This indicates that the transfer function is in the form of a lag-1 pulse function. There is a sharp (not decayed) response after lag 1 which indicates a transfer function of the form:

$$Y_t = \omega X_t + U_t$$

NB: The lack of intercept is due to the first differencing, and the lack of denominator terms due to the pulse intervention

With the subsequent estimation run, SAS provides us the estimated model of:

$$Y_t = (1.1118 - 0.14899L)X_{t-1} + U_t$$

Where

$$U_t = (1 - 0.95978L)\epsilon_t$$

The ARIMA Procedure

Moving Average Factors

Factor 1: 1 - 0.95978 B**(1)

Input Number 1

Input Variable end
Shift 1
Period(s) of Differencing 1

Numerator Factors

Factor 1: 1.1118 - 0.14899 B**(1)

The forecast values for avg are given as:

Forecasts for variable avg

0bs	Forecast	Std Error	95% Confidence	ce Limits
145	0.5924	0.0051	0.5824	0.6024
146	0.5922	0.0104	0.5719	0.6125
147	0.5923	0.0135	0.5658	0.6188