AS.280.347 CLASS 1.2

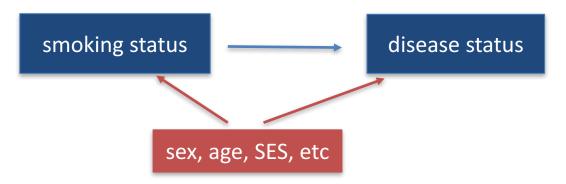
- Look at data displays!
- Review of logistic regression

Module 1: Smoking and risk of disease

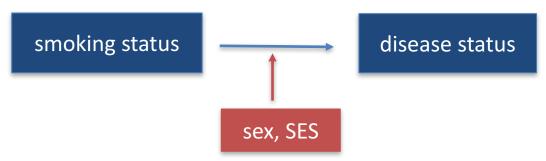
- Question 1.1 (Q1.1): How does the risk of disease compare for smokers and otherwise similar non-smokers?
- Question 1.2 (Q1.2): Does the contribution of smoking to the risk of disease vary by sex or SES?
- To address each question, we want:
 - a data display
 - a statistical analysis
- We will answer these questions using data from the National Medical Expenditures Survey (NMES)

Module 1: Smoking and risk of disease

 Question 1.1 (Q1.1): How does the risk of disease compare for smokers and otherwise similar non-smokers?



 Question 1.2 (Q1.2): Does the contribution of smoking to the risk of disease vary by sex or SES?



Today's agenda

- Group discussion and critiques of NMES data displays to address Q1.1:
 - How does the risk of disease compare for smokers and otherwise similar non-smokers?
- Plans to improve displays
- Review of logistic regression

Questions for discussion

- What are the characteristics of effective displays?
- How can the current displays that address Q1.1 be improved?
- How can the process of working together be improved?
- What statistical analysis will effectively use the NMES data to address Q1.1?
 - Multivariable logistic regression
 - Propensity scores (next week)
- What displays and statistical analysis will address Question 1.2 (Q1.2):
 - Does the contribution of smoking to the risk of disease vary by sex or SES?

In Public Health Biostatistics we used logistic regression to estimate the risk of infant mortality as a function of gestational age, parity and other factors

- Used for binary outcome variables
 - Ex: infant mortality (1=infant died, 0=infant survived)
- Models the log odds of the outcome:
 - $\log(odds \ of \ Y = 1) = \beta_0 + \beta_1 X$
 - Ex: $\log(odds \ of \ death) = \beta_0 + \beta_1 \cdot (gestational \ age)$
- We transform to get probability/risk:
 - $\log(odds \ of \ Y = 1) = \beta_0 + \beta_1 X$
 - $(odds \ of \ Y = 1) = e^{\beta_0 + \beta_1 X}$
 - $P(Y = 1) = \frac{e^{\beta_0 + \beta_1 X}}{1 + e^{\beta_0 + \beta_1 X}}$

Is the mortality risk (or odds) higher for twins than singleton

births?

	Singleton	Twin	Total
Survived	8899	187	9086
Died	526	71	597
Total	9425	258	9683

- Odds of death for twins: 71/187 = .38
- Odds of death for singletons: 526/8899 = .059
- Odds ratio of death for twins as compared to singletons:

OR = (odds for twins)/(odds for sing) = (71/187)/(526/8899) = 6.42

• Log odds ratio: $log_e(OR) = log(6.42) = 1.86$

```
log(odds \ of \ death) = \beta_0 + \beta_1 \cdot twin
  twin = \begin{cases} 1 & if \ twin \\ 0 & if \ singleton \end{cases}
> model1 = glm(death ~ twins, family=binomial(link="logit"))
> summary(model1)
   Coefficients:
                   Estimate Std. Error z value Pr(>|z|)
    (Intercept) -2.82839 0.04487 -63.03 <2e-16 ***
          1.85996 0.14644 12.70 <2e-16 ***
    twins
   Null deviance: 4483.1 on 9682 degrees of freedom
   Residual deviance: 4361.6 on 9681 degrees of freedom
   AIC: 4365.6
   Number of Fisher Scoring iterations: 5
```

```
> model1 = glm(death ~ twins, family=binomial(link="logit"))
> summary(model1)
                                           test
                            SE(\beta)
                                           statistic p-value
   Coefficients:
                  Estimate Std. Error z value Pr(>|z|)
   (Intercept) -2.82839 0.04487 -63.03 <2e-16 ***
   twins
             1.85996 0.14644 12.70 <2e-16 ***
   Null deviance: 4483.1 on 9682 degrees of freedom
   Residual deviance: 4361.6 on 9681 degrees of freedom
   AIC: 4365.6
   Number of Fisher Scoring iterations: 5
    1.86 = \beta_1 = log(OR)
    The log odds of death, comparing twins to singleton births, is 1.86.
    6.42 = e^{1.86} = e^{\beta_1} = 0R
```

The odds of death for twins is 6.42 times the odds of death for singleton births.

- Does the odds of death decrease with increasing gestational age? $log(odds\ of\ death) = \beta_0 + \beta_1 \cdot (gestational\ age)$
- $\log(odds \ of \ death \ | \ ga = 41 \ weeks) = \beta_0 + \beta_1 \cdot (41)$
- $\log(odds \ of \ death \ | \ ga = 40 \ weeks) = \beta_0 + \beta_1 \cdot (40)$
- Difference:

$$\log(odds \mid ga = 41) - \log(odds \mid ga = 40)$$
$$= (\beta_0 + 41\beta_1) - (\beta_0 + 40\beta_1) = \beta_1$$

Log odds ratio:

$$\log(OR) = \log\left(\frac{odds \mid ga=41}{odds \mid ga=40}\right)$$
$$= \log(odds \mid ga=41) - \log(odds \mid ga=40) = \beta_1$$

• Odds ratio: $OR = e^{\log(OR)} = e^{\beta_1}$

```
> model2 = glm(death ~ gestage, family=binomial(link="logit"))
> summary(model2)
   Coefficients:
              Estimate Std. Error z value Pr(>|z|)
   (Intercept) 2.3274 0.3943 5.902 3.59e-09 ***
           -0.1359 0.0108 -12.584 < 2e-16 ***
   gestage
   Null deviance: 4483.1 on 9682 degrees of freedom
   Residual deviance: 4328.0 on 9681 degrees of freedom
   AIC: 4332
```

$$-0.1359 = \beta_1 = log(OR)$$

An additional week of gestational age is associated with a decrease of .14 in the log odds of death.

$$0.87 = e^{-.1359} = e^{\beta_1} = OR$$

An additional week of gestational age is associated with a 13% decrease in the odds of infant death.

- How could we account for any possible confounding variables in a logistic regression analysis?
 - We could include potential confounding variables as covariates in our analysis using multivariable logistic regression:

```
\log(odds \ of \ death) = \beta_0 + \beta_1 \cdot (gestational \ age) + \beta_2 \cdot twin
```

- We interpret the regression coefficients in a multivariable model as ceteris paribus – holding all other things equal
- $-\beta_1 = \log(OR)$ for a one-unit change in gestational age, **holding twin** status constant
- $-\beta_2 = \log(OR)$ comparing twins to singleton births, **holding** gestational age constant

Assignment 1.2

- Q1.1: How does the risk of disease compare for smokers and otherwise similar non-smokers?
 - Improve your data display to answer this question.
 - Fit a logistic regression model to answer this question. Interpret your coefficients and significance tests to answer the question: what does this model say about Q1.1?
 - Work together in groups!
 - Submit your display in R markdown through Blackboard by Sunday @ midnight.
- Consider completing any/all of these available DataCamp modules:
 - Introduction to R
 - Reporting with R Markdown: Authoring R Markdown Reports
 - Reporting with R Markdown: Embedding Code
 - Reporting with R Markdown: Compiling Reports
 - Introduction to the Tidyverse
- Next week in class we will again start with presentations/critiques of your displays.