

Example: Demand for Cigarettes

- How much will a hypothetical cigarette tax reduce cigarette consumption? To answer this, we need the elasticity of demand for cigarettes

Panel data:

- Annual cigarette consumption and average prices paid (including tax)
- 48 continental US states, 1985-1995

$$\ln(Q_{it}^{cigarettes}) = \alpha_i + \beta_1 \ln(P_{it}^{cigarettes}) + u_{it}$$

- Why is the OLS estimator of β_1 likely to be biased?
- Proposed instrumental variable:
 $Z_i = \text{general sales tax per pack in the state} = SalesTax_i$
- Do you think this instrument is plausibly valid?
- Relevant? $\text{corr}(SalesTax_i, \ln(P_{it}^{cigarettes})) \neq 0$?
- Exogenous? $\text{corr}(SalesTax_i, u_{it}) = 0$?
- For now, use data from 1995 only.

STATA Example: Cigarette demand, First stage

Instrument = $Z = r_{taxso}$ = general sales tax (real \$/pack)

```
. reg Xlragvprs Zrtaxso if year==1995, r;
```

Regression with robust standard errors

Number of obs = 48
F(1, 46) = 40.39
Prob > F = 0.0000
R-squared = 0.4710
Root MSE = .09394

lragvprs		Robust		t	P> t	[95% Conf. Interval]	
		Coef.	Std. Err.				
-----+-----							
rtaxso		.0307289	.0048354	6.35	0.000	.0209956	.0404621
_cons		4.616546	.0289177	159.64	0.000	4.558338	4.674755

```
. predict X-hatlragvphat; Now we have the predicted values from the 1st stage
```

Second stage

```
      Y      X-hat  
. reg lpackpc lravphat if year==1995, r;
```

Regression with robust standard errors

```
Number of obs =      48  
F( 1,      46) =    10.54  
Prob > F      =    0.0022  
R-squared     =    0.1525  
Root MSE     =    .22645
```

lpackpc		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
lravphat		<i>-1.083586</i>	<i>.3336949</i>	-3.25	0.002	-1.755279 - .4118932
_cons		9.719875	1.597119	6.09	0.000	6.505042 12.93471

- These coefficients are the TSLS estimates
- The standard errors are wrong because they ignore the fact that the first stage was estimated

Combined into a single command:

```
      Y           X           Z  
. ivregress 2sls lpackpc (lavgprs = rtaxso) if year==1995, vce(robust);
```

Instrumental variables (2SLS) regression

Number of obs = 48
Wald chi2(1) = 12.05
Prob > chi2 = 0.0005
R-squared = 0.4011
Root MSE = .18635

lpackpc	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
lavgprs	-1.083587	.3122035	-3.47	0.001	-1.695494	-.471679
_cons	9.719876	1.496143	6.50	0.000	6.78749	12.65226

Instrumented: lavgprs *This is the endogenous regressor*
Instruments: rtaxso *This is the instrumental variable*

Panel data set

- Annual cigarette consumption, average prices paid by end consumer (including tax), personal income
- 48 continental US states, 1985-1995

Estimation strategy

- Having panel data allows us to control for unobserved state-level characteristics that enter the demand for cigarettes, as long as they don't vary over time
- But we still need to use IV estimation methods to handle the simultaneous causality bias that arises from the interaction of supply and demand.

Fixed-effects model of cigarette demand

$$\ln(Q_{it}^{cigarettes}) = \alpha_i + \beta_1 \ln(P_{it}^{cigarettes}) + \beta_2 \ln(Income_{it}) + u_{it}$$

- $i = 1, \dots, 48, t = 1985, 1986, \dots, 1995$
- α_i reflects unobserved omitted factors that vary across states but not over time, e.g. attitude towards smoking
- Still, $\text{corr}(\ln(P_{it}^{cigarettes}), u_{it})$ is plausibly nonzero because of supply/demand interactions
- Estimation strategy:
 - Use panel data regression methods to eliminate α_i
 - Use TSLS to handle simultaneous causality bias

The “changes” method (when $T=2$)

- One way to model long-term effects is to consider 10-year changes, between 1985 and 1995
- Rewrite the regression in “changes” form:

$$\begin{aligned}\ln(Q_{i1995}^{cigarettes}) - \ln(Q_{i1985}^{cigarettes}) \\&= \beta_1[\ln(P_{i1995}^{cigarettes}) - \ln(P_{i1985}^{cigarettes})] \\&\quad + \beta_2[\ln(Income_{i1995}) - \ln(Income_{i1985})] \\&\quad + (u_{i1995} - u_{i1985})\end{aligned}$$

- Must create “10-year change” variables, for example:
10-year change in log price = $\ln(P_{i1995}) - \ln(P_{i1985})$
- Then estimate the demand elasticity by TSLS using 10-year changes in the instrumental variables

We’ll take this approach

Use TSLS to estimate the demand elasticity by using the “10-year changes” specification

```
. ivregress 2sls Y dlpckpc W dlperinc (X dlavgprs = Z drtaxso) , r;
```

IV (2SLS) regression with robust standard errors

```
Number of obs =      48
F(   2,      45) =    12.31
Prob > F       =    0.0001
R-squared      =    0.5499
Root MSE      =    .09092
```

dlpckpc	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
dlavgprs	-.9380143	.2075022	-4.52	0.000	-1.355945	-.5200834
dlperinc	.5259693	.3394942	1.55	0.128	-.1578071	1.209746
_cons	.2085492	.1302294	1.60	0.116	-.0537463	.4708446

Instrumented: dlavgprs

Instruments: dlperinc drtaxso

NOTE:

- All the variables - Y, X, W, and Z's - are in 10-year changes
- Estimated elasticity = $-.94$ ($SE = .21$) - surprisingly elastic!
- Income elasticity small, not statistically different from zero
- Must check whether the instrument is relevant...

Check instrument relevance: compute first-stage F

```
. reg dlavgprs drtaxso dlperinc , r;
```

Regression with robust standard errors

Number of obs = 48
F(2, 45) = 16.84
Prob > F = 0.0000
R-squared = 0.5146
Root MSE = .06334

dlavgprs	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
drtaxso	.0254611	.0043876	5.80	0.000	.016624	.0342982
dlperinc	-.2241037	.2188815	-1.02	0.311	-.6649536	.2167463
_cons	.5321948	.0295315	18.02	0.000	.4727153	.5916742

```
. test drtaxso;
```

```
( 1) drtaxso = 0
```

F(1, 45) = 33.67
Prob > F = 0.0000

First stage $F = 33.7 > 10$ so instrument is not weak

*We didn't need to run "test" here
because with $m=1$ instrument, the
 F -statistic is the square of the
 t -statistic, that is,
 $5.80*5.80 = 33.67$*

Can we check instrument exogeneity? No... $m = k$

What about two instruments (cig-only tax, sales tax)?

```
      Y      W      X      Z1      Z2  
. ivregress 2sls dlpackpc dlperinc (dlavgprs = drtaxso drtax) , vce(r);
```

Instrumental variables (2SLS) regression

Number of obs = 48
Wald chi2(2) = 45.44
Prob > chi2 = 0.0000
R-squared = 0.5466
Root MSE = .08836

dlpackpc	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	

dlavgprs	-1.202403	.1906896	-6.31	0.000	-1.576148	-.8286588
dlperinc	.4620299	.2995177	1.54	0.123	-.1250139	1.049074
_cons	.3665388	.1180414	3.11	0.002	.1351819	.5978957

Instrumented: dlavgprs

Instruments: dlperinc drtaxso drtax

drtaxso = general sales tax only

drtax = cigarette-specific tax only

Estimated elasticity is -1.2, even more elastic than using general sales tax only

With $m > k$, we can test the overidentifying restrictions

Test the overidentifying restrictions

```
. predict e, resid;           Computes predicted values for most recently
                                estimated regression (the previous TSLS regression)
. reg e drtaxso drtax dlperinc;    Regress e on Z's and W's
```

Source	SS	df	MS	Number of obs =	48
Model	.037769176	3	.012589725	F(3, 44) =	1.64
Residual	.336952289	44	.007658007	Prob > F =	0.1929
Total	.374721465	47	.007972797	R-squared =	0.1008
				Adj R-squared =	0.0395
				Root MSE =	.08751

e	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
drtaxso	.0127669	.0061587	2.07	0.044	.000355	.0251789
drtax	-.0038077	.0021179	-1.80	0.079	-.008076	.0004607
dlperinc	-.0934062	.2978459	-0.31	0.755	-.6936752	.5068627
_cons	.002939	.0446131	0.07	0.948	-.0869728	.0928509

```
. test drtaxso drtax;

( 1) drtaxso = 0
( 2) drtax = 0

F( 2, 44) = 2.47
Prob > F = 0.0966

Compute J-statistic, which is m*F,
where F tests whether coefficients on
the instruments are zero
so J = 2 * 2.47 = 4.93
** WARNING - this uses the wrong d.f. **
```

The correct degrees of freedom for the J -statistic is $m-k$:

- $J = mF$, where F = the F -statistic testing the coefficients on Z_{1i}, \dots, Z_{mi} in a regression of the TSLS residuals against $Z_{1i}, \dots, Z_{mi}, W_{1i}, \dots, W_{mi}$.
- Under the null hypothesis that all the instruments are exogenous, J has a chi-squared distribution with $m-k$ degrees of freedom
- Here, $J = 4.93$, distributed chi-squared with d.f. = 1; the 5% critical value is 3.84, so reject at 5% sig. level.
- In STATA:

```
. dis "J-stat = " r(df)*r(F) " p-value = " chiprob(r(df)-1,r(df)*r(F));
J-stat = 4.9319853 p-value = .02636401
```

Check instrument relevance: compute first-stage F

```
      x      z1      z2      w  
. reg dlavgprs drtaxso drtax dlperinc , r;
```

Regression with robust standard errors

Number of obs = 48
F(3, 44) = 66.68
Prob > F = 0.0000
R-squared = 0.7779
Root MSE = .04333

dlavgprs	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

drtaxso	.013457	.0031405	4.28	0.000	.0071277	.0197863
drtax	.0075734	.0008859	8.55	0.000	.0057879	.0093588
dlperinc	-.0289943	.1242309	-0.23	0.817	-.2793654	.2213767
_cons	.4919733	.0183233	26.85	0.000	.4550451	.5289015

```
. test drtaxso drtax;
```

```
( 1) drtaxso = 0
```

```
( 2) drtax = 0
```

F(2, 44) = 88.62
Prob > F = 0.0000

88.62 > 10 so instruments aren't weak

Tabular summary of these results:

TABLE 10.1 Two Stage Least Squares Estimates of the Demand for Cigarettes Using Panel Data for 48 U.S. States

Dependent variable: $\ln(Q_{i,1995}^{cigarettes}) - \ln(Q_{i,1985}^{cigarettes})$

Regressor	(1)	(2)	(3)
$\ln(P_{i,1995}^{cigarettes}) - \ln(P_{i,1985}^{cigarettes})$	-0.94** (0.21)	-1.34** (0.23)	-1.20** (0.20)
$\ln(Inc_{i,1995}) - \ln(Inc_{i,1985})$	0.53 (0.34)	0.43 (0.30)	0.46 (0.31)
Intercept	0.21 (0.13)	0.45** (0.14)	0.37** (0.12)
Instrumental variable(s)	Sales tax	Cigarette-specific tax	Both sales tax and cigarette-specific tax
First-stage F -statistic	33.70	107.20	88.60
Overidentifying restrictions J -test and p -value	-	-	4.93 (0.026)

These regressions were estimated using data for 48 U.S. states (48 observations on the ten-year differences). The data are described in Appendix 10.1. The J -test of overidentifying restrictions is described in Key Concept 10.6 (its p -value is given in parentheses), and the first-stage F -statistic is described in Key Concept 10.5. Individual coefficients are statistically significant at the *5% level or **1% significance level.

How should we interpret the J-test rejection?

- J-test rejects the null hypothesis that both the instruments are exogenous
- This means that either $rtaxso$ is endogenous, or $rtax$ is endogenous, or both!
- The J-test doesn't tell us which! You must exercise judgment...
- Why might $rtax$ (cig-only tax) be endogenous?
 - Political forces: history of smoking or lots of smokers (political pressure for low cigarette taxes)
 - If so, cig-only tax is endogenous
- This reasoning doesn't apply to general sales tax
- → use just one instrument, the general sales tax

The Demand for Cigarettes: Summary of Empirical Results

- Use the estimated elasticity based on TSLS with the general sales tax as the only instrument:
Elasticity = $-.94$, $SE = .21$
- This elasticity is surprisingly large (not inelastic) – a 1% increase in prices reduces cigarette sales by nearly 1%.
This is much more elastic than conventional wisdom in the health economics literature.
- This is a long-run (ten-year change) elasticity. *What would you expect a short-run (one-year change) elasticity to be – more or less elastic?*