

**MATH4205/MATH7620 Probability Theory and Stochastic Processes 2025-2026, Semester 1**

**Assignment 3**

Name	Student ID	Marks
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**Instructions:**

- Due date: 11:59 PM 7 Nov 2025 (Fri)
- 20 marks will be deducted for every 24 hours (rounded up) of late submission
- Submit a soft copy (in PDF) to Moodle
- Layout the intermediate steps systematically

1. (20 marks) Let  $\{X_i \mid i \geq 1\}$  be *i.i.d.* random variables with  $P\{X_i = 1\} = P\{X_i = -1\} = 1/2$ . The martingale gambling strategy is: start with stake  $B_1 = 1$ . After a loss, double the next stake. Formally,  $B_{i+1} = 2B_i$  if  $X_i = -1$ . The game continues until the first win, then the game stops. Define the wealth process  $W_0 = 0$  and, for  $n \geq 1$ ,  $W_n = \sum_{i=1}^n B_i X_i$ . Let  $T$  be the stopping time, i.e., the “first win.” Note that under this strategy, if the first  $n$  outcomes are all losses, then  $W_n = -(1 + 2 + \dots + 2^{n-1}) = -(2^n - 1)$ . If  $X_{n+1} = 1$ , then  $W_{n+1} = 2^n - (2^n - 1) = 1$ .
  - Show that  $\{W_n \mid n \geq 0\}$  is a martingale with respect to  $\{X_i \mid i \geq 1\}$ .
  - Following a), the martingale stopping theorem should apply to  $\{W_n \mid n \geq 0\}$ . Compute  $E(W_T)$ . Is  $E(W_T) = E(W_0)$ ? Briefly explain why this conclusion does NOT contradict the Martingale Stopping Theorem. You may refer to lecture notes section 4.3.
2. (20 marks) Consider a random walk on a line which at each step either goes right 1 with probability  $p$  or left 1 with probability  $q = 1 - p$ . Let  $S_n$  be the position after  $n$  steps and let  $S_0 = 0$  be the initial position. Show that  $\left(\frac{q}{p}\right)^{S_n}$ ,  $n \geq 1$ , is a martingale.
3. (20 marks) Let  $\{B(t) \mid t \geq 0\}$  be the standard Brownian motion. Let  $\{X(t) \mid t \geq 0\}$  be a Brownian motion with drift 0.1 and variance parameter 0.09. Give the answer in terms of the CDF  $\Phi(x)$  of the standard normal random variable  $Z$  when appropriate.
  - Determine  $P\{T_{-1} < T_1 < T_{-2}\}$ .
  - Determine  $P\{\max_{0 \leq s \leq 2} B(s) < 0.5\}$ .
  - Determine the conditional distribution of  $X(1)$  given that  $X(2) = 1$ .
  - Determine  $E[e^{X(2)-X(1)}]$ .
4. (20 marks) Suppose that the price of a stock changes according to the following geometric Brownian motion:
 
$$X(t) = 50e^{0.2B(t)+0.1t}.$$
  - What is the probability that the price will exceed \$55 when  $t = 1$ ?
  - Find  $E[X(2)|X(1) = 52]$ .
5. (20 marks) Let  $\{B(t) \mid t \geq 0\}$  be the standard Brownian motion.
  - Show that  $\{Y(t) \mid t \geq 0\}$  is a martingale with respect to  $\{B(t) \mid t \geq 0\}$  when  $Y(t) = B(t)^2 - t$ .
  - Let  $T$  to be the first time that  $B(t)$  either reaches  $-A$  or  $B$ . What is  $E(T)$ ?

Assignment 3.

1. a) To show  $\{W_n\}$  is a martingale, we need to verify  $W_n$  is  $\mathcal{F}_n$ -measurable

$$E[W_n] < \infty \text{ and } E[W_{n+1} | \mathcal{F}_n] = W_n.$$

$W_0 = 0$ , and for  $n \geq 1$ ,  $W_n = \sum_{i=1}^n B_i X_i$  where  $B_1 = 1$ ,  $B_{i+1} = 2B_i$

if  $X_i = -1$  (lose), then  $B_{i+1} = 2B_i$

if  $X_i = 1$  (win), Thus,  $B_{i+1}$  is  $\mathcal{F}_n$ -measurable, and  $W_n$  is  $\mathcal{F}_n$ -measurable

as a sum of measurable terms.

for fixed  $n$ ,  $|W_n| \leq \sum_{i=1}^n 2^i = 2^n - 1$ . so  $E|W_n| < \infty$ .

$W_{n+1} = W_n + B_{n+1} X_{n+1}$ , then

$$E[W_{n+1} | \mathcal{F}_n] = W_n + B_{n+1} E[X_{n+1} | \mathcal{F}_n] = W_n + B_{n+1} \cdot 0 = W_n.$$

Thus,  $\{W_n\}$  is a martingale.

b).  $T$  is the stopping time, so  $X_T = 1$ ,  $X_i = -1$  for  $i < T$ .

Also,  $B_i = 2^{i-1}$  for  $i = 1, 2, \dots, T$ , Thus

$$W_T = \sum_{i=1}^T 2^{i-1}(-1) + 2^T \cdot 1 = 1. \text{ so } E(W_T) = 1.$$

but  $E(W_0) = 0$ . so  $E(W_0) \neq E(W_T)$

in this case,  $T$  isn't bounded,  $P(T > n) = (\frac{1}{2})^n > 0$  for all  $n$

so it doesn't contradict with the theorem.

2.  $S_{n+1} = S_n + X_{n+1}$ , where  $X_{n+1}$  is the increment at step  $n+1$ , independent

$$\text{of } \{S_0, S_1, \dots, S_n\} \text{ with } X_{n+1} = \begin{cases} 1 & P \\ 0 & 1-P=q \end{cases}$$

$$E\left[\left(\frac{q}{p}\right)^{S_{n+1}} \mid S_0, S_1, \dots, S_n\right] = E\left[\left(\frac{q}{p}\right)^{S_n + X_{n+1}} \mid S_n\right] = \left(\frac{q}{p}\right)^{S_n} \cdot E\left[\left(\frac{q}{p}\right)^{X_{n+1}} \mid S_n\right].$$

$$\begin{aligned} E\left[\left(\frac{q}{p}\right)^{X_{n+1}}\right] &= P(X_{n+1}=1) \cdot \left(\frac{q}{p}\right)^1 + P(X_{n+1}=-1) \cdot \left(\frac{q}{p}\right)^{-1} \\ &= P \cdot \frac{q}{p} + q \cdot \frac{p}{q} = q+p = 1-p+p = 1. \end{aligned}$$

$$\text{so } E\left[\left(\frac{q}{p}\right)^{S_{n+1}} \mid S_0, \dots, S_n\right] = \left(\frac{q}{p}\right)^{S_n} \cdot 1 = \left(\frac{q}{p}\right)^{S_n}$$

thus,  $\left(\frac{q}{p}\right)^{S_n}$ ,  $n \geq 1$  is a martingale.

3. a) The probability of a standard Brown motion hitting a boundary "a" before a

boundary "b" is  $\frac{|b|}{|a|+|b|}$ , so  $P(T_1 < T_2) = \frac{|1|}{|1|+|1|} = \frac{1}{2}$

$$P(T_1 < T_2 \mid X_0 = -1) = \frac{|-1|}{|1|+|1|} = \frac{1}{2}, \text{ thus } P(T_1 < T_1 < T_2) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$$

$$\text{b) let } M_t = \max_{0 \leq s \leq t} B(s), \quad P(M_t \leq a) = 2 \underline{\Phi}\left(\frac{a}{\sqrt{t}}\right) - 1 = 2 \underline{\Phi}\left(\frac{0.5}{\sqrt{2}}\right) - 1 = 2 \underline{\Phi}\left(\frac{\sqrt{2}}{4}\right) - 1.$$

c).  $X(t)$  is Brown motion with  $\mu=0.1$  and  $\sigma^2=0.09$ . the increment

$$X(2) - X(1) \sim N(\mu \cdot 1, \sigma^2 \cdot 1) = N(0.1, 0.09).$$

$X(1) \mid X(2)=1$  is normal with  $\mu_1$  and  $\sigma_1^2$ .

$$\mu_1 = X(0) + \frac{t}{T}(X(T) - X(0)) = \frac{1}{2} \cdot 1 = 0.5$$

$$\sigma_1^2 = \sigma^2 t \left(1 - \frac{t}{T}\right) = 0.09 \cdot 1 \cdot \left(1 - \frac{1}{2}\right) = 0.045.$$

thus  $X(1) \mid X(2)=1 \sim N(0.5, 0.045)$ .

d) let  $Y = X(2) - X(1) \sim N(0.1, 0.09)$ , for  $Y \sim N(0.1, 0.09)$ .

$$E[e^Y] = E[e^{X(2) - X(1)}] = e^{0.1 + \frac{0.09}{2}} = e^{0.145}$$

4. a).  $X(1) = 50e^{0.2B(1)+0.1}$ , we need  $P(X(1) > 55)$

$$50e^{0.2B(1)+0.1} > 55 \Rightarrow e^{0.2B(1)+0.1} > 1.1 \Rightarrow 0.2B(1)+0.1 > \ln(1.1)$$

$$B(1) > \frac{\ln(1.1) - 0.1}{0.2}, \text{ since } B(1) \sim N(0, 1).$$

$$\text{so } P(X(1) > 55) = P(B(1) > \frac{\ln(1.1) - 0.1}{0.2}) = \Phi\left(\frac{0.1 - \ln(1.1)}{0.2}\right)$$

$$\text{b)} X(1) = 52, 50e^{0.2B(1)+0.1} = 52 \Rightarrow B(1) = \frac{\ln(1.04) - 0.1}{0.2}$$

$$B(2) = B(1) + (B(2) - B(1)), B(2) - B(1) \sim N(0, 1)$$

$$E[X(2) | X(1) = 52] = E[50e^{0.2B(2)+0.2} | X(1) = 52]$$

$$= E[50e^{0.2[B(1) + (B(2) - B(1))] + 0.2} | X(1) = 52]$$

$$= 50 E[e^{0.2(B(2) - B(1))}] \cdot E[50e^{0.2B(1)+0.2} | X(1) = 52]$$

$$= 50 \cdot e^{0 + \frac{1}{2}(0-2)^2 \cdot 1} \cdot 1.04 \cdot e^{0.1}$$

$$= 52 \cdot e^{0.12}$$

5. a) Since  $Y(t)$  is function of  $B(t)$ , so it's  $\mathcal{F}$ -measurable

$$E[B(t)^2] = t, \text{ so } E[Y(t)] = 0. \quad \text{Var}(Y(t)) = E[(B(t)^2 - t)^2]$$

$$= E[B(t)^4] - 2t E[B(t)^2] + t^2 = 3t^2 - 2t^2 + t^2 = 2t^2$$

so ~~E|Y(t)|~~  $E|Y(t)| < \infty$  for each  $t \geq 0$ .

$$\text{for } 0 \leq s < t, E[Y(t) | \mathcal{F}_s] = E[B(t)^2 - t | \mathcal{F}_s] = E[(B(t) - B(s) + B(s))^2 | \mathcal{F}_s]$$

$$= E[(B(t) - B(s))^2 | \mathcal{F}_s] + 2B(s)E[B(t) - B(s) | \mathcal{F}_s] + B(s)^2 - t$$

since  $B(t) - B(s)$  is independent of  $\mathcal{F}_s$  with  $E(B(t) - B(s)) = 0$ .

$$E[(B(t) - B(s))^2] = t - s.$$

$$(t-s) + 2B(s) \cdot 0 + B(s)^2 - t = B(s)^2 - s = Y(s).$$

so  $\{Y(t)\}$  is a martingale

b)  $Y(t)$  is a martingale, apply optional stopping theorem.

$$E[Y(T)] = E[B(T)^2 - T] = E[B(T)^2] - E[T] = 0.$$

$$\text{So } E[T] = E[B(T)^2].$$

$$\text{P(hit B first)} = \frac{A}{A+B}, \quad \text{P(hit A first)} = \frac{B}{A+B}.$$

$$E[B(T)^2] = \frac{A}{A+B} \cdot B^2 + \frac{B}{A+B} \cdot A^2 = \frac{AB^2 + A^2B}{A+B} = AB$$

$$\text{Thus, } E[T] = AB.$$