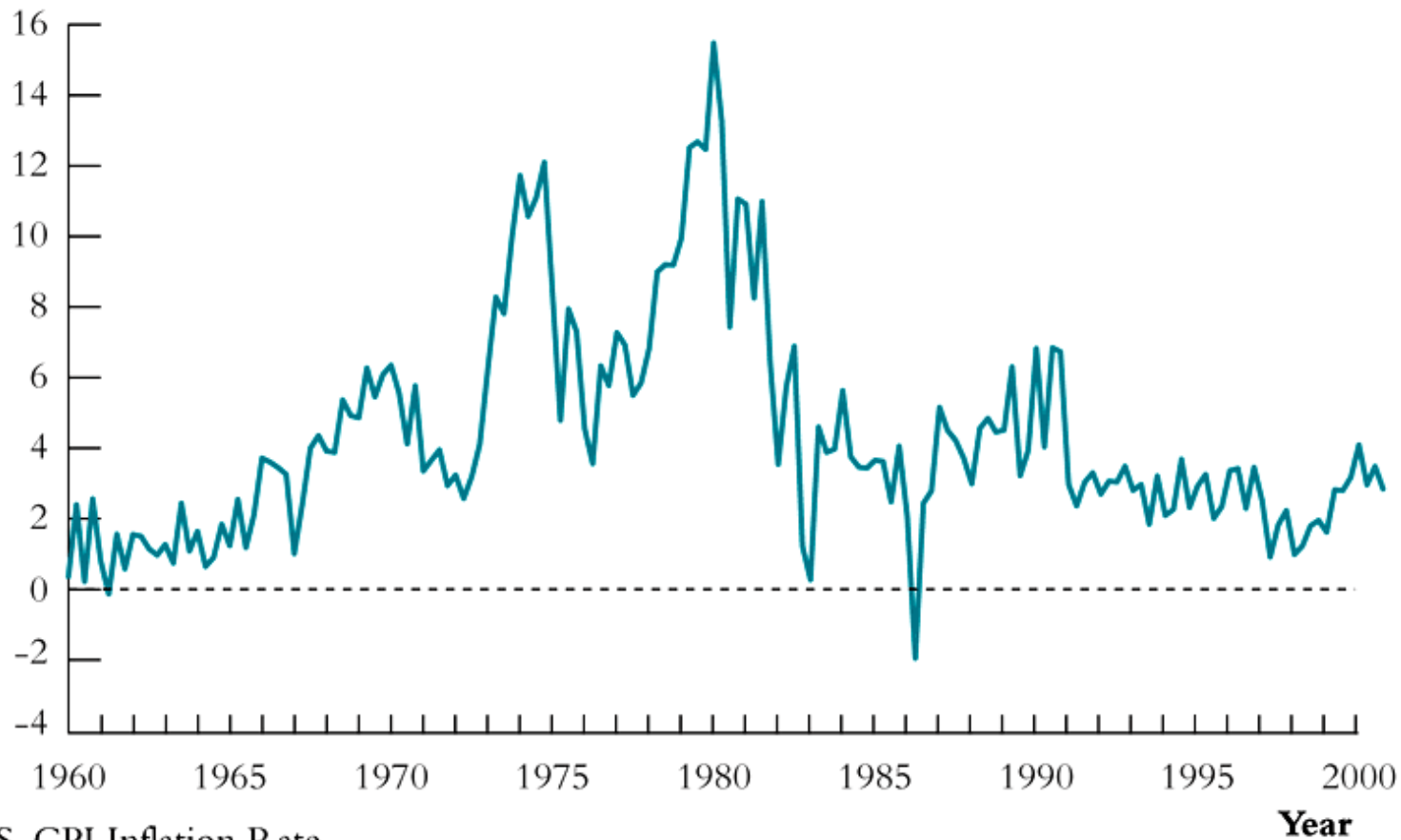


Example #1 of time series data: US rate of inflation

FIGURE 12.1 Inflation and Unemployment in the United States, 1960–1999

Percent per Annum

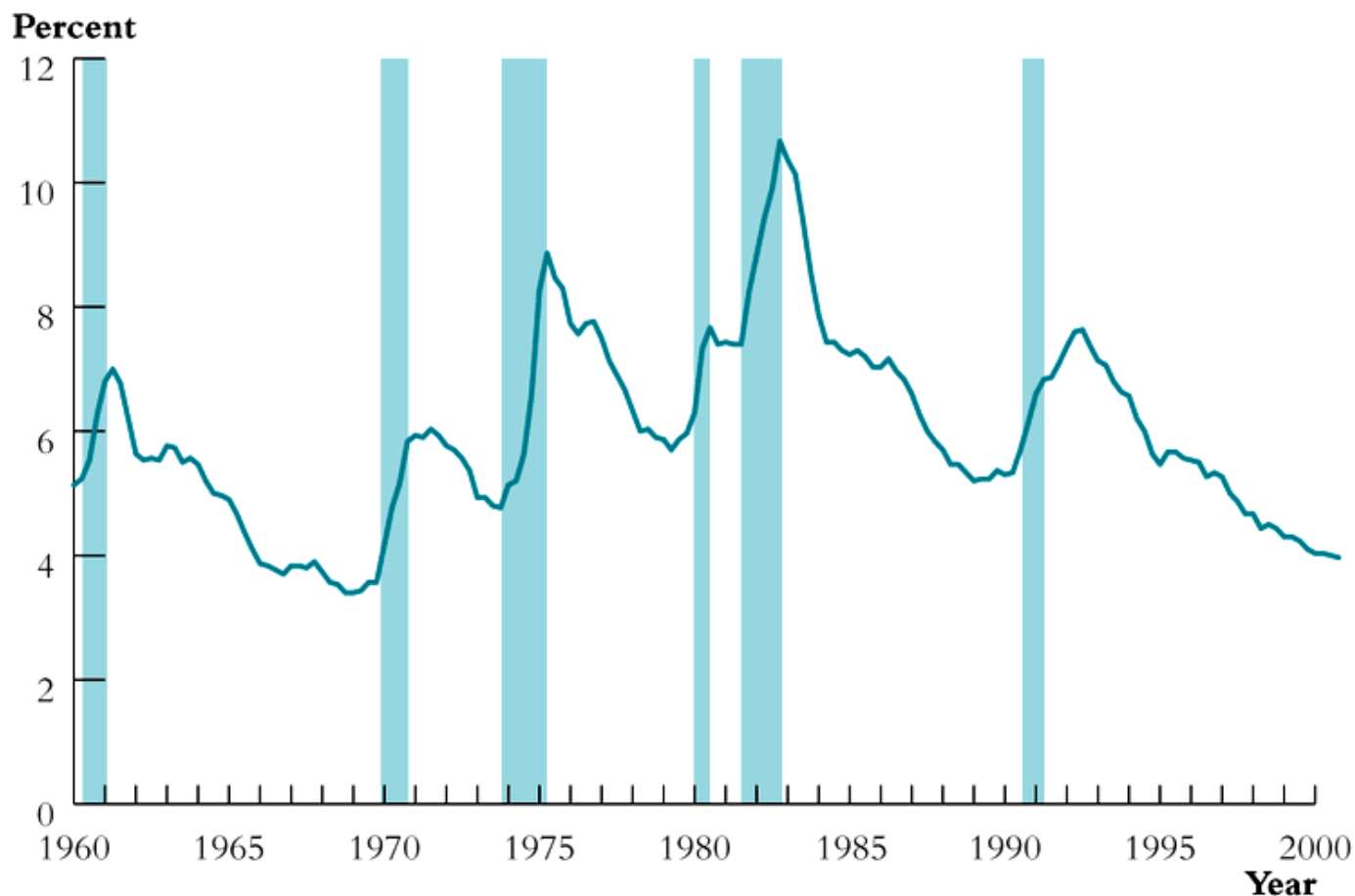


(a) U.S. CPI Inflation Rate

Price inflation in the United States (Figure 12.1a) drifted upwards from 1960 until 1980, and then fell sharply during the early 1980s. The unemployment rate in the United States (Figure 12.1b) rises during recessions (the shaded episodes) and falls during expansions.

Example #2: US rate of unemployment

FIGURE 12.1 Inflation and Unemployment in the United States, 1960–1999



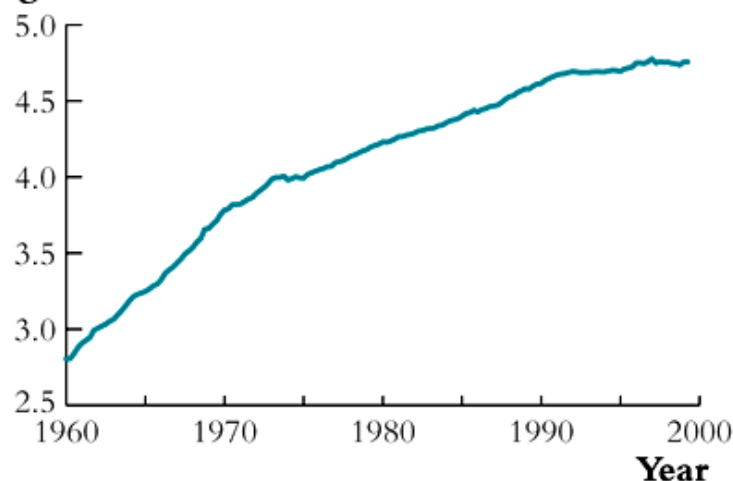
(b) U.S. Unemployment Rate

Price inflation in the United States (Figure 12.1a) drifted upwards from 1960 until 1980, and then fell sharply during the early 1980s. The unemployment rate in the United States (Figure 12.1b) rises during recessions (the shaded episodes) and falls during expansions.

More examples of time series & transformations, ctd.

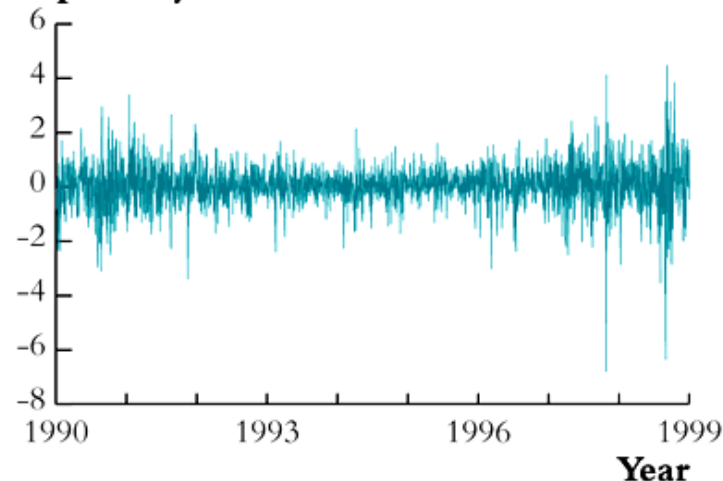
FIGURE 12.2 Four Economic Time Series

Logarithm



(c) Logarithm of Real GDP in Japan

Percent per Day



(d) Percentage Changes in Daily Values of the NYSE Composite Stock Index

The four time series have markedly different patterns. The Federal Funds Rate (Figure 12.2a) has a pattern similar to price inflation. The exchange rate between the U.S. dollar and the British pound (Figure 12.2b) shows a discrete change after the 1972 collapse of the Bretton Woods system of fixed exchange rates. The logarithm of real GDP in Japan (Figure 12.2c) shows relatively smooth growth, although the growth rate decreases in the 1970s and again in the 1990s. The daily returns on the NYSE stock price index (Figure 12.2d) are essentially unpredictable, but its variance changes: this series shows “volatility clustering.”

Example: AR(1) model of inflation – STATA

First, let STATA know you are using time series data

```
generate time=q(1959q1)+_n-1;
```

*_n is the observation no.
So this command creates a new variable
time that has a special quarterly
date format*

```
format time %tq;
```

Specify the quarterly date format

```
sort time;
```

Sort by time

```
tsset time;
```

*Let STATA know that the variable time
is the variable you want to indicate the
time scale*

Example: AR(1) model of inflation – STATA, ctd.

```
gen lcpi = log(cpi);
```

variable cpi is already in memory

```
gen inf = 400*(lcpi[_n]-lcpi[_n-1]);
```

quarterly rate of inflation at an annual rate

```
corrgram inf , noplot lags(8);
```

computes first 8 sample autocorrelations

AG	AC	PAC	Q	Prob>Q
-----	-----	-----	-----	-----
	0.8459	0.8466	116.64	0.0000
	0.7663	0.1742	212.97	0.0000
	0.7646	0.3188	309.48	0.0000
	0.6705	-0.2218	384.18	0.0000
	0.5914	0.0023	442.67	0.0000
	0.5538	-0.0231	494.29	0.0000
	0.4739	-0.0740	532.33	0.0000
	0.3670	-0.1698	555.3	0.0000

```
gen inf = 400*(lcpi[_n]-lcpi[_n-1])
```

This syntax creates a new variable, inf, the "nth" observation of which is 400 times the difference between the nth observation on lcpi and the "n-1"th observation on lcpi, that is, the first difference of lcpi

Example: AR(1) model of inflation – STATA, ctd

Syntax: L.d.inf is the first lag of d.inf ;
d.inf is the first difference of inf

```
. reg d.inf L.d.inf if tin(1962q1,1999q4), r;
```

Regression with robust standard errors

Number of obs = 152
F(1, 150) = 3.96
Prob > F = 0.0484
R-squared = 0.0446
Root MSE = 1.6619

dinf		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
dinf	L1	-.2109525	.1059828	-1.99	0.048	-.4203645	-.0015404
_cons		.0188171	.1350643	0.14	0.889	-.2480572	.2856914

```
if tin(1962q1,1999q4)
```

STATA time series syntax for using only observations between 1962q1 and 1999q4 (inclusive).

This requires defining the time scale first, as we did above

Example: AR(4) model of inflation – STATA

```
. reg dinf L(1/4).d.inf if tin(1962q1,1999q4), r;
```

Regression with robust standard errors

Number of obs = 152
 F(4, 147) = 6.79
 Prob > F = 0.0000
 R-squared = 0.2073
 Root MSE = 1.5292

d.inf		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

inf							
	L1D	-.2078575	.09923	-2.09	0.038	-.4039592	-.0117558
	L2D	-.3161319	.0869203	-3.64	0.000	-.4879068	-.144357
	L3D	.1939669	.0847119	2.29	0.023	.0265565	.3613774
	L4D	-.0356774	.0994384	-0.36	0.720	-.2321909	.1608361
_cons		.0237543	.1239214	0.19	0.848	-.2211434	.268652

NOTES

- *L(1/4).d.inf is A convenient way to say "use lags 1-4 of d.inf as regressors"*
- *L1,...,L4 refer to the first, second,... 4th lags of d.inf*

Example: AR(4) model of inflation – STATA, ctd.

```
. dis "Adjusted Rsquared = " _result(8); result(8) is the rbar-squared  
Adjusted Rsquared = .18576822 of the most recently run regression
```

```
test L2D. L3D. L4D. ; L2.d.inf is the second lag of d.inf, etc.
```

```
( 1)  L2D. inf = 0.0  
( 2)  L3D. inf = 0.0  
( 3)  L4D. inf = 0.0
```

```
F( 3, 147) = 6.43  
Prob > F = 0.0004
```

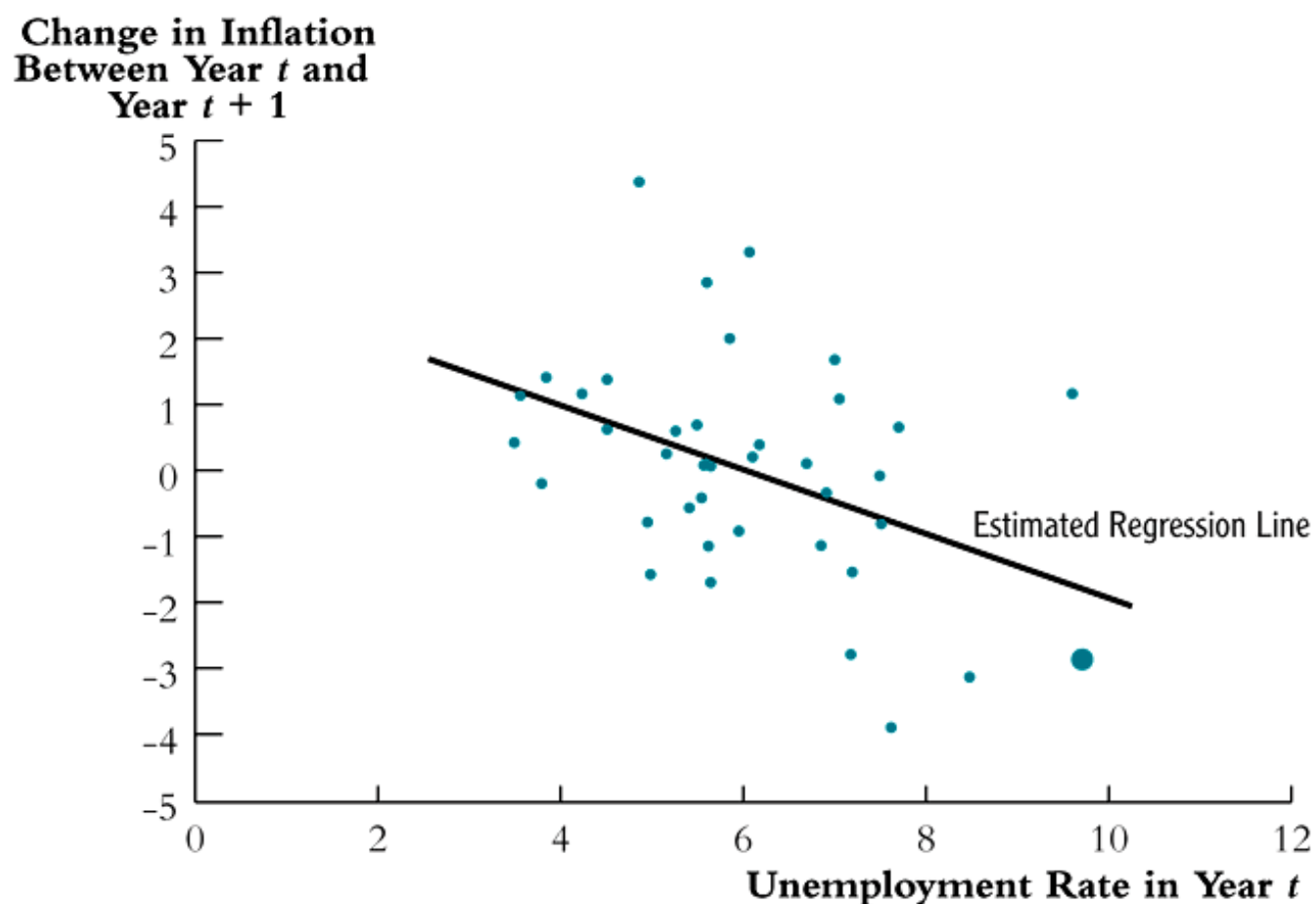

Example: lagged unemployment and inflation

- According to the “Phillips curve” says that if unemployment is above its equilibrium, or “natural,” rate, then the rate of inflation will increase.
- That is, ΔInf_t should be related to lagged values of the unemployment rate, with a negative coefficient
- The rate of unemployment at which inflation neither increases nor decreases is often called the “non-accelerating rate of inflation” unemployment rate: the NAIRU
- Is this relation found in US economic data?
Can this relation be exploited for **forecasting** inflation?

The empirical “Phillips Curve”

FIGURE 12.3 Scatterplot of Change in Inflation Between Year t and Year $t + 1$ vs. the Unemployment Rate in Year t

In 1982, the U.S. unemployment rate was 9.7% and the rate of inflation in 1983 fell by 2.9% (the large dot). In general, high values of the unemployment rate in year t tend to be followed by decreases in the rate of price inflation in the next year, year $t + 1$, with a correlation of -0.40 .



Example: d.inf and unem – STATA

```
. reg d.inf L(1/4).d.inf L(1/4).unem if tin(1962q1,1999q4), r;
```

Regression with robust standard errors

Number of obs = 152
 F(8, 143) = 7.99
 Prob > F = 0.0000
 R-squared = 0.3802
 Root MSE = 1.371

D.inf		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
inf							
	L1D.	-.3629871	.0926338	-3.92	0.000	-.5460956	-.1798786
	L2D.	-.3432017	.100821	-3.40	0.001	-.5424937	-.1439096
	L3D.	.0724654	.0848729	0.85	0.395	-.0953022	.240233
	L4D.	-.0346026	.0868321	-0.40	0.691	-.2062428	.1370377
unem							
	L1	-2.683394	.4723554	-5.68	0.000	-3.617095	-1.749692
	L2	3.432282	.889191	3.86	0.000	1.674625	5.189939
	L3	-1.039755	.8901759	-1.17	0.245	-2.799358	.719849
	L4	.0720316	.4420668	0.16	0.871	-.8017984	.9458615
_cons		1.317834	.4704011	2.80	0.006	.3879961	2.247672

Example: ADL(4,4) model of inflation – STATA, ctd.

```
. dis "Adjusted Rsquared = " _result(8);  
Adjusted Rsquared = .34548812
```

```
. test L2D.inf L3D.inf L4D.inf;
```

```
( 1)  L2D.inf = 0.0  
( 2)  L3D.inf = 0.0  
( 3)  L4D.inf = 0.0
```

```
      F(   3,   143) =    4.93  
      Prob > F =    0.0028
```

The extra lags of d.inf are signif.

```
. test L1.unem L2.unem L3.unem L4.unem;
```

```
( 1)  L.unem = 0.0  
( 2)  L2.unem = 0.0  
( 3)  L3.unem = 0.0  
( 4)  L4.unem = 0.0
```

```
      F(   4,   143) =    8.51  
      Prob > F =    0.0000
```

The lags of unem are significant

The null hypothesis that the coefficients on the lags of the unemployment rate are all zero is rejected at the 1% significance level using the F-statistic